

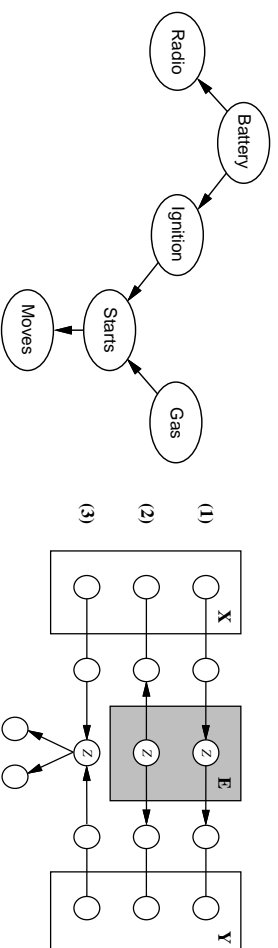
Introduction to Artificial Intelligence

Inference in belief networks

Chapter 15.2-4 + new

Dieter Fox

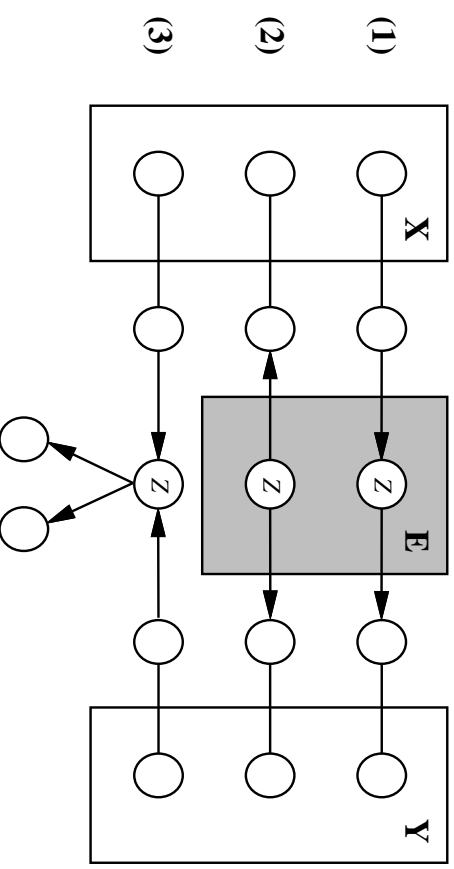
Example



1. $E = \text{Ignition}$ d-separates *Gas* and *Radio*
2. $E = \text{Battery}$ d-separates *Gas* and *Radio*
3. *Gas* and *Radio* are independent given no evidence, but *Gas* and *Radio* are dependent given $E = \text{Starts}$ or $E = \text{Moves}$.

D-Separation

Nodes X are **independent** of nodes Y given E , when every undirected path from a node in X to a node in Y is d-separated by E .



Inference, outline

- ◇ Exact inference by enumeration
- ◇ Exact inference by variable elimination
- ◇ Approximate inference by stochastic simulation

Inference by enumeration

Slightly intelligent way to sum out variables from the joint without actually constructing its explicit representation

Simple query on the burglary network:

$$\begin{aligned} & \mathbf{P}(B|J = true, M = true) \\ &= \mathbf{P}(B, J = true, M = true) / \mathbf{P}(J = true, M = true) \\ &= \alpha \mathbf{P}(B, J = true, M = true) \\ &= \alpha \sum_e \sum_a \mathbf{P}(B, e, a, J = true, M = true) \end{aligned}$$

Rewrite full joint entries using product of CPT entries:

$$\begin{aligned} & \mathbf{P}(B = true | J = true, M = true) \\ &= \alpha \sum_e \sum_a \mathbf{P}(B = true | e) \mathbf{P}(a | B = true, e) \mathbf{P}(J = true | a) \mathbf{P}(M = true | a) \\ &= \alpha \mathbf{P}(B = true) \sum_e \mathbf{P}(e) \sum_a \mathbf{P}(a | B = true, e) \mathbf{P}(J = true | a) \mathbf{P}(M = true | a) \end{aligned}$$

