

Joint Distributions

CSE 312 26Su

Lecture 17

Announcements

Homework 5 due Wednesday

Quiz 5 on Friday

CLT joint dist.
-polling

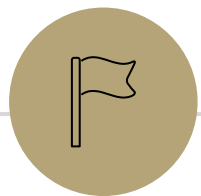
Where Are We?

Last time:

- Polling

Today:

- Joint Distributions
 - Joint Support
 - Joint PMF / PDF
 - Joint CDF
 - Joint Expectation
- Conditional Distributions
 - Conditional Expectation
- Law of Total Expectation
- Covariance



Multiple Random Variables

Analyzing Multiple Random Variables

So far, we've pretty much only analyzed one random variable at a time
We've *worked* with multiple random variables (e.g., the sum of them) but haven't really analyzed their relationships except for independence

$$Z = X + Y$$

Today, we will talk about analyzing the relations between multiple RVs

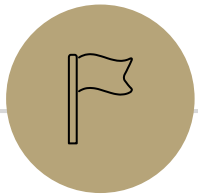
$$X \sim \text{Bin}(n, p)$$

$$P(X \leq k)$$

Analyzing *Multiple* Random Variables

Examples:

- **Economics:** Analyzing the relationship between stock returns and volume helps in understanding market behavior and making investment decisions
- • **Healthcare:** Understanding how different factors (like blood pressure and cholesterol levels) affects the probability of medical conditionals
- • **Machine learning:** understanding how different features contribute to predicting a label and improving model accuracy (what features are most important)
- • **Recommendation systems:** Understanding relationship between user preferences and item features to improve recommendation systems, or other factors like age and choices of products



Joint Distributions

We'll start with the discrete case

We have two **discrete** random variables X and Y
(that may or may not be independent)

We have two **discrete** random variables X and Y
(that may or may not be independent)

Support Ω_X

Joint Support/Range - $\Omega_{X,Y}$

$\Omega_{X,Y}$ is the set of all possible pairs of values X and Y can be together

$$> \Omega_{X,Y} = \{(\underbrace{a}_{\substack{\uparrow \\ x}}, \underbrace{b}_{\substack{\uparrow \\ y}}) : \underbrace{p_{X,Y}(a, b)}_{\text{joint pmf}} > 0\} \subseteq \underbrace{\Omega_X}_{\text{support of } X} \times \underbrace{\Omega_Y}_{\text{support of } Y}$$

We have two discrete random variables X and Y
(that may or may not be independent)

Joint **Support/Range** - $\Omega_{X,Y}$

$\Omega_{X,Y}$ is the set of all possible **pairs** of values X and Y can be together

$$> \Omega_{X,Y} = \{(a, b) : p_{X,Y}(a, b) > 0\} \subseteq \Omega_X \times \Omega_Y$$

Joint **PMF** (probability mass function) - $p_{X,Y}(a, b)$

$p_{X,Y}(a, b)$ defines probabilities of X and Y being certain values at once

$$> p_{X,Y}(a, b) = \mathbb{P}(X = a \cap Y = b) = \mathbb{P}(X = a, Y = b)$$

should be defined for *all* values of a and b

Normalization Property: $\sum_{(a,b) \in \Omega_{X,Y}} p_{X,Y}(a, b) = 1$

$$\sum_{x \in \Omega_X} p_X(x) = 1$$

We have two **discrete** random variables X and Y
(that may or may not be independent)

Joint **Support/Range** - $\Omega_{X,Y}$

$\Omega_{X,Y}$ is the set of all possible **pairs** of values X and Y can be together

$$> \Omega_{X,Y} = \{(a, b) : p_{X,Y}(a, b) > 0\} \subseteq \Omega_X \times \Omega_Y$$

Joint **PMF** (*probability mass function*) - $p_{X,Y}(a, b)$

$p_{X,Y}(a, b)$ defines probabilities of X and Y being certain values at once

$$> p_{X,Y}(a, b) = \mathbb{P}(X = a \cap Y = b) = \mathbb{P}(X = a, Y = b)$$

should be defined for **all** values of a and b

Joint **CDF** (*cumulative distribution function*) - $F_{X,Y}(a, b)$

$$> F_{X,Y}(a, b) = \mathbb{P}(\underline{X \leq a} \cap \underline{Y \leq b}) = \mathbb{P}(X \leq a, Y \leq b)$$

should be defined for **all** values of a and b

We have two **discrete** random variables X and Y
(that may or may not be independent)

(Joint) Independence

X and Y are independent if:

> $p_{X,Y}(a, b) = \underline{p_X(a)} \cdot p_Y(b)$ for all $(a, b) \in \underline{\Omega_{X,Y}}$ ✓

> $\Omega_{X,Y} = \Omega_X \times \Omega_Y$ ✓

if you find that $\Omega_{X,Y} \neq \Omega_X \times \Omega_Y$, that is enough to show that they aren't independent

This is the same definition we've seen before for independence of random variables
 $\mathbb{P}(X = a, Y = b) = \mathbb{P}(X = a) \cdot \mathbb{P}(Y = b)$

Example: Weird Dice

Roll 2 fair 4-sided dice independently

Let X be the value of the 1st dice

Let Y be the value of the 2nd dice

$$\underline{\Omega_X} = \underline{\Omega_Y} = \{1, 2, 3, 4\}$$

The joint support is $\underline{\Omega_{X,Y}} = \underline{\Omega_X \times \Omega_Y}$
because all combinations are possible

What is the joint PMF?

$$\underline{p_{X,Y}}(a, b) = \begin{cases} 1/16 & \text{if } (a, b) \in \Omega_{X,Y} \\ 0 & \text{o.w.} \end{cases}$$



$p_{X,Y}$	$Y=1$	$Y=2$	$Y=3$	$Y=4$
$X=1$	$\frac{1}{4} \cdot \frac{1}{4} = \frac{1}{16}$	$\frac{1}{16}$	$\frac{1}{16}$	$\frac{1}{16}$
$X=2$	...	...	...	...
$X=3$				
$X=4$				

Example: Weird Dice

Roll 2 fair 4-sided dice independently

Let X be the value of the 1st dice

Let Y be the value of the 2nd dice

$$\Omega_X = \Omega_Y = \{1, 2, 3, 4\}$$

The joint support is $\Omega_{X,Y} = \Omega_X \times \Omega_Y$
because all combinations are possible

What is the joint PMF?

$$p_{X,Y}(a, b) = \begin{cases} 1/16 & \text{if } a, b \in \Omega_{X,Y} \\ 0 & \text{otherwise} \end{cases}$$

