

Final Reference Sheet

Counting

Sum Rule

If you are choosing one thing between n options in one group and m in another group with no overlap, the total number of options is: $n + m$.

Product Rule

If you have a sequential process, where step 1 has n_1 options, step 2 has n_2 options, ..., step k has n_k options, and you choose one from each step, the total number of possibilities is $n_1 \cdot n_2 \cdot \dots \cdot n_k$

n Factorial

$$n! = n \cdot (n-1) \cdot (n-2) \cdot \dots \cdot 1$$

We only define $n!$ for natural numbers n . As a convention, we define: $0! = 1$

k -permutation

The number of k -element sequences of distinct symbols from a universe of n symbols is:

$$P(n, k) = n \cdot (n-1) \cdot \dots \cdot (n-k+1) = \frac{n!}{(n-k)!}$$

Edge cases: $P(n, n) = n!$, $P(n, 0) = 1$, $P(n, k)$ for $k < 0$ or $k > n$ is undefined.

k -combination

The number of k -element subsets from a set of n symbols is:

$$C(n, k) = \binom{n}{k} = \frac{P(n, k)}{k!} = \frac{n!}{k! \cdot (n-k)!}$$

Edge cases: $\binom{n}{0} = \binom{n}{n} = 1$; $P(n, k)$ for $k < 0$ or $k > n$ is undefined.

Combination Facts

Symmetry of combinations:

$$\binom{n}{k} = \binom{n}{n-k}$$

Pascal's Rule:

$$\binom{n}{k} = \binom{n-1}{k-1} + \binom{n-1}{k}$$

Binomial Theorem

Let $x, y \in \mathbb{R}$ and $n \in \mathbb{N}$ positive integer. Then:

$$(x+y)^n = \sum_{i=0}^n \binom{n}{i} x^i y^{n-i}$$

Principle of Inclusion-Exclusion (PIE)

2 events: $|A \cup B| = |A| + |B| - |A \cap B|$

3 events: $|A \cup B \cup C| = |A| + |B| + |C| - |A \cap B| - |A \cap C| - |B \cap C| + |A \cap B \cap C|$

k events: singles + doubles + triples - quads + ...

Pigeonhole Principle

If there are n pigeons we want to put into k holes (where $n > k$), then at least one pigeonhole must contain at least $\lceil \frac{n}{k} \rceil$ pigeons.

Stars and Bars

To pick n objects from k groups (where order doesn't matter and every element of each group is indistinguishable), use the formula:

$$\binom{n+k-1}{k-1}$$

Probability

Sample Space

The sample space is the set Ω of all possible outcomes of an experiment.

Event

An event $E \subseteq \Omega$ is a subset of possible outcomes (i.e. a subset of Ω).

Mutually Exclusive Events

Two events E, F are mutually exclusive if they cannot happen simultaneously (they are disjoint subsets of the sample space). In notation,

$$E \cap F = \emptyset$$

Probability Space

A (discrete) probability space is a pair (Ω, \mathbb{P}) , where Ω is the sample space and $\mathbb{P} : \Omega \rightarrow [0, 1]$ is the probability measure such that:

- $\mathbb{P}(x) \geq 0$ for all x
- $\sum_{\omega \in \Omega} \mathbb{P}(\omega) = 1$
- If $E, F \subseteq \Omega$ and $E \cap F = \emptyset$ then $\mathbb{P}(E \cup F) = \mathbb{P}(E) + \mathbb{P}(F)$

Probability Consequences (Corollaries)

Complementation: $\mathbb{P}(E^c) = 1 - \mathbb{P}(E)$

Monotonicity: If $E \subseteq F$, then $\mathbb{P}(E) \leq \mathbb{P}(F)$

Inclusion-exclusion: $\mathbb{P}(E \cup F) = \mathbb{P}(E) + \mathbb{P}(F) - \mathbb{P}(E \cap F)$

Uniform Probability Space

If every outcome in your sample space is equally likely then,

$$\mathbb{P}(E) = \frac{|E|}{|\Omega|}$$

Conditional Probability

For an event B , with $\mathbb{P}(B) > 0$, the "Probability of A conditioned on B " is

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$$

$\mathbb{P}(A|B)$ is undefined when $\mathbb{P}(B) = 0$.

