

My office hrs:

Monday immediately
after lecture.

Linear Regression

Recap

- Learning is...
 - Collect some data
 - E.g., coin flips

Data $\{x_i\}$

Recap

- Learning is...
 - Collect some data
 - E.g., coin flips
 - Choose a hypothesis class or model
 - E.g., binomial



Recap

- Learning is...
 - Collect some data
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 - Choose a hypothesis class or model
 - E.g., binomial
 - Choose a loss function
 - E.g., data likelihood

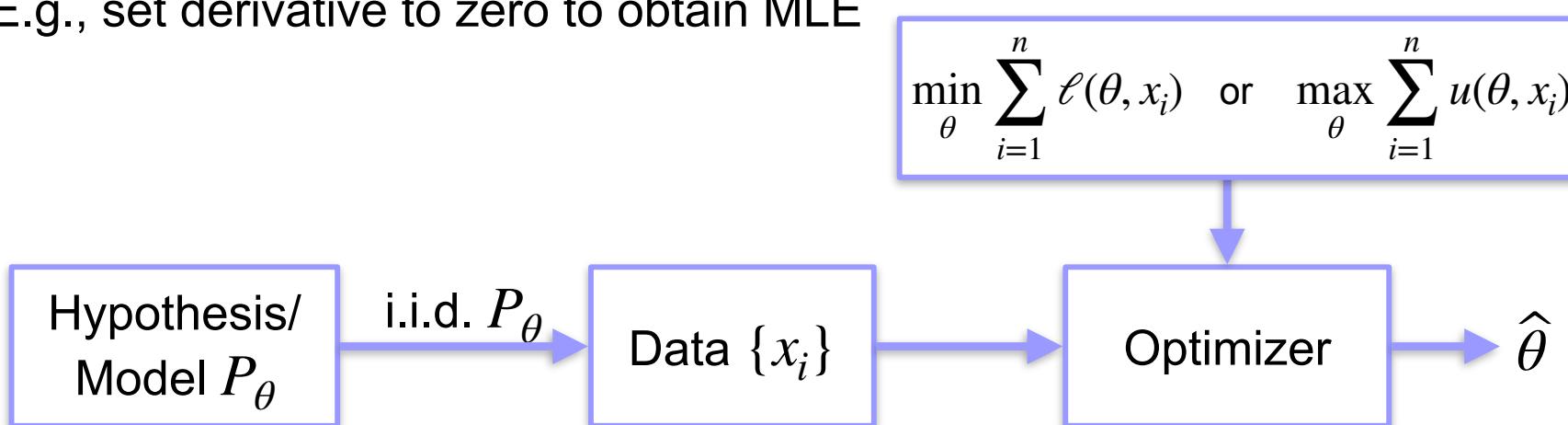
$$\min_{\theta} \sum_{i=1}^n \ell(\theta, x_i) \quad \text{or} \quad \max_{\theta} \sum_{i=1}^n u(\theta, x_i)$$



Recap

- Learning is...
 - Collect some data
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 - Choose an optimization procedure
 - E.g., set derivative to zero to obtain MLE

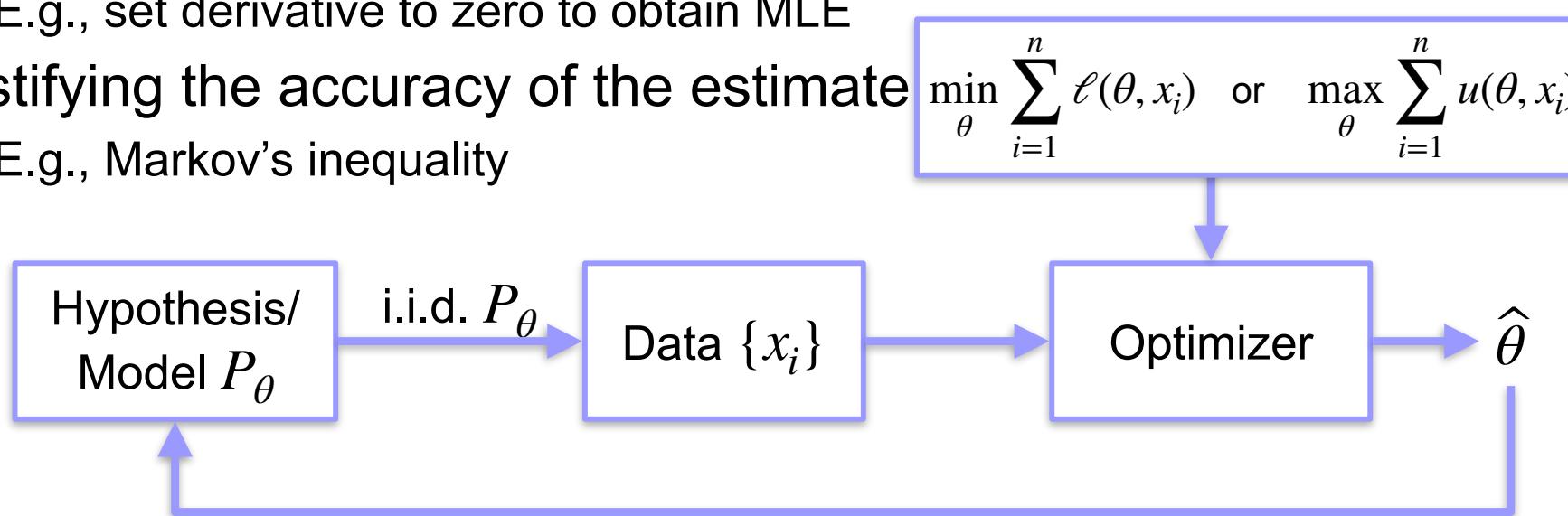
$$\min_{\theta} \sum_{i=1}^n \ell(\theta, x_i) \quad \text{or} \quad \max_{\theta} \sum_{i=1}^n u(\theta, x_i)$$



