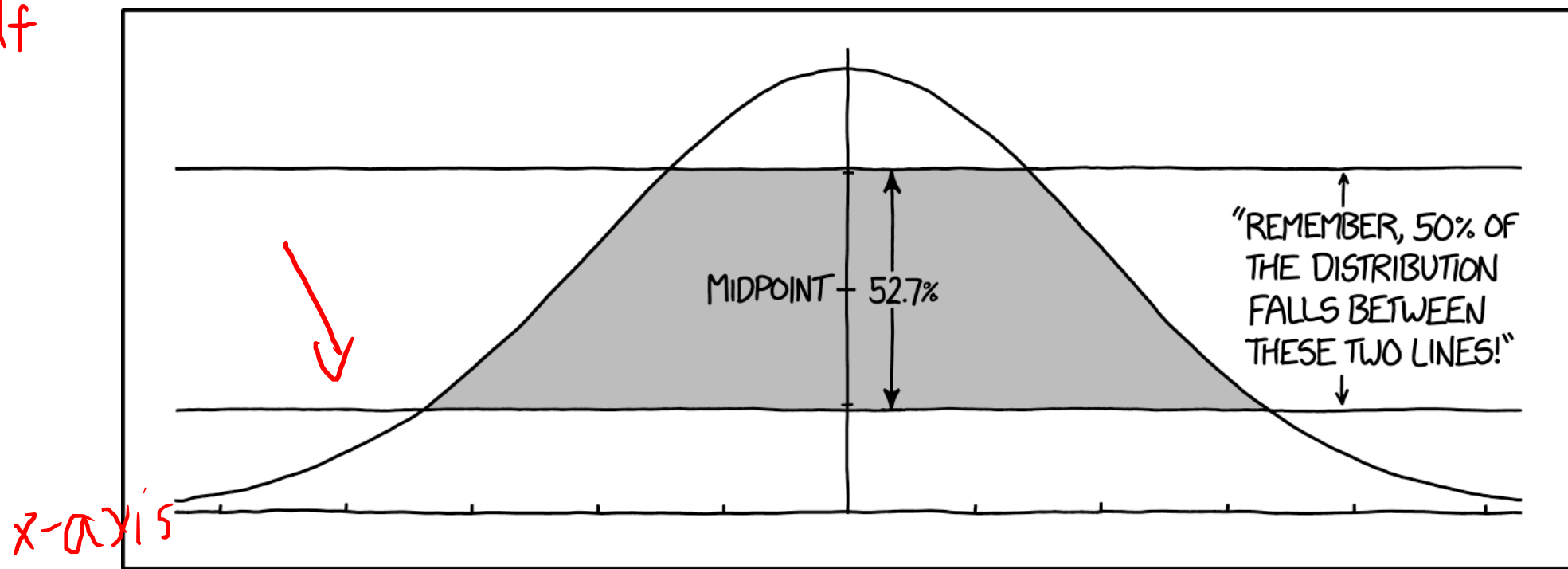


pdf



HOW TO ANNOY A STATISTICIAN

# Zoo of Continuous Random Variables

CSE 312 26Su

Lecture 14

# Where Are We?

Last time:

- Continuous Random Variables
  - • Probability Density Function
    - Cumulative Distribution Function
    - Expectation and Variance
- Comparing Discrete vs. Continuous
- Continuous Uniform Distribution

pdf  $\xrightarrow{\text{integral}}$  cdf  
 $\xleftarrow{\text{derivative}}$

Today:

- Zoo of continuous random variables
  - 
  - Exponential distribution
  - Normal distribution

# **Discrete Random Variables**

IN

The support has *finite* or countably infinite values

*e.g., number of successes, number of trials till success, attendance at a class are all discrete because they take on a set of finite or countably infinite values*

Some random experiments have uncountably-infinite sample spaces

> *How long until the next bus shows up?*

> *Throwing a dart on a board (what location does the dart land?)*

# ↪ **Continuous Random Variables**

IR

Random variables with a support of uncountably-infinite values

> *e.g., RVs that take on any **real** number in some interval(s) like distance, height, time, etc.*

# ***Discrete*** RVs

# ***Continuous*** RVs

# ***Discrete* RVs**

Support is finite/countably infinite (e.g. integers)

# ***Continuous* RVs**

Support is uncountably infinite (e.g., real numbers)

# ***Discrete*** RVs

Support is finite/countably infinite (e.g. *integers*)

# ***Continuous*** RVs

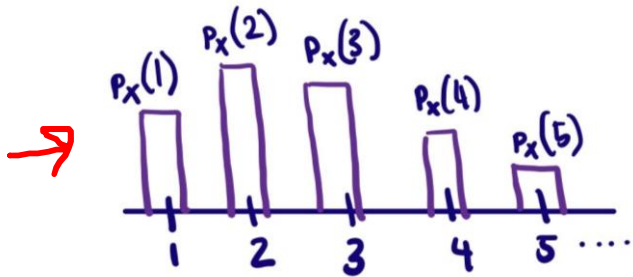
Support is uncountably infinite (e.g., *real numbers*)

$\mathbb{P}(X = k) = \frac{1}{\infty} = 0$  so we don't use PMF. instead...

# Discrete RVs

**Support** is finite/countably infinite (e.g. *integers*)

**Probability mass function**  $p_X(k)$  gives probability of each value in support



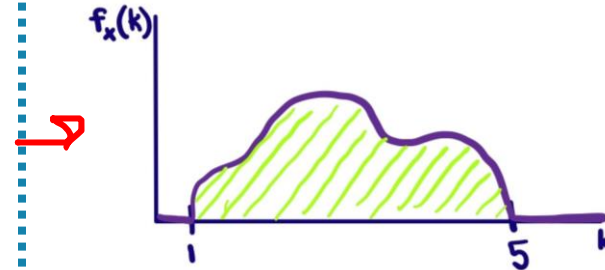
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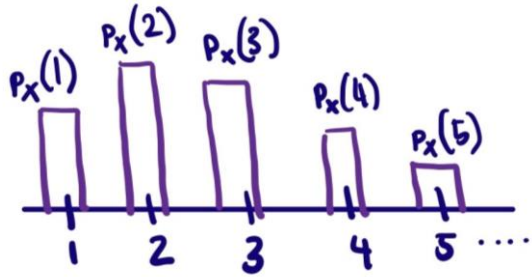
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$$\underline{f_X(k) \geq 0}$$

*Unif(0, 0.5)*

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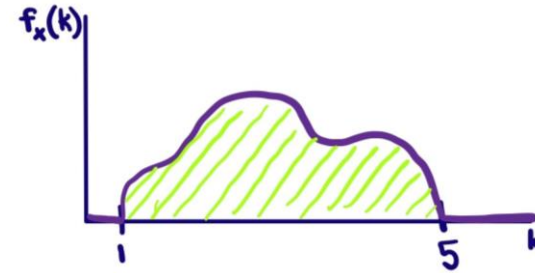
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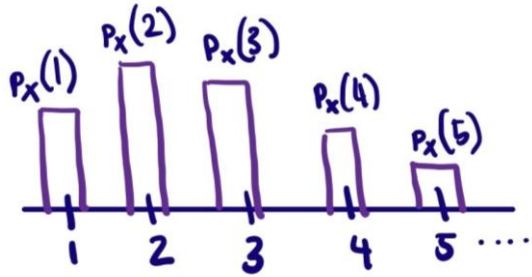
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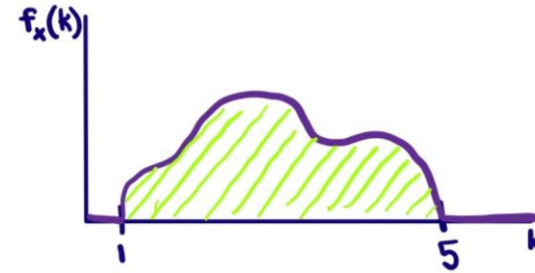
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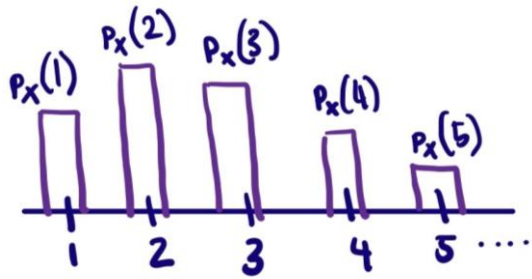
Sum up the probabilities of values  $\leq k$

Integrate over values  $\leq k$ :  $F_X(k) = \int_{-\infty}^k f_X(z) dz$

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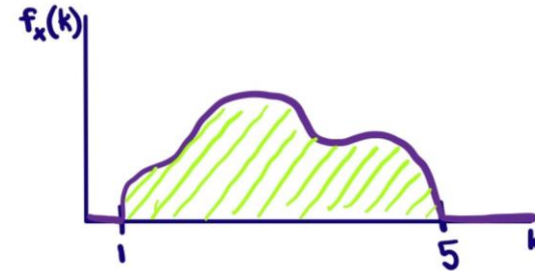
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$$\mathbb{E}[g(X)] = \sum_{k \in \Omega_X} (g(k) \cdot p_X(k))$$

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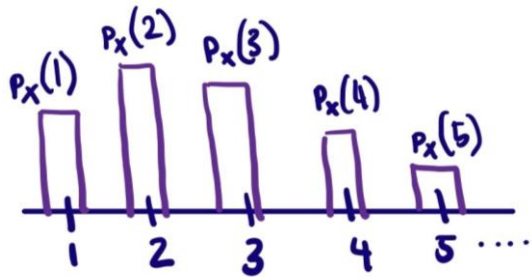
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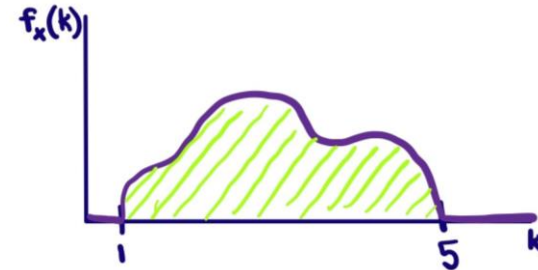
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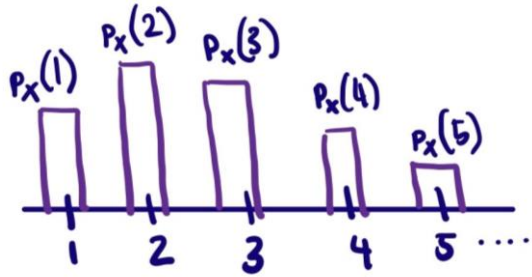
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**Variance** is  $Var(X) = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$

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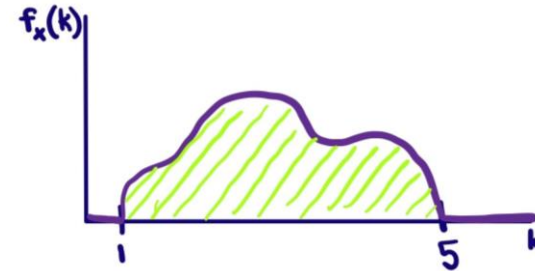
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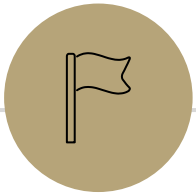
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**Variance** is  $Var(X) = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$

Linearity of expectation and properties of expectation and variance applies in both!