Bayes' Rule

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A)\mathbb{P}(A)}{\mathbb{P}(B)}$$

Partition

Non-empty events A_1, \dots, A_n partition the sample space Ω if:

- **Exhaustive:** $A_1 \cup A_2 \cup \dots \cup A_n = \Omega$ (they cover the entire sample space).
- **Pairwise Mutually Exclusive:** For all $i \neq j$, $A_i \cap A_j = \emptyset$ (none of them overlap)

Law of Total Probability (LTP)

Let A_1, A_2, \dots, A_n partition Ω .

For any event E :

$$\mathbb{P}(E) = \sum_{i=1}^n \mathbb{P}(E \cap A_i) = \sum_{i=1}^n \mathbb{P}(E|A_i)\mathbb{P}(A_i)$$

Independence

Two events A, B are independent if

$$\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

If A, B both have non-zero probabilities, then they are independent if any of the following equivalent statements hold:

1. $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$
2. $\mathbb{P}(A|B) = \mathbb{P}(A)$
3. $\mathbb{P}(B|A) = \mathbb{P}(B)$

Chain Rule

$$\mathbb{P}(A_1 \cap A_2 \cap \dots \cap A_n) =$$

$$\mathbb{P}(A_1) \cdot \mathbb{P}(A_2|A_1) \cdot \mathbb{P}(A_3|A_1 \cap A_2) \cdots \mathbb{P}(A_n|A_1 \cap \dots \cap A_{n-1})$$

Conditional Independence

We say A and B are conditionally independent on C

$$\mathbb{P}(A \cap B|C) = \mathbb{P}(A|C) \cdot \mathbb{P}(B|C)$$

Pairwise Independence

Events A_1, A_2, \dots, A_n are pairwise independent if for all i, j

$$\mathbb{P}(A_i \cap A_j) = \mathbb{P}(A_i) \cdot \mathbb{P}(A_j)$$

Mutual Independence

Events A_1, A_2, \dots, A_n are mutually independent if

$$\mathbb{P}(A_{i_1} \cap A_{i_2} \cap \dots \cap A_{i_k}) = \mathbb{P}(A_{i_1}) \cdot \mathbb{P}(A_{i_2}) \cdots \mathbb{P}(A_{i_k})$$

For every subset $\{i_1, i_2, \dots, i_k\}$ of $\{1, 2, \dots, n\}$.

Random Variables (Discrete and Continuous)

Random Variable (RV)

A random variable (RV) X is a numeric function of the outcome $X : \Omega \rightarrow \mathbb{R}$. $X(\omega)$ is the summary of the outcome ω . The set of possible values X can take on is its range/support, denoted Ω_X .

Probability Mass Function (PMF) (Discrete)

For a discrete RV X , it assigns probabilities to values in its range. That is $p_X : \Omega_X \rightarrow [0, 1]$, where $p_X(k) = P(X = k)$.

Expectation (Discrete)

The "expectation" (or "expected value") of a random variable X is:

$$\mathbb{E}[X] = \sum_{k \in \Omega_X} k \cdot p_X(k)$$

LOTUS - Expectation of A Function (Discrete)

For a RV X and function g :

$$\mathbb{E}[g(X)] = \sum_{b \in \Omega_X} g(b) \cdot p_X(b)$$

Linearity of Expectation (LoE)

For any random variables X, Y (possibly dependent):

$$\mathbb{E}[aX + bY + c] = a\mathbb{E}[X] + b\mathbb{E}[Y] + c$$

Variance

$$\text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

Standard Deviation (SD)

$$\sigma_X = \sqrt{\text{Var}(X)}$$

Property of Variance

$$\text{Var}(aX + b) = a^2 \text{Var}(X)$$

Probability Density Function (PDF) (Continuous)

The PDF of a continuous RV X is the function $f_X : \mathbb{R} \rightarrow \mathbb{R}$, such that

1. $f_X(z) \geq 0$ for all $z \in \mathbb{R}$
2. $\int_{-\infty}^{\infty} f_X(t) dt = 1$

Furthermore, $\mathbb{P}(a \leq X \leq b) = \int_a^b f_X(w) dw$

Cumulative Distribution Function (CDF)

The CDF of ANY random variable (discrete or continuous) is $F_X(t) = \mathbb{P}(X \leq t)$.