Recap

- Learning is...
 - Collect some data
 - E.g., coin flips
 - Choose a hypothesis class or model
 - E.g., binomial
 - Choose a loss function
 - E.g., data likelihood
 - Choose an optimization procedure
 - E.g., set derivative to zero to obtain MLE
 - Justifying the accuracy of the estimate
 - E.g., Markov's inequality

$$\min_{\theta} \sum_{i=1}^n \ell(\theta, x_i) \quad \text{or} \quad \max_{\theta} \sum_{i=1}^n u(\theta, x_i)$$



Linear Regression

Hypothesis
Model

Real-valued
predictions
(supervised)

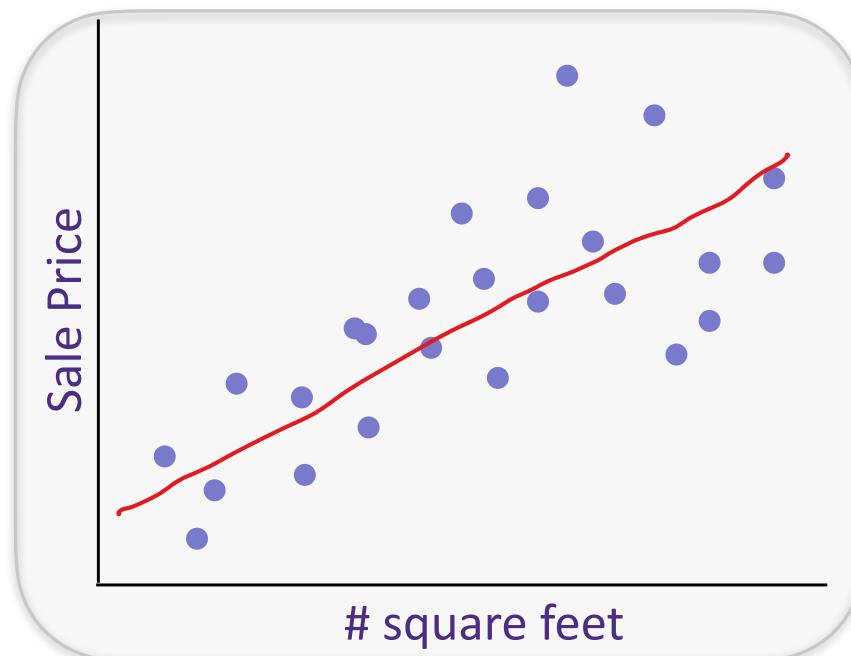
The regression problem, 1-dimensional

You want to sell your house that is 2,500 sq.ft.

Q. What is the right price?

Collect past sales data on [zillow.com](#):

$y = \text{House sale price}$ and $x = \{\# \text{ sq. ft.}\}$



Training Data: $x_i \in \mathbb{R}$ $y_i \in \mathbb{R}$
 $\{(x_i, y_i)\}_{i=1}^n$

\downarrow \rightarrow label / sale price
square footage

Process

1. Decide on a model/hypothesis class

assume house sale price is a linear function of square feet.

