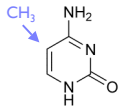
CSEP 590 A
Lecture 6
Markov Models and Hidden Markov Models

DNA Methylation

CpG - 2 adjacent nts, same strand (not Watson-Crick pair; "p" mnemonic for the phosphodiester bond of the DNA backbone)

C of CpG is often (70-80%) methylated in mammals i.e., CH3 group added (both strands)



cytosine

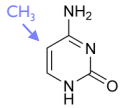
Why? Generally silences transcription.
X-inactivation, imprinting, repression of mobile elements, some cancers, aging, and *developmental differentiation*

How? DNA methyltransferases convert hemi- to fully-methylated

Major exception: promoters of housekeeping genes

"CpG Islands"

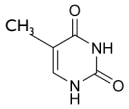
Methyl-C mutates to T relatively easily



cytosine

Net: CpG is less common than expected genome-wide:
f(CpG) < f(C)*f(G)

BUT in promoter (& other) regions, CpG remain unmethylated, so CpG → TpG less likely there: makes "CpG Islands"; often mark gene-rich regions



thymine

CpG Islands

CpG Islands
- More CpG than elsewhere
- More C & G than elsewhere, too
- Typical length: few 100 to few 1000 bp

Questions
- Is a short sequence (say, 200 bp) a CpG island or not?
- Given long sequence (say, 10-100kb), find CpG islands?

Markov & Hidden Markov Models

References:

Durbin, Eddy, Krogh and Mitchison, "Biological Sequence Analysis", Cambridge, 1998

Rabiner, "A Tutorial on Hidden Markov Models and Selected Application in Speech Recognition," Proceedings of the IEEE, v 77 #2,Feb 1989, 257-286

Independence

A key issue: All models we've talked about so far assume *independence* of nucleotides in different positions - definitely unrealistic.

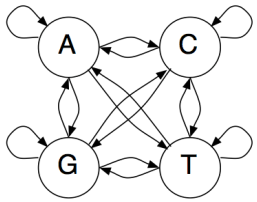
Markov Chains

A sequence $x_1, x_2, \ldots$ of random variables is a *k-th order Markov chain* if, for all *i*, i^{th} value is independent of all but the previous *k* values:

$$P(x_i \mid x_1, x_2, \ldots, x_{i-1}) = P(x_i \mid x_{i-k}, x_{i-k+1}, \ldots, x_{i-1})$$

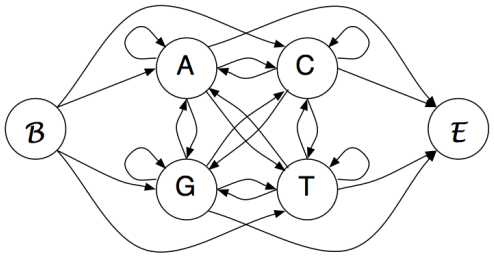
Example 1: Uniform random ACGT } 0th order
Example 2: Weight matrix model
Example 3: ACGT, but ↓ Pr(G following C) } 1st order

A Markov Model (1st order)



States: A,C,G,T
Emissions: corresponding letter
Transitions: $a_{st} = P(x_i = t \mid x_{i-1} = s)$ ←1st order

A Markov Model (1st order)



States: A,C,G,T
Emissions: corresponding letter
Transitions: $a_{st} = P(x_i = t \mid x_{i-1} = s)$
ℬegin/ℰnd states

Pr of emitting sequence *x*

$$
\begin{aligned}
x &= x_1\, x_2\, \ldots\, x_n \\
P(x) &= P(x_1, x_2, \ldots, x_n) \\
&= P(x_1) \cdot P(x_2 \mid x_1) \cdots P(x_n \mid x_{n-1}, \ldots, x_1) \\
&= P(x_1) \cdot P(x_2 \mid x_1) \cdots P(x_n \mid x_{n-1}) \\
&= P(x_1) \prod_{i=1}^{n-1} a_{x_i, x_{i+1}} \\
&= \prod_{i=0}^{n-1} a_{x_i, x_{i+1}} \quad \text{(with Begin state)}
\end{aligned}
$$

Training

Max likelihood estimates for transition probabilities are just the frequencies of transitions when emitting the training sequences

E.g., from 48 CpG islands in 60k bp:

+	A	C	G	T
A	0.180	0.274	0.426	0.120
C	0.171	0.368	0.274	0.188
G	0.161	0.339	0.375	0.125
T	0.079	0.355	0.384	0.182

−	A	C	G	T
A	0.300	0.205	0.285	0.210
C	0.322	0.298	0.078	0.302
G	0.248	0.246	0.298	0.208
T	0.177	0.239	0.292	0.292

Discrimination/Classification

Log likelihood ratio of CpG model vs background model

$$
S(x) = \log \frac{P(x|\text{model}+)}{P(x|\text{model}-)} = \sum_{i=1}^{L} \log \frac{a^{+}_{x_{i-1}x_i}}{a^{-}_{x_{i-1}x_i}} = \sum_{i=1}^{L} \beta_{x_{i-1}x_i}
$$