$p_{X,Y}$	$Y=1$	$Y=2$	$Y=3$	$Y=4$
$X=1$	1/16	1/16	1/16	1/16
$X=2$	1/16	1/16	1/16	1/16
$X=3$	1/16	1/16	1/16	1/16
$X=4$	1/16	1/16	1/16	1/16

Example: Weird Dice

x, y – 2nd die roll
1st die roll

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$

Let W be the max value $W = \max(X, Y)$

$$\Omega_U = \Omega_W = \{1, 2, 3, 4\}$$

The joint support is:

$$\Omega_{U,W} = \{(u, w) \in \Omega_U \times \Omega_W; u \leq w\}$$

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$				
$U=2$				
$U=3$				
$U=4$				

Example: Weird Dice

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$

Let W be the max value $W = \max(X, Y)$

$$\Omega_U = \Omega_W = \{1, 2, 3, 4\}$$

The joint support is:

$$\Omega_{U,W} = \{(u, w) \in \Omega_U \times \Omega_W; u \leq w\}$$

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$				
$U=2$				
$U=3$				
$U=4$				

Example: Weird Dice

e.g. roll $X=2, Y=4$ then $U=2, W=4$ | $X=4, Y=2$ then $U=2, W=4$

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$

Let W be the max value $W = \max(X, Y)$

$$\Omega_U = \Omega_W = \{1, 2, 3, 4\}$$

The joint support is:

$$\Omega_{U,W} = \{(u, w) \in \Omega_U \times \Omega_W; u \leq w\}$$

What is the joint PMF?

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	$1/16$	$2/16$	$2/16$	$2/16$
$U=2$	0	$1/16$	$2/16$	$2/16$
$U=3$	0	0	$1/16$	$2/16$
$U=4$	0	0	0	$1/16$

$$p_{U,W}(u, w) = \begin{cases} 1/16 & \text{if } (u, w) \in \Omega_{u, w} \text{ and } u = w \\ 2/16 & \text{if } (u, w) \in \Omega_{u, w} \text{ and } u < w \\ 0 & \text{o.w.} \end{cases}$$

Example: Weird Dice

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$

Let W be the max value $W = \max(X, Y)$

$$\Omega_U = \Omega_W = \{1, 2, 3, 4\}$$

The joint support is:

$$\Omega_{U,W} = \{(u, w) \in \Omega_U \times \Omega_W; u \leq w\}$$

What is the joint PMF?

$$p_{U,W}(u, w) = \begin{cases} 1/16 & \text{if } (u, w) \in \Omega_{U,W} \text{ and } u = w \\ 2/16 & \text{if } (u, w) \in \Omega_{U,W} \text{ and } u < w \\ 0 & \text{otherwise} \end{cases}$$

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16

Example: Weird Dice

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$

Let W be the max value $W = \max(X, Y)$

$$\Omega_U = \Omega_W = \{1, 2, 3, 4\}$$

The joint support is:

$$\Omega_{U,W} = \{(u, w) \in \Omega_U \times \Omega_W; u \leq w\}$$

What is the joint PMF?

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16

$$p_{U,W}(u, w) = \begin{cases} 1/16 & \text{if } (u, w) \in \Omega_{U,W} \text{ and } u = w \\ 2/16 & \text{if } (u, w) \in \Omega_{U,W} \text{ and } u < w \\ 0 & \text{otherwise} \end{cases}$$

U and W are NOT independent because $\Omega_{U,W} \neq \Omega_U \times \Omega_W$

Example: Weird Dice

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$. Let W be the max value $W = \max(X, Y)$

Suppose I didn't know how to compute $p_U(u) = \mathbb{P}(U = u)$ directly.
Can we find it from the joint PMF $p_{U,W}(u, w)$?

We know $\mathbb{P}(U = 1 \cap W = w)$ for all w ...

$$\mathbb{P}(U = 1)$$

$$= \mathbb{P}(U=1 \cap W=1) + \\ \mathbb{P}(U=1 \cap W=2) + \\ \mathbb{P}(U=1 \cap W=3) + \\ \mathbb{P}(U=1 \cap W=4)$$

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16

Example: Weird Dice

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$. Let W be the max value $W = \max(X, Y)$

*Suppose I didn't know how to compute $p_U(u) = \mathbb{P}(U = u)$ directly.
Can we find it from the joint PMF $p_{U,W}(u, w)$?*

We know $\mathbb{P}(U = 1 \cap W = w)$ for all w ...

$$\begin{aligned}\mathbb{P}(U = 1) &= \mathbb{P}(U = 1 \cap W = 1) + \mathbb{P}(U = 1 \cap W = 2) \\ &\quad + \mathbb{P}(U = 1 \cap W = 3) + \mathbb{P}(U = 1 \cap W = 4) \\ &= \frac{1}{16} + \frac{2}{16} + \frac{2}{16} + \frac{2}{16} = \frac{7}{16}\end{aligned}$$

Use LoTP because the events $W=1, W=2, W=3, W=4$ partition the sample space

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16

Example: Weird Dice

$$\Omega_{\underline{u}} = \Omega_{\underline{w}} = \{1, 2, 3, 4\}$$

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$. Let W be the max value $W = \max(X, Y)$

Suppose I didn't know how to compute $p_U(u) = \mathbb{P}(U = u)$ directly.

Can we find it from the joint PMF $p_{U,W}(u, w)$?

In general...

$$\begin{aligned} \mathbb{P}(U = \underline{u}) &= \mathbb{P}(U = u \cap W = 1) + \mathbb{P}(U = u \cap W = 2) \\ &\quad + \mathbb{P}(U = u \cap W = 3) + \mathbb{P}(U = u \cap W = 4) \\ &= \sum_{w \in \Omega_W} \underline{p_{U,W}(u, w)} \end{aligned}$$

Use LoTP because the events $W=1, W=2, W=3, W=4$ partition the sample space

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16

Example: Weird Dice

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$. Let W be the max value $W = \max(X, Y)$

*Suppose I didn't know how to compute $p_U(u) = \mathbb{P}(U = u)$ directly.
Can we find it from the joint PMF $p_{U,W}(u, w)$?*

In general...

$$\mathbb{P}(U = u) = \sum_{w \in \Omega_W} p_{U,W}(u, w)$$

Use LoTP because the events $W=1, W=2, W=3, W=4$ partition the sample space

$p_U(u)$ is the "marginal" PMF for U
(because we "marginalized" ~~W~~)

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16

Marginal Distribution

If you have the joint PMF for two random variables, you can find the PMF of one of them by **using the law of total probability, partitioning on the values of the other random variable:**

$$p_X(a) = \sum_{b \in \Omega_Y} p_{X,Y}(a, b)$$

^ it's the same PMF we've talked about before, the "marginal" is just there to indicate it was derived from a joint PMF

