

Lecture 1: Random Walks

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**Disclaimer:** *These notes have not been subjected to the usual scrutiny reserved for formal publications.*

In this note we study random walks and mixing time in weighted undirected non-regular graphs  $G$ . The notation that we define here will be used later to study high dimensional random walks on simplicial complexes.

Given an undirected (weighted) graph  $G = (V, E)$  with weight function  $w : E \rightarrow \mathbb{R}_{\geq 0}$ , let  $A \in \mathbb{R}_{\geq 0}^{V \times V}$  be the adjacency matrix of  $G$ .

Consider a natural probability distribution on edges of  $G$  where for any edge  $\{u, v\}$

$$\pi_1(\{u, v\}) = \frac{w(\{u, v\})}{\sum_e w(e)}.$$

This distribution naturally induces a distribution on vertices; given a random edge  $e$ , we randomly drop one of the endpoints of  $e$ :

$$\pi_0(v) = \sum_e \mathbb{P}_{\pi_1}[e] \mathbb{P}[v|e] = \sum_{e:v \in e} \frac{1}{2} \pi_1(e) = \frac{\sum_{u \sim v} \frac{1}{2} w(\{u, v\})}{\sum_e w(e)} = \frac{d_w(v)}{\sum_u d_w(u)}.$$

Conversely, suppose we want to sample an edge  $e = \{u, v\} \sim \pi_1$ . One way is to first sample a vertex  $v \sim \pi_0$  and then sample an edge  $e|v$ , i.e., among all edge incident to  $v$  choose one proportional to its weight.

$$\mathbb{P}[\{u, v\}|u] := \frac{\pi_1\{u, v\}}{\sum_{x \sim u} \pi_1\{u, x\}}$$

We claim this process chooses  $e$  with probability  $\pi_1(e)$ , because,

$$\begin{aligned} \mathbb{P}[e] &= \mathbb{P}_{\pi_0}[v] \mathbb{P}[e|v] + \mathbb{P}_{\pi_0}[u] \mathbb{P}[e|u] \\ &= \pi_0(v) \frac{\pi_1(e)}{\sum_{f \sim v} \pi_1(f)} + \pi_0(u) \frac{\pi_1(e)}{\sum_{f \sim u} \pi_1(f)} \\ &= \pi_0(v) \frac{w(e)}{\sum_{f \sim v} w(f)} + \pi_0(u) \frac{w(e)}{\sum_{f \sim u} w(f)} \\ &= \pi_0(v) \frac{w(e)}{d_w(v)} + \pi_0(u) \frac{w(e)}{d_w(u)} = \frac{\pi(e)}{2} + \frac{\pi(e)}{2} = \pi(e). \end{aligned}$$

**Fact 1.1.** *Summarizing the above observations,*

- *To choose  $v \sim \pi_0$ , we can first choose  $e \sim \pi_1$  and then drop one endpoint.*
- *To choose  $e \sim \pi_1$ , we can first choose  $v \sim \pi_0$  and then choose  $e|v$ .*

## 1.1 Random Walk Operator

We can look at the simple random walk operator on  $G$  that says if I am at vertex  $u$ , I choose an edge incident to  $u$  with probability proportional to its weight. Let

$$\mathbb{P}[u \rightarrow v] = \mathbb{P}[\{u, v\}|u].$$

For a function  $f : V \rightarrow \mathbb{R}$ , we have

$$Pf(v) := \mathbb{E}_{\{u,v\}|v} [f(u)] = \sum_u \mathbb{P}[v \rightarrow u] f(u).$$

The following reversibility property will also be used crucially.

$$P(u, v)\pi_0(u) = \pi_0(u)\mathbb{P}[\{u, v\}|u] = \pi_1(\{u, v\})\mathbb{P}[u|\{u, v\}] = \pi_1(\{u, v\})\mathbb{P}[v|\{u, v\}] = P(v, u)\pi_0(v) \quad (1.1)$$

We equip the linear space  $\mathbb{R}^V$  with the following inner product: For two vectors  $f, g : V \rightarrow \mathbb{R}$ ,

$$\langle f, g \rangle = \mathbb{E}_{v \sim \pi_0} f(v)g(v) = \sum_v \pi_0(v) f(v)g(v).$$

This naturally defines a norm, where for any such function  $f$ ,  $\|f\| = \sqrt{\langle f, f \rangle}$ .