β	A	C	G	T
A	−0.740	0.419	0.580	−0.803
C	−0.913	0.302	1.812	−0.685
G	−0.624	0.461	0.331	−0.730
T	−1.169	0.573	0.393	−0.679

CpG Island Scores

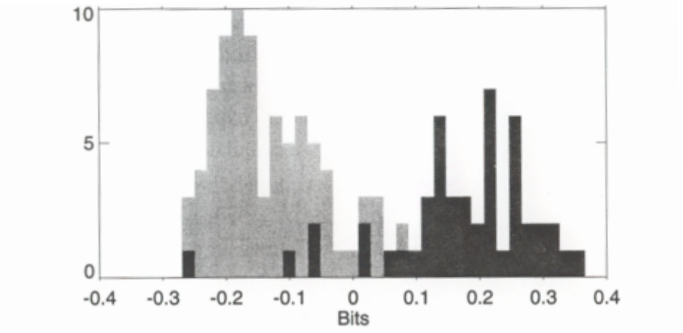
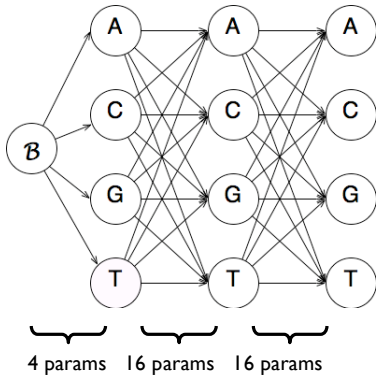


Figure 3.2 *The histogram of the length-normalised scores for all the sequences. CpG islands are shown with dark grey and non-CpG with light grey.*

Aside: 1st Order "WMM"



Questions

Q1: Given a *short* sequence, is it more likely from feature model or background model? Above

Q2: Given a *long* sequence, where are the features in it (if any)

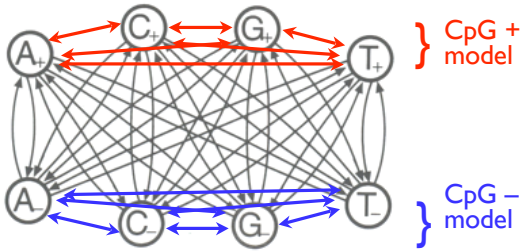
Approach 1: score 100 bp (e.g.) windows

Pro: simple

Con: arbitrary, fixed length, inflexible

Approach 2: *combine* +/- models.

Combined Model



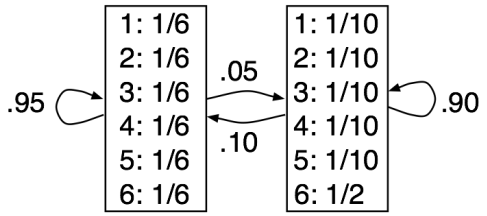
Emphasis is "Which (hidden) state?" not "Which model?"

Hidden Markov Models (HMMs)

States:	$1, 2, 3, \ldots$
Paths:	sequences of states $\pi = (\pi_1, \pi_2, \ldots)$
Transitions:	$a_{k,l} \quad = P(\pi_i = l \mid \pi_{i-1} = k)$
Emissions:	$e_k(b) = P(x_i = b \mid \pi_i = k)$
Observed data:	emission sequence
Hidden data:	state/transition sequence

The Occasionally Dishonest Casino

1 fair die, 1 "loaded" die, occasionally swapped



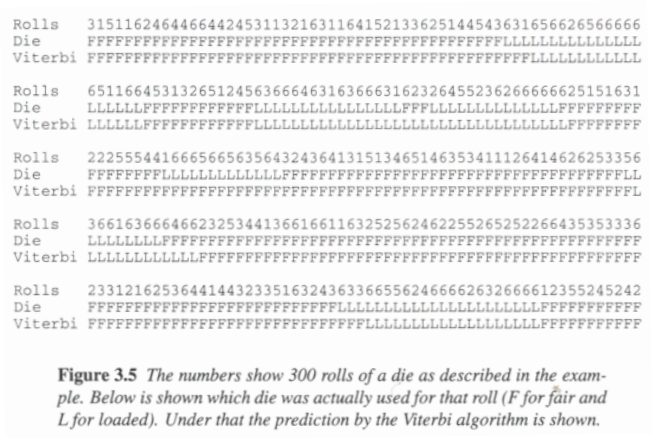


Figure 3.5 *The numbers show 300 rolls of a die as described in the example. Below is shown which die was actually used for that roll (F for fair and L for loaded). Under that the prediction by the Viterbi algorithm is shown.*

