CSEP 590B Spring 2011

5 – Motifs: Representation & Discovery

Outline

Previously: Learning from data

MLE: Max Likelihood Estimators

EM: Expectation Maximization (MLE w/hidden data)

These Slides:

Bio: Expression & regulation

Expression: creation of gene products

Regulation: when/where/how much of each gene product; complex and critical

Comp: using MLE/EM to find regulatory motifs in biological sequence data

Gene Expression & Regulation

Gene Expression

Recall a gene is a DNA sequence for a protein To say a gene is expressed means that it is transcribed from DNA to RNA the mRNA is processed in various ways is exported from the nucleus (eukaryotes) is translated into protein A key point: not all genes are expressed all the time, in all cells, or at equal levels

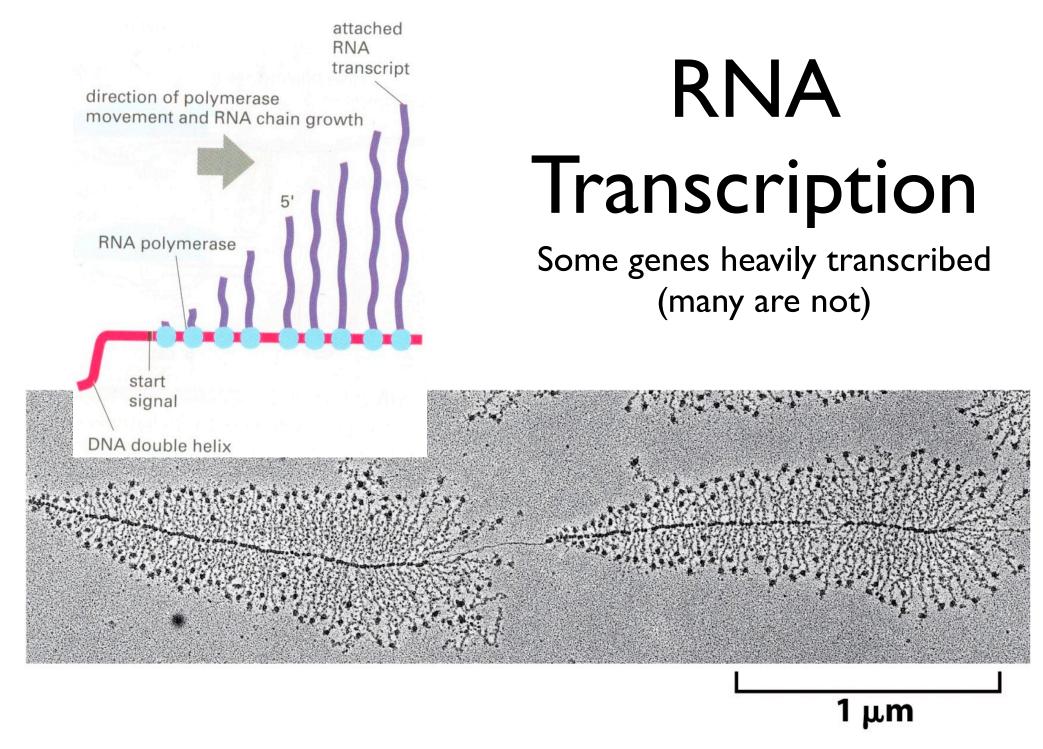


Figure 6-9 Molecular Biology of the Cell 5/e (© Garland Science 2008)

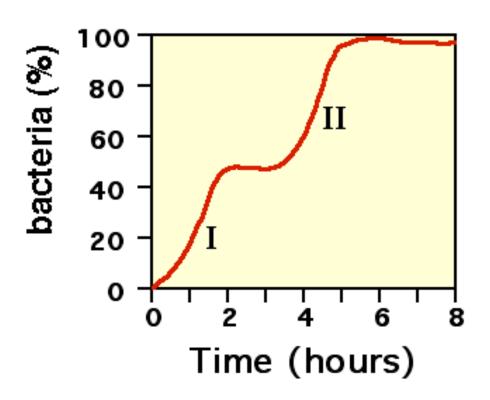
Regulation

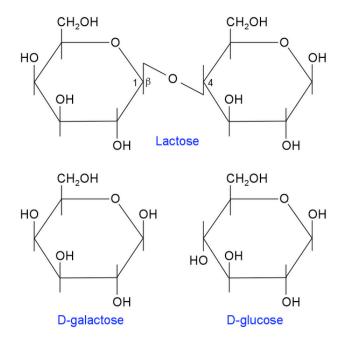
In most cells, pro- or eukaryote, easily a 10,000-fold difference between least- and most-highly expressed genes

Regulation happens at all steps. E.g., some genes are highly transcribed, some are not transcribed at all, some transcripts can be sequestered then released, or rapidly degraded, some are weakly translated, some are very actively translated, ...

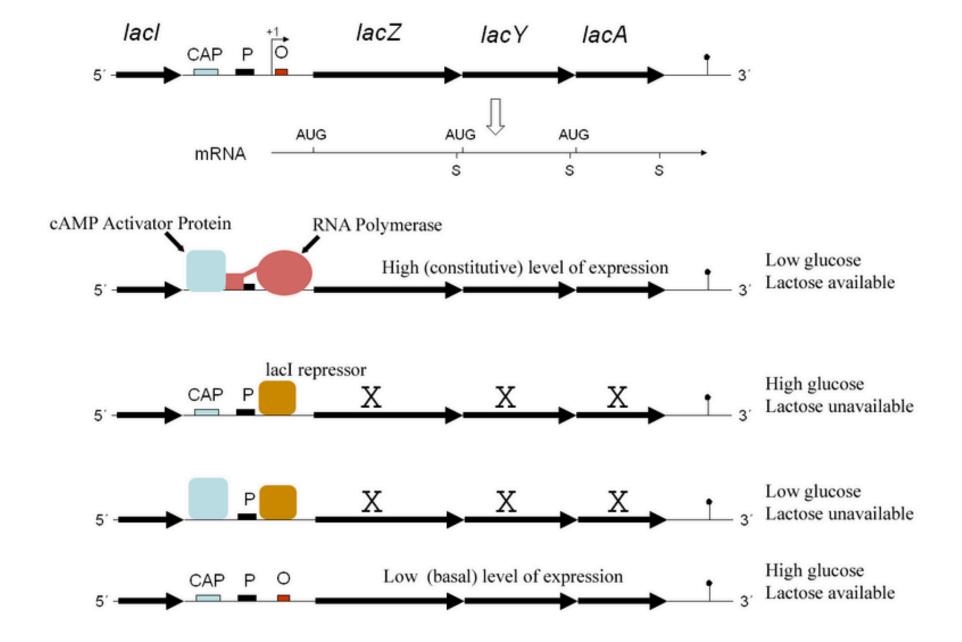
Below, focus on 1st step only: transcriptional regulation

E. coli growth on glucose + lactose





The lac Operon and its Control Elements



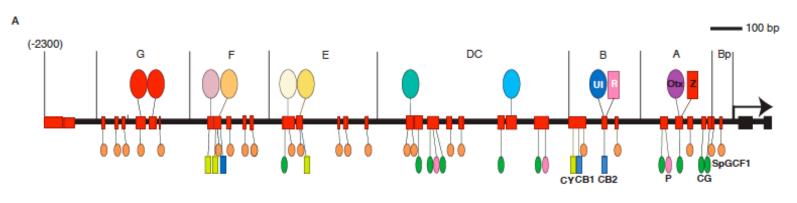
1965 Nobel Prize

Physiology or Medicine

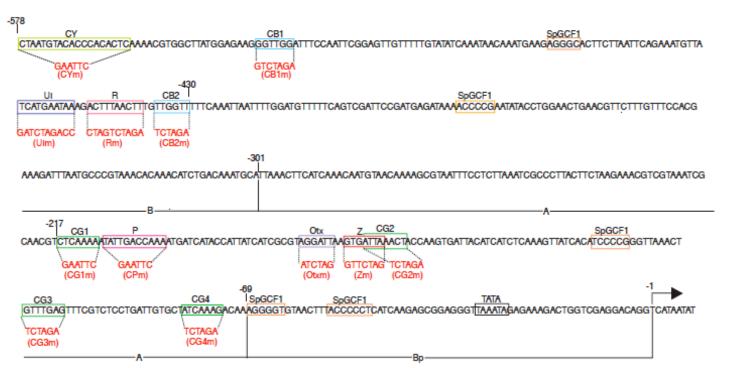
François Jacob, Jacques Monod, André Lwoff

