CSE 574

Finite State Machines for Information Extraction

Today and Friday

- Dan @ IUI on the Canary Islands
- I am presenting

• Topics

- HMMs, Conditional Random Fields
- Inference and Learning

Landscape of IE Techniques: <u>Models</u>



Each model can capture words, formatting, or both

Slides from Cohen & McCallum

1/18/2008 3:12 PM

Finite State Models



Graphical Models



• Factor Graphs



 $p(\mathbf{x}) = \frac{1}{Z} \prod_{A} \Psi_{A}(\mathbf{x}_{A})$ $A \subset \{x_{1}, \dots, x_{K}\}$ Ψ_{A} factor function

Recap: Naïve Bayes

- Assumption: features independent given label
- Generative Classifier
 - Model joint distribution p(x,y)

$$p(y, \mathbf{x}) = p(y) \prod_{k=1}^{K} p(x_k | y)$$



- Inference

$$p(y|\mathbf{x}) = p(y) \prod_{k=1}^{K} p(x_k|y) \frac{1}{p(\mathbf{x})}$$

- Learning: counting
- Example



Labels of neighboring words dependent!

> Need to consider sequence!

Hidden Markov Models



Generative Sequence Model

- 2 assumptions to make joint distribution tractable
 - 1. Each state depends only on its immediate predecessor.
 - 2. Each observation depends only on current state.



Hidden Markov Models



• Generative Sequence Model $p(\mathbf{y}, \mathbf{x}) = \prod_{t=1}^{T} p(y_t | y_{t-1}) p(x_t | y_t)$



Graphical Model

Model Parameters

- Start state probabilities $p(y_1) := p(y_1|y_0)$
- Transition probabilities $p(y_t|y_{t-1})$
- Observation probabilities $p(x_t|y_t)$

IE with Hidden Markov Models

Given a sequence of observations:

Yesterday Pedro Domingos spoke this example sentence.







Find the most likely state sequence: (Viterbi) $\arg \max_{\bar{s}} P(\bar{s}, \bar{o})$



Any words said to be generated by the designated "person name" state extract as a person name:

Person name: Pedro Domingos

Slide by Cohen & McCallum

IE with Hidden Markov Models

- For sparse extraction tasks :
- Separate HMM for each type of target
- Each HMM should
 - Model entire document
 - Consist of *target* and *non-target* states
 - Not necessarily fully connected



Information Extraction with HMMs

• Example - Research Paper Headers



HMM Example: "Nymble"

Task: Named Entity Extraction

start-ofsentence (Five other name classes) Other

Train on ~500k words of news wire text.

Results:

Case	Language	F1.
Mixed	English	93%
Upper	English	91%
Mixed	Spanish	90%

<u>Transition</u> probabilities	<u>Observation</u> probabilities
$p(y_t y_{t-1}, x_{t-1})$	$p(x_t y_t, y_{t-1})$
	or $p(x_t y_t,x_{t-1})$
Back-off to:	Back-off to:
$p(y_t y_{t-1})$	$p(x_t y_t)$
$p(y_t)$	$p(x_t)$

[Bikel, et al 1998],

[BBN "IdentiFinder"]

Other examples of shrinkage for HMMs in IE: [Freitag and McCallum '99]

Slide by Cohen & McCallum

A parse of a sequence

Given a sequence $x = x_1 \dots x_N$, A <u>parse</u> of o is a sequence of states $y = y_1, \dots, y_N$



Slide by Serafim Batzoglou

Question #1 - Evaluation

GIVEN

A sequence of observations $x_1 x_2 x_3 x_4 \dots x_N$ A trained HMM

 $\Theta = (p(y_t|y_{t-1}), p(x_t|y_t), p(y_1))$

QUESTION

How likely is this sequence, given our HMM ? $P(x, \theta)$

Why do we care?

Need it for learning to choose among competing models!

Question #2 - Decoding

GIVEN

A sequence of observations $x_1 x_2 x_3 x_4 \dots x_N$ A trained HMM

 $\Theta = (p(y_t|y_{t-1}), p(x_t|y_t), p(y_1))$

QUESTION

How dow we choose the corresponding parse (state sequence) $y_1 y_2 y_3 y_4 \dots y_N$, which "best" explains $x_1 x_2 x_3 x_4 \dots x_N$?

There are several reasonable optimality criteria: single optimal sequence, average statistics for individual states, ...

Question #3 - Learning

GIVEN

A sequence of observations $x_1 x_2 x_3 x_4 \dots x_N$ QUESTION

How do we learn the model parameters $\theta = (p(y_t|y_{t-1}) \quad p(x_t|y_t), \quad p(y_1))$ to maximize $P(x, \lambda)$?