Example: Weird Dice

Roll 2 fair 4-sided dice independently

Let U be the min value $U = \min(X, Y)$. Let W be the max value $W = \max(X, Y)$

*Suppose I didn't know how to compute $p_U(u) = \mathbb{P}(U = u)$ directly.
Can we find it from the joint PMF $p_{U,W}(u, w)$?*


$$p_U(u) = \begin{cases} \frac{7}{16} & \text{if } u = 1 \\ \frac{5}{16} & \text{if } u = 2 \\ \frac{3}{16} & \text{if } u = 3 \\ \frac{1}{16} & \text{if } u = 4 \\ 0 & \text{otherwise} \end{cases}$$

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16

Joint Expectation

Joint Expectations

For a function $g(X, Y)$, the expectation can be written in terms of the joint PMF.

$$\mathbb{E}[g(X, Y)] = \sum_{(a, b) \in \Omega_{X, Y}} g(a, b) \cdot \underline{p_{X, Y}(a, b)}$$

Handwritten annotations: Red arrows point from X and Y in $g(X, Y)$ to a and b in (a, b) . A red arrow points from $(a, b) \in \Omega_{X, Y}$ to the word "lotus" written below it.

Same ideas as before!

Examples of joint functions:

$$g(X, Y) = \underline{X + Y}, g(X, Y) = XY, g(X, Y) = X^Y, g(X, Y) = \binom{X}{Y}, \text{ etc.}$$

Expectation of a function of two RVs

What's $\mathbb{E}[UW]$ for U, W from the last slide?

Expectation of a function of two RVs

What's $\mathbb{E}[UW]$ for U, W from the last slide?

$$\{1, 2, 3, 4\} = \Omega_U = \Omega_W$$

$$\sum_{u \in \Omega_U} \sum_{w \in \Omega_W} \underline{uw} \cdot \underline{p_{U,W}(u, w)}$$

$$= \underline{1 \cdot 1} \cdot \frac{1}{16} + \underline{1 \cdot 2} \cdot \frac{2}{16} + 1 \cdot 3 \cdot \frac{2}{16} + 1 \cdot 4 \cdot \frac{2}{16} + 2 \cdot 2 \cdot \frac{1}{16} + 2 \cdot 3 \cdot \frac{2}{16} + 2 \cdot 4 \cdot \frac{2}{16} + \underline{3 \cdot 3} \cdot \frac{1}{16} + \underline{3 \cdot 4} \cdot \frac{2}{16} + \underline{4 \cdot 4} \cdot \frac{1}{16}$$

$$= \frac{100}{16} = \frac{25}{4} = \underline{6.25}$$

Conditional Expectation

side
tangent

→ Waaaaaay back when, we said conditioning on an event creates a new probability space, with all the laws holding.

So we can define things like “conditional expectations” which is the expectation of a random variable in that new probability space.

$$\rightarrow \mathbb{E}[X|E] = \sum_{x \in \Omega_X} x \cdot \mathbb{P}(X = x|E)$$

$$\mathbb{E}[X|Y = y] = \sum_{x \in \Omega_X} x \cdot \mathbb{P}(X = x|Y = y)$$

Conditional Expectations

All your favorite theorems are still true.

For example, linearity of expectation still holds

$$\mathbb{E}[(aX + bY + c) | E] = a\mathbb{E}[X|E] + b\mathbb{E}[Y|E] + c$$

↑
condition on E

Law of Total Expectation

Let $A_1, A_2, \dots, A_k$ be a partition of the sample space, then

$$\boxed{\mathbb{E}[X]} = \sum_{i=1}^n \underbrace{\mathbb{E}[X|A_i]}_{\text{conditional exp.}} \underbrace{\mathbb{P}(A_i)}_{\text{total exp.}}$$

total
exp.

Let X, Y be discrete random variables, then

$$\mathbb{E}[X] = \sum_{y \in \Omega_Y} \mathbb{E}[X|Y = y] \mathbb{P}(Y = y)$$

Similar in form to law of total probability, and the proof goes that way as well.

LTE

You will flip 2 (independent, fair coins). Call the number of heads X . Then (independently of the coin flips) draw an exponential random variable Y from the distribution $\text{Exp}(X + 1)$.

What is $\mathbb{E}[Y]$?

LTE

$$\downarrow \Omega_X = \{0, 1, 2\}$$

You will flip 2 (independent, fair coins). Call the number of heads X . Then (independently of the coin flips) draw an exponential random variable Y from the distribution $\text{Exp}(X + 1)$.

What is $\mathbb{E}[Y]$?

$$\text{Expecd. of } X \sim \text{Exp}(\lambda) = \frac{1}{\lambda}$$

$$\mathbb{E}[Y] =$$

$$= \mathbb{E}[Y|X=0]\mathbb{P}(X=0) + \mathbb{E}[Y|X=1]\mathbb{P}(X=1) + \mathbb{E}[Y|X=2]\mathbb{P}(X=2)$$

$$= \mathbb{E}[Y|X=0] \cdot \frac{1}{4} + \mathbb{E}[Y|X=1] \cdot \frac{1}{2} + \mathbb{E}[Y|X=2] \cdot \frac{1}{4}$$

$$= \frac{1}{0+1} \cdot \frac{1}{4} + \frac{1}{1+1} \cdot \frac{1}{2} + \frac{1}{2+1} \cdot \frac{1}{4} = \frac{7}{12}$$

$$\rightarrow Y|X=0 \sim \text{Exp}(0+1)$$

$$Y|X=1 \sim \text{Exp}(1+1)$$

LTE formula

Different dice

Roll two fair dice independently.
Let U be the minimum of the two rolls and W be the maximum

conditional pmf

What is $\mathbb{P}(U = 2 | W = 3)$?