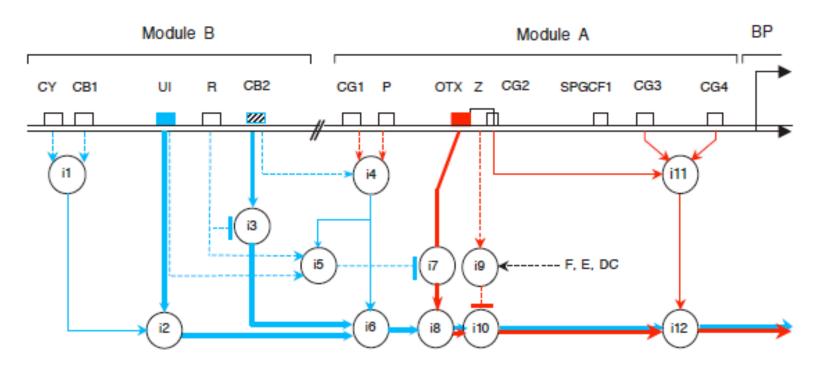


Sea Urchin - Endo 16

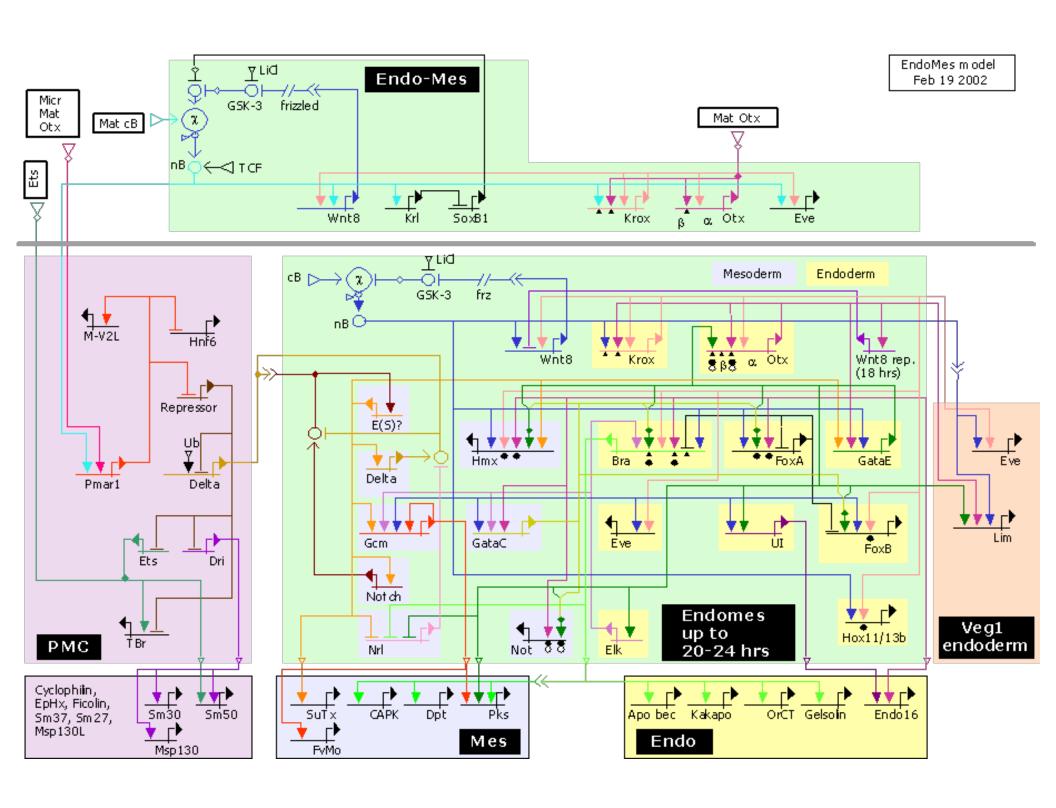


В





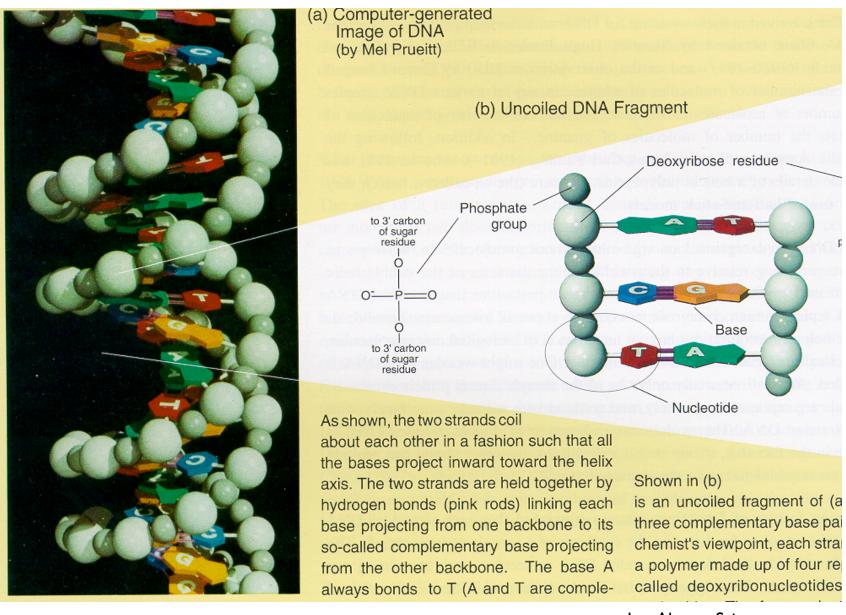
if CY & CB1	i1 = 1	if i5 = 0	i7 = OTX(t)
else	i1 = 0.5	else	i7 = 0
			i8 = i6 + i7
	i2 = i1 • UI(t)		
		if (For Eor DC) & Z	i9 = 1
		else	i9 = 0
if R	i3 = CB2(t)		
else	$i3 = k \cdot CB2(t)$	if i9=1	i10 = 0
	(1 <k<2)< td=""><td>else</td><td>i10 = i8</td></k<2)<>	else	i10 = i8
if P & CG1 & CB2	i4 = 2		
else	i4 = 0	if (CG2 & CG3 & CG4)	i11 = 2
		else	i11 =1
if UI (t)>threshold &R & i4‡0	i5 = 1		i12 = i11·i10
			112 = 111-110
else	i5 = 0		
	i6 = i4 · (i2+i3)		



DNA Binding Proteins

A variety of DNA binding proteins (so-called "transcription factors"; a significant fraction, perhaps 5-10%, of all human proteins) modulate transcription of protein coding genes

The Double Helix



In the groove

Different
patterns of
potential H
bonds at
edges of
different base
pairs,
accessible
esp. in major
groove

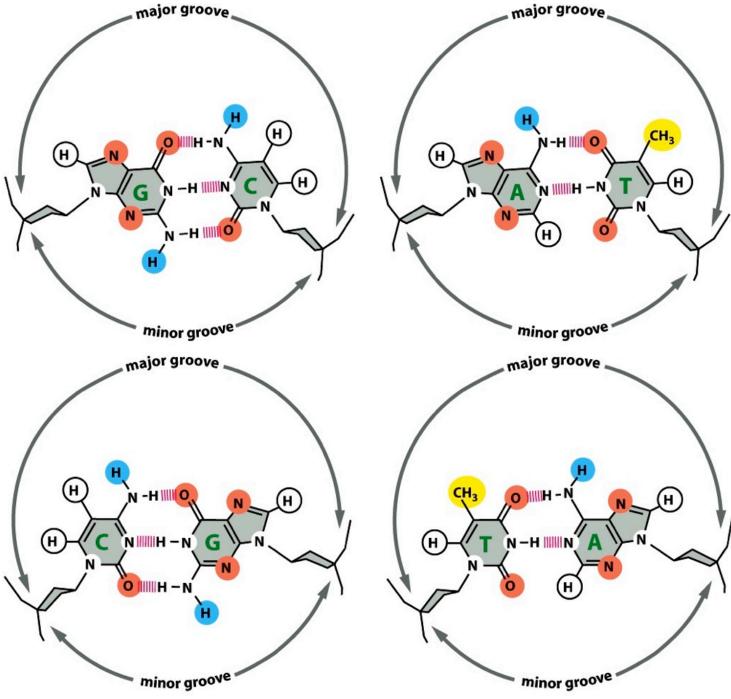


Figure 7-7 Molecular Biology of the Cell 5/e (© Garland Science 2008)

Helix-Turn-Helix DNA Binding Motif

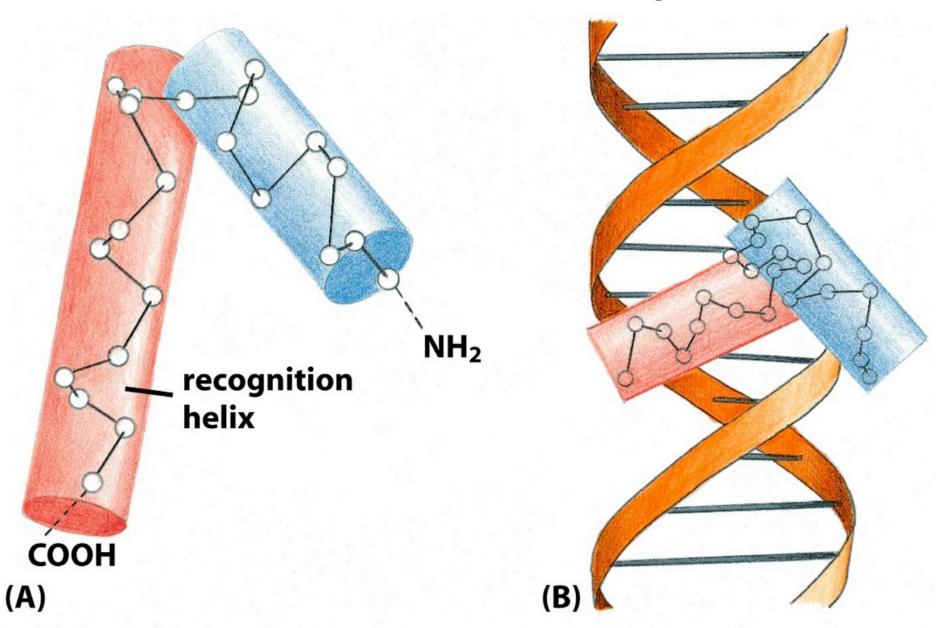


Figure 7-10 Molecular Biology of the Cell 5/e (© Garland Science 2008)