# **Zoo of Continuous Random variables**

and get practice with working with continuous random variables!

# Continuous Zoo

*This zoo defines common patterns for continuous random variables and gives us the PDF, CDF, expectation, and variance, so we don't have to compute it every time!*

  $X \sim \text{Unif}(a, b)$

$$f_X(k) = \frac{1}{b-a} \text{ for } a \leq k \leq b$$

$$F_X(k) = \frac{k-a}{b-a} \text{ if } a \leq k < b$$

$$\mathbb{E}[X] = \frac{a+b}{2}$$

$$\text{Var}(X) = \frac{(b-a)^2}{12}$$

  $X \sim \text{Exp}(\lambda)$

$$f_X(k) = \lambda e^{-\lambda k} \text{ for } k \geq 0$$

$$F_X(k) = 1 - e^{-\lambda k} \text{ if } k \geq 0$$

$$\mathbb{E}[X] = \frac{1}{\lambda}$$

$$\text{Var}(X) = \frac{1}{\lambda^2}$$

  $X \sim \mathcal{N}(\mu, \sigma^2)$

$$f_X(k) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

$$F_X(k) = \Phi\left(\frac{k-\mu}{\sigma}\right)$$

$$\mathbb{E}[X] = \mu$$

$$\text{Var}(X) = \sigma^2$$

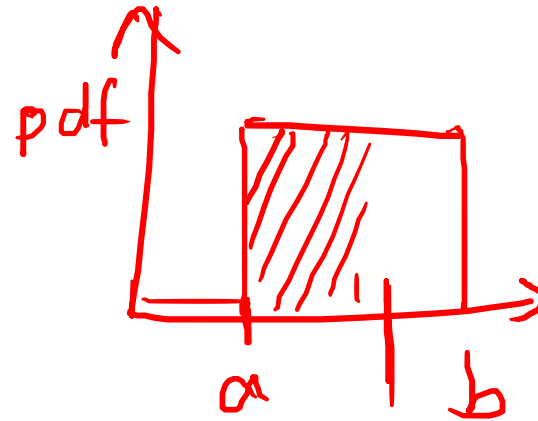
It's a smaller zoo, but it's just as much fun! :D

# Continuous Uniform Distribution

$X \sim \text{Unif}(a, b)$  (uniform real number between  $a$  and  $b$ )

$$\text{PDF: } f_X(k) = \begin{cases} \frac{1}{b-a} & \text{if } a \leq k \leq b \\ 0 & \text{otherwise} \end{cases}$$

$$\text{CDF: } F_X(k) = \begin{cases} 0 & \text{if } k < a \\ \frac{k-a}{b-a} & \text{if } a \leq k \leq b \\ 1 & \text{if } k \geq b \end{cases}$$



$$\begin{aligned} P(X \leq b) \\ = P(X < b) \end{aligned}$$

$$\text{Expectation: } \mathbb{E}[X] = \frac{a+b}{2}$$

$$\text{Variance: } \text{Var}(X) = \frac{(b-a)^2}{12}$$

diff. from discrete

# Exponential Distribution

goal  
time until event  
occurs

How much **time** till an event occurs?

e.g., seconds till thunder, time till the first customer

This sounds very similar to a geometric distribution!

- > Geometric random variable is the number of trials till success (discrete).
- > Exponential random variable is time (a real number, continuous) till success

*With the geometric distribution, we said trials must be independent ->  
"If the flip 1 is tails, the coin doesn't remember it was tails, you've made no progress"*

↪ *Here, waiting must not make the event happen any sooner ->  
"If we don't get success in the first 3.87sec, chances of seeing success doesn't change"*

This means **memorylessness**!  $\mathbb{P}(X \geq k + 1 | X \geq 1) = \mathbb{P}(Y \geq k)$

# Exponential Distribution- **CDF**

eg. hour  
day

$X \sim \text{Exp}(\lambda)$  is time till the first event. Average of  $\lambda$  events per time unit.

It would be hard to come up with the PDF directly here, so start with CDF.

We want  $F_X(t) = \mathbb{P}(X \leq t) = 1 - \mathbb{P}(X > t)$  complement

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*What distribution do we know about the also deals with time?*

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Poisson random variable gives us the number of events in a unit of time

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Poisson random variable gives us the number of events in a unit of time

*What Poisson are we waiting on, and what event for it tells you that  $X > t$ ?  
what must be true about the number of successes in a certain time interval?*

# Exponential Distribution- **CDF**

→  $X \sim \text{Exp}(\lambda)$  is time till the first event. Average of  $\lambda$  events per time unit.

It would be hard to come up with the PDF directly here, so start with CDF.

$$\text{We want } F_X(t) = \mathbb{P}(X \leq t) = 1 - \mathbb{P}(X > t)$$

*What distribution do we know about the also deals with time...Poisson!*

Poisson random variable gives us the number of events in a unit of time

*What Poisson are we waiting on, and what event for it tells you that  $X > t$ ?  
there must be 0 events in the first  $t$  time units*

$Y \sim \text{Poi}(\lambda t)$  (average  $\lambda t$  events in  $t$  time units).

$$\text{Then, } \mathbb{P}(\underline{X} > \underline{t}) = \mathbb{P}(\underline{Y} = \underline{0})$$

*Poisson is discrete  
modeling # of events  
Exp is cont.  
modeling time*

# Exponential Distribution- **CDF**

Exp and Poi  
are related

$X \sim \text{Exp}(\lambda)$  is time till the first event. Average of  $\lambda$  events per time unit.

We want  $F_X(t) = \mathbb{P}(X \leq t) = 1 - \mathbb{P}(X > t)$

Let  $Y \sim \text{Poi}(\lambda t)$  (average  $\lambda t$  events in  $t$  time units).

Then,  $\mathbb{P}(X > t) = \mathbb{P}(Y = 0)$

"its take more than  $t$  time units for first event" = "0 successes in the first  $t$  time units"

Putting it all together...

$$\begin{aligned} F_X(t) &= 1 - \mathbb{P}(X > t) \\ &= 1 - \mathbb{P}(Y = 0) \\ &= 1 - e^{-\lambda t} \frac{(\lambda t)^0}{0!} \\ &= \underline{1 - e^{-\lambda t}} \end{aligned}$$

$Y \sim \text{Poi}(\lambda t)$

$$F_X(\underline{k}) = \begin{cases} 1 - e^{-\lambda \underline{k}} & \text{if } k \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

# Exponential Distribution- **PDF**

Now we know the CDF:  $F_X(k) = \begin{cases} 1 - e^{-\lambda k} & \text{if } k \geq 0 \\ 0 & \text{otherwise} \end{cases}$

What's the PDF (probability density function)?

$$\begin{aligned} f_Y(t) &= \frac{d}{dt} (1 - e^{-\lambda t}) = 0 - \lambda \cdot (-e^{-\lambda t}) \\ &= \begin{cases} \lambda e^{-\lambda t} & \text{if } t \geq 0 \\ 0 & \text{o.w.} \end{cases} \end{aligned}$$