+ Noise/error

2. Find the function/model/hypothesis which explains/fits the data best

3. Use function to make prediction on new examples

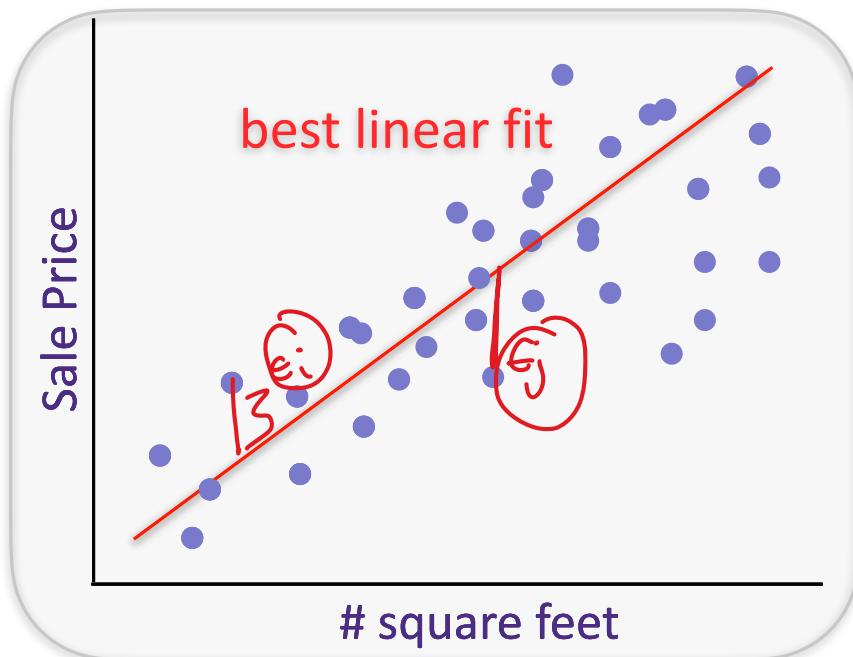
How much should you put your house on the market?

Fit a function to our data, 1-dimension

Given past sales data on [zillow.com](#), predict:

y = House sale price from

x = {# sq. ft.}



Fixing $w = 10$
 $\epsilon_i = 10 \cdot x_i - y_i$

1. Training Data:

$$\{(x_i, y_i)\}_{i=1}^n \quad x_i \in \mathbb{R} \quad y_i \in \mathbb{R}$$

2. Hypothesis/Model: linear

$$y_i = w \cdot x_i + \epsilon_i \quad \text{Noise}$$
$$y_i - w x_i = \epsilon_i$$

3. Measure of good fit: ℓ_2 -loss

$$\min_{w \in \mathbb{R}} \sum_{i=1}^n (y_i - w x_i)^2 = \sum_{i=1}^n \epsilon_i^2$$

The regression problem, d -dimensions

Given past sales data on [zillow.com](#), predict:

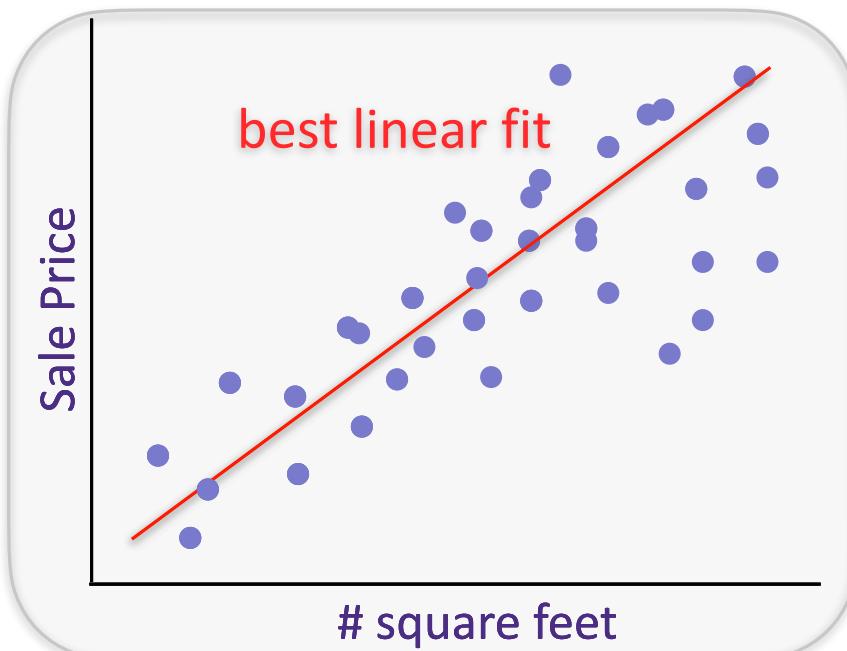
y = House sale price from

$x = \{\# \text{ sq. ft., zip code, date of sale, etc.}\}$

$w = (w_0, w_1, \dots, w_d)$

years old

construction



1. Training Data:

$$\{(x_i, y_i)\}_{i=1}^n$$

2. Hypothesis/Model: linear

$$y_i = w^T x_i + \epsilon_i$$

3. Measure of good fit: ℓ_2 -loss

$$\min_{w \in \mathbb{R}^d} \sum_{i=1}^n (y_i - w^T x_i)^2 = \sum_{i=1}^n \epsilon_i^2$$