$$= \frac{\mathbb{P}(U=2 \cap W=3)}{\mathbb{P}(W=3)} = \frac{2/16}{5/16} = \frac{2}{5}$$

$$p_{U|W}(2|3) = \frac{2}{5}$$

$$p_{U|W}(2|3) = P(U=2 | W=3)$$

marginal pmf

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16

Different dice

Find these values

$$\underline{p_{W|U}(2|1)} = \frac{P(W=2 \cap U=1)}{P(U=1)}$$

$$\underline{p_{U|W}(1|2)} =$$

$$\underline{p_{U|W}(4|1)} =$$

$$\frac{P(U=4 \cap W=1)}{P(W=1)} = 0$$

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16

$$\frac{2/16}{7/16}$$

1-2 min

$$\frac{2/16}{3/16}$$

Different dice

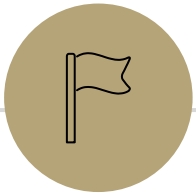
Find these values

$$p_{W|U}(2|1) = \frac{p_{W,U}(2,1)}{p_U(1)} = \frac{2/16}{7/16} = \frac{2}{7}$$

$$p_{U|W}(1|2) = \frac{p_{U,W}(1,2)}{p_W(2)} = \frac{2/16}{3/16} = \frac{2}{3}$$

$$p_{U|W}(4|1) = \frac{p_{U,W}(4,1)}{p_W(1)} = \frac{0}{1/16} = 0$$

$p_{U,W}$	$W=1$	$W=2$	$W=3$	$W=4$
$U=1$	1/16	2/16	2/16	2/16
$U=2$	0	1/16	2/16	2/16
$U=3$	0	0	1/16	2/16
$U=4$	0	0	0	1/16



Joint Distributions

For continuous joint distributions, everything will look very similar but we *replace summations with integrals*, and use a *density function instead of the PMF*

pdf

We have two continuous random variables X and Y
(that may or may not be independent)

Joint **Support/Range** - $\Omega_{X,Y}$

$$\Omega_{X,Y} = \{(a, b) : f_{X,Y}(a, b) > 0\} \subseteq \Omega_X \times \Omega_Y$$

Joint **PDF** - $f_{X,Y}(a, b)$

$f_{X,Y}(a, b)$ defined for all $(a, b) \in \mathbb{R} \times \mathbb{R}$

Joint **CDF** - $F_{X,Y}(a, b)$

$F_{X,Y}(a, b) = \mathbb{P}(X \leq a, Y \leq b)$
defined for all $(a, b) \in \mathbb{R} \times \mathbb{R}$

Normalization Property:

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx dy = 1$$

→ Joint **Independence** ✓

> $f_{X,Y}(a, b) = f_X(a) \cdot f_Y(b)$ for all $(a, b) \in \Omega_{X,Y}$

> $\Omega_{X,Y} = \Omega_X \times \Omega_Y$

Joint **Expectation**

→ $\mathbb{E}[g(X, Y)] =$
 $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X,Y}(x, y) dx dy$

Marginal **PDF** - $f_X(x), f_Y(y)$

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$$
$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx$$

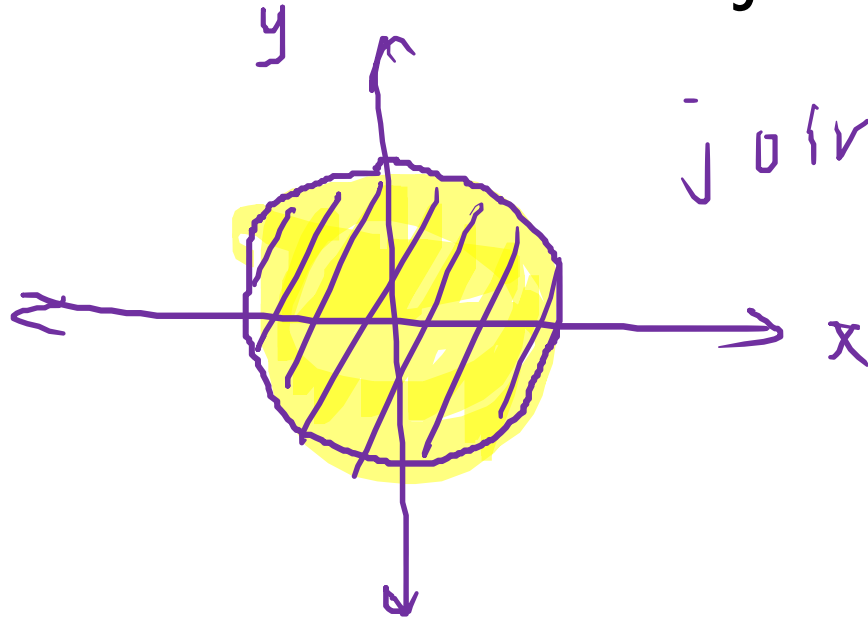
Notice we're integrating (summing) over what the other RV can be

Example: Darts

We throw a dart uniformly at random onto a circle of radius r centered around the origin. X and Y are the x and y coordinates of the point the dart lands at.

origin

Let's find and sketch the joint range $\Omega_{X,Y}$



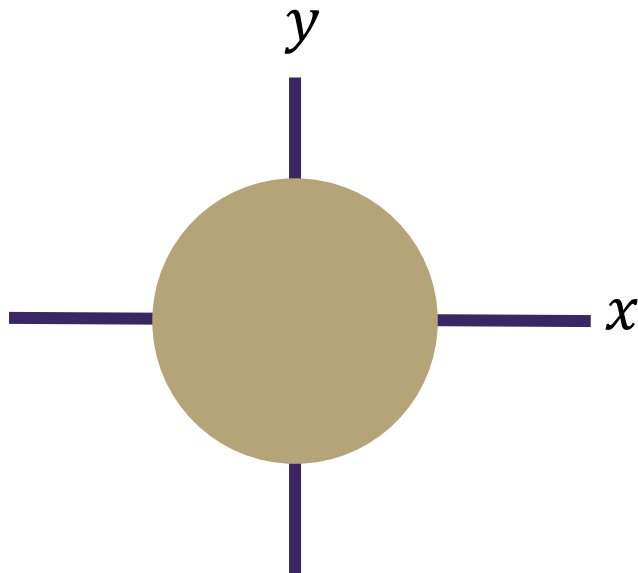
joint range

$$\Omega_{X,Y} = \left\{ (x,y) \in \mathbb{R}^2 ; x^2 + y^2 \leq r^2 \right\}$$

Example: Darts

We throw a dart uniformly at random onto a circle of radius r centered around the version. X and Y are the x and y coordinates of the point the dart lands at.