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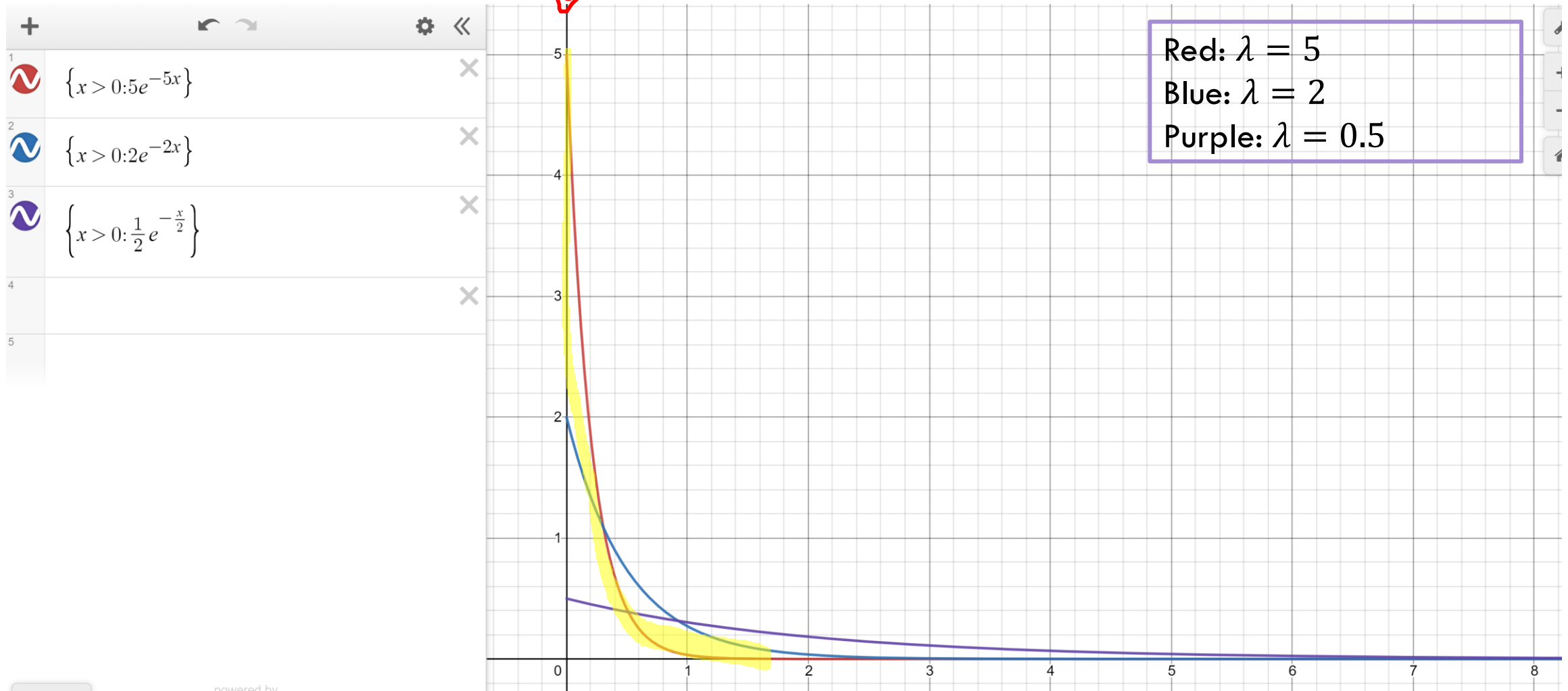
$$f_Y(t) = \frac{d}{dt} (1 - e^{-\lambda t}) = 0 - \frac{d}{dt} (e^{-\lambda t}) = \lambda e^{-\lambda t}.$$

For  $t \geq 0$  it's that expression

For  $t < 0$  it's just 0.

# Exponential Distribution- **PDF**

$$\lambda e^{-\lambda t}$$



# Memorylessness of exp.

$$\begin{aligned}\mathbb{P}(X \geq k+1 | X \geq 1) &= \frac{\mathbb{P}(X \geq k+1 \cap X \geq 1)}{\mathbb{P}(X \geq 1)} = \frac{\mathbb{P}(X \geq k+1)}{1 - (1 - e^{-\lambda \cdot 1})} \\ &= \frac{e^{-\lambda(k+1)}}{e^{-\lambda}} = e^{-\lambda k}\end{aligned}$$

CDF

What about  $\mathbb{P}(X \geq k)$  (without conditioning on the first step)?

$$1 - (1 - e^{-\lambda k}) = e^{-\lambda k}$$

$$\mathbb{P}(X \geq k+1 | X \geq 1) = \mathbb{P}(X \geq k)$$

It's the same!!!

More generally, for an exponential rv  $X$ ,  $\mathbb{P}(X \geq s + t | X \geq s) = \mathbb{P}(X \geq t)$

# Side note

cont.

$$P(X < 1) = P(X \leq 1)$$

I hid a trick in that algebra,

$$\underline{P(X \geq 1) = 1 - P(X < 1) = 1 - P(X \leq 1)}$$

The first step is the complementary law.

The second step is using that  $\underline{\int_1^1 f_X(z) dz = 0}$

$$P(X = k) = 0$$

cont. rvs  
only!

\* In general, for continuous random variables we can switch out  $\leq$  and  $<$  without anything changing.

We can't make those switches for discrete random variables.

$$\Omega_Z = \{1, 2, 3, 4\} \quad \underline{P(Z \leq 3) = P(Z=1) + P(Z=2) + P(Z=3)}$$

$$P(Z < 3) = P(Z=1) + P(Z=2)$$

# Exponential Distribution- **Expectation**

$X \sim \text{Exp}(\lambda)$  is time till the first event. Average of  $\lambda$  events per time unit.

$$\mathbb{E}[X] = \int_{-\infty}^{\infty} z \cdot f_X(z) dz$$

$$= \int_0^{\infty} z \cdot \lambda e^{-\lambda z} dz$$

$$\text{Let } u = z; dv = \lambda e^{-\lambda z} dz \text{ (} v = -e^{-\lambda z} \text{)}$$

Don't worry about the derivation  
(it's here if you're interested;  
you're not responsible for the  
derivation. Just the value.

$$\text{Integrate by parts: } -ze^{-\lambda z} - \int -e^{-\lambda z} dz = -ze^{-\lambda z} - \frac{1}{\lambda} e^{-\lambda z}$$

$$\text{Definite Integral: } -ze^{-\lambda z} - \frac{1}{\lambda} e^{-\lambda z} \Big|_{z=0}^{\infty} = \left( \lim_{z \rightarrow \infty} -ze^{-\lambda z} - \frac{1}{\lambda} e^{-\lambda z} \right) - \left( 0 - \frac{1}{\lambda} \right)$$

$$\text{By L'Hopital's Rule } \left( \lim_{z \rightarrow \infty} -\frac{z}{e^{\lambda z}} - \frac{1}{\lambda e^{\lambda z}} \right) - \left( 0 - \frac{1}{\lambda} \right) = \left( \lim_{z \rightarrow \infty} -\frac{1}{\lambda e^{\lambda z}} \right) + \frac{1}{\lambda} = \frac{1}{\lambda}$$

# Exponential Distribution- **Variance**

$X \sim \text{Exp}(\lambda)$  is time till the first event. Average of  $\lambda$  events per time unit.

$$\text{Var}(X) = E[X^2] - (E[X])^2$$

$$= \int_{-\infty}^{\infty} z^2 \cdot f_X(z) dz - \left(\frac{1}{\lambda}\right)^2$$

= ...after a bunch of calculus

$$= \frac{1}{\lambda^2}$$

Don't worry about the actual calculus here as well 😊

# Exponential Distribution

$$\underline{X \sim \text{Exp}(\lambda)}$$

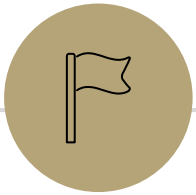
Parameter  $\lambda \geq 0$  is the average number of events in a unit of time.

$$\text{PDF: } f_X(k) = \begin{cases} \lambda e^{-\lambda k} & \text{if } k \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

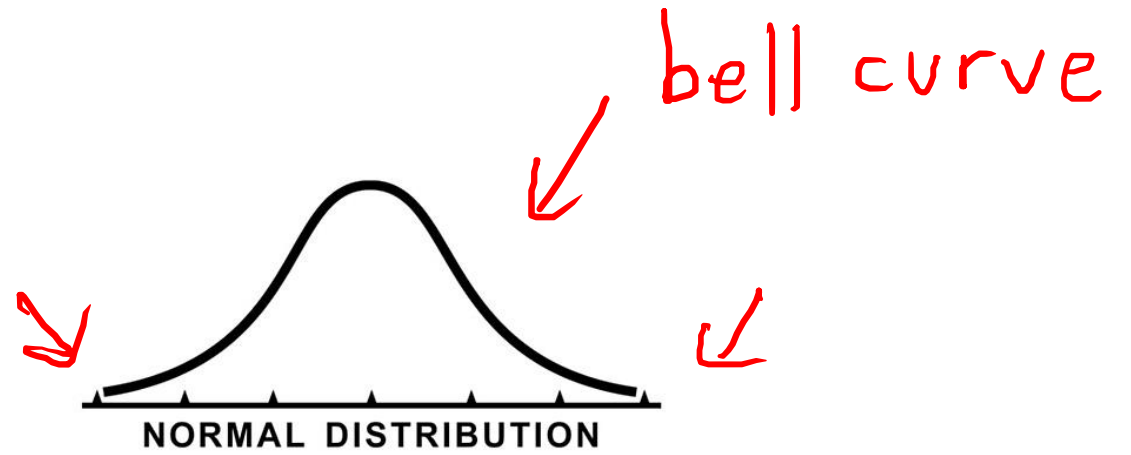
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$$\text{Expectation: } \mathbb{E}[X] = \frac{1}{\lambda} \quad \checkmark$$

$$\text{Variance: } \text{Var}(X) = \frac{1}{\lambda^2} \quad \checkmark$$



# Normal Distribution



# Normal Random Variable (AKA Gaussian)

There's not a single scenario that follows a normal distribution...  
But we're going to see that it shows up in a lot of real world situations!

A normal random variable  $X \sim \mathcal{N}(\mu, \sigma^2)$  has two parameters:

- $\mu$  =  $\mathbb{E}[X]$  is the mean
- $\sigma^2$  =  $\text{Var}(X)$  is the variance ( $\sigma = \sqrt{\text{Var}(X)}$  is *standard deviation*)

and follows this *probability density function* (a bell curve!):

$$f_X(k) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(k-\mu)^2}{2\sigma^2}}$$

# Let's take a closer look at that PDF...

$X \sim \mathcal{N}(\mu, \sigma^2)$  iff  $X$  follows the following PDF:

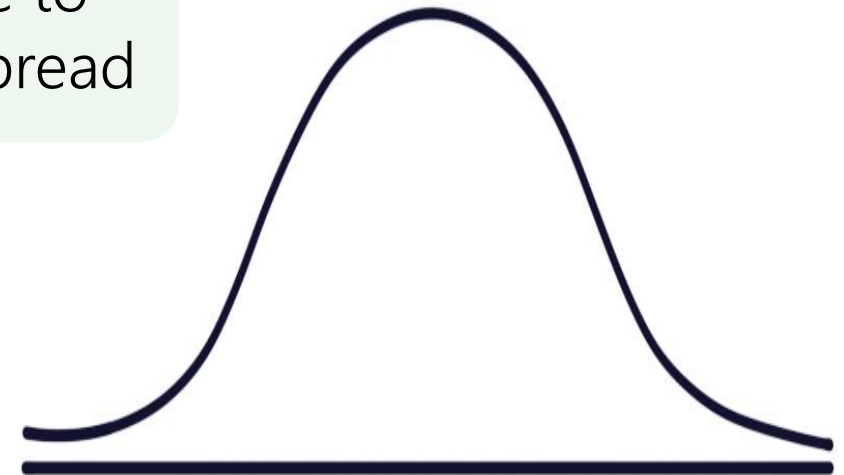
$$f_X(k) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(k-\mu)^2}{2\sigma^2}}$$

constant for normalization

symmetric around the mean  $\mu$

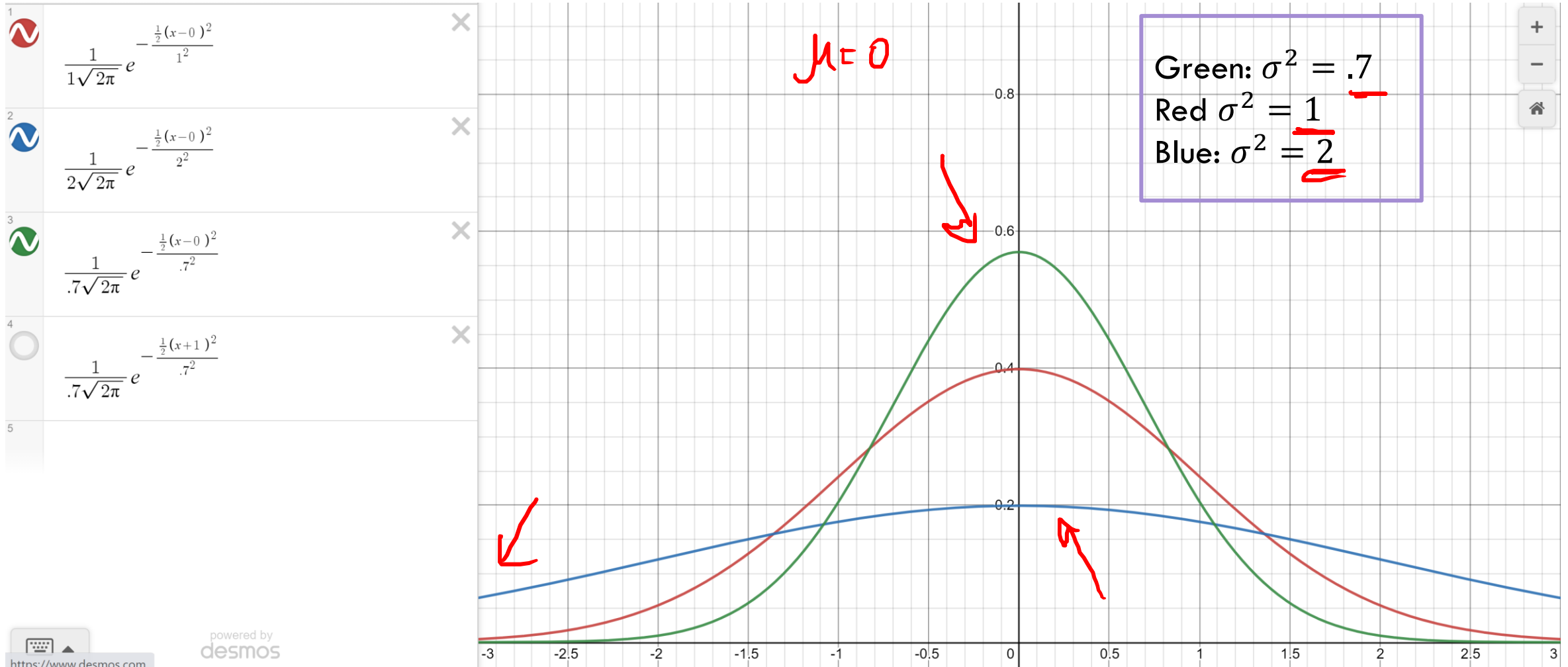
variance to control spread

exponential term for tails

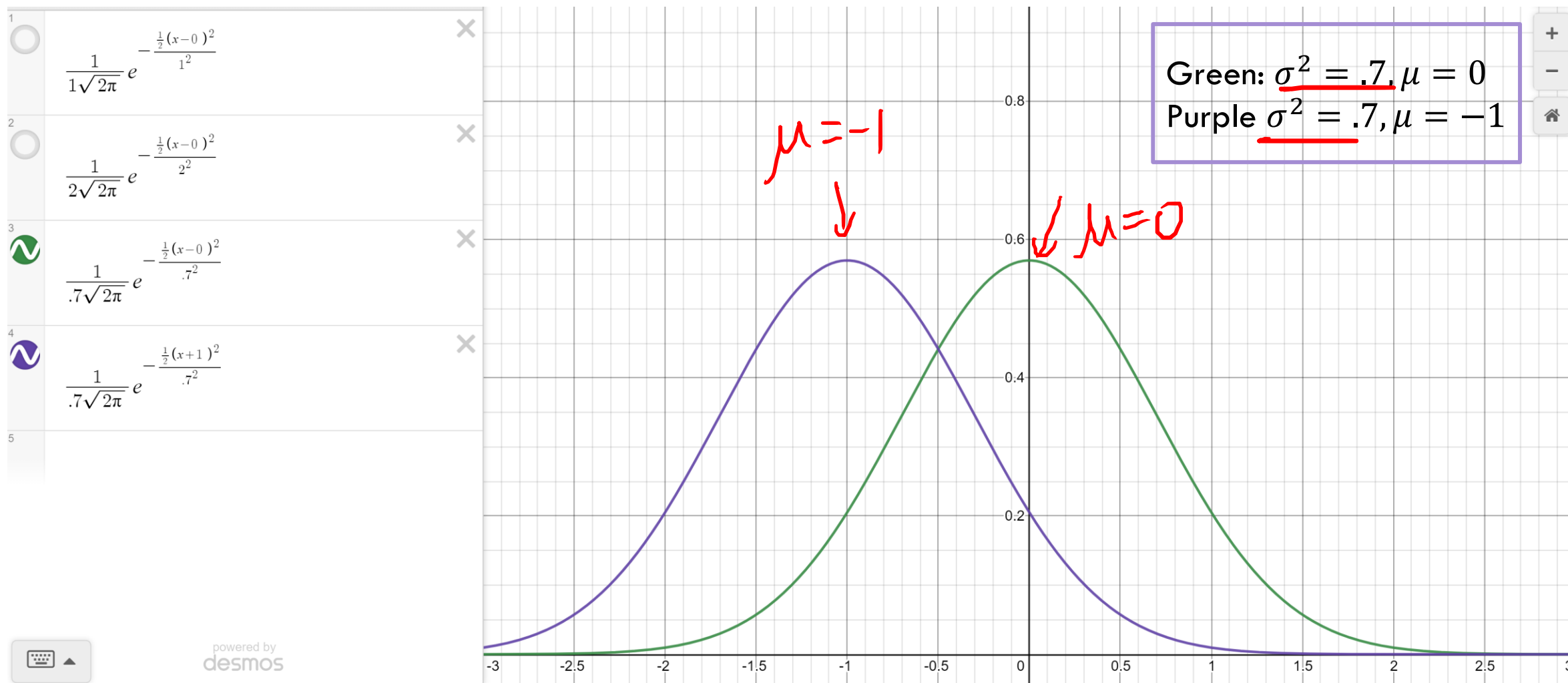


# Changing the variance

$$\frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$



# Changing the mean



# Closure of Normals Under Scale and Shift

When we scale a normal (multiplying by a constant) or shift it (adding a constant) we get a normal random variable back!

If  $X \sim \mathcal{N}(\mu, \sigma^2)$

Then for  $Y = aX + b$ ,  $Y \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$

intuitively: we are just stretching and squishing the distribution – it's symmetric without major disruptions it still follows the same general shape

Normals are unique in that you get a NORMAL back.

If you multiply a binomial by  $3/2$  you don't get a binomial (it's support isn't even integers!)

$$X \sim \text{Bin}(n, p)$$

$$Y = \frac{3}{2} \cdot X$$

$$\Omega_X = \{1, 2, 3, \dots, n\}$$

$$\Omega_Y = \frac{3}{2}, 3, \dots$$

# Closure of Normals *Under Addition*

When we add two independent normal random variables, you get another normal random variable.

If  $X \sim \mathcal{N}(\mu_X, \sigma_X^2)$  and  $Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$  and  $X$  and  $Y$  are independent,

Then, for  $Z = aX + bY + c$ ,  $Z \sim \mathcal{N}(a\mu_X + b\mu_Y + c, a^2\sigma_X^2 + b^2\sigma_Y^2)$

$$\text{var}(aX + bY + c) = \text{var}(aX) + \text{var}(bY) = a^2\text{var}(X)$$

Normals are unique in that you get a NORMAL back.  $+ b^2\text{var}(Y)$

↘  
*The sum of two dice rolls (sum of two uniform distributions) does not follow a uniform distribution*

# Ok...what about the CDF?

There is no closed form for the CDF ☹

So how *can* we find the values  $F_X(k) = \mathbb{P}(X \leq k)$ ?

And for finding the probability of  $X$  being in other ranges, we certainly don't want to bother integrating over that PDF...