H-T-H Dimers

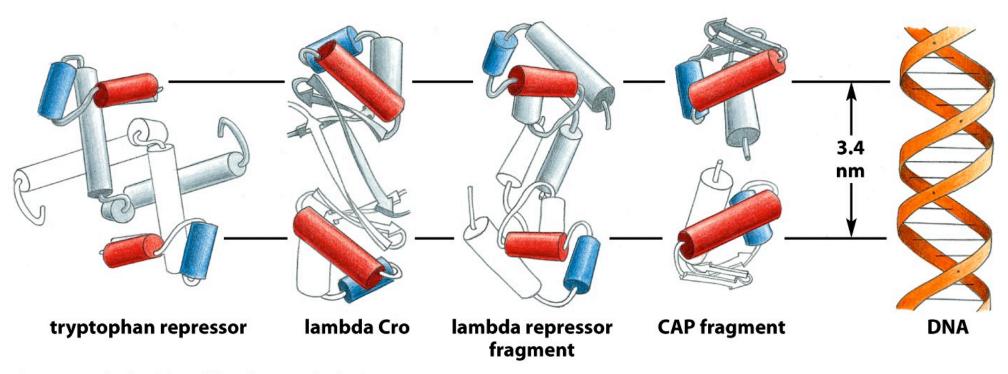


Figure 7-11 Molecular Biology of the Cell 5/e (© Garland Science 2008)

Bind 2 DNA patches, ~ I turn apart Increases both specificity and affinity

Zinc Finger Motif HOOC His 19 COOH СООН (**B**) Arg Arg His (A) NH_2 Arg Arg G Arg (B)

Figure 7-15 Molecular Biology of the Cell Sie (© Garland Science 2008)

Leucine Zipper Motif

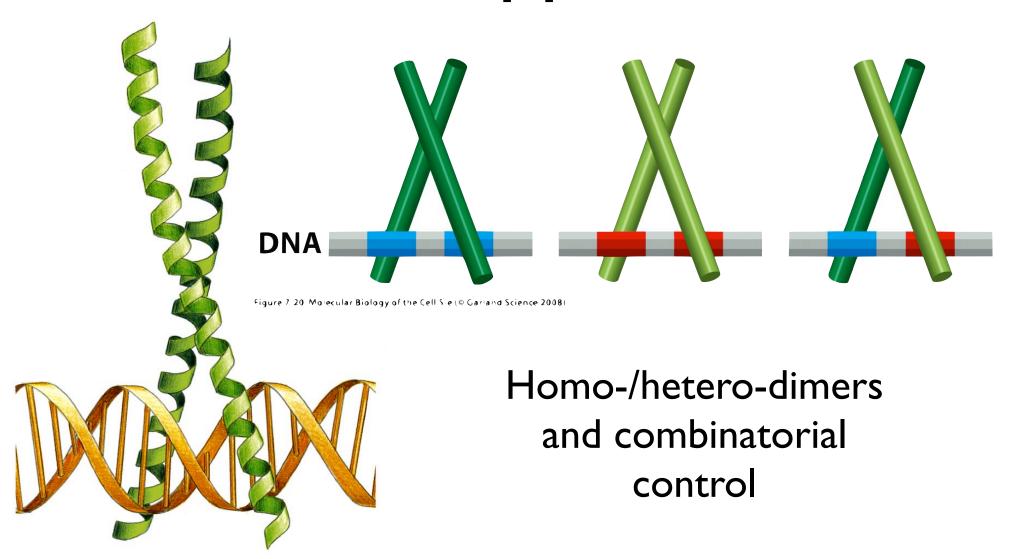


Figure 7-19 Molecular Biology of the Cell 5/e (© Garland Science 2008)



http://www.rcsb.org/pdb/explore/jmol.do?structureId=IMDY&bionumber=I

Some Protein/DNA interactions well-understood

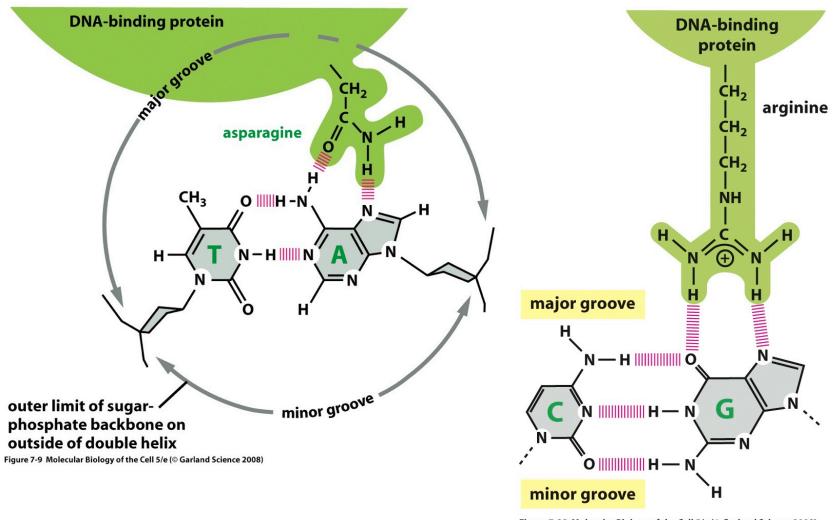
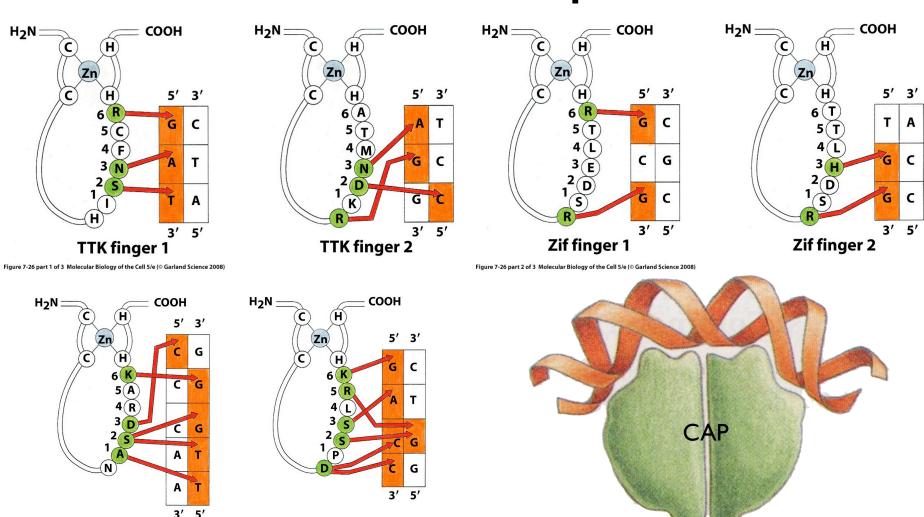


Figure 7-25 Molecular Biology of the Cell 5/e (© Garland Science 2008)

But the overall DNA binding "code" still defies prediction



GLI finger 5

GLI finger 4Figure 7-26 part 3 of 3 Molecular Biology of the Cell 5/e (© Garland Science 2008)

Summary

Proteins can bind DNA to regulate gene expression (i.e., production of other proteins & themselves)

This is widespread

Complex combinatorial control is possible

Sequence Motifs

Motif: "a recurring salient thematic element"

Last few slides described structural motifs in proteins

Equally interesting are the DNA sequence motifs to which these proteins bind - e.g., one leucine zipper dimer might bind (with varying affinities) to dozens or hundreds of similar sequences

DNA binding site summary

Complex "code"

Short patches (4-8 bp)

Often near each other (I turn = 10 bp)

Often reverse-complements

Not perfect matches

E. coli Promoters

"TATA Box" ~ 10bp upstream of transcription start

How to define it? TACGAT

Consensus is TATAAT TAAAAT

BUT all differ from it

GATAAT

Allow k mismatches? TATGAT

Equally weighted? TATGTT

Wildcards like R,Y? ({A,G}, {C,T}, resp.)

E. coli Promoters

- "TATA Box" consensus TATAAT
- ~10bp upstream of transcription start Not exact: of 168 studied (mid 80's)
 - nearly all had 2/3 of TAxyzT
 - 80-90% had all 3
 - 50% agreed in each of x,y,z
 - no perfect match
- Other common features at -35, etc.