Let's find and sketch the joint range $\Omega_{X,Y}$



$$\Omega_{X,Y} = \{(x, y) \in \mathbb{R} \times \mathbb{R}; x^2 + y^2 \leq r^2\}$$

Example: Darts

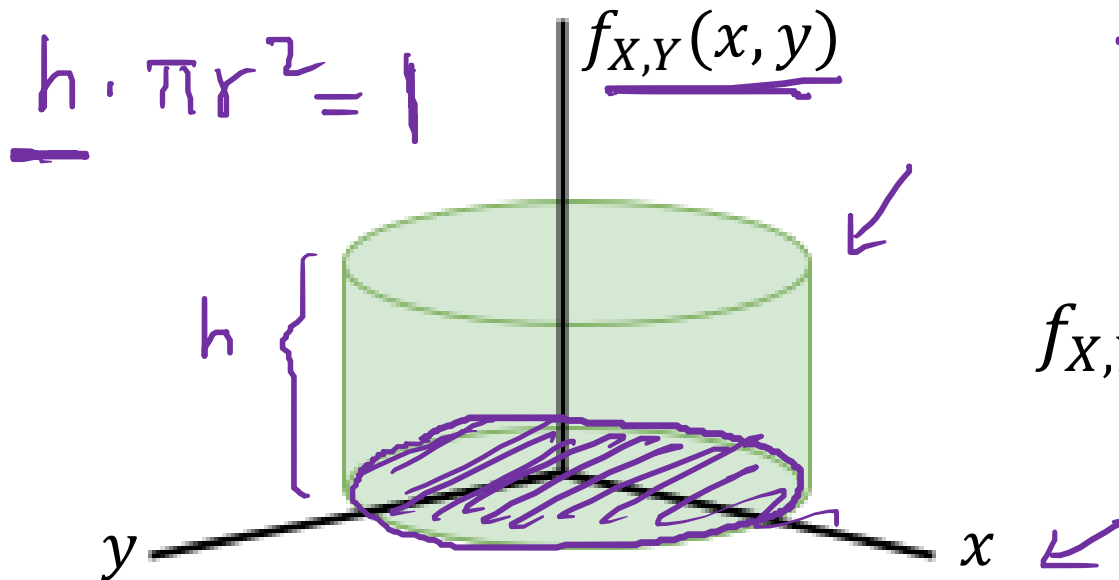
joint PDF

We throw a dart uniformly at random onto a circle of radius r centered around the version. X and Y are the x and y coordinates of the point the dart lands at.

Let's find and sketch the **joint PDF** $f_{X,Y}(x, y)$

normalization

integrate pdf over
support = 1



$$f_{X,Y}(x, y) = \begin{cases} \frac{1}{\pi r^2} \\ 0 \end{cases}$$

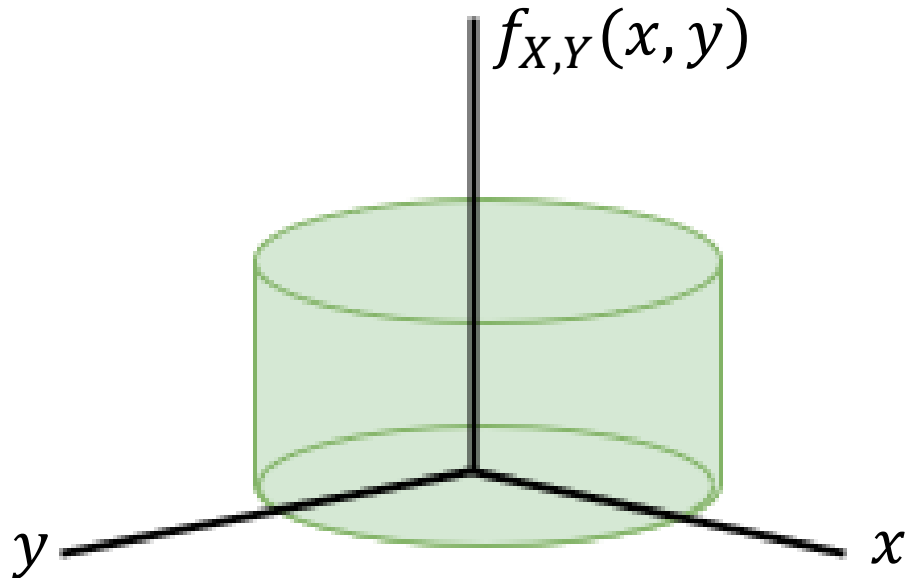
$$x^2 + y^2 \leq r^2$$

$$0 \leq x, y$$

Example: Darts

We throw a dart uniformly at random onto a circle of radius r centered around the version. X and Y are the x and y coordinates of the point the dart lands at.

Let's find and sketch the **joint PDF** $f_{X,Y}(x, y)$



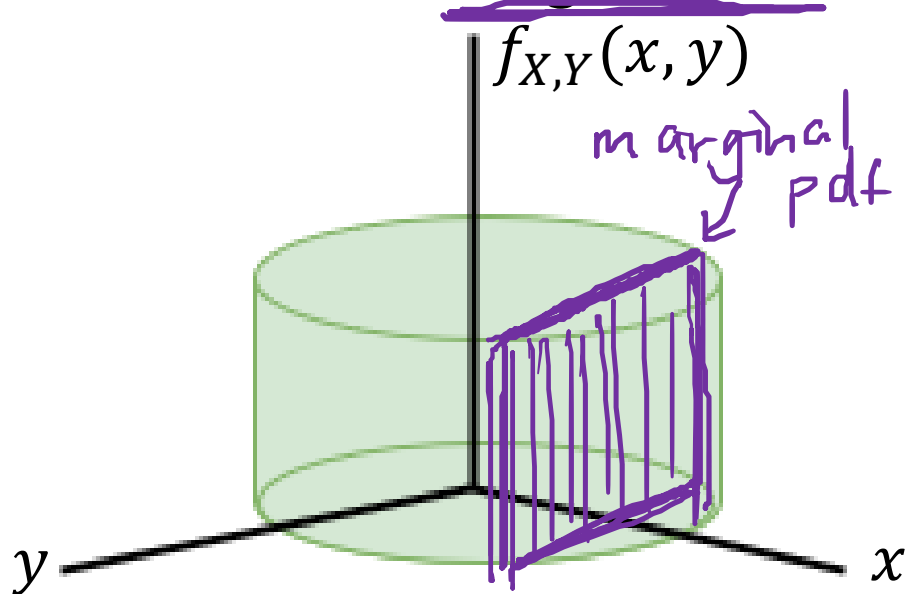
$$f_{X,Y}(x, y) = \begin{cases} \frac{1}{\pi r^2} & x^2 + y^2 \leq r^2 \\ 0 & \text{otherwise} \end{cases}$$

Example: Darts

We throw a dart uniformly at random onto a circle of radius r centered around the version. X and Y are the x and y coordinates of the point the dart lands at.

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{\pi r^2} & x^2 + y^2 \leq r^2 \\ 0 & \text{otherwise} \end{cases}$$

What is the marginal PDF for X and Y ?