The regression problem in matrix notation

Data:

$$\mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} \quad \mathbf{X} = \begin{bmatrix} \mathbb{E} \mathbb{R}_{\geq 0} \\ x_1^T \\ \vdots \\ x_n^T \end{bmatrix}$$

$n \left[\begin{array}{ccc|c} x_{11} & \cdots & x_{1d} & \\ \hline & x_1^T & & \\ & x_n^T & & \end{array} \right]$

d : # of features/size of the input
n : # of examples/datapoints

The regression problem in matrix notation

Data:

$$\mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} \quad \mathbf{X} = \begin{bmatrix} \mathbf{x}_1^T \\ \vdots \\ \mathbf{x}_n^T \end{bmatrix}$$

d : # of features/size of the input
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Linear Model:

$$y_1 = \mathbf{x}_1^T \mathbf{w} + \epsilon_1$$

$$y_2 = \mathbf{x}_2^T \mathbf{w} + \epsilon_2$$

⋮

⋮

$$y_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$$

$$\mathbf{y} = \mathbf{X}\mathbf{w} + \mathbf{\epsilon}$$

Diagram illustrating the linear model equation $\mathbf{y} = \mathbf{X}\mathbf{w} + \mathbf{\epsilon}$ using red annotations:

- Dimensions: $\mathbf{y} \in \mathbb{R}^n$, $\mathbf{X} \in \mathbb{R}^{n \times d}$, $\mathbf{w} \in \mathbb{R}^d$, $\mathbf{\epsilon} \in \mathbb{R}^n$.
- Matrix multiplication: $\mathbf{X}\mathbf{w}$ is shown as n vectors of length d being summed.
- Vector addition: $\mathbf{y} = \mathbf{X}\mathbf{w} + \mathbf{\epsilon}$ is shown as the sum of $\mathbf{X}\mathbf{w}$ and $\mathbf{\epsilon}$.

The regression problem in matrix notation

Data:

$$\mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} \quad \mathbf{X} = \begin{bmatrix} \mathbf{x}_1^T \\ \vdots \\ \mathbf{x}_n^T \end{bmatrix}$$

d : # of features/size of the input
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Linear Model:

$$y_1 = \mathbf{x}_1^T \mathbf{w} + \epsilon_1$$

$$\mathbf{y} = \mathbf{X} \mathbf{w} + \boldsymbol{\epsilon}$$

$$y_2 = \mathbf{x}_2^T \mathbf{w} + \epsilon_2$$

.

.

$$y_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$$

ℓ_2 -Loss:

$$\widehat{\mathbf{w}}_{\text{LS}} = \arg \min_{\mathbf{w} \in \mathbb{R}^d} \sum_{i=1}^n (y_i - \mathbf{x}_i^T \mathbf{w})^2$$

this is also known as **Least Squares** solution

ℓ_2 -norm of a vector:
(also known as Euclidean norm)

$$\|\boldsymbol{\epsilon}\|_2 = \sqrt{\epsilon_1^2 + \epsilon_2^2 + \cdots + \epsilon_d^2}$$

it follows that

$$\sum_{i=1}^d \epsilon_i^2 = \|\boldsymbol{\epsilon}\|_2^2 = \boldsymbol{\epsilon}^T \boldsymbol{\epsilon}$$

The regression problem in matrix notation

Data:

$$\mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} \quad \mathbf{X} = \begin{bmatrix} \mathbf{x}_1^T \\ \vdots \\ \mathbf{x}_n^T \end{bmatrix}$$

d : # of features/size of the input
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Linear Model:

$$y_1 = \mathbf{x}_1^T \mathbf{w} + \epsilon_1 \quad \mathbf{y} = \mathbf{X} \mathbf{w} + \boldsymbol{\epsilon}$$

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⋮

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$$y_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$$

ℓ_2 -norm of a vector:

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it follows that

$$\sum_{i=1}^d \epsilon_i^2 = \|\boldsymbol{\epsilon}\|_2^2 = \boldsymbol{\epsilon}^T \boldsymbol{\epsilon}$$

ℓ_2 -Loss: $\widehat{\mathbf{w}}_{\text{LS}} = \arg \min_{\mathbf{w} \in \mathbb{R}^d} \sum_{i=1}^n (y_i - \mathbf{x}_i^T \mathbf{w})^2$

$$\begin{aligned} &= \arg \min_{\mathbf{w}} \|\mathbf{y} - \mathbf{X} \mathbf{w}\|_2^2 \\ &= \arg \min_{\mathbf{w}} (\mathbf{y} - \mathbf{X} \mathbf{w})^T (\mathbf{y} - \mathbf{X} \mathbf{w}) \end{aligned}$$