TATA Box Frequencies

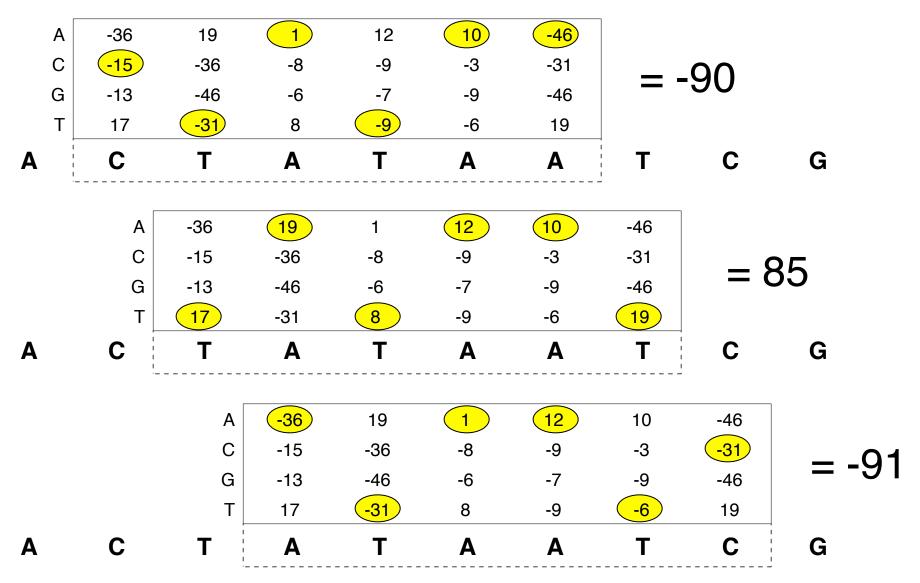
pos base	1	2	3	4	5	6
Α	2	95	26	59	51	1
С	9	2	14	13	20	3
G	10	1	16	15	13	0
Т	79	3	44	13	17	96

TATA Scores

A "Weight Matrix Model" or "WMM"

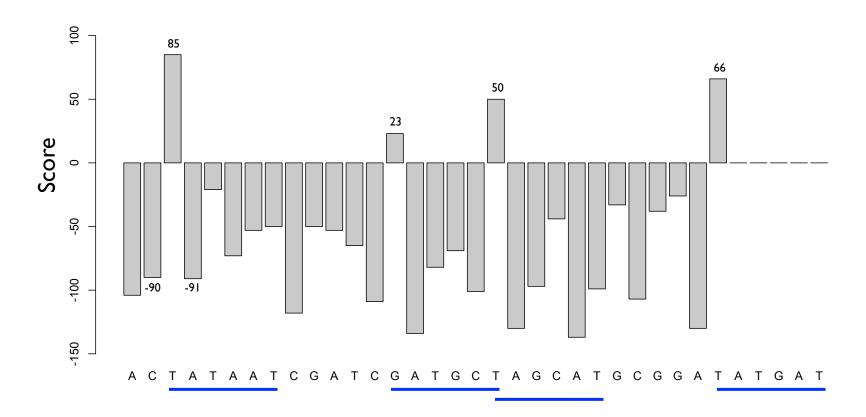
pos base	1	2	3	4	5	6
Α	-36	19	1	12	10	-46
С	-15	-36	-8	-9	- 3	-31
G	-13	-46	-6	-7	-9	-46(?)
T	17	-31	8	-9	-6	19

Scanning for TATA

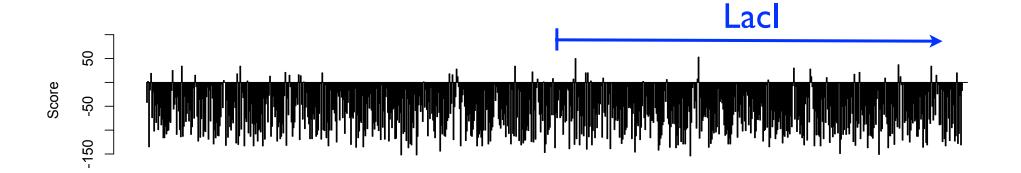


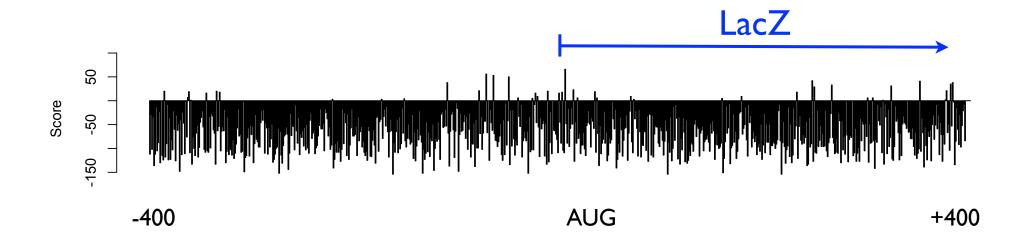
Stormo, Ann. Rev. Biophys. Biophys Chem, 17, 1988, 241-263

Scanning for TATA



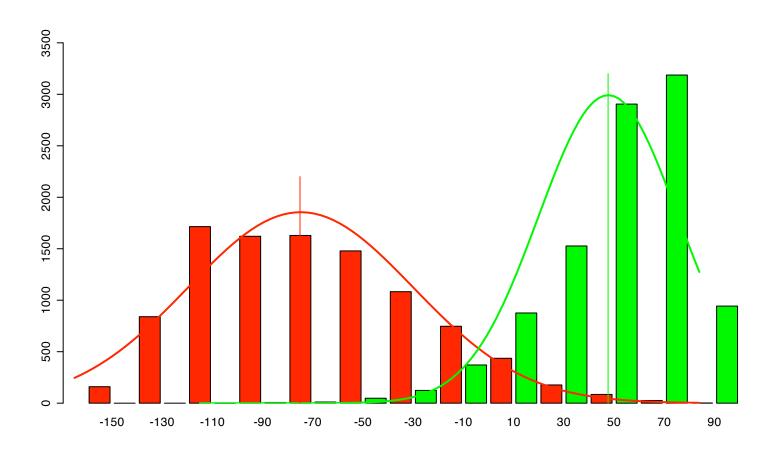
TATA Scan at 2 genes





Score Distribution

(Simulated)



Weight Matrices: Statistics

Assume:

 $f_{b,i}$ = frequency of base b in position i in TATA

 f_b = frequency of base b in all sequences

Log likelihood ratio, given $S = B_1 B_2 ... B_6$:

$$\log \left(\frac{P(S|\text{ "tata"})}{P(S|\text{ "non-tata"})} \right) = \log \frac{\prod_{i=1}^6 f_{B_i,i}}{\prod_{i=1}^6 f_{B_i}} = \sum_{i=1}^6 \log \frac{f_{B_i,i}}{f_{B_i}}$$

Assumes independence

Neyman-Pearson

Given a sample $x_1, x_2, ..., x_n$, from a distribution $f(...|\Theta)$ with parameter Θ , want to test hypothesis $\Theta = \theta_1$ vs $\Theta = \theta_2$.

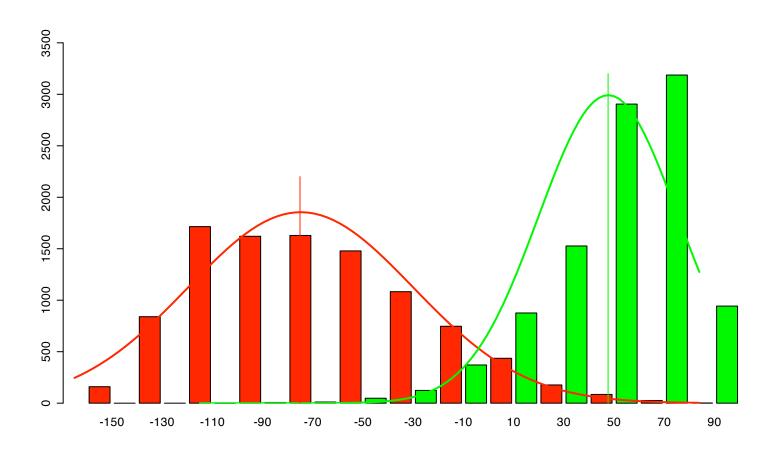
Might as well look at likelihood ratio:

$$\frac{f(x_1, x_2, ..., x_n | \theta_1)}{f(x_1, x_2, ..., x_n | \theta_2)} > \tau$$

(or log likelihood ratio)

Score Distribution

(Simulated)



What's best WMM?

Given, say, 168 sequences $s_1, s_2, ..., s_k$ of length 6, assumed to be generated at random according to a WMM defined by 6 x (4-1) parameters θ , what's the best θ ?

E.g., what's MLE for θ given data $s_1, s_2, ..., s_k$?

Answer: like coin flips or dice rolls, count frequencies per position (see HW?)