# We have a table with precomputed values!

AKA the "z-table", "phi-table"

| z   | 0.00    | 0.01    | 0.02    | 0.03    | 0.04    | 0.05    | 0.06    | 0.07    | 0.08    | 0.09    |
|-----|---------|---------|---------|---------|---------|---------|---------|---------|---------|---------|
| 0.0 | 0.5     | 0.50399 | 0.50798 | 0.51197 | 0.51595 | 0.51994 | 0.52392 | 0.5279  | 0.53188 | 0.53586 |
| 0.1 | 0.53983 | 0.5438  | 0.54776 | 0.55172 | 0.55567 | 0.55962 | 0.56356 | 0.56749 | 0.57142 | 0.57535 |
| 0.2 | 0.57926 | 0.58317 | 0.58706 | 0.59095 | 0.59483 | 0.59871 | 0.60257 | 0.60642 | 0.61026 | 0.61409 |
| 0.3 | 0.61791 | 0.62172 | 0.62552 | 0.6293  | 0.63307 | 0.63683 | 0.64058 | 0.64431 | 0.64803 | 0.65173 |
| 0.4 | 0.65542 | 0.6591  | 0.66276 | 0.6664  | 0.67003 | 0.67364 | 0.67724 | 0.68082 | 0.68439 | 0.68793 |
| 0.5 | 0.69146 | 0.69497 | 0.69847 | 0.70194 | 0.7054  | 0.70884 | 0.71226 | 0.71566 | 0.71904 | 0.7224  |
| 0.6 | 0.72575 | 0.72907 | 0.73237 | 0.73565 | 0.73891 | 0.74215 | 0.74537 | 0.74857 | 0.75175 | 0.7549  |
| 0.7 | 0.75804 | 0.76115 | 0.76424 | 0.7673  | 0.77035 | 0.77337 | 0.77637 | 0.77935 | 0.7823  | 0.78524 |
| 0.8 | 0.78814 | 0.79103 | 0.79389 | 0.79673 | 0.79955 | 0.80234 | 0.80511 | 0.80785 | 0.81057 | 0.81327 |
| 0.9 | 0.81594 | 0.81859 | 0.82121 | 0.82381 | 0.82639 | 0.82894 | 0.83147 | 0.83398 | 0.83646 | 0.83891 |
| 1.0 | 0.84134 | 0.84375 | 0.84614 | 0.84849 | 0.85083 | 0.85314 | 0.85543 | 0.85769 | 0.85993 | 0.86214 |
| 1.1 | 0.86433 | 0.8665  | 0.86864 | 0.87076 | 0.87286 | 0.87493 | 0.87698 | 0.879   | 0.881   | 0.88298 |
| 1.2 | 0.88493 | 0.88686 | 0.88877 | 0.89065 | 0.89251 | 0.89435 | 0.89617 | 0.89796 | 0.89973 | 0.90147 |
| 1.3 | 0.9032  | 0.9049  | 0.90658 | 0.90824 | 0.90988 | 0.91149 | 0.91309 | 0.91466 | 0.91621 | 0.91774 |
| 1.4 | 0.91924 | 0.92073 | 0.9222  | 0.92364 | 0.92507 | 0.92647 | 0.92785 | 0.92922 | 0.93056 | 0.93189 |
| 1.5 | 0.93319 | 0.93448 | 0.93574 | 0.93699 | 0.93822 | 0.93943 | 0.94062 | 0.94179 | 0.94295 | 0.94408 |
| 1.6 | 0.9452  | 0.9463  | 0.94738 | 0.94845 | 0.9495  | 0.95053 | 0.95154 | 0.95254 | 0.95352 | 0.95449 |
| 1.7 | 0.95543 | 0.95637 | 0.95728 | 0.95818 | 0.95907 | 0.95994 | 0.9608  | 0.96164 | 0.96246 | 0.96327 |
| 1.8 | 0.96407 | 0.96485 | 0.96562 | 0.96638 | 0.96712 | 0.96784 | 0.96856 | 0.96926 | 0.96995 | 0.97062 |
| 1.9 | 0.97128 | 0.97193 | 0.97257 | 0.9732  | 0.97381 | 0.97441 | 0.975   | 0.97558 | 0.97615 | 0.9767  |
| 2.0 | 0.97725 | 0.97778 | 0.97831 | 0.97882 | 0.97932 | 0.97982 | 0.9803  | 0.98077 | 0.98124 | 0.98169 |
| 2.1 | 0.98214 | 0.98257 | 0.983   | 0.98341 | 0.98382 | 0.98422 | 0.98461 | 0.985   | 0.98537 | 0.98574 |
| 2.2 | 0.9861  | 0.98645 | 0.98679 | 0.98713 | 0.98745 | 0.98778 | 0.98809 | 0.9884  | 0.9887  | 0.98899 |
| 2.3 | 0.98928 | 0.98956 | 0.98983 | 0.9901  | 0.99036 | 0.99061 | 0.99086 | 0.99111 | 0.99134 | 0.99158 |
| 2.4 | 0.9918  | 0.99202 | 0.99224 | 0.99245 | 0.99266 | 0.99286 | 0.99305 | 0.99324 | 0.99343 | 0.99361 |
| 2.5 | 0.99379 | 0.99396 | 0.99413 | 0.9943  | 0.99446 | 0.99461 | 0.99477 | 0.99492 | 0.99506 | 0.9952  |
| 2.6 | 0.99534 | 0.99547 | 0.9956  | 0.99573 | 0.99585 | 0.99598 | 0.99609 | 0.99621 | 0.99632 | 0.99643 |
| 2.7 | 0.99653 | 0.99664 | 0.99674 | 0.99683 | 0.99693 | 0.99702 | 0.99711 | 0.9972  | 0.99728 | 0.99736 |
| 2.8 | 0.99744 | 0.99752 | 0.9976  | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.99865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999   |

We have a table containing values for the CDF of the standard normal random variable  $Z \sim \mathcal{N}(0,1)$

$$\mu=0 \quad \sigma^2=1$$

$$P(Z < 0.23) = \Phi(0.23)$$

Yes, we're going to use a table in 2026 🤖  
Mainly for consistency in this class.

(In the real world, we have programming libraries like Python's scipy: stats.norm.cdf)

# We have a table with precomputed values!

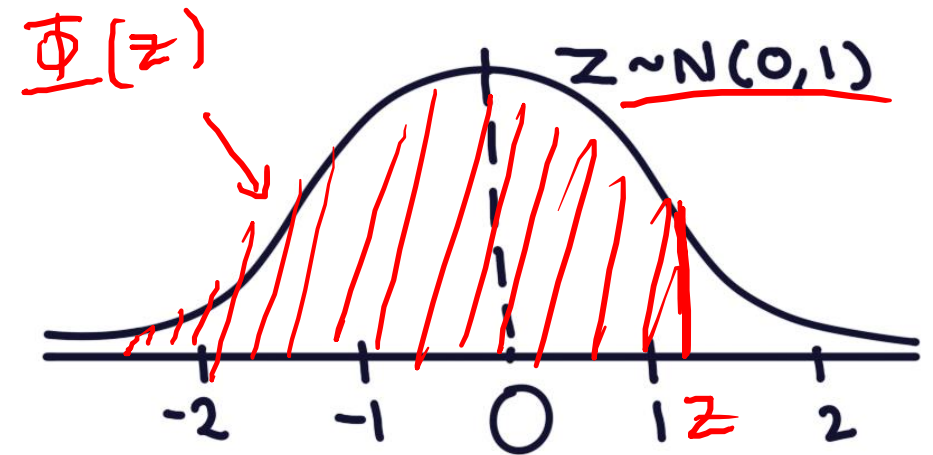
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| 2.8 | 0.99744 | 0.99752 | 0.9976  | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
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We have a table containing values for the CDF of the **standard normal random variable**  $Z \sim \mathcal{N}(0,1)$

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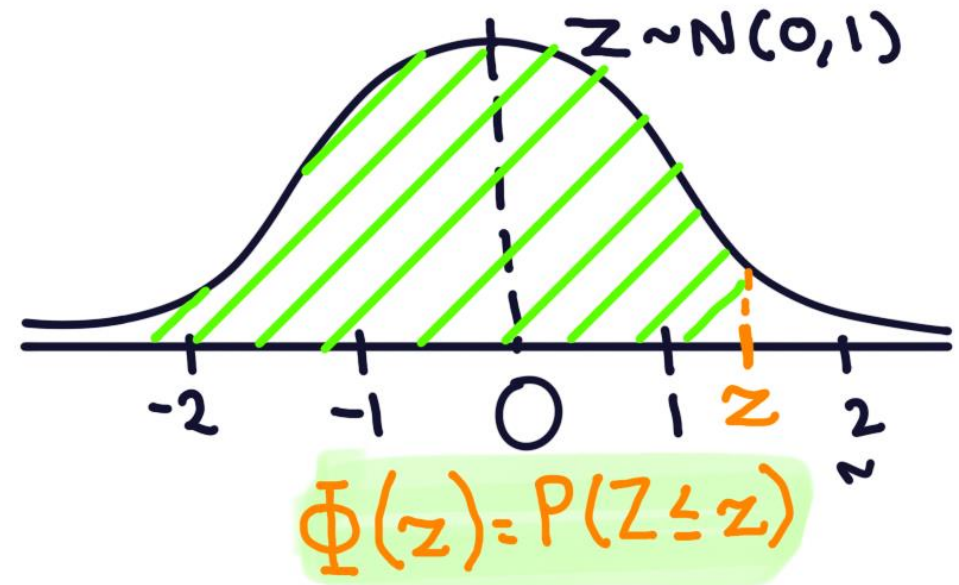
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| 2.1 | 0.98214 | 0.98257 | 0.983   | 0.98341 | 0.98382 | 0.98422 | 0.98461 | 0.985   | 0.98537 | 0.98574 |
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| 2.8 | 0.99744 | 0.99752 | 0.9976  | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
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We have a table containing values for the CDF of the standard normal random variable  $Z \sim \mathcal{N}(0,1)$

>  $\Phi$  is a function for CDF of  $\mathcal{N}(0,1)$

>  $\Phi(z) = F_Z(z) = \mathbb{P}(Z \leq z)$



# But how to go from $\mathcal{N}(\mu, \sigma^2)$ to $\mathcal{N}(0,1)$ ?

We have a table for the values of  $\mathcal{N}(0,1)$ . How to use this for  $\mathcal{N}(\mu, \sigma^2)$ ?

go to  $\mu=0, \sigma^2=1$

We will **standardize**  $X$ ! If we have  $X \sim \mathcal{N}(\mu, \sigma^2)$

1. Subtract  $\mu$  to shift the distribution to have **mean of 0**

$$\mathbb{E}[X - \mu] = \mathbb{E}[X] - \mu = \mu - \mu = 0$$

2. Divide by  $\sigma$  to squish/stretch the distribution to have **variance of 1**

$$\text{Var}\left(\frac{X - \mu}{\sigma}\right) = \frac{1}{\sigma^2} \text{Var}(X - \mu) = \frac{1}{\sigma^2} \text{Var}(X) = \frac{1}{\sigma^2} \sigma^2 = 1$$

$Z = \frac{X - \mu}{\sigma}$  is a standard normal random variable:  $Z \sim \mathcal{N}(0,1)$

# Computing Probabilities of Normal RVs

1. Write the probability we're interested in in terms of the CDF
2. Standardize the normal random variable:  $Z = \frac{X - \mu}{\sigma}$
3. Round the "z-score"(s) to the hundredths place.
4. Look up the value(s) in the table

$$Z \sim N(0, 1)$$

$$P(Z \leq k)$$

$$\rightarrow \Phi(k) = \dots$$

z-score

# Practice!

1. CDF notation
2. Standardize
3. Get z-score
4. LOOKUP

Let  $X \sim \mathcal{N}(\overset{\mu}{5}, \overset{\sigma^2}{4})$ . What is  $\mathbb{P}(X \leq 9)$ ?