Using the above fact, it follows that  $P$  is self-adjoint with respect to the above inner product, i.e., for any two  $f, g : V \rightarrow \mathbb{R}$ ,

$$\begin{aligned} \langle f, Pg \rangle &= \mathbb{E}_{v \sim \pi_0} [f(v)Pg(v)] \\ &= \mathbb{E}_{v \sim \pi_0} [f(v)\mathbb{E}_{\{u,v\}|v} [g(u)]] \\ &= \mathbb{E}_{\{u,v\} \sim \pi_1} \mathbb{E}_{v|\{u,v\}} f(v)g(u) = \mathbb{E}_{\{u,v\} \sim \pi_1} \mathbb{E}_{v|\{u,v\}} f(u)g(v) \end{aligned}$$

Similarly, we can show that  $\langle Pf, g \rangle$  is also equal to the RHS.

**Fact 1.2.**  $\lambda_1 = 1$  and  $\lambda_n \geq -1$ .

*Proof.* First, observe that the all-ones function  $\mathbf{1}$  is an eigenfunction,

$$P\mathbf{1} = \mathbf{1}.$$

Second, we show  $\lambda_i \leq 1$  for all  $i$ . For any eigenfunction  $f : V \rightarrow \mathbb{R}$ , with eigenvalue  $\lambda$ , i.e.,  $Pf = \lambda f$ , we claim that  $\lambda \leq 1$ . Say  $u = \operatorname{argmax}_v |f(v)|$ . Then,

$$\lambda f(u) = Pf(u) = \mathbb{E}_{\{u,v\}|u} f(v) \leq \mathbb{E}_{\{u,v\}|u} |f(v)| \leq \mathbb{E}_{\{u,v\}|u} |f(u)| = |f(u)|.$$

In the second inequality we used  $|f(v)| \leq |f(u)|$  for all  $v \in V$ . So, we have  $\lambda \leq 1$  as desired. Also, observe from the same inequality that  $|\lambda| \leq 1$  as desired.  $\square$

**Fact 1.3.** *The matrix  $P$  has  $n$ -real eigenvalues with  $n$ -orthonormal eigenfunctions.*

*Proof.* The proof is fairly general and holds for any self-adjoint matrix with respect to an inner-product. First, Suppose  $Pf = \lambda f$  for  $f \in \mathbb{C}^n$  and  $\lambda \in \mathbb{C}$ . And, recall for any function  $f$ ,  $\langle f, f \rangle = \mathbb{E}_u f(u)\overline{f(u)} \in \mathbb{R}$ . Then,

$$\|Pf\|^2 = \langle Pf, Pf \rangle = \langle P^2 f, f \rangle = \lambda^2 \|f\|^2.$$

So,  $\lambda^2 \in \mathbb{R}$ .

On the other hand, suppose  $P$  has two eigenfunctions  $f, g \in \mathbb{R}^n$  with eigenvalues  $\lambda_f \leq \lambda_g$ ; then,

$$\langle \lambda_f f, g \rangle = \langle Pf, g \rangle = \langle f, Pg \rangle = \langle f, \lambda_g g \rangle.$$

Therefore, since  $\lambda_f \neq \lambda_g$  we must have  $\langle f, g \rangle = 0$ .  $\square$

It follows by the spectral theorem that  $P$  has  $n = |V|$  eigenvalues  $\lambda_1 \geq \dots \geq \lambda_n$  with corresponding orthonormal eigenfunctions  $f_1, \dots, f_n$  such that for any  $1 \leq i < j \leq n$ ,

$$\langle f_i, f_j \rangle = 0,$$

and for any  $i$ ,  $\|f_i\| = 1$ .