Inference by variable elimination

Enumeration is inefficient: repeated computation

e.g., computes $\mathbf{P}(J = true | a) \mathbf{P}(M = true | a)$ for each value of e

Variable elimination: carry out summations right-to-left, storing intermediate results (**factors**) to avoid recomputation

$$\begin{aligned} & \mathbf{P}(B | J = true, M = true) \\ &= \alpha \underbrace{\mathbf{P}(B)}_B \sum_e \underbrace{\mathbf{P}(e)}_E \sum_a \underbrace{\mathbf{P}(a | B, e)}_A \underbrace{\mathbf{P}(J = true | a)}_J \underbrace{\mathbf{P}(M = true | a)}_M \\ &= \alpha \mathbf{P}(B) \sum_e \mathbf{P}(e) \sum_a \mathbf{P}(a | B, e) \mathbf{P}(J = true | a) f_M(a) \\ &= \alpha \mathbf{P}(B) \sum_e \mathbf{P}(e) \sum_a \mathbf{P}(a | B, e) f_J(a) f_M(a) \\ &= \alpha \mathbf{P}(B) \sum_e \mathbf{P}(e) \sum_a f_A(a, b, e) f_J(a) f_M(a) \\ &= \alpha \mathbf{P}(B) \sum_e \mathbf{P}(e) f_{AJM}(b, e) \text{ (sum out } A) \\ &= \alpha \mathbf{P}(B) f_{EAJM}(b) \text{ (sum out } E) \\ &= \alpha f_B(b) \times f_{EAJM}(b) \end{aligned}$$

Enumeration algorithm

Exhaustive depth-first enumeration: $O(n)$ space, $O(d^n)$ time

<p>ENUMERATE_{ASK}(X, e, bn) returns a distribution over X</p> <p>inputs: X, the query variable</p> <p>e, evidence specified as an event</p> <p>bn, a belief network specifying joint distribution $\mathbf{P}(X_1, \dots, X_n)$</p> <p>$\mathbf{Q}(X) \leftarrow$ a distribution over X</p> <p>for each value x_i of X do</p> <p> extend e with value x_i for X</p> <p> $\mathbf{Q}(x_i) \leftarrow$ ENUMERATE_{ALL}($\text{Vars}[bn], e$)</p> <p>return NORMALIZE($\mathbf{Q}(X)$)</p> <p>ENUMERATE_{ALL}(vars, e) returns a real number</p> <p>if EMPT_Y?(vars) then return 1.0</p> <p>else do</p> <p> $Y \leftarrow$ FIRST(vars)</p> <p> if Y has value y in e then return $\mathbf{P}(y P_a(Y)) \times$ ENUMERATE_{ALL}(REST(vars), e)</p> <p> else return $\sum_y \mathbf{P}(y P_a(Y)) \times$ ENUMERATE_{ALL}(REST(vars), e_y)</p> <p> where e_y is e extended with $Y = y$</p>	
---	--

Complexity of exact inference

Singly connected networks (or **polytrees**):

- any two nodes are connected by at most one (undirected) path
- time and space cost of variable elimination are $O(d^k n)$

Multiply connected networks:

- can reduce 3SAT to exact inference \Rightarrow NP-hard
- equivalent to *counting* 3SAT models \Rightarrow #P-complete

Inference by stochastic simulation

Basic idea:

- 1) Draw N samples from a sampling distribution S
- 2) Compute an approximate posterior probability \hat{P}
- 3) Show this converges to the true probability P

Outline:

- Sampling from an empty network
- Rejection sampling: reject samples disagreeing with evidence
- Likelihood weighting: use evidence to weight samples
- MCMC: sample from a stochastic process whose stationary distribution is the true posterior

Sampling from an empty network contd.