①  $\mathbb{P}(X \leq 9)$

②  $= P\left(\frac{X-5}{\sqrt{4}} \leq \frac{9-5}{\sqrt{4}}\right)$

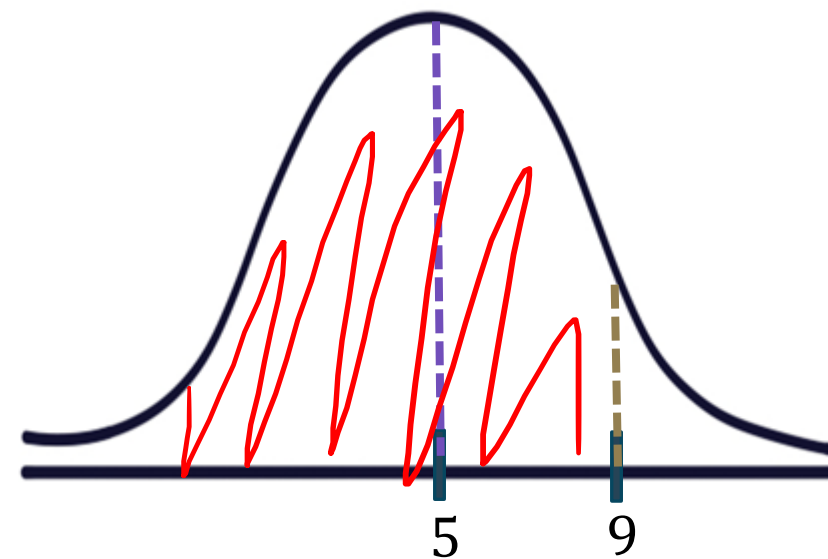
$= P\left(Z \leq \frac{9-5}{2}\right)$

③  $= P(Z \leq 2)$

$= \Phi(2)$

We use  $\Phi(z)$  to mean  $F_Z(z)$  where  $Z \sim \mathcal{N}(0,1)$ .

$$Z = \frac{X - \mu}{\sigma}$$



# Practice!

We use  $\Phi(z)$  to mean  $F_Z(z)$  where  $Z \sim \mathcal{N}(0,1)$ .

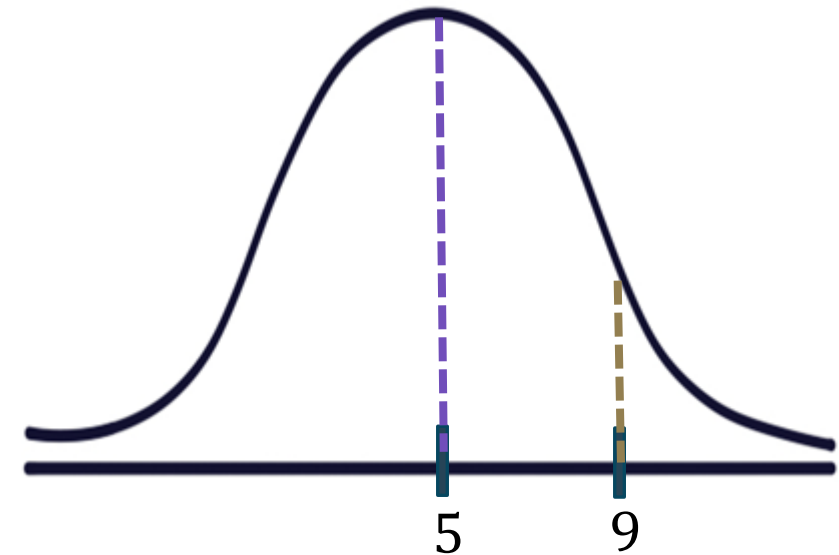
Let  $X \sim \mathcal{N}(5,4)$ . What is  $\mathbb{P}(X \leq 9)$ ?

$$\mathbb{P}(X \leq 9)$$

$$= \mathbb{P}\left(\frac{Y-5}{2} \leq \frac{9-5}{2}\right) \text{ standardize (algebra on both sides)}$$

$$= \mathbb{P}\left(Z \leq \frac{9-5}{2}\right) \text{ where } Z \sim \mathcal{N}(0,1).$$

$$= \mathbb{P}\left(Z \leq \frac{9-5}{2}\right) = \Phi(2.00) = \underline{0.97725}$$



# Practice!

Let  $X \sim \mathcal{N}(5, 4)$ . What is  $\mathbb{P}(X > 9)$ ?

$$\mathbb{P}(X > 9)$$

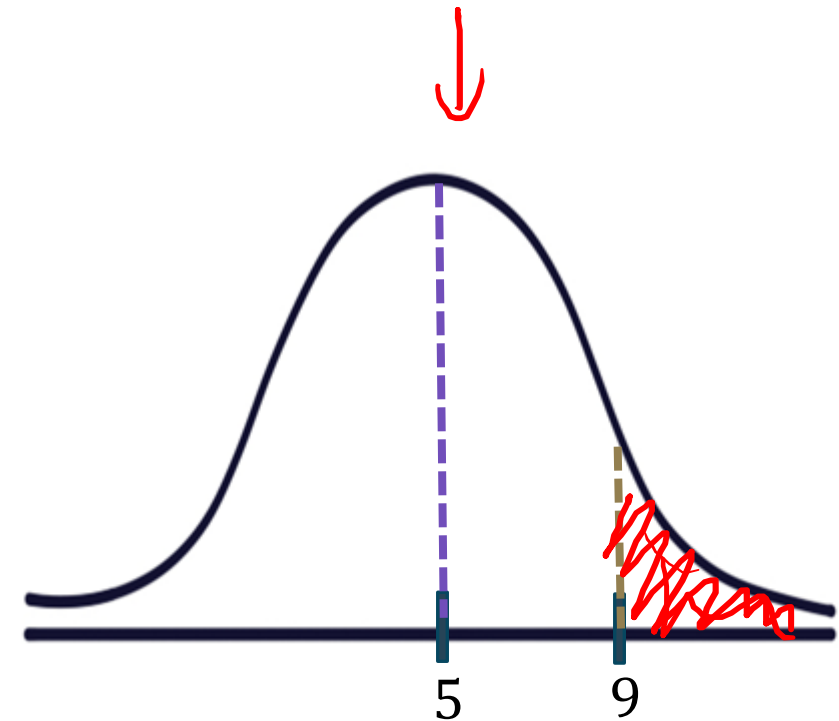
$$= 1 - \mathbb{P}(X \leq 9)$$

$$= 1 - \mathbb{P}\left(\frac{X - 5}{2} \leq \frac{9 - 5}{2}\right)$$

$$= 1 - \Phi(z)$$

We use  $\Phi(z)$  to mean  $F_Z(z)$  where  $Z \sim \mathcal{N}(0, 1)$ .

$$\Phi = \mathbb{P}(Z \leq z)$$



# Practice!

We use  $\Phi(z)$  to mean  $F_Z(z)$  where  $Z \sim \mathcal{N}(0,1)$ .

Let  $X \sim \mathcal{N}(5,4)$ . What is  $\mathbb{P}(X > 9)$ ?

$$\mathbb{P}(X > 9) = 1 - \mathbb{P}(X \leq 9)$$

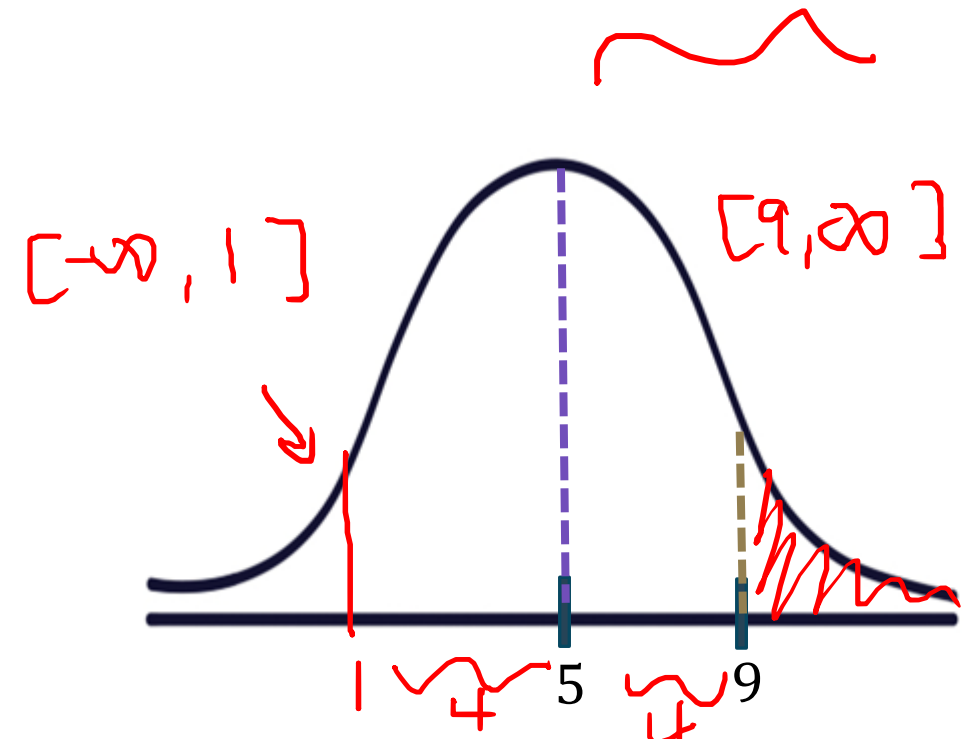
$$= 1 - \mathbb{P}\left(\frac{Y-5}{2} \leq \frac{9-5}{2}\right) \text{ standardize (algebra on both sides)}$$