The regression problem in matrix notation

$$f(\hat{w}_{\text{LS}}) = \arg \min_{w \in \mathbb{R}^d} (\underline{y} - \underline{\mathbf{X}w})^T (\underline{y} - \underline{\mathbf{X}w})$$

Set gradient w.r.t. w to zero to find the minima:

$$\begin{aligned}\nabla_w f &= 2 \Omega^T (\Omega \gamma + \beta) \\ &= -2 \mathbf{X}^T (-\mathbf{X}w + y) \\ &= 2 \mathbf{X}^T \mathbf{X}w - 2 \mathbf{X}^T y \\ \Rightarrow \mathbf{X}^T \mathbf{X}w &= \mathbf{X}^T y \\ \Rightarrow w &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T y\end{aligned}$$

A few reminders on vector calculus

- Gradient of a function:

$$\nabla_w f(w) = \begin{bmatrix} \frac{df(w)}{dw_1} \\ \frac{df(w)}{dw_2} \\ \vdots \\ \frac{df(w)}{dw_d} \end{bmatrix} \quad \begin{aligned}f(\gamma) &= \gamma^T \gamma \\ \nabla_\gamma f &= 2 \gamma \\ f(\gamma) - (\Omega \gamma)^T (\Omega \gamma) &= \nabla_\gamma f = 2 \Omega \Omega^T \gamma \\ f(\gamma) &= (\Omega \gamma + \beta)^T \gamma\end{aligned}$$

- Example:

$$\begin{aligned}f(w) &= w^T w \implies \nabla_w f(w) = 2w \\ f(w) &= (Aw)^T (Aw) \implies \nabla_w f(w) = 2AA^T w \\ f(w) &= (Aw + b)^T (Aw + b) \implies \nabla_w f(w) = 2A^T(Aw + b)\end{aligned}$$

$$\Rightarrow \nabla_f f = 2 \Omega^T (\Omega \gamma + \beta)$$

The regression problem in matrix notation

$$\begin{aligned}\hat{w}_{LS} &= \arg \min_w \|\mathbf{y} - \mathbf{X}w\|_2^2 \\ &= \arg \min_w (\mathbf{y} - \mathbf{X}w)^T (\mathbf{y} - \mathbf{X}w) \\ &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}\end{aligned}$$

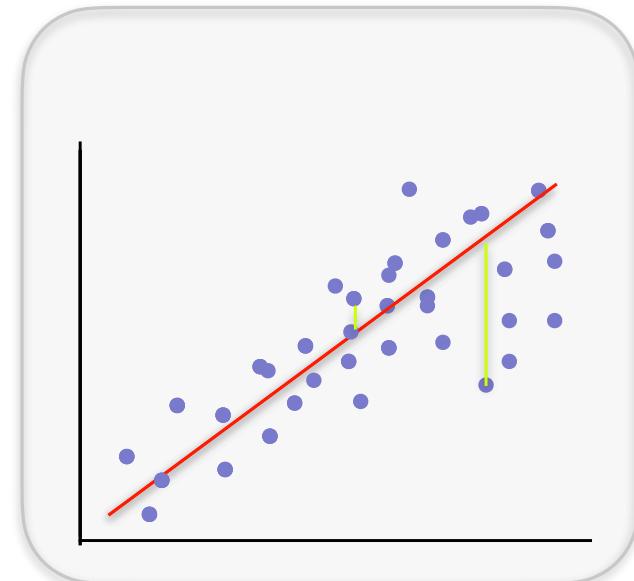
“Closed form” solution!

The regression problem in matrix notation

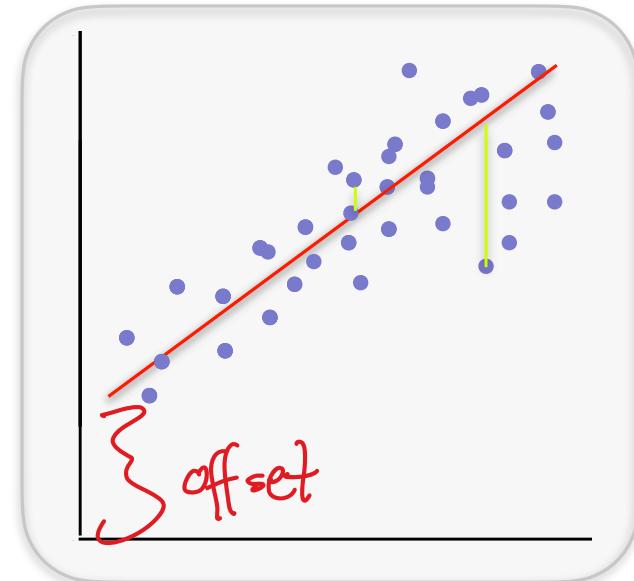
Linear model: $y_i = x_i^T w + \epsilon_i$

Least squares solution:

$$\begin{aligned}\hat{w}_{LS} &= \arg \min_w \|\mathbf{y} - \mathbf{X}w\|_2^2 \\ &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}\end{aligned}$$



What about an offset
(a.k.a intercept)?