Weight Matrices: Chemistry

Experiments show ~80% correlation of log likelihood weight matrix scores to measured binding energy of RNA polymerase to variations on TATAAT consensus [Stormo & Fields]

Another WMM example

8 Sequences:

ATG

ATG

ATG

ATG

ATG

GTG

GTG

TTG

Log-Likelihood Ratio:

$$\log_2 \frac{f_{x_i,i}}{f_{x_i}}, \ f_{x_i} = \frac{1}{4}$$

Freq.	Col I	Col 2	Col 3
Α	0.625	0	0
С	0	0	0
G	0.250	0	
T	0.125		0

LLR	Col I	Col 2	Col 3
Α	1.32	-8	-8
С	8	-8	-8
G	0	-8	2.00
T	-1.00	2.00	-8

Non-uniform Background

- E. coli DNA approximately 25% A, C, G, T
- M. jannaschi 68% A-T, 32% G-C

LLR from previous example, assuming

$$f_A = f_T = 3/8$$

 $f_C = f_G = 1/8$

LLR	Col I	Col 2	Col 3
Α	0.74	8	-8
С	-8	-8	-8
G	1.00	-8	3.00
Т	-1.58	1.42	-8

e.g., G in col 3 is 8 x more likely via WMM than background, so (\log_2) score = 3 (bits).

Relative Entropy

AKA Kullback-Liebler Distance/Divergence, AKA Information Content

Given distributions P, Q

$$H(P||Q) = \sum_{x \in \Omega} P(x) \log \frac{P(x)}{Q(x)} \ge \mathbf{0}$$

Notes:

Let
$$P(x) \log \frac{P(x)}{Q(x)} = 0$$
 if $P(x) = 0$ [since $\lim_{y \to 0} y \log y = 0$]

Undefined if
$$0 = Q(x) < P(x)$$

WMM: How "Informative"? Mean score of site vs bkg?

For any fixed length sequence x, let

P(x) = Prob. of x according to WMM

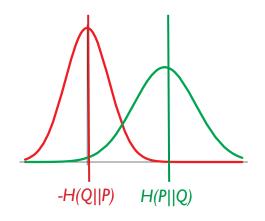
Q(x) = Prob. of x according to background Relative Entropy:

$$H(P||Q) = \sum_{x \in \Omega} P(x) \log_2 \frac{P(x)}{Q(x)}$$

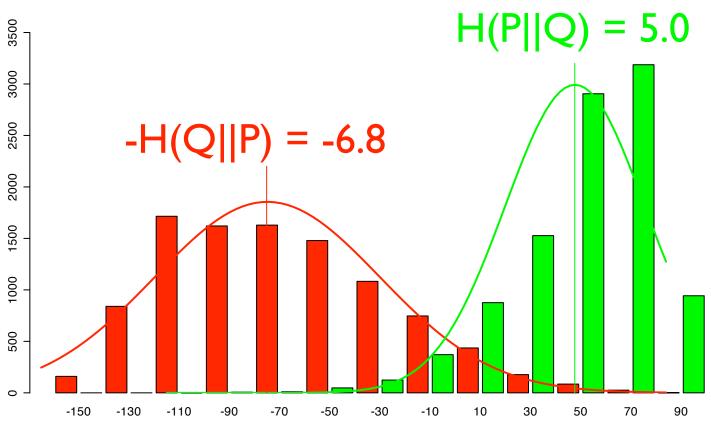
H(P||Q) is expected log likelihood score of a sequence randomly chosen from WMM;

-H(Q||P) is expected score of Background

Expected score difference: H(P||Q) + H(Q||P)



WMM Scores vs Relative Entropy



On average, foreground model scores > background by 11.8 bits (score difference of 118 on 10x scale used in examples above).

For a WMM:

$$H(P||Q) = \sum_{i} H(P_i||Q_i)$$

where P_i and Q_i are the WMM/background distributions for column i.

Proof: exercise

Hint: Use the assumption of independence between WMM columns

WMM Example, cont.

Freq.	Col I	Col 2	Col 3		
Α	0.625	0	0		
С	0	0	0		
G	0.250	0	Ī		
Т	0.125	I	0		

Uniform

LLR	Col I	Col 2	Col 3	
Α	1.32	8	8	
С	8	8	8	
G	0	8	2.00	
Т	-1.00	2.00	8	
RelEnt	0.70	2.00	2.00	4.

Non-uniform

LLR	Col I	Col 2	Col 3
Α	0.74	8	-∞
С	8	8	-∞
G	1.00	-8	3.00
Т	-1.58	1.42	-∞
RelEnt	0.51	1.42	3.00

Pseudocounts

Are the $-\infty$'s a problem?

Certain that a given residue *never* occurs in a given position? Then $-\infty$ just right

Else, it may be a small-sample artifact

Typical fix: add a pseudocount to each observed count—small constant (e.g., .5, I)

Sounds ad hoc; there is a Bayesian justification

WMM Summary

- Weight Matrix Model (aka Position Weight Matrix, PWM, Position Specific Scoring Matrix, PSSM, "possum", 0th order Markov model)
- Simple statistical model assuming independence between adjacent positions
- To build: count (+ pseudocount) letter frequency per position, log likelihood ratio to background
- To scan: add LLRs per position, compare to threshold
- Generalizations to higher order models (i.e., letter frequency per position, conditional on neighbor) also possible, with enough training data

How-to Questions

Given aligned motif instances, build model? Frequency counts (above, maybe w/ pseudocounts)

Given a model, find (probable) instances Scanning, as above

Given unaligned strings thought to contain a motif, find it? (e.g., upstream regions of coexpressed genes)

Hard ... rest of lecture.

Motif Discovery

Unfortunately, finding a site of max relative entropy in a set of unaligned sequences is NP-hard [Akutsu]

Motif Discovery: 4 example approaches

Brute Force

Greedy search

Expectation Maximization

Gibbs sampler

Brute Force

Input:

Motif length L, plus sequences s_1 , s_2 , ..., s_k (all of length n+L-1, say), each with one instance of an unknown motif

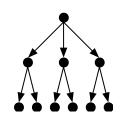
Algorithm:

Build all k-tuples of length L subsequences, one from each of $s_1, s_2, ..., s_k$ (n^k such tuples)

Compute relative entropy of each

Pick best

Brute Force, II



Input:

Motif length L, plus seqs s_1 , s_2 , ..., s_k (all of length n+L-1, say), each with one instance of an unknown motif

Algorithm in more detail:

Build singletons: each len L subseq of each s_1 , s_2 , ..., s_k (nk sets)

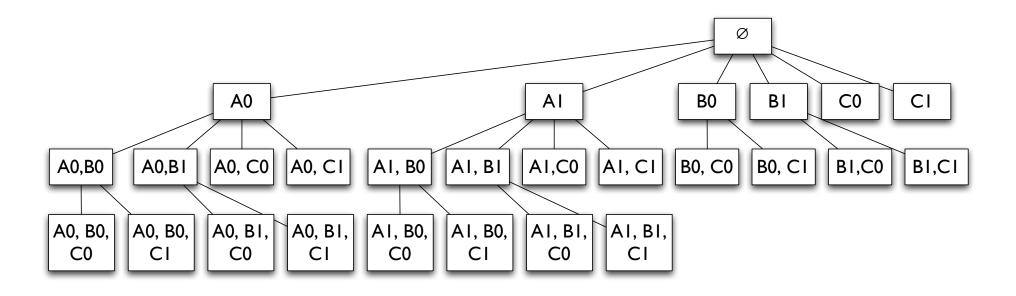
Extend to pairs: len L subseqs of each pair of seqs $\binom{n^2\binom{k}{2}}{n^2}$ sets)

Then triples: len L subseqs of each triple of seqs $(n^3 \binom{k}{3})$ sets)

Repeat until all have k sequences $(n^k \binom{k}{k})$ sets)

Compute relative entropy of each; pick best

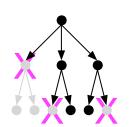
Example



Three sequences (A, B, C), each with two possible motif positions (0, I)

Greedy Best-First

[Hertz, Hartzell & Stormo, 1989, 1990]



Input:

Sequences $s_1, s_2, ..., s_k$; motif length L;

"breadth" d, say d = 1000

Algorithm:

As in brute, but discard all but best d relative entropies at each stage

7=P

Expectation Maximization [MEME, Bailey & Elkan, 1995]

Input (as above):

Sequence s_1 , s_2 , ..., s_k ; motif length l; background model; again assume one instance per sequence (variants possible)

Algorithm: EM

Visible data: the sequences

Hidden data: where's the motif

$$Y_{i,j} = \begin{cases} 1 & \text{if motif in sequence } i \text{ begins at position } j \\ 0 & \text{otherwise} \end{cases}$$

Parameters θ : The WMM

MEME Outline

Typical EM algorithm:

Parameters θ^t at t^{th} iteration, used to estimate where the motif instances are (the hidden variables)

Use those estimates to re-estimate the parameters θ to maximize likelihood of observed data, giving θ^{t+1} Repeat

Key: given a few good matches to best motif, expect to pick more

Expectation Step

(where are the motif instances?)

$$\widehat{Y}_{i,j} = E(Y_{i,j} \mid s_i, \theta^t) \xrightarrow{\mathbb{P}^{(0)^{l+1}} \cdot P^{(1)}}$$

$$= P(Y_{i,j} = 1 \mid s_i, \theta^t) \xrightarrow{\mathbb{P}(Y_{i,j} = 1 \mid \theta^t)}$$

$$= P(s_i \mid Y_{i,j} = 1, \theta^t) \frac{P(Y_{i,j} = 1 \mid \theta^t)}{P(s_i \mid \theta^t)}$$

$$= cP(s_i \mid Y_{i,j} = 1, \theta^t)$$

$$= c' \prod_{k=1}^{l} P(s_{i,j+k-1} \mid \theta^t)$$

Sequence i

where c' is chosen so that $\sum_{j} \widehat{Y}_{i,j} = 1$.