Inferring hidden stuff

Joint probability of a given path π & emission sequence *x*:

$$P(x, \pi) = a_{0,\pi_1} \prod_{i=1}^{n} e_{\pi_i}(x_i) \cdot a_{\pi_i, \pi_{i+1}}$$

But π is hidden; what to do? Some alternatives:

Most probable single path

$$\pi^* = \arg\max_{\pi} P(x, \pi)$$

Sequence of most probable states

$$\hat{\pi}_i = \arg\max_{k} P(\pi_i = k \mid x)$$

The Viterbi Algorithm: The most probable path

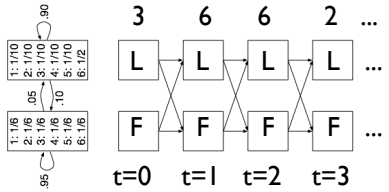
Viterbi finds: $\pi^* = \arg\max_{\pi} P(x, \pi)$

Possibly there are 10^{99} paths of prob 10^{-99}

More commonly, one path dominates others. (If not, other approaches may be preferable.)

Key problem: exponentially many paths π

Unrolling an HMM



Conceptually, sometimes convenient

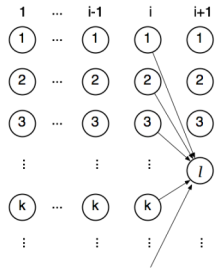
Note exponentially many paths

Viterbi

$v_l(i) =$ probability of the most probable path emitting $x_1, x_2, \ldots, x_i$ and ending in state l

Initialize:

$$v_l(0) = \begin{cases} 1 & \text{if } l = \text{Begin state} \\ 0 & \text{otherwise} \end{cases}$$



General case:

$$v_l(i+1) = e_l(x_{i+1}) \cdot \max_k (v_k(i)\, a_{k,l})$$

Viterbi Traceback

Above finds *probability* of best path

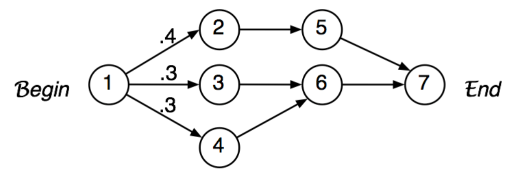
To find the path itself, trace *backward* to the state *k* attaining the max at each stage

```
Rolls   315116246446644245311321631164152133625144543631656626566666
Die     FFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFLLLLLLLLLLLLLLLL
Viterbi FFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFLLLLLLLLLLLLL

Rolls   651166453132651245636664631636663162326455236266666625151631
Die     LLLLLLFFFFFFFFFFFFLLLLLLLLLLLLLLLLFFFLLLLLLLLLLLLLLFFFFFFFFF
Viterbi LLLLLLFFFFFFFFFFFFLLLLLLLLLLLLLLLLLLLLLLLLLLLLLLLLLFFFFFFFFF

Rolls   222555441666566563564324364131513465146353411126414626253356
Die     FFFFFFFFLLLLLLLLLLLLLFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFLL
Viterbi FFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFL

Rolls   366163666466232534413661661163252562462255265252266435353336
Die     LLLLLLLLFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFF
Viterbi LLLLLLLLLLLLFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFFF

Rolls   233121625364414432335163243633665562466662632666612355245242
Die     FFFFFFFFFFFFFFFFFFFFFFFFFFLLLLLLLLLLLLLLLLLLLLLFFFFFFFFFFF
Viterbi FFFFFFFFFFFFFFFFFFFFFFFFFFFFFLLLLLLLLLLLLLLLLLLLLFFFFFFFFFFF
```

Figure 3.5 *The numbers show 300 rolls of a die as described in the example. Below is shown which die was actually used for that roll (F for fair and L for loaded). Under that the prediction by the Viterbi algorithm is shown.*

Is Viterbi "best"?

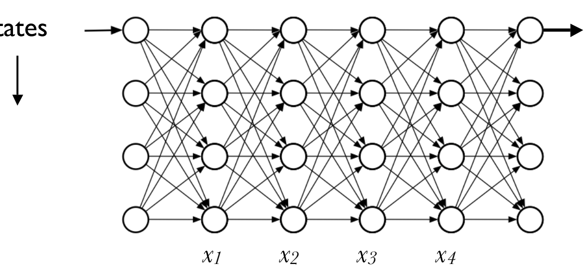
Viterbi finds $\pi^* = \arg\max_{\pi} P(x, \pi)$



Most probable (Viterbi) path goes through 5, but most probable state at 2nd step is 6
(I.e., Viterbi is not the only interesting answer.)