Probability that $P_{PRIORSAMPLE}$ generates a particular event

$$S_{PS}(x_1 \dots x_n) = \prod_{i=1}^n P(x_i | Parents(X_i)) = P(x_1 \dots x_n)$$

i.e., the true prior probability

Let $N_{PS}(\mathbf{Y} = \mathbf{y})$ be the number of samples generated for which $\mathbf{Y} = \mathbf{y}$, for any set of variables \mathbf{Y} .

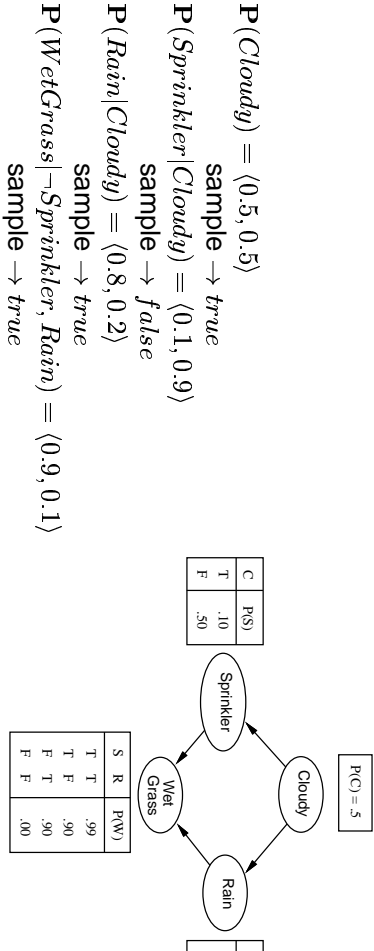
Then $\hat{P}(\mathbf{Y} = \mathbf{y}) = N_{PS}(\mathbf{Y} = \mathbf{y})/N$ and

$$\begin{aligned} \lim_{N \rightarrow \infty} \hat{P}(\mathbf{Y} = \mathbf{y}) &= \sum_h S_{PS}(\mathbf{Y} = \mathbf{y}, \mathbf{H} = \mathbf{h}) \\ &= \sum_h P(\mathbf{Y} = \mathbf{y}, \mathbf{H} = \mathbf{h}) \\ &= P(\mathbf{Y} = \mathbf{y}) \end{aligned}$$

That is, estimates derived from $P_{PRIORSAMPLE}$ are **consistent**

Sampling from an empty network

function $P_{PRIORSAMPLE}(bn)$ returns an event sampled from $P(X_1, \dots, X_n)$ specified
x ← an event with n elements
for $i = 1$ to n do
 x_i ← a random sample from $P(X_i | Parents(X_i))$
return x



Rejection sampling

$\hat{P}(X|e)$ estimated from samples agreeing with e

function REJECTIONSMPLING(X, e, bn, N) returns an approximation to $P(X|e)$
 $N[X] \leftarrow$ a vector of counts over X , initially zero
for $j = 1$ to N do
 $x \leftarrow P_{PRIORSAMPLE}(bn)$
 if x is consistent with e then
 $N[x] \leftarrow N[x] + 1$ where x is the value of X in x
return NORMALIZE($N[X]$)

E.g., estimate $P(Rain|Sprinkler = true)$ using 100 samples
27 samples have $Sprinkler = true$

Of these, 8 have $Rain = true$ and 19 have $Rain = false$.

$\hat{P}(Rain|Sprinkler = true) = \text{NORMALIZE}(\langle 8, 19 \rangle) = \langle 0.296, 0.704 \rangle$

Similar to a basic real-world empirical estimation procedure

Analysis of rejection sampling

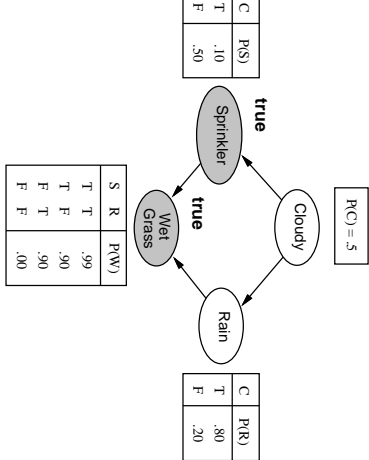
$$\begin{aligned} \hat{P}(X|e) &= \alpha N_{PS}(X, e) && \text{(algorithm defn.)} \\ &= N_{PS}(X, e) / N_{PS}(e) && \text{(normalized by } N_{PS}(e)) \\ &\approx P(X, e) / P(e) && \text{(property of } P_{\text{PRIORSAMPLE}}) \\ &= P(X|e) && \text{(defn. of conditional probability)} \end{aligned}$$