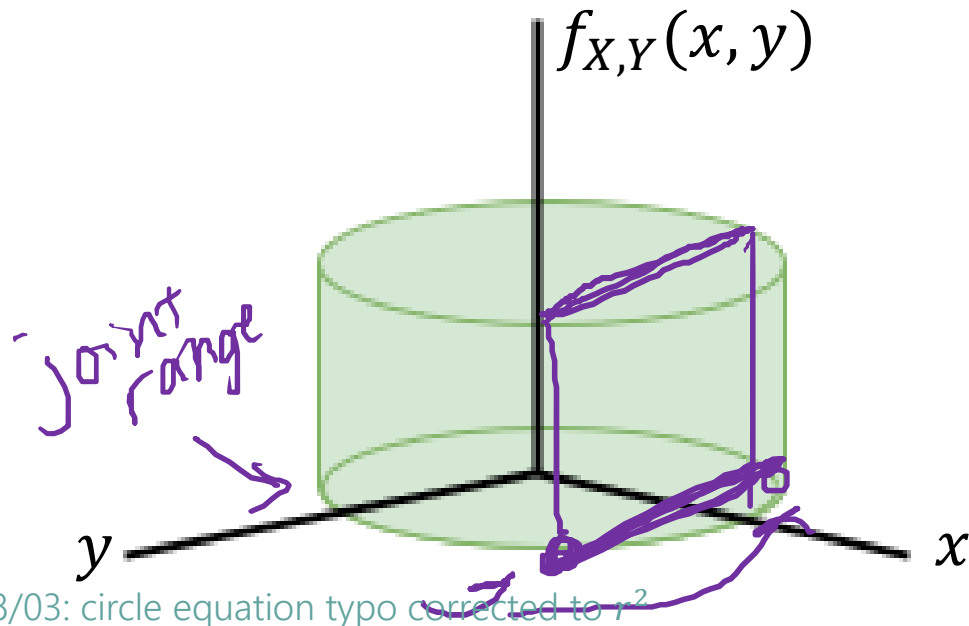
$$\underline{f_X(a)} = \int_{-\infty}^{\infty} f_{X,Y}(x,y) \underline{dy}$$
$$\underline{f_Y(b)} = \int_{-\infty}^{\infty} f_{X,Y}(x,y) \underline{dx}$$

Example: Darts

We throw a dart uniformly at random onto a circle of radius r centered around the version. X and Y are the x and y coordinates of the point the dart lands at.

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{\pi r^2} & x^2 + y^2 \leq r^2 \\ 0 & \text{otherwise} \end{cases}$$

What is the marginal PDF for X and Y ?



$$f_X(a) = \int_{-\infty}^{\infty} f_{X,Y}(a,y) dy$$

$$f_Y(b) = \int_{-\infty}^{\infty} f_{X,Y}(x,b) dx$$

Example: Darts

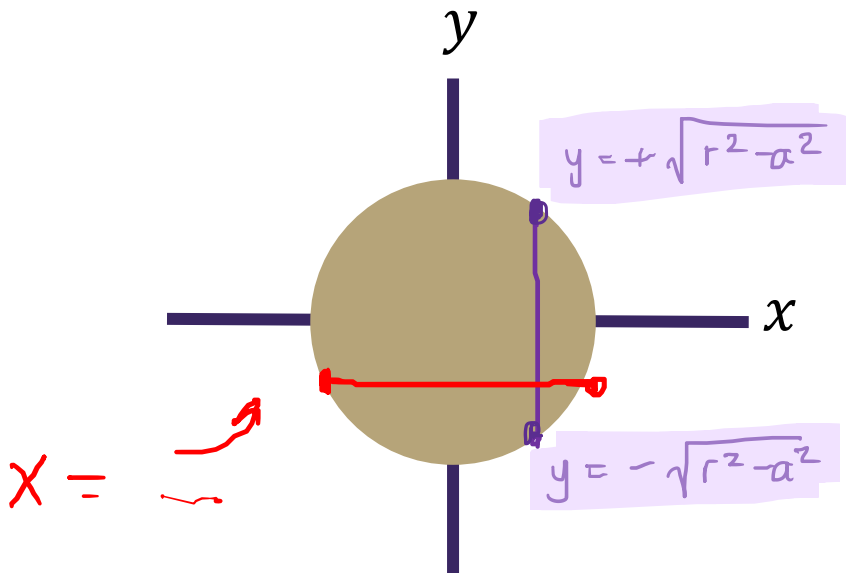
We throw a dart uniformly at random onto a circle of radius r centered around the version. X and Y are the x and y coordinates of the point the dart lands at.

fix $x=a$

$$\begin{aligned} x^2 + y^2 &= r^2 \\ a^2 + y^2 &= r^2 \end{aligned} \rightarrow y^2 = r^2 - a^2$$

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{\pi r^2} & x^2 + y^2 \leq r^2 \\ 0 & \text{otherwise} \end{cases}$$

What is the marginal PDF for X and Y ?



$$f_X(a) = \int_{-\infty}^{\infty} f_{X,Y}(a,y) dy = \int_{-\sqrt{r^2-a^2}}^{\sqrt{r^2-a^2}} \frac{1}{\pi r^2} dy$$

$$f_Y(b) = \int_{-\infty}^{\infty} f_{X,Y}(x,b) dx = \int_{-\sqrt{r^2-b^2}}^{\sqrt{r^2-b^2}} \frac{1}{\pi r^2} dx$$

Example: Darts

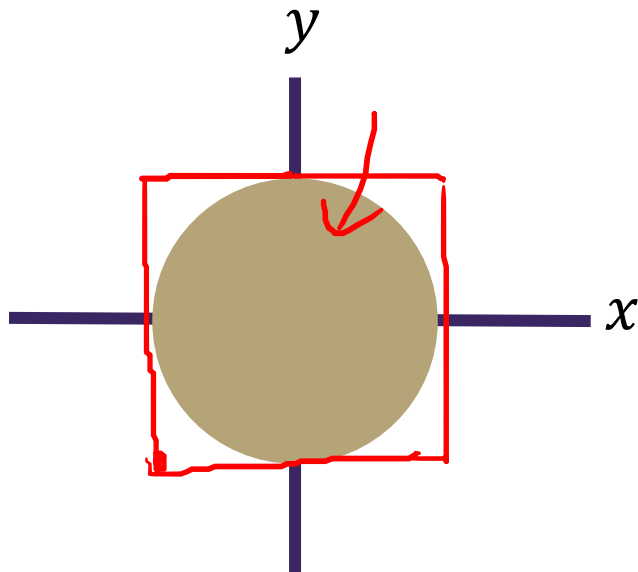
$$\mathcal{D}_X = [-1, 1]$$
$$\mathcal{D}_Y = [-1, 1]$$

We throw a dart uniformly at random onto a circle of radius r centered around the version. X and Y are the x and y coordinates of the point the dart lands at.

$$X = \underline{\hspace{2cm}}$$

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{\pi r^2} & x^2 + y^2 \leq r^2 \\ 0 & \text{otherwise} \end{cases}$$

What is the marginal PDF for X and Y ?