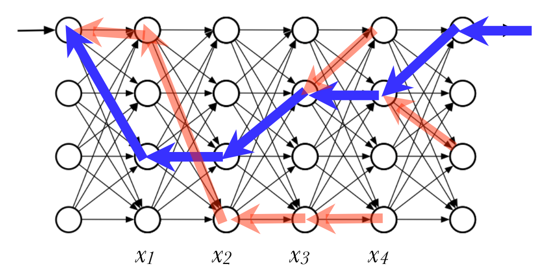
An HMM (unrolled)

States



Emissions/sequence positions

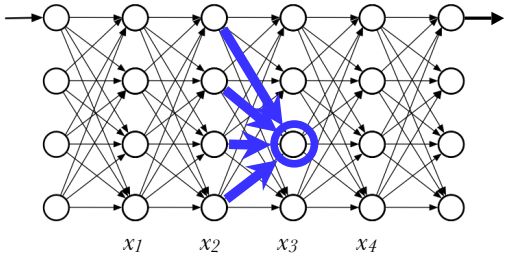
Viterbi: best path to each state



$$v_l(i+1) = e_l(x_{i+1}) \cdot \max_k (v_k(i)\, a_{k,l})$$

The Forward Algorithm

For each state/time, want *total* probability of all paths leading to it, with given emissions



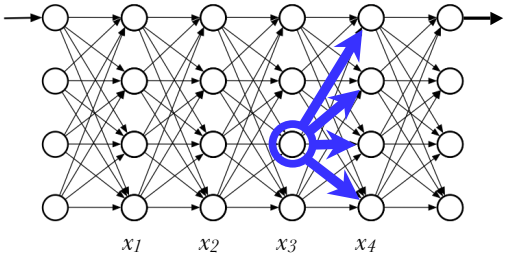
$$f_k(i) = P(x_1 \dots x_i,\ \pi_i = k)$$

$$f_l(i+1) = e_l(x_{i+1}) \sum_k f_k(i) a_{k,l}$$

$$P(x) = \sum_\pi P(x, \pi) = \sum_k f_k(n) a_{k,0}$$

The Backward Algorithm

Similar: for each state/time, want total probability of all paths from it, with given emissions, *conditional* on that state.



$$b_k(i) \triangleq P(x_{i+1} \cdots x_n \mid \pi_i = k)$$

$$b_k(i) = \sum_l a_{k,l}\ e_l(x_{i+1})\ b_l(i+1)$$

$$b_k(n) = a_{k,0}$$

In state *k* at step *i* ?

$$P(x,\ \pi_i = k)$$

$$= P(x_1, \dots, x_i,\ \pi_i = k) \cdot P(x_{i+1}, \dots, x_n \mid x_1, \dots, x_i,\ \pi_i = k)$$

$$= P(x_1, \dots, x_i,\ \pi_i = k) \cdot P(x_{i+1}, \dots, x_n \mid \pi_i = k)$$

$$= f_k(i) \cdot b_k(i)$$

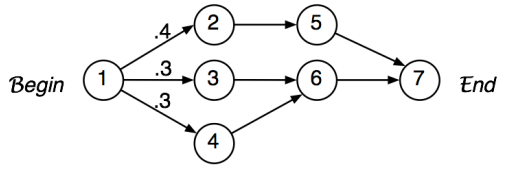
$$P(\pi_i = k \mid x) = \frac{P(x, \pi_i = k)}{P(x)} = \frac{f_k(i) \cdot b_k(i)}{P(x)}$$

Posterior Decoding, I

Alternative 1: what's the most likely state at step *i*?

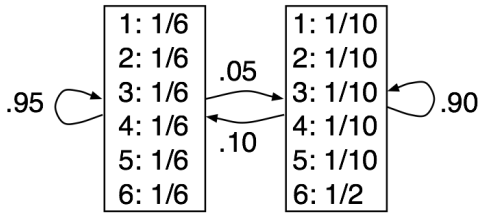
$$\hat{\pi}_i = \arg\max_k P(\pi_i = k \mid x)$$

Note: the sequence of most likely states ≠ the most likely sequence of states. May not even be legal!