If X is *continuous*, $F_X(t) = \mathbb{P}(X \leq t) = \int_{-\infty}^t f_X(w) dw$ for all $t \in \mathbb{R}$. Further, $\frac{d}{du} F_X(u) = f_X(u)$

Expectation (Continuous)

$$\mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx$$

LOTUS - Expectation of A Function (Continuous)

For a RV X and function g :

$$\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx$$

Multiplicativity of expectation

For any independent random variables X, Y :

$$\mathbb{E}[XY] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$$

Independence (Random Variables)

Random variables X and Y are independent if for all $x \in \Omega_X$ and all $y \in \Omega_Y$:

$$\mathbb{P}(X = x, Y = y) = \mathbb{P}(X = x) \cdot \mathbb{P}(Y = y)$$

Variance Adds for Independent RVs

If X, Y are independent, then

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$$

Zoo of Random Variables

Bernoulli/Indicator Random Variable

$X \sim \text{Ber}(p)$ iff X has PMF:

$$p_X(k) = \begin{cases} p, & k = 1 \\ 1 - p, & k = 0 \end{cases}$$

$\mathbb{E}[X] = p$ and $\text{Var}(X) = p(1-p)$.

Negative Binomial Random Variable

$X \sim \text{NegBin}(r, p)$ iff X has PMF:

$$p_X(k) = \binom{k-1}{r-1} (1-p)^{k-r} p^r, k \in \Omega_X = \{r, r+1, \dots\}$$

$\mathbb{E}[X] = \frac{r}{p}$ and $\text{Var}(X) = n \frac{r(1-p)}{p^2}$.

X is the number of independent trials till the r 'th success.

Binomial Random Variable

$X \sim \text{Bin}(n, p)$ iff X has PMF

$$p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}, k \in \Omega_X = \{0, 1, \dots, n\}$$

$\mathbb{E}[X] = np$ and $\text{Var}(X) = np(1-p)$.

X is the number of successes in n independent trials,

Poisson Random Variable

$X \sim \text{Poi}(\lambda)$ iff X has PMF:

$$p_X(k) = e^{-\lambda} \frac{\lambda^k}{k!}, k \in \Omega_X = \{0, 1, 2, \dots\}$$

$\mathbb{E}[X] = \lambda$ and $\text{Var}(X) = \lambda$

X is the number of successes seen in a time interval where λ is the average number of successes in the time interval.

Uniform Random Variable (Discrete)

$X \sim \text{Unif}(a, b)$, for integers $a \leq b$, iff X has PMF:

$$p_X(k) = \frac{1}{b-a+1}, k \in \Omega_X = \{a, a+1, \dots, b\}$$

$\mathbb{E}[X] = \frac{a+b}{2}$ and $\text{Var}(X) = \frac{(b-a)(b-a+1)}{12}$

X is an integer between a and b uniformly at random.

Uniform Random Variable (Continuous)

$X \sim \text{Unif}(a, b)$ iff X has PDF:

$$f_X(x) = \begin{cases} \frac{1}{b-a} & \text{if } x \in \Omega_X = [a, b] \\ 0 & \text{otherwise} \end{cases}$$

$\mathbb{E}[X] = \frac{a+b}{2}$ $\text{Var}(X) = \frac{(b-a)^2}{12}$ $F_X(k) = \frac{k-a}{b-a}$ if $a \leq k < b$

X is a real number uniformly at random between a and b

Geometric Random Variable

$X \sim \text{Geo}(p)$ iff X has PMF:

$$p_X(k) = (1-p)^{k-1} p, k \in \Omega_X = \{1, 2, 3, \dots\}$$

$\mathbb{E}[X] = \frac{1}{p}$ and $\text{Var}(X) = \frac{1-p}{p^2}$ and $F_X(k) = 1 - (1-p)^k$

X is the number of independent trials until the first success.

Exponential Random Variable

$X \sim \text{Exp}(\lambda)$ iff X has PDF:

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \in \Omega_X = [0, \infty) \\ 0 & \text{otherwise} \end{cases}$$

$\mathbb{E}[X] = \frac{1}{\lambda}$ and $\text{Var}(X) = \frac{1}{\lambda^2}$

$F_X(x) = 1 - e^{-\lambda x}$ for $x \geq 0$.