Solution to #1: Evaluation

Given observations $x = x_1 \dots x_N$ and HMM θ , what is p(x)?

Naïve: enumerate every possible state sequence $y=y_1$ y_N Probability of x and given particular y $p(\mathbf{x}|\mathbf{y}) = \prod p(x_t|y_t)$ **2T** multiplications Probability of particular y per sequence $p(\mathbf{y}) = \prod p(y_t | y_{t-1})$ Summing over all possible state sequences we get For small HMMs $p(\mathbf{x}) = \sum p(\mathbf{x}|\mathbf{y})p(\mathbf{y})$ T=10, N=10 all \mathbf{y} there are 10 N^T state sequences! billion sequences!

Solution to #1: Evaluation

Use Dynamic Programming:

Define forward variable

$$\alpha_t(i) = P(\mathbf{x}_1 \mathbf{x}_2 \dots \mathbf{x}_t, y_t = S_i)$$

probability that at time t

- the state is y_i
- the partial observation sequence x=x1 ...x+ has been omitted

Solution to #1: Evaluation

• Use Dynamic Programming

$$p(\mathbf{x}) = \sum_{\mathbf{y}} p(\mathbf{y}) p(\mathbf{x} | \mathbf{y})$$

=
$$\sum_{\mathbf{y}} \prod_{t=1..T} p(y_t | y_{t-1}) p(x_t | y_t)$$

=
$$\sum_{y_T} \sum_{y_{T-1}} p(y_T, y_{T-1}, x_T) \sum_{y_{T-2}} p(y_{T-1} | y_{T-2}) p(x_{T-1} | y_{T-1}) \sum_{y_{T-3}} \cdots$$

- Cache and reuse inner sums
- Define forward variables

$$\alpha_t(i) := P(\mathbf{x}_1 \mathbf{x}_2 \dots \mathbf{x}_t, y_t = S_i)$$

probability that at time t

- the state is $y_t = S_i$
- the partial observation sequence $x=x_1 x_+$ has been omitted

The Forward Algorithm

$$\alpha_t(i) := P(x_1 x_2 \dots x_t, y_t = S_i)$$

INITIALIZATION

$$\alpha_1(i) = p(y_1 = S_i)p(x_1|y_1)$$

INDUCTION

$$\alpha_t(i) = p(x_1 x_2 \dots x_t, y_t = S_i) = \sum_{j \in S} \alpha_{t-1}(j) p(y_t = S_i | y_{t-1} = S_j) p(x_t | y_t)$$

TERMINATION

$$p(\mathbf{x}) = \sum_{j \in S} \alpha_T(j)$$

K = |S| #states N length of sequence

The Forward Algorithm $\alpha_t(i) := P(\mathbf{x}_1 \mathbf{x}_2 \dots \mathbf{x}_t, y_t = S_i)$ $\alpha_{t-1}(1)$ S_1 S_1 P gr 11 S3 gr $\alpha_{t-1}(2)$ S_2 S_2 $I(y_t = \frac{S_3}{y_{t-1}} = \frac{S_2}{S_2})$ IN ST $\alpha_{t-1}(3)$ $\alpha_t(3)$ S_3 S_3 $p(y_t = S_3 | y_{t-1} = S_3)$ $p(y_t = S_3 | y_{t-1} = S_3)$ $p(y_t = S_3 | y_{t-1} = S_3)$ $\alpha_{t-1}(N)$ $p(x_t = o_t | y_t = S_3)$ S_N S_N t - 1t o_t

The Backward Algorithm



The Backward Algorithm

$$\beta_t(i) := P(y_t = S_i, x_{t+1}x_{t+2}...x_T)$$

INITIALIZATION

$$\beta_T(i) = 1$$

INDUCTION

$$\begin{aligned}
\beta_t(i) &= p(y_t = S_i, x_{t+1}, x_{t+2} \dots x_T) \\
&= \sum_{j \in S} p(y_{t+1} = S_j | y_t = S_i) p(x_{t+1} | y_{t+1}) \beta_{t+1}(j)
\end{aligned}$$

TERMINATION

$$p(\mathbf{x}) = \sum_{j \in S} p(y_1 = S_j) p(x_1 | y_1) \beta_1(j)$$



Solution to #2 - Decoding

Given $x=x_1 x_N$ and HMM θ , what is "best" parse $y_1 y_N$?