$$= 1 - \mathbb{P}\left(Z \leq \frac{9-5}{2}\right) \text{ where } Z \sim \mathcal{N}(0,1).$$

$$= 1 - \mathbb{P}\left(Z \leq \frac{9-5}{2}\right) = 1 - \Phi(2.00)$$

$$= 1 - 0.97725 = 0.02275$$

$$P(X > 9) = P(X < 1)$$



positive

| $z$ | 0.00    | 0.01    | 0.02    | 0.03    | 0.04    | 0.05    | 0.06    | 0.07    | 0.08    | 0.09    |
|-----|---------|---------|---------|---------|---------|---------|---------|---------|---------|---------|
| 0.0 | 0.5     | 0.50399 | 0.50798 | 0.51197 | 0.51595 | 0.51994 | 0.52392 | 0.5279  | 0.53188 | 0.53586 |
| 0.1 | 0.53983 | 0.5438  | 0.54776 | 0.55172 | 0.55567 | 0.55962 | 0.56356 | 0.56749 | 0.57142 | 0.57535 |
| 0.2 | 0.57926 | 0.58317 | 0.58706 | 0.59095 | 0.59483 | 0.59871 | 0.60257 | 0.60642 | 0.61026 | 0.61409 |
| 0.3 | 0.61791 | 0.62172 | 0.62552 | 0.6293  | 0.63307 | 0.63683 | 0.64058 | 0.64431 | 0.64803 | 0.65173 |
| 0.4 | 0.65542 | 0.6591  | 0.66276 | 0.6664  | 0.67003 | 0.67364 | 0.67724 | 0.68082 | 0.68439 | 0.68793 |
| 0.5 | 0.69146 | 0.69497 | 0.69847 | 0.70194 | 0.7054  | 0.70884 | 0.71226 | 0.71566 | 0.71904 | 0.7224  |
| 0.6 | 0.72575 | 0.72907 | 0.73237 | 0.73565 | 0.73891 | 0.74215 | 0.74537 | 0.74857 | 0.75175 | 0.7549  |
| 0.7 | 0.75804 | 0.76115 | 0.76424 | 0.7673  | 0.77035 | 0.77337 | 0.77637 | 0.77935 | 0.7823  | 0.78524 |
| 0.8 | 0.78814 | 0.79103 | 0.79389 | 0.79673 | 0.79955 | 0.80234 | 0.80511 | 0.80785 | 0.81057 | 0.81327 |
| 0.9 | 0.81594 | 0.81859 | 0.82121 | 0.82381 | 0.82639 | 0.82894 | 0.83147 | 0.83398 | 0.83646 | 0.83891 |
| 1.0 | 0.84134 | 0.84375 | 0.84614 | 0.84849 | 0.85083 | 0.85314 | 0.85543 | 0.85769 | 0.85993 | 0.86214 |
| 1.1 | 0.86433 | 0.8665  | 0.86864 | 0.87076 | 0.87286 | 0.87493 | 0.87698 | 0.879   | 0.881   | 0.88298 |
| 1.2 | 0.88493 | 0.88686 | 0.88877 | 0.89065 | 0.89251 | 0.89435 | 0.89617 | 0.89796 | 0.89973 | 0.90147 |
| 1.3 | 0.9032  | 0.9049  | 0.90658 | 0.90824 | 0.90988 | 0.91149 | 0.91309 | 0.91466 | 0.91621 | 0.91774 |
| 1.4 | 0.91924 | 0.92073 | 0.9222  | 0.92364 | 0.92507 | 0.92647 | 0.92785 | 0.92922 | 0.93056 | 0.93189 |
| 1.5 | 0.93319 | 0.93448 | 0.93574 | 0.93699 | 0.93822 | 0.93943 | 0.94062 | 0.94179 | 0.94295 | 0.94408 |
| 1.6 | 0.9452  | 0.9463  | 0.94738 | 0.94845 | 0.9495  | 0.95053 | 0.95154 | 0.95254 | 0.95352 | 0.95449 |
| 1.7 | 0.95543 | 0.95637 | 0.95728 | 0.95818 | 0.95907 | 0.95994 | 0.9608  | 0.96164 | 0.96246 | 0.96327 |
| 1.8 | 0.96407 | 0.96485 | 0.96562 | 0.96638 | 0.96712 | 0.96784 | 0.96856 | 0.96926 | 0.96995 | 0.97062 |
| 1.9 | 0.97128 | 0.97193 | 0.97257 | 0.9732  | 0.97381 | 0.97441 | 0.975   | 0.97558 | 0.97615 | 0.9767  |
| 2.0 | 0.97725 | 0.97778 | 0.97831 | 0.97882 | 0.97932 | 0.97982 | 0.9803  | 0.98077 | 0.98124 | 0.98169 |
| 2.1 | 0.98214 | 0.98257 | 0.983   | 0.98341 | 0.98382 | 0.98422 | 0.98461 | 0.985   | 0.98537 | 0.98574 |
| 2.2 | 0.9861  | 0.98645 | 0.98679 | 0.98713 | 0.98745 | 0.98778 | 0.98809 | 0.9884  | 0.9887  | 0.98899 |
| 2.3 | 0.98928 | 0.98956 | 0.98983 | 0.9901  | 0.99036 | 0.99061 | 0.99086 | 0.99111 | 0.99134 | 0.99158 |
| 2.4 | 0.9918  | 0.99202 | 0.99224 | 0.99245 | 0.99266 | 0.99286 | 0.99305 | 0.99324 | 0.99343 | 0.99361 |
| 2.5 | 0.99379 | 0.99396 | 0.99413 | 0.9943  | 0.99446 | 0.99461 | 0.99477 | 0.99492 | 0.99506 | 0.9952  |
| 2.6 | 0.99534 | 0.99547 | 0.9956  | 0.99573 | 0.99585 | 0.99598 | 0.99609 | 0.99621 | 0.99632 | 0.99643 |
| 2.7 | 0.99653 | 0.99664 | 0.99674 | 0.99683 | 0.99693 | 0.99702 | 0.99711 | 0.9972  | 0.99728 | 0.99736 |
| 2.8 | 0.99744 | 0.99752 | 0.9976  | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.99865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999   |

# More practice

Let  $X \sim \mathcal{N}(3, 2)$ .

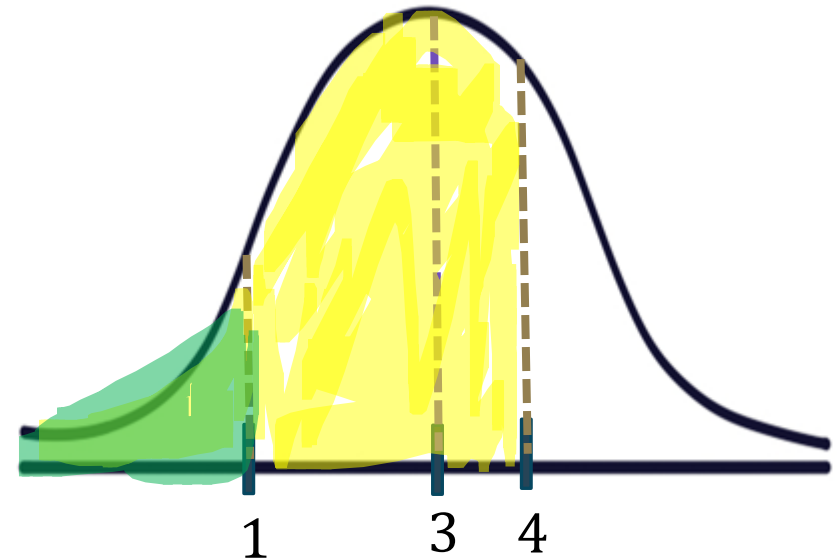
What is the probability that  $1 \leq X \leq 4$ ?

$$\mathbb{P}(1 \leq X \leq 4)$$

$$= P(X \leq 4) - P(X \leq 1)$$

$$= P\left(\frac{X-3}{\sqrt{2}} \leq \frac{4-3}{\sqrt{2}}\right) - P\left(\frac{X-3}{\sqrt{2}} \leq \frac{1-3}{\sqrt{2}}\right)$$

$$= P\left(Z \leq \frac{1}{\sqrt{2}}\right) - P\left(Z \leq \underline{\underline{-\frac{2}{\sqrt{2}}}}\right)$$

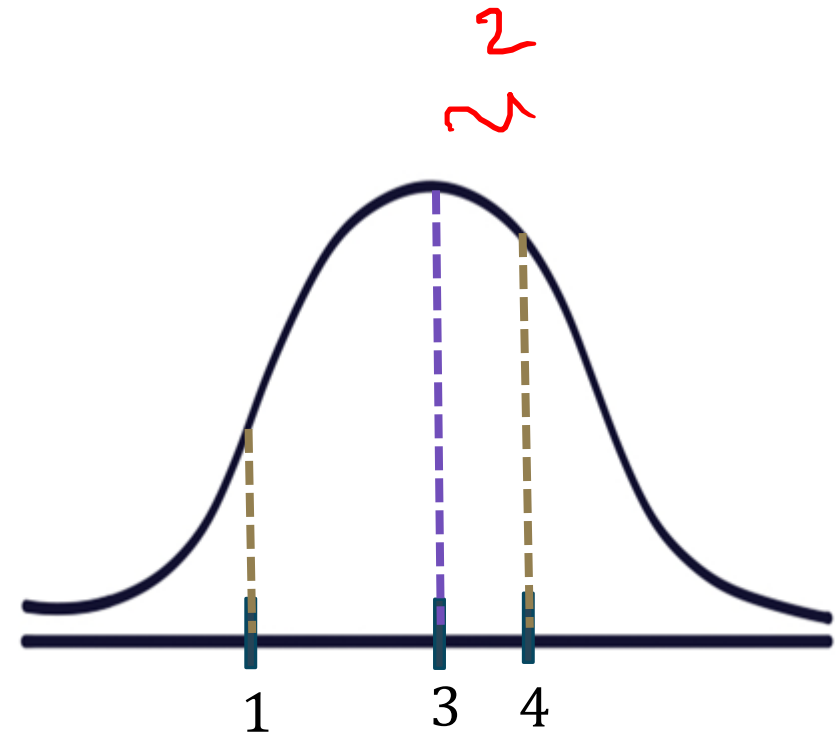


# More practice

Let  $X \sim \mathcal{N}(3, 2)$ .