The Occasionally Dishonest Casino

1 fair die, 1 "loaded" die, occasionally swapped



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Posterior Decoding

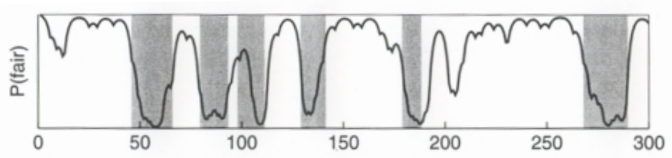


Figure 3.6 *The posterior probability of being in the state corresponding to the fair die in the casino example. The x axis shows the number of the roll. The shaded areas show when the roll was generated by the loaded die.*

Posterior Decoding, II

Alternative 1: what's most likely state at step *i* ?

$$\hat{\pi}_i = \arg\max_k P(\pi_i = k \mid x)$$

Alternative 2: given some function *g(k)* on states, what's its expectation. E.g., what's probability of "+" model in CpG HMM (*g(k)*=1 iff *k* is "+" state)?

$$G(i \mid x) = \sum_k P(\pi_i = k \mid x) \cdot g(k)$$

CpG Islands again

Data: 41 human sequences, totaling 60kbp, including 48 CpG islands of about 1kbp each

Viterbi:	Post-process:
Found 46 of 48	46/48
plus 121 "false positives"	67 false pos

Posterior Decoding:

same 2 false negatives	46/48
plus 236 false positives	83 false pos

(merge within 500; discard < 500)

Training

Given model topology & training sequences, learn transition and emission probabilities

If π known, then MLE is just frequency observed in training data

$$a_{k,l} = \frac{\text{count of } k \to l \text{ transitions}}{\text{count of } k \to \text{anywhere transitions}}$$

$$e_k(b) = \ldots$$

← + pseudocounts?

If π hidden, then use EM:
given π, estimate θ; given θ estimate π. } 2 ways

Viterbi Training

given π, estimate θ; given θ estimate π

Make initial estimates of parameters θ
Find Viterbi path π for each training sequence
Count transitions/emissions on those paths, getting new θ
Repeat

Not rigorously optimizing desired likelihood, but still useful & commonly used.
(Arguably good if you're doing Viterbi decoding.)

Baum-Welch Training

given θ, estimate π ensemble; then re-estimate θ

$$P(\pi_i = k,\, \pi_{i+1} = l \mid x, \theta)$$

$$= \frac{f_k(i \mid \theta)\, a_{k,l}\, e_l(x_{i+1})\, b_l(i+1 \mid \theta)}{P(x \mid \theta)}$$

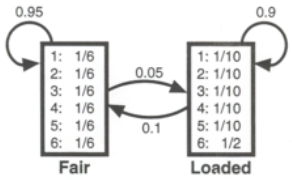
Estimated # of $k \to l$ transitions $\hat{A}_{k,l}$

$$= \sum_{\text{training seqs } x^j} \sum_i P(\pi_i = k,\, \pi_{i+1} = l \mid x^j, \theta)$$

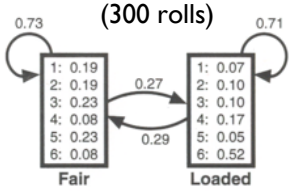
New estimate $\hat{a}_{k,l} = \frac{\hat{A}_{k,l}}{\sum_l \hat{A}_{k,l}}$

Emissions: similar

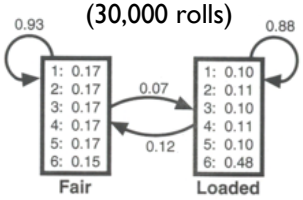
True Model



B-W Learned Model (300 rolls)



B-W Learned Model (30,000 rolls)



Log-odds per roll

True model 0.101 bits

300-roll est. 0.097 bits

30k-roll est. 0.100 Bits

(NB: overfitting)

HMM Summary

Viterbi – best single path (max of products)

Forward – Sum over all paths (sum of products)

Backward – similar

Baum-Welch – Training via EM and forward/backward (aka the forward/backward algorithm)

Viterbi training – also "EM", but Viterbi-based

HMMs in Action: Pfam

Proteins fall into families, both across & within species

Ex: Globins, GPCRs, Zinc Fingers, Leucine zippers,...

Identifying family very useful: suggests function, etc.