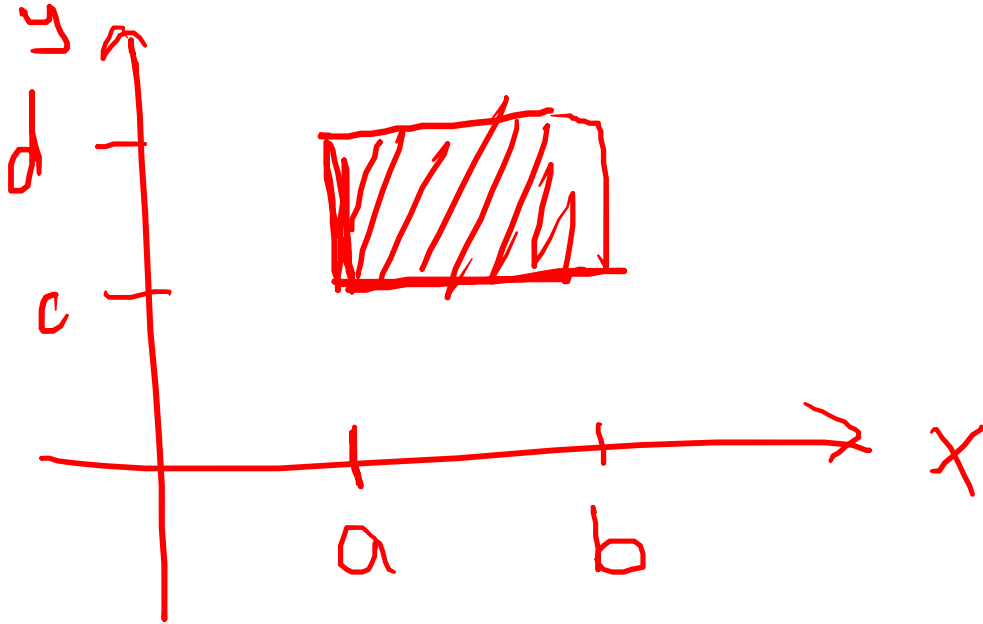
$$f_X(a) = \int_{-\infty}^{\infty} f_{X,Y}(a,y) dy = \int_{-\sqrt{r^2 - a^2}}^{\sqrt{r^2 - a^2}} \frac{1}{\pi r^2} dy = \frac{2\sqrt{r^2 - a^2}}{\pi r^2}$$

$$f_Y(b) = \int_{-\infty}^{\infty} f_{X,Y}(x,b) dx = \int_{-\sqrt{r^2 - b^2}}^{\sqrt{r^2 - b^2}} \frac{1}{\pi r^2} dx = \frac{2\sqrt{r^2 - b^2}}{\pi r^2}$$

Joint Continuous Probabilities

Just like we've done with PDFs for single random variables...

$$\mathbb{P}(a \leq X \leq b \cap c \leq Y \leq d) = \int_a^b \int_c^d f_{X,Y}(x,y) dy dx$$



Analogue for continuous

Everything we saw today has a continuous version.

There are “no surprises”—replace pmf with pdf and sums with integrals.



	Discrete	Continuous
Joint PMF/PDF	$p_{X,Y}(x, y) = P(X = x, Y = y)$	$f_{X,Y}(x, y) \neq P(X = x, Y = y)$
Joint CDF	$F_{X,Y}(x, y) = \sum_{t \leq x} \sum_{s \leq y} p_{X,Y}(t, s)$	$F_{X,Y}(x, y) = \int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(t, s) ds dt$
Normalization	$\sum_x \sum_y p_{X,Y}(x, y) = 1$	$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx dy = 1$
Marginal PMF/PDF	$p_X(x) = \sum_y p_{X,Y}(x, y)$	$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$
Expectation	$E[g(X, Y)] = \sum_x \sum_y g(x, y) p_{X,Y}(x, y)$	$E[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X,Y}(x, y) dx dy$
Conditional PMF/PDF	$p_{X Y}(x y) = \frac{p_{X,Y}(x, y)}{p_Y(y)}$	$f_{X Y}(x y) = \frac{f_{X,Y}(x, y)}{f_Y(y)}$
Conditional Expectation	$E[X Y = y] = \sum_x x p_{X Y}(x y)$	$E[X Y = y] = \int_{-\infty}^{\infty} x f_{X Y}(x y) dx$
Independence	$\forall x, y, p_{X,Y}(x, y) = p_X(x) p_Y(y)$	$\forall x, y, f_{X,Y}(x, y) = f_X(x) f_Y(y)$

A note on independence

The definition of independence says X, Y independent if and only if $p_{X,Y}(x, y) = p_X(x)p_Y(y)$ or $f_{X,Y}(x, y) = f_X(x)f_Y(y)$ (as appropriate)

There's often a nice shortcut. If X, Y are independent then joint support of X, Y (denoted $\Omega_{X,Y}$) must be $\Omega_X \times \Omega_Y$.

Joint support is $\{(x, y): p_{X,Y}(x, y) > 0\}$.

Often easier to verify dependence when those are different (especially in the continuous case).

But note this is a single implication not an if-and-only-if.

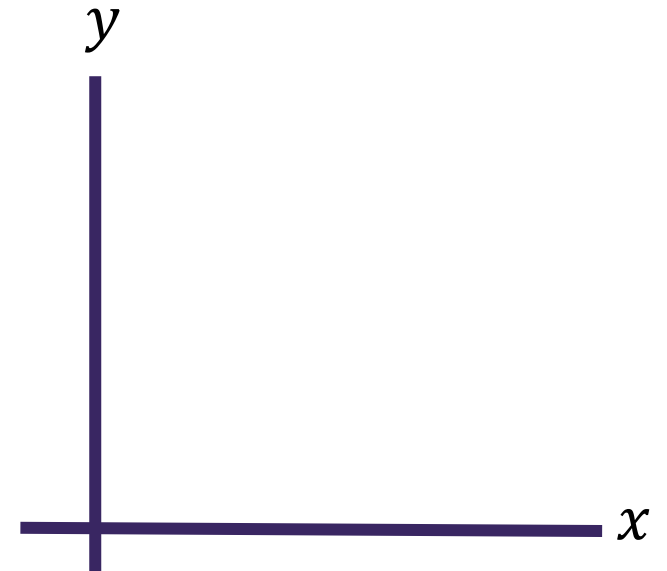
Example: Continuous Servers

The time until server 1 crashes is $X \sim \text{Exp}(u)$, and the time until server 2 crashes is $Y \sim \text{Exp}(v)$. Both servers are independent of each other.