X is the time till the first success (λ is same as in Poisson)

Hypergeometric Random Variable

$X \sim \text{HypGeo}(N, K, n)$ iff X has PMF:

$$p_X(k) = \frac{\binom{K}{k} \binom{N-K}{n-k}}{\binom{N}{n}},$$

$$k \in \Omega_X = \{\max\{0, n+K-N\}, \dots, \min\{K, n\}\}$$

$\mathbb{E}[X] = \frac{nK}{N}$ and $\text{Var}(X) = n \frac{K(N-K)(N-n)}{N^2(N-1)}$.

X is the number of success in a sample of size n from a set of N elements, K of which are successes

Normal (Gaussian, "bell curve") Random Variable

$X \sim \mathcal{N}(\mu, \sigma^2)$ iff X has PDF:

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}}, x \in \Omega_X = \mathbb{R}$$

$\mathbb{E}[X] = \mu$ and $\text{Var}(X) = \sigma^2$. random variable is typically denoted Z and has mean 0 and variance 1.

Scale/shift: If $X \sim \mathcal{N}(\mu, \sigma^2)$, then $aX + b \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$. E.g., standard normal: $\frac{X-\mu}{\sigma} \sim \mathcal{N}(0, 1)$.

Addition: If $X \sim \mathcal{N}(\mu_X, \sigma_X^2)$, $Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$ are independent,

$$aX + bY + c \sim \mathcal{N}(a\mu_X + b\mu_Y + c, a^2\sigma_X^2 + b^2\sigma_Y^2)$$

Central Limit Theorem

The Central Limit Theorem (CLT)

Let X_1, X_2, \dots, X_n be i.i.d. RVs with mean μ and variance σ^2 . Let $Y_n = \frac{X_1 + X_2 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$. As $n \rightarrow \infty$, the CDF of Y_n converges to the CDF of $\mathcal{N}(0, 1)$

Joint Distributions

Joint PMFs The joint PMF of discrete RVs X and Y is:

$$p_{X,Y}(a, b) = \mathbb{P}(X = a, Y = b)$$

Their joint range is

$$\Omega_{X,Y} = \{(c, d) : p_{X,Y}(c, d) > 0\} \subseteq \Omega_X \times \Omega_Y$$

Note that $\sum_{(s,t) \in \Omega_{X,Y}} p_{X,Y}(s, t) = 1$.

Joint PDFs The joint PDF of continuous RVs X and Y is:

$$f_{X,Y}(a, b) \geq 0$$

Their joint range is

$$\Omega_{X,Y} = \{(c, d) : f_{X,Y}(c, d) > 0\} \subseteq \Omega_X \times \Omega_Y$$

Note that $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(u, v) du dv = 1$.

Multi-dimensional LOTUS (Discrete) If $g : \mathbb{R}^2 \rightarrow \mathbb{R}$ is a function, then

$$\mathbb{E}[g(X, Y)] = \sum_{x \in \Omega_X} \sum_{y \in \Omega_Y} g(x, y) p_{X,Y}(x, y)$$

Multi-dimensional LOTUS (Continuous) If $g : \mathbb{R}^2 \rightarrow \mathbb{R}$ is a function, then

$$\mathbb{E}[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(s, t) f_{X,Y}(s, t) ds dt$$

Marginal PMFs Let X, Y be discrete random variables. The marginal PMF of X is:

$$p_X(a) = \sum_{b \in \Omega_Y} p_{X,Y}(a, b)$$

Marginal PDFs Let X, Y be continuous random variables. The marginal PDF of X is:

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$$

Independence of RVs (Discrete) Discrete RVs X, Y are **independent**, written $X \perp Y$, if for all $x \in \Omega_X$ and $y \in \Omega_Y$:

$$p_{X,Y}(x, y) = p_X(x)p_Y(y)$$

Independence of RVs (Continuous) Continuous RVs X, Y are independent, written $X \perp Y$, if for all $x \in \Omega_X$ and $y \in \Omega_Y$,

$$f_{X,Y}(x, y) = f_X(x)f_Y(y)$$

Probabilities

$$\mathbb{P}(a \leq X < b, c \leq Y \leq d) = \sum_{x=a}^b \sum_{y=c}^d p_{X,Y}(x, y)$$

Probabilities

$$\mathbb{P}(a \leq X < b, c \leq Y \leq d) = \int_a^b \int_c^d f_{X,Y}(x, y) dy dx$$