Maximization Step

(what is the motif?)

Find θ maximizing expected value:

$$Q(\theta \mid \theta^{t}) = E_{Y \sim \theta^{t}} [\log P(s, Y \mid \theta)]$$

$$= E_{Y \sim \theta^{t}} [\log \prod_{i=1}^{k} P(s_{i}, Y_{i} \mid \theta)]$$

$$= E_{Y \sim \theta^{t}} [\sum_{i=1}^{k} \log P(s_{i}, Y_{i} \mid \theta)]$$

$$= E_{Y \sim \theta^{t}} [\sum_{i=1}^{k} \sum_{j=1}^{|s_{i}|-l+1} Y_{i,j} \log P(s_{i}, Y_{i,j} = 1 \mid \theta)]$$

$$= E_{Y \sim \theta^{t}} [\sum_{i=1}^{k} \sum_{j=1}^{|s_{i}|-l+1} Y_{i,j} \log (P(s_{i} \mid Y_{i,j} = 1, \theta) P(Y_{i,j} = 1 \mid \theta))]$$

$$= \sum_{i=1}^{k} \sum_{j=1}^{|s_{i}|-l+1} E_{Y \sim \theta^{t}} [Y_{i,j}] \log P(s_{i} \mid Y_{i,j} = 1, \theta) + C$$

$$= \sum_{i=1}^{k} \sum_{j=1}^{|s_{i}|-l+1} \widehat{Y}_{i,j} \log P(s_{i} \mid Y_{i,j} = 1, \theta) + C$$

M-Step (cont.)

$$Q(\theta \mid \theta^t) = \sum_{i=1}^k \sum_{j=1}^{|s_i|-l+1} \widehat{Y}_{i,j} \log P(s_i \mid Y_{i,j} = 1, \theta) + C$$

Exercise: Show this is maximized by "counting" letter frequencies over all possible motif instances, with counts weighted by $\widehat{Y}_{i,j}$, again the "obvious" thing.

$$s_1: \ \mathsf{ACGGATT}...$$
 $s_k: \ \mathsf{GC}...\mathsf{TCGGAC}$
 $\widehat{Y}_{1,1} \ \mathsf{ACGG}$
 $\widehat{Y}_{1,2} \ \mathsf{CGGA}$
 $\widehat{Y}_{1,3} \ \mathsf{GGAT}$
 $\vdots \ \vdots$
 $\widehat{Y}_{k,l-1} \ \mathsf{CGGA}$
 $\widehat{Y}_{k,l} \ \mathsf{GGAC}$

Initialization

- I. Try every motif-length substring, and use as initial θ a WMM with, say, 80% of weight on that sequence, rest uniform
- 2. Run a few iterations of each
- 3. Run best few to convergence

(Having a supercomputer helps):

http://meme.sdsc.edu/

Another Motif Discovery Approach The Gibbs Sampler

Lawrence, et al. "Detecting Subtle Sequence Signals: A Gibbs Sampling Strategy for Multiple Sequence Alignment," Science 1993

Sigma-37	223	TIDIATYTONK	SQKETGDILGISQMHVSR	LORKAVKKLR	240	A25944	
SpoIIIC			TORETAKELGISRSYVSR		111	A28627	
NahR			RVSITAENLGLTQPAVSN		39	A32837	
Antennapedia			RRIEIAHALCLTERQIKI		343	A23450	
NtrC (Brady.)			NQIRAADLIGLNRNTLRK		466	B26499	
DicA			TORSI AKAL KISHVSVSQ	-	39	B24328	(BVECDA)
MerD	5		TVSRIALDAGVSVHIVRD		22	C29010	
Fis	73	•	NOTRA ALMMGINRGTLRK		90	A32142	(DNECFS)
MAT a1	99	FRRKQSLNSK	EKEEVAKKOGITPLQVRV	WFINKRMRSK	116	A90983	(JEBY1)
Lambda cII	25	SALLNKIAML	GTEKTAEAVGVDKSQISR	WKRDWIPKFS	42	A03579	(QCBP2L)
Crp (CAP)	169	THPDGMQIKI	TRQEIGQIVGCSRETVGR	ILKMLEDQNL	186	A03553	(QRECC)
Lambda Cro	15	ITLKDYAMRF	GQTKTAKDLGVYQSAINK	AIHAGRKIFL	32	A03577	(RCBPL)
P22 Cro	12	YKKDVIDHFG	TQRAVAKALGISDAAVSQ	WKÉVIPEKDA	29	A25867	(RGBP22)
AraC	196	ISDHLADSNF	DIASVAQHVCLSPSRLSH	LFRQQLGISV	213	A03554	(RGECA)
Fnr	196	FSPREFRLTM	TRGDIGNYLGLTVETISR	LLGRFQKSGM	213	A03552	(RGECF)
HtpR	252	ARWLDEDNKS	TLQEL ADRYGVSAERVRQ	LEKNAMKKLR	269	A00700	(RGECH)
NtrC (K.a.)	444	LTTALRHTQG	HKQEA ARLLGWGRNTLTR	KLKELGME	461	A03564	(RGKBCP)
CytR	11	MKAKKQETAA	TMKDVALKAKVSTATVSR	ALMNPDKVSQ	28	A24963	(RPECCT)
DeoR	23	LQELKRSDKL	HLKDAALLGVSEMTIRR	DLNNHSAPVV	40	A24076	(RPECDO)
GalR	3	MA	TIKDVARLAGVSVATVSR	VINNSPKASE	20	A03559	(RPECG)
LacI	5	MKPV	TLYDVAEYAGVSYQTVSR	VVNQASHVSA	22	A03558	(RPECL)
TetR	26	LLNEVGIEGL	TTRKI AQKLGVEQPTLYW	HVKNKRALLD	43	A03576	(RPECTN)
TrpR	67	IVEELLRGEM	SQRELKNELGAGIATITR	GSNSLKAAPV	84	A03568	(RPECW)
NifA			VQAKAARLLGMTPRQVAY	-	512	s02513	
SpoliG	205	RFGLVGEEEK	TQKDVADMMGISQSYISR	LEKRIIKRLR	222	s07337	
, Pin	160		PRQKVAIIYDVGVSTLYK		177	s07958	
PurR	. 3		TIKDVAKRANVSTTTVSH		20	S08477	
EbgR	3		TLKDIA I EAGVSLATVSR		20	s09205	
LexA			TRAEIA QRLGFRSPNAAE		44	S11945	
P22 cI	25	SSILNRIAIR	GQRKVADALGINESQISR		42	B25867	(Z1BPC2)
			****	***			
			6 10				

В								Posit	ion i	n site								
	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18
Arg	94	222	265	137	9	9	137	137	9	9	9	52	222	94	94	9	265	606
Lys	9	133	442	380	9	71	380	194	9	133	9	9	71	9	9	9	71	256
Glu	53	9	96	401	9	9	140	140	9	9	9	53	140	140	9	9	9	53
Asp	67	9	9	473	9	9	299	125	9	67	9	67	67	9	9	9	9	67
Gln	9	600	224	9	9	9	224	9	9	9	9	9	278	63	278	9	9	170
His	240	9	´ 9	9	9	9	125	125	9	9	9	9	125	125	125	9	9	240
Asn	168	9	9	9	9	9	168	89	9	89	9	248	9	168	89	9	89	89
Ser	117	9	117	117	9	9	9	9	9	9	9	819	63	387	63	9	819	9
Gly	151	9	56	9	9	151	9	9	9	1141	9	151	9	56	9	9	56	9
Ala	.9	9	112	43	181	901	43	181	215	9	43	9	43	181	112	. 43	78	9
Thr	915	130	130	9	251	9	9	9	9	9	9	311	130	70	855	و ً	130	9
Pro	76	9	9	9	9	9	9	9	9	9	9	9	210	210	9	9	9,	9
Cys	9	9	9	9	9	9	9	9	295	581	295	9	9	9	9	9	. 9	9
Val	58	107	9	9	500	9	9	. 9	156	9	598	9	205	58	9	746	9	58
Leu	9	121	9	9	149	9	93	149	458	9	149	9	37	37	9	177	9	9
Ile	9	166	114	61	323	9	114	166	9	9	427	9.	61	9	61	427	9	61
Met	9	104	9	9	9	9	9	198	198	9	104	9	9	198	9	9	9	9
Tyr	9	9	136	9	· 9	9	9	262	262	9	9	136	136	9	262	9	262	136
Phe	9	9	9	9	9	9	9	9	9	9	108	9	9	9	9	9	9	9
${ t Trp}$	9	9	9	9	9	9	9	9	9	9	366	9	9	9	9	9	9	366