Several optimal solutions

• 1. States which are individually most likely:

$$P(y_t = S_i | \mathbf{x}) = \frac{\alpha_t(i)\beta_t(i)}{P(\mathbf{x})} = \frac{\alpha_t(i)\beta_t(i)}{\sum_{i=1}^N \alpha_t(i)\beta_t(i)}$$

most likely state y_{+}^{*} at time t is then

$$y_t^* = \arg \max_{1 \le i \le N} P(y_t = S_i | \mathbf{x})$$

But some transitions may have 0 probability!

Solution to #2 - Decoding

Given $x=x_1 x_N$ and HMM θ , what is "best" parse $y_1 y_N$?

- Several optimal solutions
- 1. States which are individually most likely
- 2. Single best state sequence

$$y^* = argmax_y P(x, y)$$

Again, we can use dynamic programming!

The Viterbi Algorithm

DEFINE

$$\delta_t(i) = \max_{y_1, y_2, \dots, y_{t-1}} P(y_1, y_2, \dots, y_{t-1}, y_t = i, o_1, o_2, \dots, o_t | \lambda)$$

INITIALIZATION

$$\delta_1(i) = p(y_1 = S_i)p(x_1|y_1 = S_i)$$

INDUCTION

$$\delta_t(j) = \max_{i \in S} \delta_{t-1}(i) p(y_t = S_j | y_{t-1} = S_i) p(x_t | y_t = S_j)$$

TERMINATION



The Viterbi Algorithm



Remember:

 $\delta_{k}(i)$ = probability of most likely state seq ending with state S.

Slides from Serafim Batzoglou

The Viterbi Algorithm



Solution to #3 - Learning

- Given $x_1 \dots x_N$, how do we learn $\theta = (p(y_t|y_{t-1}), p(x_t|y_t), p(y_1))$ to maximize P(x)?
- Unfortunately, there is no known way to analytically find a global maximum θ * such that θ * = arg max P(o | θ)

- But it is possible to find a local maximum; given an initial model $\theta,$ we can always find a model θ' such that

 $P(o | \theta') \ge P(o | \theta)$

Solution to #3 - Learning

• Use hill-climbing

- Called the forward-backward (or Baum/Welch) algorithm
- Idea
 - Use an initial parameter instantiation
 - Loop
 - Compute the forward and backward probabilities for given model parameters and our observations
 - Re-estimate the parameters
 - Until estimates don't change much

- The forward-backward algorithm is an instance of the more general EM algorithm
 - The E Step:
 Compute the forward and backward probabilities for given model parameters and our observations
 - The M Step: Re-estimate the model parameters

Chicken & Egg Problem

- If we knew the actual sequence of states
 - It would be easy to learn transition and emission probabilities
 - But we can't observe states, so we don't!

• If we knew transition & emission probabilities

- Then it'd be easy to estimate the sequence of states (Viterbi)
- But we don't know them!



Input Looks Like



We Want to Predict



Chicken & Egg

Note that coloring instances would be easy *if* we knew Gausians....



Chicken & Egg

And finding the Gausians would be easy If we knew the coloring



- Pretend we do know the parameters
 - Initialize randomly: set $\theta_1 = ?$; $\theta_2 = ?$



- Pretend we do know the parameters
 Initialize randomly
- [E step] Compute probability of instance having each possible value of the hidden variable



- Pretend we do know the parameters
 Initialize randomly
- [E step] Compute probability of instance having each possible value of the hidden variable



- Pretend we do know the parameters
 Initialize randomly
- [E step] Compute probability of instance having each possible value of the hidden variable
 [M step] Treating each instance as *fractionally* having both values compute the new parameter values



ML Mean of Single Gaussian

$$U_{ml} = argmin_u \sum_i (x_i - u)^2$$



.01 .03 .05 .07 .09

[M step] Treating each instance as fractionally having both values compute the new parameter values



 [E step] Compute probability of instance having each possible value of the hidden variable



Slide by Daniel S. Weld

- [E step] Compute probability of instance having each possible value of the hidden variable
 - [M step] Treating each instance as fractionally having both values compute the new parameter values



- [E step] Compute probability of instance having each possible value of the hidden variable
 - [M step] Treating each instance as fractionally having both values compute the new parameter values



The Problem with HMMs

- We want more than an Atomic View of Words
- We want many arbitrary, overlapping features of words

y_t у_{t+1} У_{t-1} identity of word ends in "-ski" is capitalized is part of a noun phrase is "Wisniewski is in a list of city names is under node X in WordNet part o is in bold font noun phrase is indented Х X Х t+1 t -1 is in hyperlink anchor last person name was female next two words are "and Associates"

Finite State Models



Problems with Richer Representation and a Joint Model

These arbitrary features are not independent.