So, search & alignment are both important

One very successful approach: profile HMMs

```
Helix                AAAAAAAAAAAAAAA   BBBBBBBBBBBBBBBBCCCCCCCCCCC
HBA_HUMAN  ---------VLSPADKTNVKAAWGKVGA--HAGEYGAEALERMFLSFPTTKTYFPHF
HBB_HUMAN  --------VHLTPEEKSAVTALWGKV----NVDEVGGEALGRLLVVYPWTQRFFESF
MYG_PHYCA  ---------VLSEGEWQLVLHVWAKVEA--DVAGHGQDILIRLFKSHPETLEKFDRF
GLB3_CHITP ----------LSADQISTVQASFDKVKG------DPVGILYAVFKADPSIMAKFTQF
GLB5_PETMA PIVDTGSVAPLSAAEKTKIRSAWAPVYS--TYETSGVDILVKFFTSTPAAQEFFPKF
LGB2_LUPLU --------GALTESQAALVKSSWEEFNA--NIPKHTHRFFILVLEIAPAAKDLFS-F
GLB1_GLYDI ---------GLSAAQRQVIAATWKDIAGADNGAGVGKDCLIKFLSAHPQMAAVFG-F
Consensus             Ls....  v a W kv .  .    g . L.. f . P .   F  F

Helix          DDDDDDDEEEEEEEEEEEEEEEEEE              FFFFFFFFFFF
HBA_HUMAN  -DLS-----HGSAQVKGHGKKVADALTNAVAHV---D--DMPNALSALSDLHAHKL-
HBB_HUMAN  GDLSTPDAVMGNPKVKAHGKKVLGAFSDGLAHL---D--NLKGTFATLSELHCDKL-
MYG_PHYCA  KHLKTEAEMKASEDLKKHGVTVLTALGAILKK----K-GHHEAELKPLAQSHATKH-
GLB3_CHITP AG-KDLESIKGTAPFETHANRIVGFFSKIIGEL--P---NIEADVNTFVASHKPRG-
GLB5_PETMA KGLTTADQLKKSADVRWHAERIINAVNDAVASM--DDTEKMSMKLRDLSGKHAKSF-
LGB2_LUPLU LK-GTSEVPQNNPELQAHAGKVFKLVYEAAIQLQVTGVVVTDATLKNLGSVHVSKG-
GLB1_GLYDI SG----AS---DPGVAALGAKVLAQIGVAVSHL--GDEGKMVAQMKAVGVRHKGYGN
Consensus   .  t     .. . v..Hg kv. a    a...l    d    . a l. l    H  .

Helix        FFGGGGGGGGGGGGGGGGGG      HHHHHHHHHHHHHHHHHHHHHHHHH
HBA_HUMAN  -RVDPVNPKLLSHCLLVTLAAHLPAEFTPAVHASLDKFLASVSTVLTSKYR------
HBB_HUMAN  -HVDPENFRLLGNVLVCVLAHHFGKEFTPPVQAAYQKVVAGVANALAHKYH------
MYG_PHYCA  -KIPIKYLEFISEAIIHVLHSRHPGDFGADAQGAMNKALELFRKDIAAKYKELGYQG
GLB3_CHITP --VTHDQLNNFRAGFVSYMKAHT--DPA-GAEAAWGATLDTFFGMIFSKM-------
GLB5_PETMA -QVDPQYFKVLAAVIADTVAAG---------DAGFEKLMSMICILLRSAY--------
LGB2_LUPLU --VADAHFPVVKEAILKTIKEVVGAKWSEELNSAWTIAYDELAIVIKKEMNDAA-----
GLB1_GLYDI KHIKAQYFEPLGASLLSAMEHRIGGKMNAAAKDAWAAAYADISGALISGLQS-------
Consensus    v.   f  l . ..  ....    f   . aa. k. .    l sky
```

Alignment of 7 globins. A-H mark 8 alpha helices.
Consensus line: upper case = 6/7, lower = 4/7, dot=3/7.
Could we have a profile (aka weight matrix) w/ indels?

Profile Hmm Structure

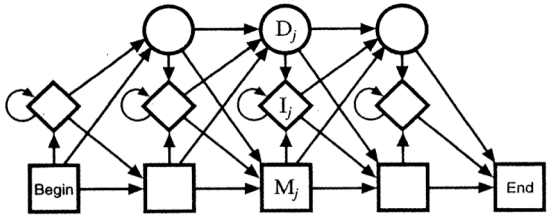


Figure 5.2 *The transition structure of a profile HMM.*

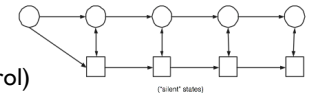
M_j: Match states (20 emission probabilities)
I_j: Insert states (Background emission probabilities)
D_j: Delete states (silent - no emission)

Silent States

Example: chain of states, can skip some

Problem: many parameters.

A solution: chain of "silent" states; fewer parameters (but less detailed control)



Algorithms: basically the same.