Some History

Geman & Geman, IEEE PAMI 1984

Hastings, Biometrika, 1970

Metropolis, Rosenbluth, Rosenbluth, Teller, & Teller, "Equations of State Calculations by Fast Computing Machines," J. Chem. Phys. 1953

Josiah Williard Gibbs, 1839-1903, American physicist, a pioneer of thermodynamics

How to Average

An old problem:

n random variables:

Joint distribution (p.d.f.):

Some function:

Want Expected Value:

$$x_1, x_2, \dots, x_k$$
 $P(x_1, x_2, \dots, x_k)$
 $f(x_1, x_2, \dots, x_k)$
 $E(f(x_1, x_2, \dots, x_k))$

How to Average

$$E(f(x_1, x_2, \dots, x_k)) = \int_{x_1} \int_{x_2} \dots \int_{x_k} f(x_1, x_2, \dots, x_k) \cdot P(x_1, x_2, \dots, x_k) dx_1 dx_2 \dots dx_k$$

Approach I: direct integration (rarely solvable analytically, esp. in high dim)

Approach 2: numerical integration (often difficult, e.g., unstable, esp. in high dim)

Approach 3: Monte Carlo integration

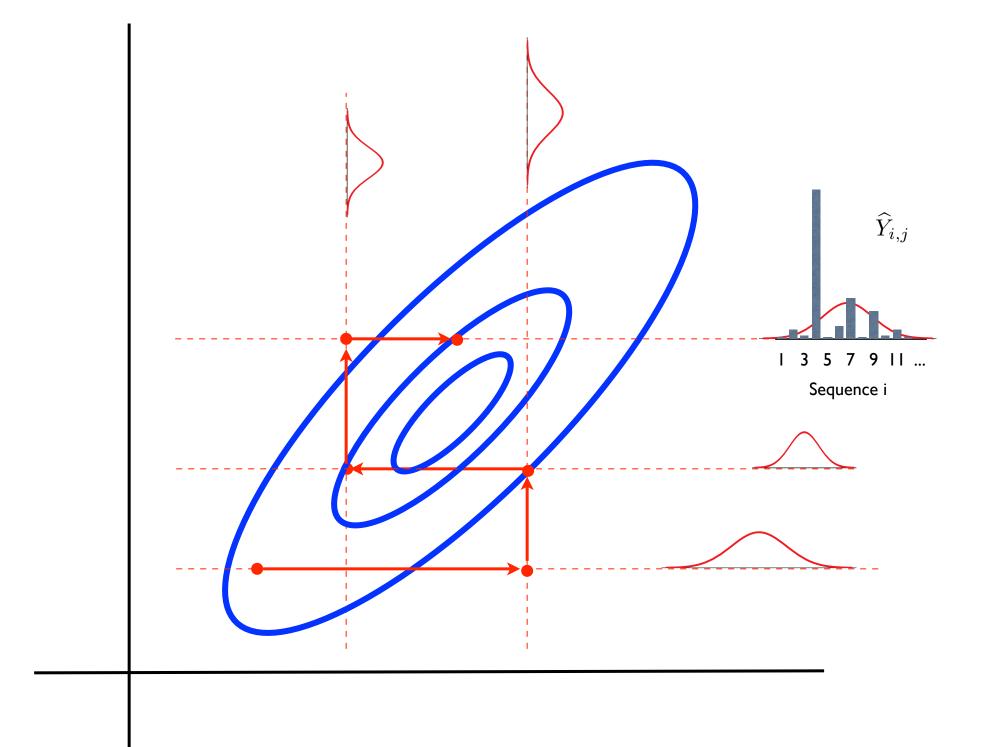
sample $\vec{x}^{(1)}, \vec{x}^{(2)}, \dots \vec{x}^{(n)} \sim P(\vec{x})$ and average:

$$E(f(\vec{x})) \approx \frac{1}{n} \sum_{i=1}^{n} f(\vec{x}^{(i)})$$

Markov Chain Monte Carlo (MCMC)

- Independent sampling also often hard, but not required for expectation
- MCMC $ec{X}_{t+1} \sim P(ec{X}_{t+1} \mid ec{X}_t)$ w/ stationary dist = P
- Simplest & most common: Gibbs Sampling $P(x_i \mid x_1, x_2, \dots, x_{i-1}, x_{i+1}, \dots, x_k)$
- Algorithm

for t = I to
$$\infty$$
 $t+1$ t for i = I to k do: $x_{t+1,i} \sim P(x_{t+1,i} \mid x_{t+1,1}, x_{t+1,2}, \dots, x_{t+1,i-1}, x_{t,i+1}, \dots, x_{t,k})$



Input: again assume sequences $s_1, s_2, ..., s_k$ with one length w motif per sequence

Motif model: WMM

Parameters: Where are the motifs?

for
$$1 \le i \le k$$
, have $1 \le x_i \le |s_i| - w + 1$

"Full conditional": to calc

$$P(x_i = j \mid x_1, x_2, \dots, x_{i-1}, x_{i+1}, \dots, x_k)$$

build WMM from motifs in all sequences except i, then calc prob that motif in i^{th} sequences occurs at j by usual "scanning" alg.

Overall Gibbs Alg

```
Randomly initialize x_i's
```

for
$$t = 1$$
 to ∞
for $i = 1$ to k
discard motif instance from s_i ;
recalc WMM from rest
for $j = 1 ... |s_i| - w + 1$

Similar to MEME, but it would average over, rather than sample from

calculate prob that ith motif is at j:

average over,
$$P(x_i = j \mid x_1, x_2, \dots, x_{i-1}, x_{i+1}, \dots, x_k)$$

pick new x_i according to that distribution

Issues

Burnin - how long must we run the chain to reach stationarity?

Mixing - how long a post-burnin sample must we take to get a good sample of the stationary distribution? In particular:

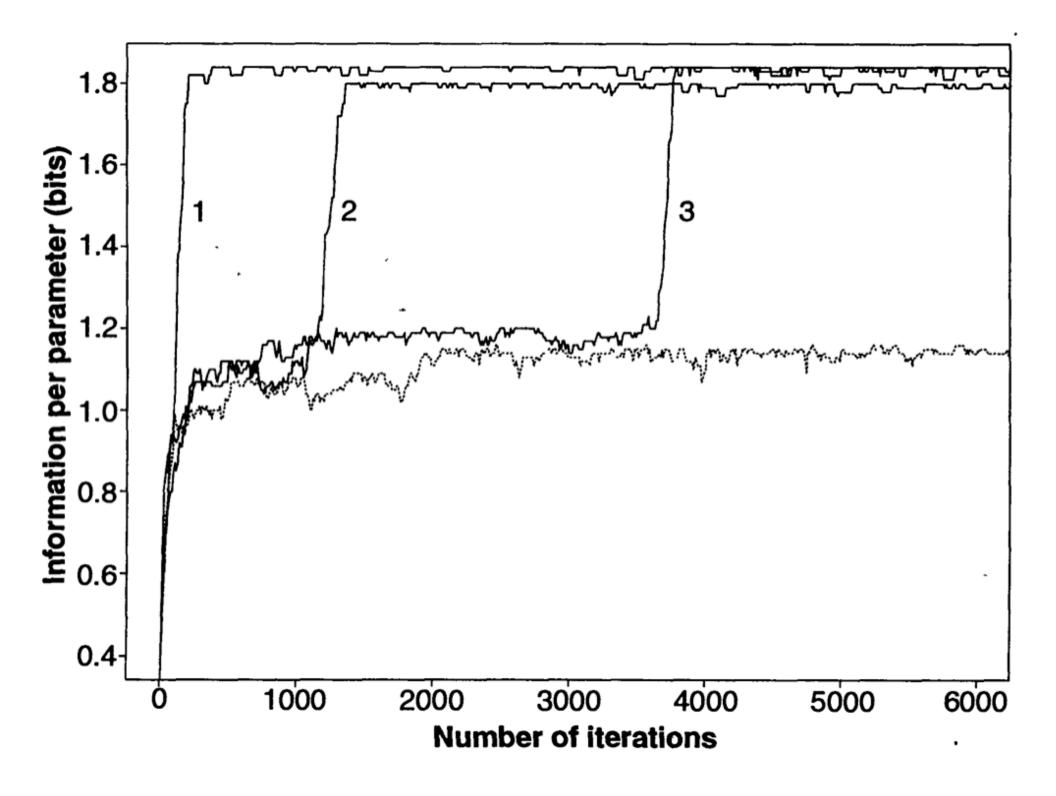
Samples are not independent; may not "move" freely through the sample space Many isolated modes