## 1.2 Mixing

**Theorem 1.4.** *For a weighted graph with eigenvalues  $\lambda_1 \geq \dots \geq \lambda_n$ , if  $\max\{|\lambda_2|, |\lambda_n|\} < 1$ , then for any vertex  $u$ ,*

$$\lim_{t \rightarrow \infty} P^t f(u) = \mathbb{E}f.$$

Letting  $f(v) = \mathbf{1}_v$ , we have  $\mathbb{E}f = \pi_0(v)$ . So, the above theorem shows that as  $t \rightarrow \infty$ , the distribution of the walk converges to  $\pi_0$ . In other words,  $\pi_0$  is the *stationary distribution* of the simple random walk on  $G$ .

The mixing time of a random walk is defined as

$$\tau_{\max} = \min \left\{ t : \|P^t(u, \cdot) - \pi_0\|_{\text{TV}} = \frac{1}{2} \sum_v |P^t \mathbf{1}_v(u) - \pi_0(v)| \leq \frac{1}{4}, \forall u \in V \right\} \quad (1.2)$$

where  $\tau_{\text{mix}}$  is the first time that the total variation distance of the distribution of the walk started at  $u$ , from the stationary distribution is at most a constant.

For a function  $f : V \rightarrow \mathbb{R}$  define

$$\|f\|_p = (\mathbb{E}_\pi f^p)^{1/p}.$$

More generally, we define the  $L_p$  mixing time of the walk as

$$\tau_{\text{mix}, p} := \min \left\{ t : \left\| \frac{P^t(u, \cdot)}{\pi} - \mathbf{1} \right\|_p \leq \frac{1}{4}, \forall u \in V \right\}$$

Note that by Cauchy-Schwarz inequality

$$\sum_v |P^t(u, v) - \pi(v)| = \mathbb{E}_\pi |P^t(u, \cdot)/\pi - \mathbf{1}| \leq \sqrt{\mathbb{E}_\pi |P^t(u, \cdot)/\pi - \mathbf{1}|^2} = \left\| \frac{P^t(u, \cdot)}{\pi} - \mathbf{1} \right\|_2$$

So, the  $L_2$ -mixing of a chain implies  $L_1$  mixing (but not vice versa). And similarly, for any  $p > 1$ ,  $L_p$  mixing implies  $L_1$  mixing.

**Theorem 1.5.** *Let  $G$  be a weighted graph with eigenvalues  $\lambda_1 \geq \dots \geq \lambda_n$ , such that  $\lambda^* := \max\{|\lambda_2|, |\lambda_n|\} < 1$ . For any  $\epsilon > 0$ , any function  $f \in V \rightarrow \mathbb{R}$  and  $t \geq \frac{\log \epsilon}{1 - \lambda^*}$*

$$\|P^t f - \mathbb{E}f\| \leq \epsilon \|f\|.$$

*Proof.* Let  $f_1, \dots, f_n$  be the orthonormal eigenfunctions of  $P$  with corresponding eigenvalues  $\lambda_1 \geq \dots \geq \lambda_n$ . Recall  $\lambda_1 = 1$  and  $f_1 = \mathbf{1}$ . Also note that  $P^t$  have exactly the same eigenfunctions with eigenvalues  $\lambda_1^t, \dots, \lambda_n^t$ . Therefore, we can write  $f = \sum_i \langle f, f_i \rangle f_i$ . First, by (??),

$$\begin{aligned} P^t f - \mathbb{E}f &= P^t \left( \sum_{i=1}^n \langle f, f_i \rangle f_i \right) - \langle f, \mathbf{1} \rangle \mathbf{1} \\ &= \sum_{i=1}^n \lambda_i^t \langle f, f_i \rangle f_i - \langle f, \mathbf{1} \rangle \mathbf{1} = \sum_{i=2}^n \lambda_i^t \langle f, f_i \rangle f_i \end{aligned}$$