Using Profile HMM's

Search
- Forward or Viterbi
- Scoring
  - Log likelihood (length adjusted)
  - Log odds vs background
  - Z scores from either

} next slides

Alignment
- Viterbi

Likelihood vs Odds Scores

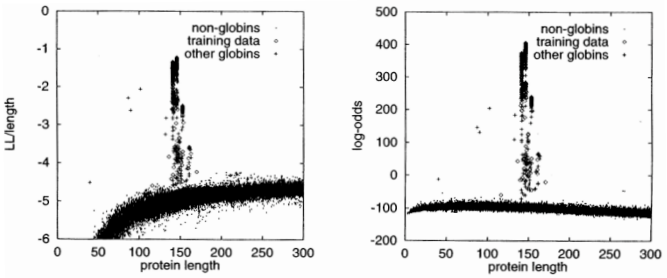


Figure 5.5 *To the left the length-normalized LL score is shown as a function of sequence length. The right plot shows the same for the log-odds score.*

Z-Scores

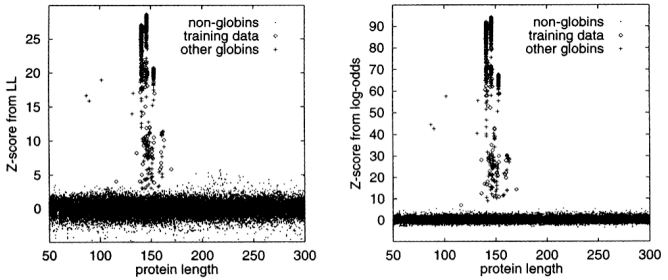


Figure 5.6 *The Z-score calculated from the LL scores (left) and the log-odds (right).*

Pfam Model Building

Hand-curated "seed" multiple alignments

Train profile HMM from seed alignment

Hand-chosen score threshold(s)

Automatic classification/alignment of all other protein sequences

7973 families in Rfam 18.0, 8/2005 (covers ~75% of proteins)

Model-building refinements

Pseudocounts (count = 0 common when training with 20 aa's)

$$e_i(a) = \frac{C_{i,a} + A \cdot q_a}{\sum_a C_{i,a} + A}, \quad A \sim 20, \; q_a = \text{ background}$$

(~50 training sequences)

Pseudocount "mixtures", e.g. separate pseudocount vectors for various contexts (hydrophobic regions, buried regions,...)

(~10-20 training sequences)

More refinements

Weighting: may need to down weight highly similar sequences to reflect phylogenetic or sampling biases, etc.

Match/insert assignment: Simple threshold, e.g. "> 50% gap ⇒ insert", may be suboptimal. Can use forward-algorithm-like dynamic programming to compute max *a posteriori* assignment.

Numerical Issues

Products of many probabilities → 0

For Viterbi: just add logs

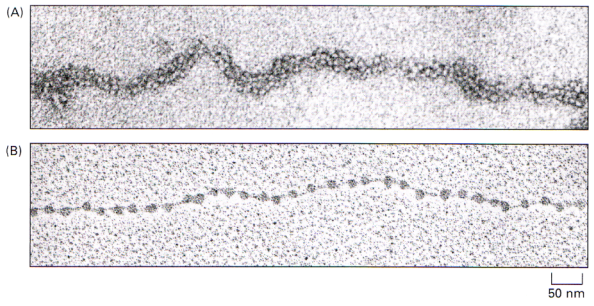
For forward/backward: also work with logs, but you need sums of products, so need "log-of-sum-of-product-of-exp-of-logs", e.g., by table/interpolation

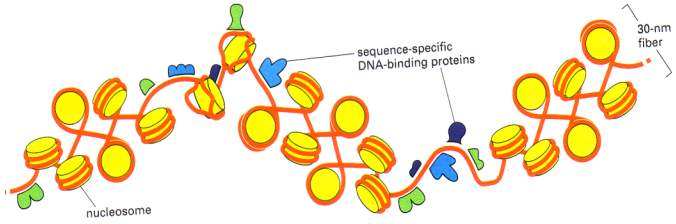
Keep high precision and perhaps scale factor

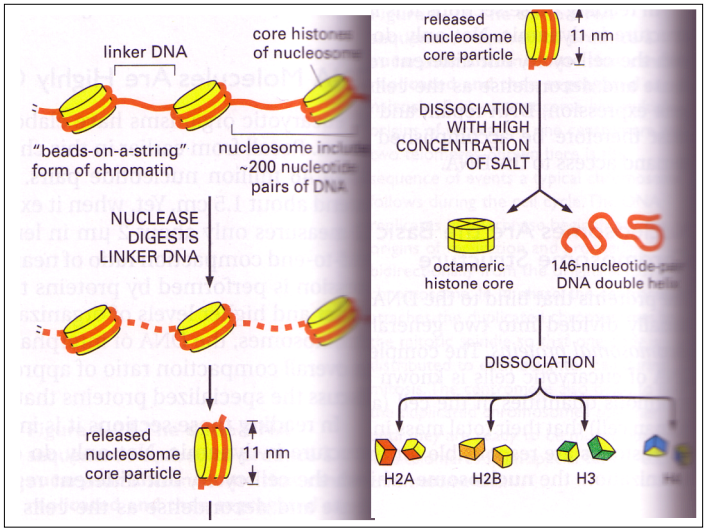
Working with log-odds also helps.