Hence rejection sampling returns consistent posterior estimates

Problem: hopelessly expensive if $P(e)$ is small

Likelihood weighting example

Estimate $P(Rain|Sprinkler = true, WetGrass = true)$



Likelihood weighting

Idea: fix evidence variables, sample only nonevidence variables, and weight each sample by the likelihood it accords the evidence

```
function WEIGHTEDSAMPLE( $bn, e$ ) returns an event and a weight
   $x \leftarrow$  an event with  $n$  elements;  $w \leftarrow 1$ 
  for  $i = 1$  to  $n$  do
    if  $X_i$  has a value  $x_i$  in  $e$ 
      then  $w \leftarrow w \times P(X_i = x_i \mid Parents(X_i))$ 
    else  $x_i \leftarrow$  a random sample from  $P(X_i \mid Parents(X_i))$ 
  return  $x, w$ 

function LIKELIHOODWEIGHTING( $X, e, bn, N$ ) returns an approximation to  $P(X|e)$ 
   $W[X] \leftarrow$  a vector of weighted counts over  $X$ , initially zero
  for  $j = 1$  to  $N$  do
     $x, w \leftarrow$  WEIGHTEDSAMPLE( $bn$ )
     $W[x] \leftarrow W[x] + w$  where  $x$  is the value of  $X$  in  $x$ 
  return NORMALIZE( $W[X]$ )
```

LW example contd.

Sample generation process:

1. $w \leftarrow 1.0$
2. Sample $P(Cloudy) = \langle 0.5, 0.5 \rangle$; say *true*
3. *Sprinkler* has value *true*, so
 $w \leftarrow w \times P(Sprinkler = true \mid Cloudy = true) = 0.1$
4. Sample $P(Rain \mid Cloudy = true) = \langle 0.8, 0.2 \rangle$; say *true*
5. *WetGrass* has value *true*, so
 $w \leftarrow w \times P(WetGrass = true \mid Sprinkler = true, Rain = true) = 0.099$

Approximate inference using MCMC

“State” of network = current assignment to all variables

Generate next state by sampling one variable given its Markov blanket

Sample each variable in turn, keeping evidence fixed

Approaches **stationary distribution**: long-run fraction of time spent in each state is exactly proportional to its posterior probability

Main computational problems:

- 1) Difficult to tell if convergence has been achieved
- 2) Can be wasteful if Markov blanket is large:

$P(Y_i | MB(Y_i))$ won't change much (law of large numbers)

Case study: Pathfinder IV

Diagnostic expert system for lymph-node diseases.

Deciding on vocabulary: 8 hours

Design topology of network: 35 hours

Make 14,000 probability assessments: 40 hours

Pathfinder now outperforms experts who were consulted during its creation!

Performance of approximation algorithms

Absolute approximation: $|P(X|\mathbf{e}) - \hat{P}(X|\mathbf{e})| \leq \epsilon$

Relative approximation: $\frac{|P(X|\mathbf{e}) - \hat{P}(X|\mathbf{e})|}{P(X|\mathbf{e})} \leq \epsilon$

Relative \Rightarrow absolute since $0 \leq P \leq 1$

Randomized algorithms may fail with probability at most δ

Polytime approximation: $\text{poly}(n, \epsilon^{-1}, \log \delta^{-1})$

Theorem (Dagum and Luby, 1993): both absolute and relative approximation for either deterministic or randomized algorithms are NP-hard for any $\epsilon, \delta < 0.5$