What is the probability server 1 crashes before server 2?

$$\mathbb{P}(X < Y) =$$

pollev.com/jolie312

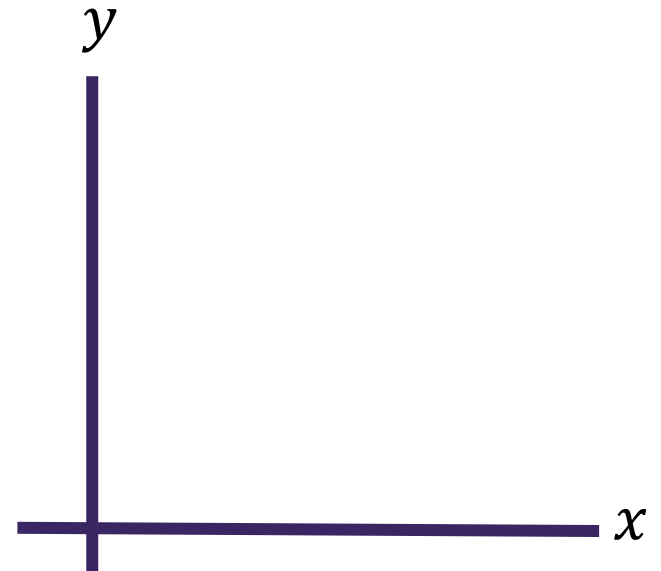


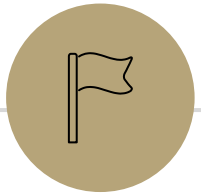
Example: Continuous Servers

The time until server 1 crashes is $X \sim \text{Exp}(u)$, and the time until server 2 crashes is $Y \sim \text{Exp}(v)$. Both servers are independent of each other.

What is the probability server 1 crashes before server 2?

$$\begin{aligned}\mathbb{P}(X < Y) &= \int_0^\infty \int_x^\infty f_{X,Y}(x,y) \, dy \, dx \\ &= \int_0^\infty \int_x^\infty f_X(x)f_Y(y) \, dy \, dx \text{ by independence}\end{aligned}$$





Covariance

Covariance

We sometimes want to measure how “intertwined” X and Y are – how much knowing about one of them will affect the other.

If X turns out “big” how likely is it that Y will be “big”, how much do they “vary together”

Covariance

$$\text{Cov}(X, Y) = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

Covariance

$$\text{Cov}(X, Y) = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

If X, Y go in the same direction

If X, Y go in the opposite directions

Covariance

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$$

That's consistent with our previous knowledge for independent variables. (for X, Y independent, $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$).

Covariance

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$$

You and your friend are playing a game, you flip a coin: if heads you pay your friend a dollar, if tails they pay you a dollar. Let X be your profit and Y be your friend's profit.

What is $\text{Var}(X + Y)$?

Before you calculate, make a prediction. What should it be?

Covariance

You and your friend are playing a game, you flip a coin: if heads you pay your friend a dollar, if tails they pay you a dollar. Let X be your profit and Y be your friend's profit.

What is $\text{Var}(X + Y)$?

$$\text{Var}(X) = \text{Var}(Y) = \mathbb{E}[X^2] - (\mathbb{E}[X])^2 = 1 - 0^2 = 1$$

$$\text{Cov}(X, Y) = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

$$\mathbb{E}[XY] = \frac{1}{2} \cdot (-1 \cdot 1) + \frac{1}{2} (1 \cdot -1) = -1$$

$$\text{Cov}(X, Y) = -1 - 0 \cdot 0 = -1.$$

$$\text{Var}(X + Y) = 1 + 1 + 2 \cdot -1 = 0$$

Covariance, Another example

Let X be a Bernoulli RV with probability p of success.

Let $Y = X$ (Y is X , not an iid copy, literally the same experiment)

Let $Z = -X$

Let W be an independent Bernoulli, identically distributed to X

Find

$\text{Cov}(X, Y)$, $\text{Cov}(X, Z)$, $\text{Cov}(X, W)$

Covariance, Another example

Let X be a Bernoulli RV with probability p of success.

Let $Y = X$ (Y is X , not an iid copy, literally the same experiment)

Let $Z = -X$

Let W be an independent Bernoulli, identically distributed to X

$$\text{Cov}(X, Y) = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

$$= (1 \cdot 1 \cdot p + 0 \cdot 0 \cdot [1 - p]) - p \cdot p$$

$$= p - p^2 = p(1 - p)$$

Hey, that's the variance of X . This is a pattern: $\text{Cov}(X, X) = \text{Var}(X)$

Covariance, Another example

Let X be a Bernoulli RV with probability p of success.

Let $Y = X$ (Y is X , not an iid copy, literally the same experiment)

Let $Z = -X$

Let W be an independent Bernoulli, identically distributed to X

$$\text{Cov}(X, Z) = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

$$= (1 \cdot -1 \cdot p + 0 \cdot -0 \cdot [1 - p]) - (p \cdot [-p])$$

$$= -p - [-p^2] = -p(1 - p)$$

General pattern: $\text{Cov}(X, -Y) = -\text{Cov}(X, Y)$

Covariance, Another example

Let X be a Bernoulli RV with probability p of success.

Let $Y = X$ (Y is X , not an iid copy, literally the same experiment)

Let $Z = -X$

Let W be an independent Bernoulli, identically distributed to X

$$\text{Cov}(X, W) = \mathbb{E}[XW] - \mathbb{E}[X]\mathbb{E}[W]$$

$$= (1 \cdot 1 \cdot p^2 + 1 \cdot 0 \cdot p[1 - p] + 0 \cdot 1 \cdot [1 - p]p + 0 \cdot 0 \cdot [1 - p]^2) - (p \cdot [p])$$

$$= (p^2) - p^2 = 0$$

General pattern: if X, Y independent $\text{Cov}(X, Y) = 0$

A Few Notes

Covariance is an un-normalized number.

It measures both how intertwined X, Y are and in some sense how much X, Y vary in the first place (if you multiply both X, Y by 2, the strength of the relationship intuitively is the same, but covariance increases).

If you want just the strength of the relationship, you probably want the "correlation coefficient": $\frac{\text{Cov}(X,Y)}{\sqrt{\text{Var}(X)\text{Var}(Y)}}$ always between -1 and 1 .

Covariance directly measures only "linear" relationships; if Y depends on X^2 , the covariance might not be as high as you expect.

If dealing with real data, look at a plot to see if you should be looking for a linear relationship in the first place.