Conditional PMF Let X, Y be discrete random variables, then

$$p_{X|Y}(x|y) = \frac{p_{X,Y}(x, y)}{p_Y(y)}$$

Conditional PDF Let X, Y be continuous random variables, then

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x, y)}{f_Y(y)}$$

Conditional Expectation Let X, Y be discrete random variables, then we define the conditional expectation of $g(X)$ given (the event that) $Y = y$ as:

$$\mathbb{E}[X|Y = y] = \sum_{x \in \Omega_X} x \cdot p_{X|Y}(x|y)$$

Conditional Expectation Let X, Y be continuous random variables, then we define the conditional expectation of $g(X)$ given (the event that) $Y = y$ as:

$$\mathbb{E}[X|Y = y] = \int_{-\infty}^{\infty} x \cdot f_{X|Y}(x|y) dx$$

Law of Total Expectation (LTE)

Let X, Y be jointly distributed discrete random variables.

$$\mathbb{E}[g(X)] = \sum_{y \in \Omega_Y} \mathbb{E}[g(X) | Y = y] p_Y(y)$$

Law of Total Expectation (LTE)

Let X, Y be jointly distributed continuous RVs.

$$\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} \mathbb{E}[g(X) | Y = y] f_Y(y) dy$$

Covariance and Variance

$$\begin{aligned} \text{Cov}(X, Y) &= \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] \\ &= \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y] \end{aligned}$$

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$$

Covariance Patterns

1. $\text{Cov}(X, X) = \text{Var}(X)$
2. $\text{Cov}(X, -Y) = -\text{Cov}(X, Y)$
3. if X, Y independent $\text{Cov}(X, Y) = 0$

Tail Bounds

Markov's Inequality Let X be a random variable supported (only) on **non-negative** numbers. For any $t > 0$:

$$\mathbb{P}(X \geq t) \leq \frac{\mathbb{E}[X]}{t}$$

Chebyshev's Inequality Let X be a random variable. For any $t > 0$:

$$\mathbb{P}(|X - \mathbb{E}[X]| \geq t) \leq \frac{\text{Var}(X)}{t^2}$$

Chernoff Bound Let $X = X_1 + X_2 + \dots + X_n$, where X_1, X_2, \dots, X_n are independent Bernoulli random variables and $\mu = \mathbb{E}[X]$. For any $0 \leq \delta \leq 1$:

$$\mathbb{P}(X \geq (1 + \delta)\mu) \leq \exp\left(-\frac{\delta^2\mu}{3}\right)$$

$$\mathbb{P}(X \leq (1 - \delta)\mu) \leq \exp\left(-\frac{\delta^2\mu}{2}\right)$$

The Union Bound Let E_1, E_2, \dots, E_n be a collection of events. Then: $\mathbb{P}(\bigcup_{i=1}^n E_i) \leq \sum_{i=1}^n \mathbb{P}(E_i)$.

Maximum Likelihood Estimation (MLE)

Likelihood (Discrete)

Let x_1, \dots, x_n be i.i.d. samples from PMF $p_X(x; \theta)$ where θ is a parameter (or vector of parameters). The **likelihood** of θ given the samples is

$$\mathcal{L}(x_1, \dots, x_n; \theta) = \prod_{i=1}^n p_X(x_i; \theta).$$

Likelihood (Continuous)

Let x_1, \dots, x_n be i.i.d. samples from PDF $f_X(x; \theta)$ where θ is a parameter (or vector of parameters). The **likelihood** of θ given the samples is

$$\mathcal{L}(x_1, \dots, x_n; \theta) = \prod_{i=1}^n f_X(x_i; \theta).$$

Maximum Likelihood Estimator (MLE)

Let x_1, \dots, x_n be i.i.d. samples from probability distribution with parameter (or vectors of parameters) θ . The **maximum likelihood estimator (MLE)** $\hat{\theta}_{MLE}$ for θ is the value of θ that maximizes the likelihood/log-likelihood: $\hat{\theta}_{MLE} = \arg \max_{\theta} \mathcal{L}(x_1, \dots, x_n; \theta) = \arg \max_{\theta} \ln \mathcal{L}(x_1, \dots, x_n; \theta)$.

Unbiased Estimators

An estimator $\hat{\theta}$ of a parameter θ is **unbiased** if $\mathbb{E}[\hat{\theta}] = \theta$.

Consistent Estimators

An estimator $\hat{\theta}$ of a parameter θ is **consistent** if $\lim_{n \rightarrow \infty} \mathbb{E}[\hat{\theta}] = \theta$. The MLE is consistent (under some very mild assumptions).