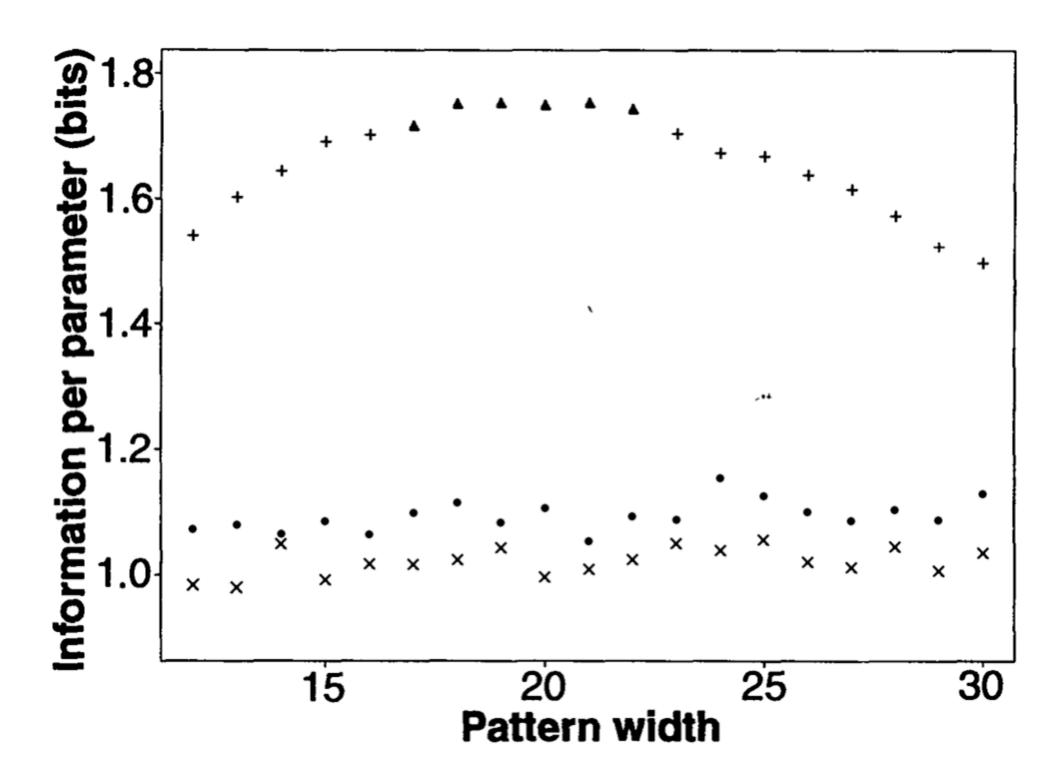
Variants & Extensions

"Phase Shift" - may settle on suboptimal solution that overlaps part of motif. Periodically try moving all motif instances a few spaces left or right.

Algorithmic adjustment of pattern width: Periodically add/remove flanking positions to maximize (roughly) average relative entropy per position

Multiple patterns per string





Assessing computational tools for the discovery of transcription factor binding sites

Martin Tompa^{1,2}, Nan Li¹, Timothy L Bailey³, George M Church⁴, Bart De Moor⁵, Eleazar Eskin⁶, Alexander V Favorov^{7,8}, Martin C Frith⁹, Yutao Fu⁹, W James Kent¹⁰, Vsevolod J Makeev^{7,8}, Andrei A Mironov^{7,11}, William Stafford Noble^{1,2}, Giulio Pavesi¹², Graziano Pesole¹³, Mireille Régnier¹⁴, Nicolas Simonis¹⁵, Saurabh Sinha¹⁶, Gert Thijs⁵, Jacques van Helden¹⁵, Mathias Vandenbogaert¹⁴, Zhiping Weng⁹, Christopher Workman¹⁷, Chun Ye¹⁸ & Zhou Zhu⁴

Methodology

13 tools

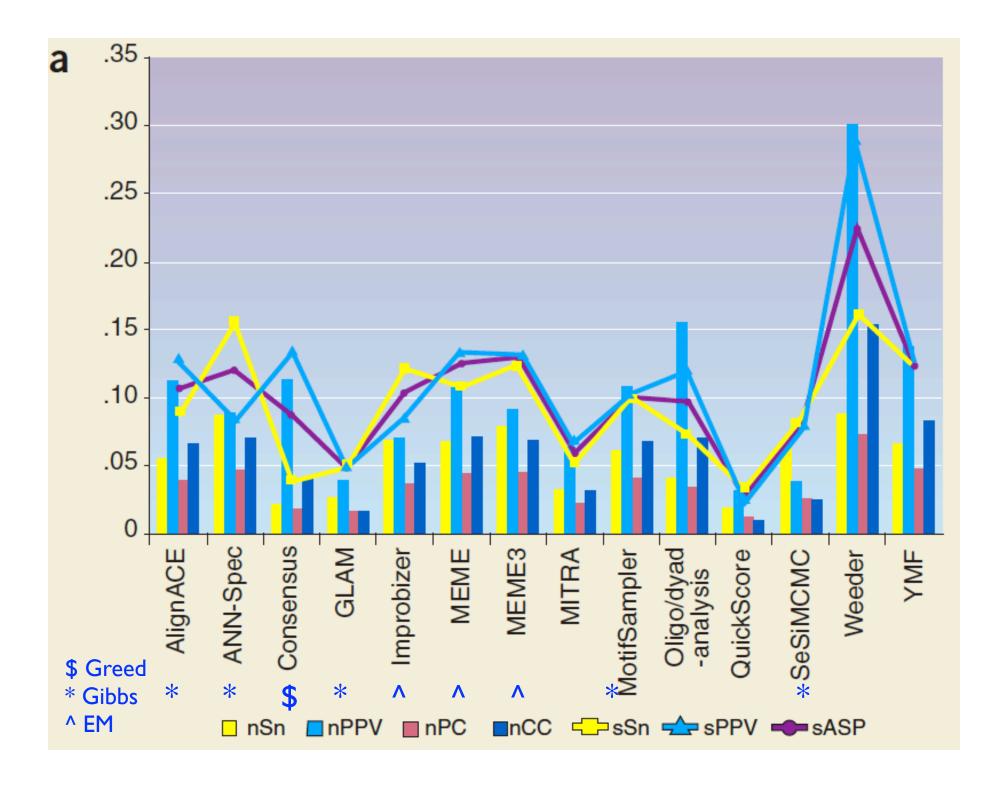
Real 'motifs' (Transfac)

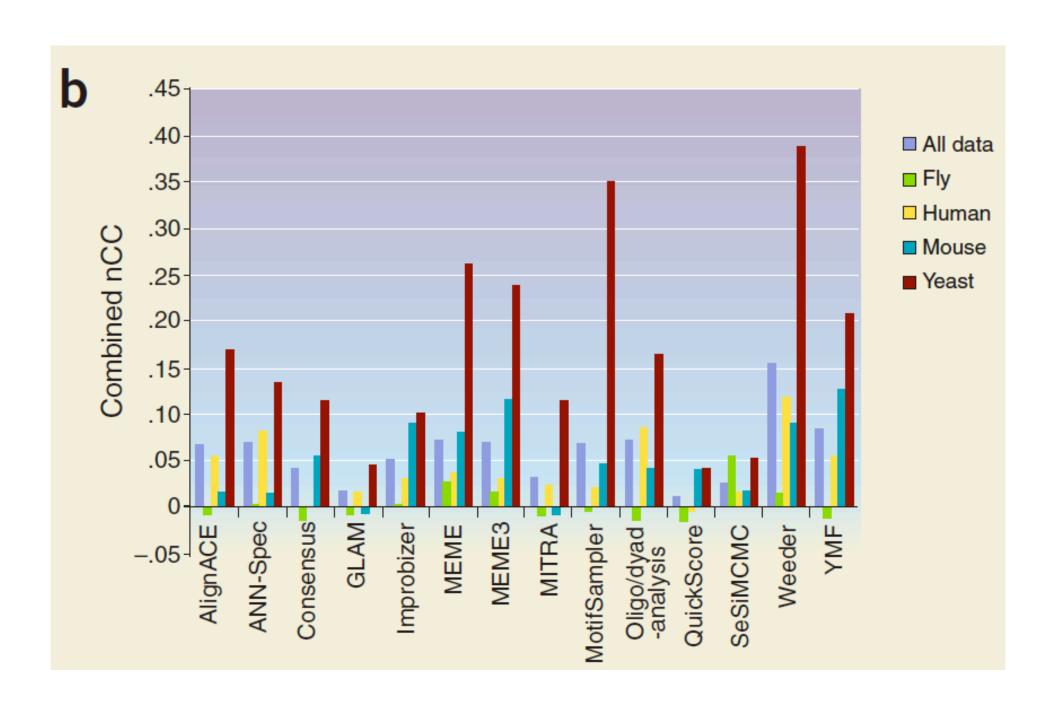
56 data sets (human, mouse, fly, yeast)

'Real', 'generic', 'Markov'

Expert users, top prediction only

"Blind" - sort of





Lessons

Evaluation is hard (esp. when "truth" is unknown)

Accuracy low

partly reflects limitations in evaluation methodology (e.g. ≤ I prediction per data set; results better in synth data)

partly reflects difficult task, limited knowledge (e.g. yeast > others)

No clear winner re methods or models

Motif Discovery Summary

Important problem: a key to understanding gene regulation

Hard problem: short, degenerate signals amidst much noise

Many variants have been tried, for representation, search, and discovery. We looked at only a few:

Weight matrix models for representation & search

Greedy, MEME and Gibbs for discovery

Still much room for improvement. Comparative genomics, i.e. cross-species comparison is very promising