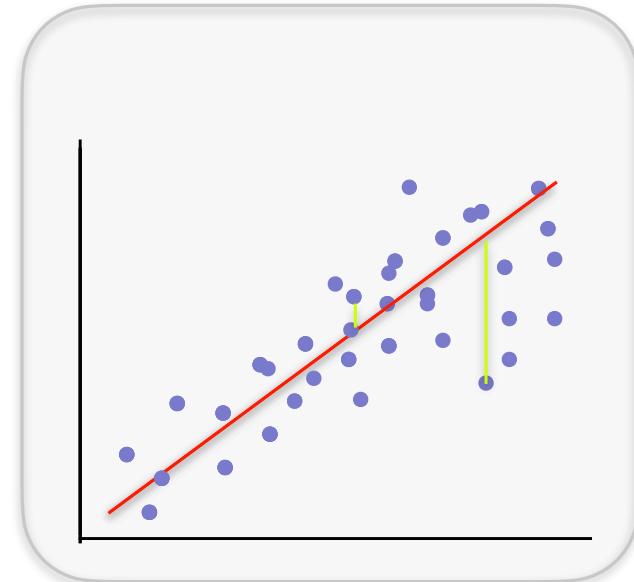


The regression problem in matrix notation

Linear model: $y_i = x_i^T w + \epsilon_i$

Least squares solution:

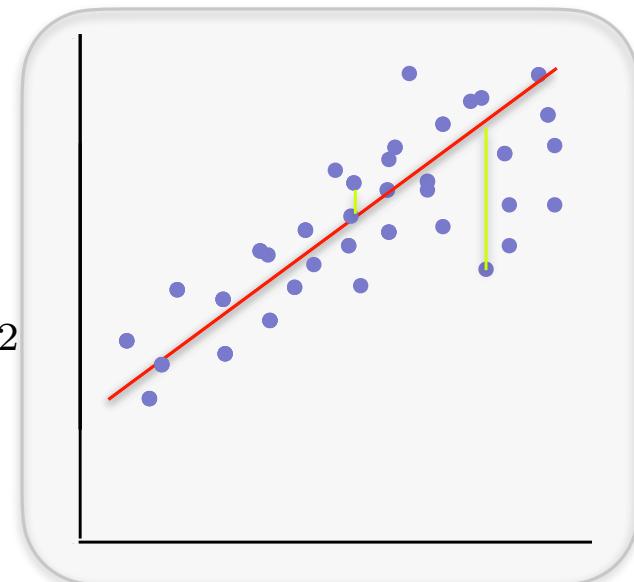
$$\begin{aligned}\hat{w}_{LS} &= \arg \min_w \|\mathbf{y} - \mathbf{X}w\|_2^2 \\ &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}\end{aligned}$$



Affine model: $y_i = x_i^T w + b + \epsilon_i$

Least squares solution:

$$\begin{aligned}\hat{w}_{LS}, \hat{b}_{LS} &= \arg \min_{w,b} \sum_{i=1}^n (y_i - (x_i^T w + b))^2 \\ &= \arg \min_{w,b} \|\mathbf{y} - (\mathbf{X}w + \underbrace{\mathbf{1}b}_{\in \mathbb{R}})\|_2^2\end{aligned}$$



Dealing with an offset

$$\begin{aligned}\hat{\mathbf{w}}_{\text{LS}}, \hat{b}_{\text{LS}} &= \arg \min_{w \in \mathbb{R}^d, b \in \mathbb{R}} \|\mathbf{y} - (\mathbf{X}w + \mathbf{1}b)\|_2^2 \\ &= \arg \min_{w \in \mathbb{R}^d, b \in \mathbb{R}} \underbrace{(\mathbf{y} - (\mathbf{X}w + \mathbf{1}b))^T(\mathbf{y} - (\mathbf{X}w + \mathbf{1}b))}_{\mathcal{L}(w, b)}\end{aligned}$$

Set gradient w.r.t. w and b to zero to find the minima:

$$\beta = \gamma^{-1} b$$

A reminder on vector calculus

$$f(w) = (\mathbf{A}w + \mathbf{b})^T(\mathbf{A}w + \mathbf{b}) \implies \nabla_w f(w) = 2\mathbf{A}^T(\mathbf{A}w + \mathbf{b})$$