Therefore, using the orthonormality of  $f_i$ 's,

$$\begin{aligned} \|P^t f - \mathbb{E}f\|^2 &= \left\langle \sum_i \lambda_i^t \langle f, f_i \rangle, \sum_j \lambda_j^t \langle f, f_j \rangle \right\rangle \\ &= \sum_{i=2}^n \lambda_i^{2t} \langle f, f_i \rangle^2 \\ &\leq \lambda^{*2t} \|f\|^2 \end{aligned}$$

where in the last inequality we used  $|\lambda_i| \leq \lambda^*$  for all  $i > 1$ . By the above inequality, for  $t \geq \frac{\log \epsilon}{1 - \lambda^*}$  we have

$$\|P^t f(u) - \mathbb{E}f\| \leq \lambda^{*t} \|f\| = (1 - (1 - \lambda^*))^t \|f\| \leq e^{-(1 - \lambda^*)t} \|f\| = \epsilon \|f\|$$

where in the second inequality we used  $1 - x \leq e^{-x}$ .  $\square$

**Corollary 1.6.** *Suppose for any function  $f : V \rightarrow \mathbb{R}$ ,  $\epsilon > 0$  and  $t \geq \frac{\log \epsilon^{-1}}{1 - \lambda^*}$   $\|P^t f - \mathbb{E}f\| \leq \epsilon \|f\|$  then,*

$$\tau_{mix,1} \leq \tau_{mix,2} \leq \max_u \frac{\log \epsilon^{-1} \pi(u)^{-1}}{1 - \lambda^*}$$

Note that in particular, if we want to bound the mixing time starting at a given vertex  $u$ , we don't need the max on the right hand side.

*Proof.* Fix a vertex  $u \in V$  and let  $f := \mathbf{1}_u / \pi_0(u)$ . Then,  $\mathbb{E}f = 1$ .

$$\|P^t f - \mathbb{E}f\|^2 = \|P^t f - (\mathbb{E}f)\mathbf{1}\|^2 = \|P^t f\|^2 - 1.$$

On the other hand,

$$\begin{aligned} \|P^t f\|^2 &= \langle P^t f, P^t f \rangle = \mathbb{E}_v P^t f(v) \cdot P^t f(v) \\ &= \sum_{f(w) \neq 0, \forall w \neq u} \mathbb{E}_{v \sim \pi_0} \frac{P^t(v, u)}{\pi(u)} \cdot \frac{P^t(v, u)}{\pi(u)} \\ &\stackrel{(1.1)}{=} \mathbb{E}_{v \sim \pi_0} \left( \frac{P^t(u, v)}{\pi(v)} \right)^2 \end{aligned}$$

Putting these together,  $\|P^t f - \mathbb{E}f\|^2 = \left\| \frac{P^t(u, \cdot)}{\pi(\cdot)} - 1 \right\|^2$ . Finally, since  $\|f\|^2 = \frac{1}{\pi(u)}$ , letting  $\epsilon$  of [Theorem 1.5](#)  $\epsilon \pi(u)$  proves the claim.  $\square$

**Theorem 1.5** allows us to estimate the probability of any event. Say  $A \subseteq V$  and we want to estimate  $\pi_0(A)$ . For  $f = \mathbf{1}_A$  we have  $\mathbb{E}f = \pi_0(A)$  and  $\langle f, f \rangle \leq \pi_0(A) < 1$ . So, we can use  $P^t f(u)$ , for a vertex  $u \in V$ , as an estimate of  $\mathbb{E}f$ . By Cauchy-Schwarz inequality,

$$\mathbb{E}_v |P^t f(v) - \mathbb{E}f| \leq \sqrt{\mathbb{E}_u |P^t f(u) - \mathbb{E}f|^2} = \|P^t f - \mathbb{E}f\|.$$

So, for  $t = \frac{\log \epsilon^{-1} \log \pi(u)^{-1}}{1 - \lambda^*}$

$$|P^t f(u) - F| \leq \frac{1}{\pi_0(u)} \mathbb{E}_v |P^t(v) - F| \stackrel{\text{Theorem 1.5}}{\leq} \epsilon \|f\|$$

Having this, we can simply estimate  $\pi_0(A)$  within an additive error of  $\epsilon$  (with probability  $1 - \delta$ ) by averaging  $k$  independent  $f(X_t)$ , where  $X_t$  is the  $t$ -th state that the simple random walk started at  $u$  lands on. Here,  $k = O(\frac{1}{\epsilon^2} \log \delta)$  and  $t$  is as defined above.