What is the probability that  $1 \leq X \leq 4$ ?

$$\begin{aligned} & \mathbb{P}(1 \leq X \leq 4) \\ &= \mathbb{P}(X \leq 4) - \mathbb{P}(X \leq 1) \\ &= \mathbb{P}\left(\frac{X-3}{\sqrt{2}} \leq \frac{4-3}{\sqrt{2}}\right) - \mathbb{P}\left(\frac{X-3}{\sqrt{2}} \leq \frac{1-3}{\sqrt{2}}\right) = \\ &= \mathbb{P}\left(Z \leq \frac{4-3}{\sqrt{2}}\right) - \mathbb{P}\left(Z \leq \frac{1-3}{\sqrt{2}}\right) \text{ where } Z \sim \mathcal{N}(0,1) \\ &= \Phi(.71) - \Phi(-1.41) \end{aligned}$$



# More practice

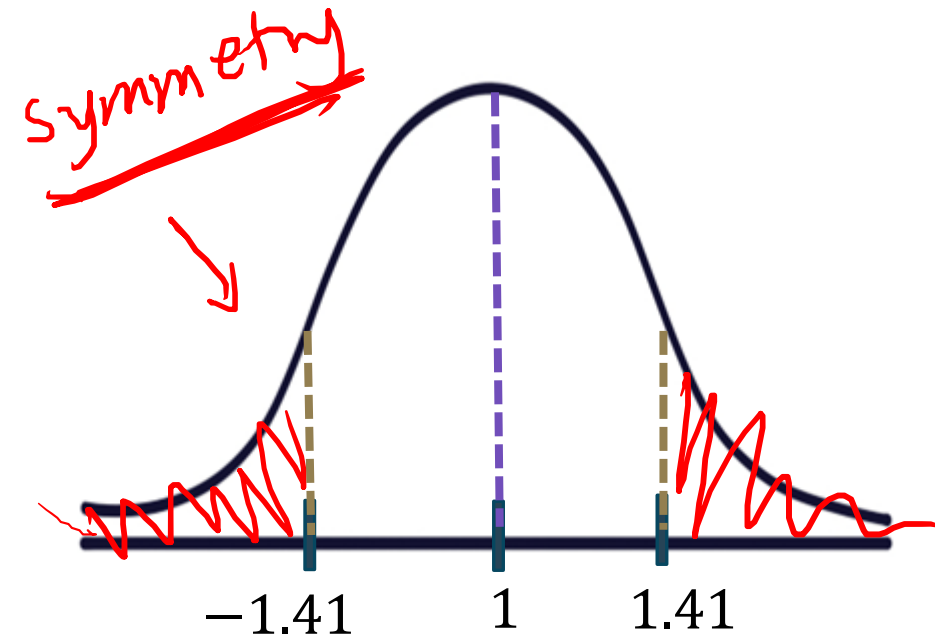
How do we find  $\Phi(-1.41) = \mathbb{P}(Z \leq -1.41)$ ? Our table only has CDF values for positive numbers! 🤖

Recall: the normal distribution is symmetric

$$\mathbb{P}(Z \leq -1.41) = \underline{\mathbb{P}(Z \geq 1.41)}$$

$$= 1 - \mathbb{P}(Z \leq 1.41)$$

$$= 1 - \underline{\Phi(1.41)}$$



# More practice

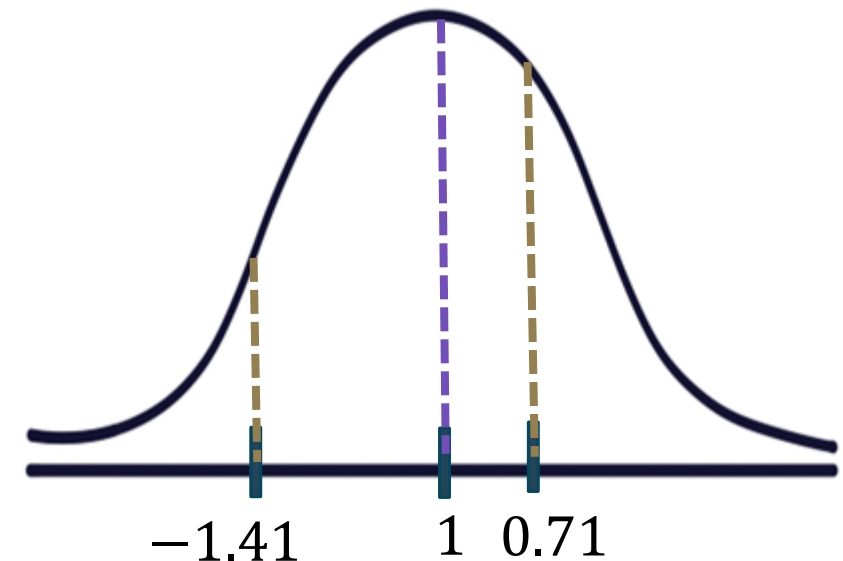
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Recall: the normal distribution is symmetric

$$\mathbb{P}(Z \leq -1.41) = \mathbb{P}(Z \geq 1.41)$$

$$= 1 - \mathbb{P}(Z \leq 1.41)$$

$$= 1 - \Phi(1.41) \text{ this is something we can do...}$$



# More practice

Let  $X \sim \mathcal{N}(3, 2)$ .

What is the probability that  $1 \leq X \leq 4$ ?

$$\mathbb{P}(1 \leq X \leq 4)$$

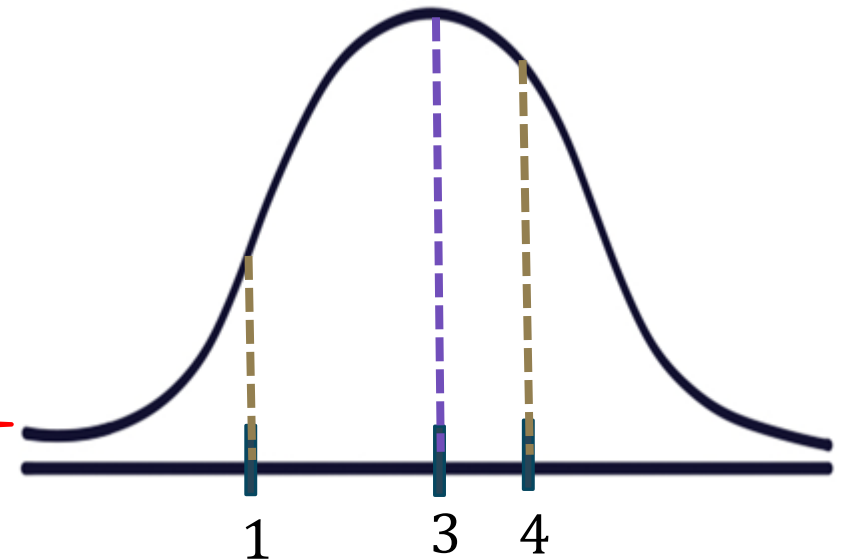
$$= \mathbb{P}(X \leq 4) - \mathbb{P}(X \leq 1)$$

$$= \mathbb{P}\left(\frac{X-3}{\sqrt{2}} \leq \frac{4-3}{\sqrt{2}}\right) - \mathbb{P}\left(\frac{X-3}{\sqrt{2}} \leq \frac{1-3}{\sqrt{2}}\right) =$$

$$= \mathbb{P}\left(Z \leq \frac{4-3}{\sqrt{2}}\right) - \mathbb{P}\left(Z \leq \frac{1-3}{\sqrt{2}}\right) \text{ where } Z \sim \mathcal{N}(0,1)$$

$$= \Phi(.71) - \Phi(-1.41) = \Phi(.71) - (1 - \Phi(1.41))$$

$$= 0.76115 - (1 - 0.92073) = \mathbf{0.68188}$$



# *"within    standard deviations from the mean"*

What's the probability of being within two standard deviations from the mean?

$$\begin{aligned} & \mathbb{P}(\mu - 2\sigma \leq X \leq \mu + 2\sigma) \\ &= \mathbb{P}(X \leq \mu + 2\sigma) - \mathbb{P}(X \leq \mu - 2\sigma) = \mathbb{P}\left(Z \leq \frac{\mu + 2\sigma - \mu}{\sigma}\right) - \mathbb{P}\left(Z \leq \frac{\mu - 2\sigma - \mu}{\sigma}\right) \\ &= \Phi(2) - \Phi(-2) \\ &= \Phi(2) - (1 - \Phi(2)) \text{ symmetry + compl.} \\ &= .97725 - (1 - .97725) = .9545 \end{aligned}$$

You'll sometimes hear statisticians refer to the "68-95-99.7 rule" which is the probability of being within 1, 2, or 3 standard deviations of the mean.

# Normal (aka Guassian)

When finding the probability of a normal random variable, draw a picture!! It can help reasoning about how we can use the CDF and the z-table to compute the desired region.

$$X \sim N(\mu, \sigma^2)$$

1. Write the probability we're interested in in terms of the CDF

2. Standardize the normal random variable:  $Z = \frac{X - \mu}{\sigma}$

3. Round the "z-score"(s) to the hundredths place.

4. Look up the value(s) in the table

# Can you spot the normal distribution?



It turns out the normal distribution appear a LOT in the real world. Like...in the gym!

Picture from reddit