Dealing with an offset

$$\hat{w}_{LS}, \hat{b}_{LS} = \arg \min_{w,b} \|\mathbf{y} - (\mathbf{X}w + \mathbf{1}b)\|_2^2$$

$$\mathbf{X}^T \mathbf{X} \hat{w}_{LS} + \hat{b}_{LS} \mathbf{X}^T \mathbf{1} = \mathbf{X}^T \mathbf{y}$$

$$\mathbf{1}^T \mathbf{X} \hat{w}_{LS} + \hat{b}_{LS} \mathbf{1}^T \mathbf{1} = \mathbf{1}^T \mathbf{y}$$

If $\mathbf{X}^T \mathbf{1} = 0$, if the features have zero mean,

$$\hat{w}_{LS} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

$$\hat{b}_{LS} = (\mathbf{1}^T \mathbf{1})^{-1} \mathbf{1}^T \mathbf{y} = \frac{1}{n} \sum_{i=1}^n y_i$$

Dealing with an offset

$$\hat{w}_{LS}, \hat{b}_{LS} = \arg \min_{w,b} \|\mathbf{y} - (\mathbf{X}w + \mathbf{1}b)\|_2^2$$

$$\mathbf{X}^T \mathbf{X} \hat{w}_{LS} + \hat{b}_{LS} \mathbf{X}^T \mathbf{1} = \mathbf{X}^T \mathbf{y}$$

$$\mathbf{1}^T \mathbf{X} \hat{w}_{LS} + \hat{b}_{LS} \mathbf{1}^T \mathbf{1} = \mathbf{1}^T \mathbf{y}$$

If $\mathbf{X}^T \mathbf{1} = 0$,

$$\hat{w}_{LS} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

$$\hat{b}_{LS} = \frac{1}{n} \sum_{i=1}^n y_i$$

In general, when $\mathbf{X}^T \mathbf{1} \neq 0$,

Dealing with an offset

$$\hat{w}_{LS}, \hat{b}_{LS} = \arg \min_{w,b} \|\mathbf{y} - (\mathbf{X}w + \mathbf{1}b)\|_2^2$$

$$\mathbf{X}^T \mathbf{X} \hat{w}_{LS} + \hat{b}_{LS} \mathbf{X}^T \mathbf{1} = \mathbf{X}^T \mathbf{y}$$

$$\mathbf{1}^T \mathbf{X} \hat{w}_{LS} + \hat{b}_{LS} \mathbf{1}^T \mathbf{1} = \mathbf{1}^T \mathbf{y}$$

If $\mathbf{X}^T \mathbf{1} = 0$,

$$\hat{w}_{LS} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

$$\hat{b}_{LS} = \frac{1}{n} \sum_{i=1}^n y_i$$

In general, when $\mathbf{X}^T \mathbf{1} \neq 0$,

$$\mu = \frac{1}{n} \mathbf{X}^T \mathbf{1}$$

$$\tilde{\mathbf{X}} = \mathbf{X} - \mathbf{1} \mu^T$$

$$\hat{w}_{LS} = (\tilde{\mathbf{X}}^T \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{X}}^T \mathbf{y}$$

$$\hat{b}_{LS} = \frac{1}{n} \sum_{i=1}^n y_i - \mu^T \hat{w}_{LS}$$

Process for linear regression with intercept

Collect data: $\mathcal{D} = \{(x_i, y_i)\}_{i=1}^n$

Decide on a **model**: $y_i = x_i^T w + b + \epsilon_i$

Choose a loss function - least squares

Pick the function which minimizes loss on data

$$\hat{w}_{LS}, \hat{b}_{LS} = \arg \min_{w,b} \sum_{i=1}^n (y_i - (x_i^T w + b))^2$$

Use function to make prediction on new examples x_{new}

$$\hat{y}_{\text{new}} = x_{\text{new}}^T \hat{w}_{LS} + \hat{b}_{LS}$$

Another way of dealing with an offset

$$\hat{w}_{LS}, \hat{b}_{LS} = \arg \min_{w,b} \|\mathbf{y} - (\mathbf{X}w + \mathbf{1}b)\|_2^2$$

reparametrize the problem as $\bar{\mathbf{X}} = [\mathbf{X}, \mathbf{1}]$ and $\bar{w} = \begin{bmatrix} w \\ b \end{bmatrix}$

$$\bar{\mathbf{X}} \bar{w} =$$

Why do we use least squares (i.e. ℓ_2 -loss)?

$$\begin{aligned}\hat{w}_{LS} &= \arg \min_w \|\mathbf{y} - \mathbf{X}w\|_2^2 \\ &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}\end{aligned}$$

$$\implies y_i \sim$$

$$\implies P(y_i; x_i, w, \sigma) =$$

Why do we use least squares (i.e. ℓ_2 -loss)?

$$\begin{aligned}\hat{w}_{LS} &= \arg \min_w \|\mathbf{y} - \mathbf{X}w\|_2^2 \\ &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}\end{aligned}$$

Consider $y_i = x_i^T w + \epsilon_i$ where $\epsilon_i \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \sigma^2)$

$$\implies y_i \sim$$

$$\implies P(y_i; x_i, w, \sigma) =$$

Why do we use least squares (i.e. ℓ_2 -loss)?