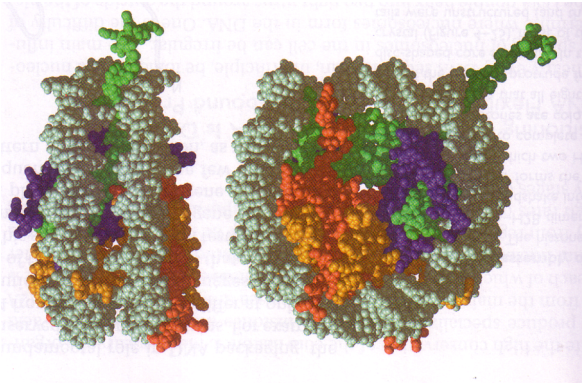
The Bio Interlude: Chromatin Codes & some DNA binding experiments

Chromatin

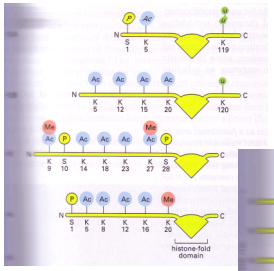




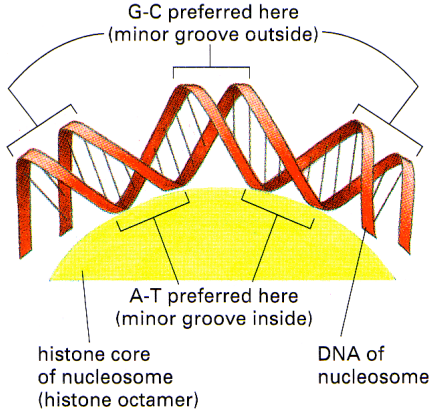




Histone Codes

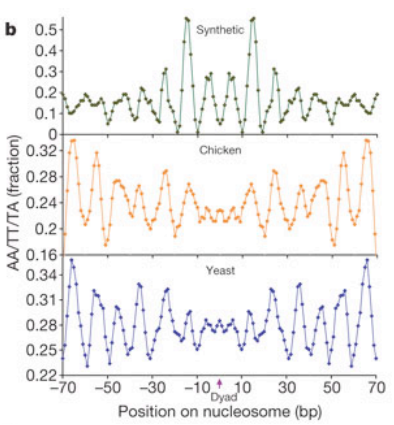


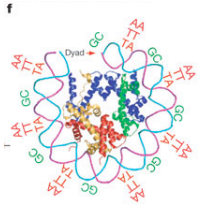
N-terminal tail	modification state	"meaning"
9 10 14 18 23 28	unmodified	gene silencing?
Ac	acetylated	gene expression
Ac	acetylated	histone deposition
Me	methylated	gene silencing/ heterochromatin
P P	phosphorylated	mitosis/meiosis
P Ac	phosphorylated/ acetylated	gene expression
P Ac Ac Me	higher-order combinations	?
Ac Ac	unmodified	gene silencing?
5 Ac 12 Ac 8 16	acetylated	histone deposition
	acetylated	gene expression



A genomic code for nucleosome positioning

Eran Segal, Yvonne Fondufe-Mittendorf, Lingyi Chen, AnnChristine Thastrom, Yair Field, Irene K. Moore, Ji-Ping Z. Wang and Jonathan Widom
doi:10.1038/nature04979 (7/19/06)

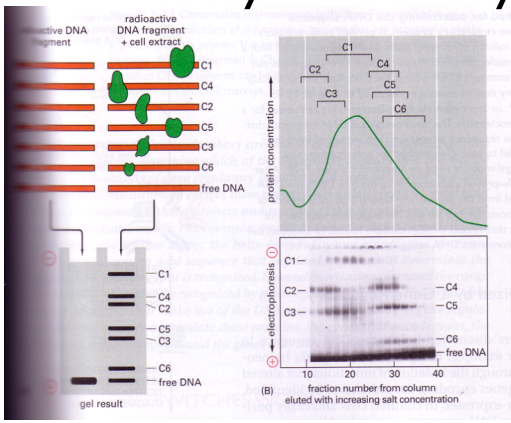




Method: ~ "1st order WMM" (as above) trained on 200 aligned nucleosome binding seqs; alt: MEME-like EM algorithm

Experimental approaches to learning DNA binding proteins & their targets

Gel Mobility Shift Assay



Chromatin Immuno-Precipitation