- Multiple levels of granularity (chars, words, phrases)
- Multiple dependent modalities (words, formatting, layout)
- Past & future

Two choices:

Model the dependencies.

Each state would have its own Bayes Net. *But we are already starved for training data!*



Ignore the dependencies.

This causes "over-counting" of evidence (ala naïve Bayes). Big problem when combining evidence, as in Viterbi!



Discriminative and Generative Models

- So far: all models generative
- Generative Models ...

model P(x,y)

 Discriminative Models ... model P(x|y)

P(x|y) does not include a model of P(x), so it does not need to model the dependencies between features!

Discriminative Models often better

- Eventually, what we care about is p(y|x)!
 - Bayes Net describes a family of joint distributions of, whose conditionals take certain form
 - But there are many other joint models, whose conditionals also have that form.
- We want to make independence assumptions among y, but not among x.

Conditional Sequence Models

 We prefer a model that is trained to maximize a conditional probability rather than joint probability:

P(y|x) instead of P(y,x):

- Can examine features, but not responsible for generating them.
- Don't have to explicitly model their dependencies.
- Don't "waste modeling effort" trying to generate what we are given at test time anyway.

Finite State Models



Linear-Chain Conditional Random Fields • From HMMs to CRFs

$$p(\mathbf{y}, \mathbf{x}) = \prod_{t=1}^{T} p(y_t | y_{t-1}) p(x_t | y_t)$$

can also be written as

$$p(\mathbf{y}, \mathbf{x}) = \frac{1}{Z} exp\left(\sum_{t} \sum_{i,j \in S} \lambda_{ij} \mathbf{1}_{\{y_t=i\}} \mathbf{1}_{\{y_t=i\}} + \sum_{t} \sum_{i \in S} \sum_{o \in O} \mu_{oi} \mathbf{1}_{\{y_t=i\}} \mathbf{1}_{\{x_t=o\}}\right)$$

(set
$$\lambda_{ij} := \log p(y' = i | y = j)$$
 , ...)

We let new parameters vary freely, so we need normalization constant Z.

$$\begin{aligned} & \text{Linear-Chain} \\ & \text{Conditional Random Fields} \end{aligned}$$

$$p(\mathbf{y}, \mathbf{x}) = \frac{1}{Z} exp \left(\sum_{t} \sum_{i,j \in S} \lambda_{ij} \mathbf{1}_{\{y_t=i\}} \mathbf{1}_{\{y_{t-1}=j\}} + \sum_{t} \sum_{i \in S} \sum_{o \in O} \mu \\ \text{Inis is a} \\ \text{Inear-chain} \\ \text{One feature per transition} \\ & \text{One feature per state-observati} \\ & \text{I}_{fij}(y, y', x_t) := \mathbf{1}_{y=i} \mathbf{1}_{y'=j} , \quad f_{io}(y, y', x_t) := \mathbf{1}_{y=i} \mathbf{1}_{x=o} \\ & p(\mathbf{y}, \mathbf{x}) = \frac{1}{Z} exp \left(\sum_{k=1}^{K} \lambda_k f_k(y_t, y_{t-1}, x_t) \right) \\ \text{One the the conditional distribution is} \end{aligned}$$

$$p(\mathbf{y}|\mathbf{x}) = \frac{p(\mathbf{y}, \mathbf{x})}{\sum_{y'} p(\mathbf{y'}, \mathbf{x})} = \frac{exp \left(\sum_{k=1}^{K} \lambda_k f_k(y_t, y_{t-1}, x_t) \right)}{\sum_{y'} exp \left(\sum_{k=1}^{K} \lambda_k f_k(y_t, y_{t-1}, x_t) \right)} \end{aligned}$$

Linear-Chain Conditional Random Fields • Conditional p(y|x) that follows from joint p(y,x) of HMM is a linear CRF with certain feature functions!

Linear-Chain Conditional Random Fields

• Definition:

A <u>linear-chain CRF</u> is a distribution that

takes the form

parameters

$$p(\mathbf{y}|\mathbf{x}) = \frac{1}{Z(\mathbf{x})} exp\left(\sum_{k=1}^{K} \lambda_k f_k(y_t, y_{t-1}, \mathbf{x}_t)\right)$$

where Z(x) is a normalization function

$$Z(x) = \sum_{\mathbf{y}} exp\left(\sum_{k=1}^{K} \lambda_k f_k(y_t, y_{t-1}, \mathbf{x}_t)\right)$$

Linear-Chain Conditional Random Fields • HMM-like linear-chain CRF



• Linear-chain CRF, in which transition score depends on the current observation