Maximum Likelihood Estimator:

$$\begin{aligned}\hat{w}_{\text{MLE}} &= \arg \max_w \log P(\{y_i\}_{i=1}^n; \{x_i\}_{i=1}^n, w, \sigma) \\ &= \arg \max_w -n \log(\sigma \sqrt{2\pi}) + \sum_{i=1}^n -\frac{(y_i - x_i^T w)^2}{2\sigma^2}\end{aligned}$$

Why do we use least squares (i.e. ℓ_2 -loss)?

Maximum Likelihood Estimator:

$$\begin{aligned}\hat{w}_{MLE} &= \arg \max_w \log P(\{y_i\}_{i=1}^n; \{x_i\}_{i=1}^n, w, \sigma) \\ &= \arg \max_w -n \log(\sigma \sqrt{2\pi}) + \sum_{i=1}^n -\frac{(y_i - x_i^T w)^2}{2\sigma^2} \\ &= \arg \min_w \sum_{i=1}^n (y_i - x_i^T w)^2\end{aligned}$$

Recall: $\hat{w}_{LS} = \arg \min_w \sum_{i=1}^n (y_i - x_i^T w)^2$

$$\boxed{\hat{w}_{LS} = \hat{w}_{MLE} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y}}$$

Recap of linear regression

Data $\{(x_i, y_i)\}_{i=1}^n$

**Minimize the loss
(Empirical Risk Minimization)**

Choose a loss
e.g., ℓ_2 -loss: $(y_i - x_i^T w)^2$

Solve $\widehat{w}_{\text{LS}} = \arg \min_w \sum_{i=1}^n (y_i - x_i^T w)^2$

**Maximize the likelihood
(MLE)**

Choose a Hypothesis class
e.g., $y_i = x_i^T w + \epsilon_i$, $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$

Maximize the likelihood,
 $\widehat{w}_{\text{MLE}} = \arg \max_w \left\{ -n \log(\sigma \sqrt{2\pi}) - \sum_{i=1}^n \frac{(y_i - x_i^T w)^2}{2\sigma^2} \right\}$

Analysis of Error under additive Gaussian noise

Let's suppose $y_i = x_i^T w^* + \epsilon_i$ and $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$, then this can be written as
 $\mathbf{y} = \mathbf{X}w^* + \boldsymbol{\epsilon}$

$$\begin{aligned}\widehat{w}_{\text{MLE}} &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \\ &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T (\mathbf{X}w^* + \boldsymbol{\epsilon}) \\ &= w^* + (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \boldsymbol{\epsilon}\end{aligned}$$

Maximum Likelihood Estimator is unbiased:

Analysis of Error under additive Gaussian noise

Let's suppose $y_i = x_i^T w^* + \epsilon_i$ and $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$, then this can be written as
 $\mathbf{y} = \mathbf{X}w^* + \boldsymbol{\epsilon}$

$$\begin{aligned}\widehat{w}_{\text{MLE}} &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \\ &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T (\mathbf{X}w^* + \boldsymbol{\epsilon}) \\ &= w^* + (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \boldsymbol{\epsilon}\end{aligned}$$

Covariance is:

Analysis of Error under additive Gaussian noise

Let's suppose $y_i = \mathbf{x}_i^T w^* + \epsilon_i$ and $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$, then this can be written as $\mathbf{y} = \mathbf{X}w^* + \boldsymbol{\epsilon}$, and the MLE is

$$\hat{w}_{\text{MLE}} = w^* + (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \boldsymbol{\epsilon}$$

This random estimate has the following distribution:

$$\mathbb{E}[\hat{w}_{\text{MLE}}] = w^*, \text{Cov}(\hat{w}_{\text{MLE}}) = \mathbb{E}[(\hat{w} - \mathbb{E}[\hat{w}])(\hat{w} - \mathbb{E}[\hat{w}])^T] = \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1}$$

$$\hat{w}_{\text{MLE}} \sim \mathcal{N}(w^*, \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1})$$

Interpretation: consider an example with $\mathbf{x} = \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ -1 & 0 \\ -1 & 0 \\ 0 & 1 \\ 0 & -1 \end{bmatrix}$

The covariance of the MLE, $\sigma^2 (\mathbf{X}^T \mathbf{X})^{-1}$, captures how each sample gives information about the unknown w^* , but each sample gives information about different (linear combination of) coordinates and of different quality/strength

Questions?
