

Section 8

Review

1) Multivariate: Discrete to Continuous:

	Discrete	Continuous
Joint PMF/PDF	$p_{X,Y}(x,y) = \mathbb{P}(X = x, Y = y)$	$f_{X,Y}(x,y) \neq \mathbb{P}(X = x, Y = y)$
Joint range/support $\Omega_{X,Y}$	$\{(x,y) \in \Omega_X \times \Omega_Y : p_{X,Y}(x,y) > 0\}$	$\{(x,y) \in \Omega_X \times \Omega_Y : f_{X,Y}(x,y) > 0\}$
Joint CDF	$F_{X,Y}(x,y) = \sum_{t \leq x, s \leq y} p_{X,Y}(t,s)$	$F_{X,Y}(x,y) = \int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(t,s) ds dt$
Normalization	$\sum_{x,y} p_{X,Y}(x,y) = 1$	$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy = 1$
Marginal PMF/PDF	$p_X(x) = \sum_y p_{X,Y}(x,y)$	$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$
Expectation	$\mathbb{E}[g(X,Y)] = \sum_{x,y} g(x,y) p_{X,Y}(x,y)$	$\mathbb{E}[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{X,Y}(x,y) dx dy$
Independence must have	$\forall x,y, p_{X,Y}(x,y) = p_X(x)p_Y(y)$ $\Omega_{X,Y} = \Omega_X \times \Omega_Y$	$\forall x,y, f_{X,Y}(x,y) = f_X(x)f_Y(y)$ $\Omega_{X,Y} = \Omega_X \times \Omega_Y$
Conditional PMF/PDF	$p_{X Y}(x y) = \frac{p_{X,Y}(x,y)}{p_Y(y)}$	$f_{X Y}(x y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$
Conditional Expectation	$\mathbb{E}[X Y = y] = \sum_x x \cdot p_{X Y}(x y)$	$\mathbb{E}[X Y = y] = \int_{-\infty}^{\infty} x f_{X Y}(x y) dx$

2) Normal (Gaussian, "bell curve"): $X \sim \mathcal{N}(\mu, \sigma^2)$ iff X has the following probability density function:

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}}, \quad x \in \mathbb{R}$$

$\mathbb{E}[X] = \mu$ and $\text{Var}(X) = \sigma^2$. The "standard normal" random variable is typically denoted Z and has mean 0 and variance 1: if $X \sim \mathcal{N}(\mu, \sigma^2)$, then $Z = \frac{X-\mu}{\sigma} \sim \mathcal{N}(0, 1)$. The CDF has no closed form, but we denote the CDF of the standard normal as $\Phi(z) = F_Z(z) = \mathbb{P}(Z \leq z)$. Note from symmetry of the probability density function about $z = 0$ that: $\Phi(-z) = 1 - \Phi(z)$.

3) Central Limit Theorem (CLT):

Let X_1, \dots, X_n be iid random variables with $\mathbb{E}[X_i] = \mu$ and $\text{Var}(X_i) = \sigma^2$. Let $X = \sum_{i=1}^n X_i$, which has $\mathbb{E}[X] = n\mu$ and $\text{Var}(X) = n\sigma^2$. Let $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$, which has $\mathbb{E}[\bar{X}] = \mu$ and $\text{Var}(\bar{X}) = \frac{\sigma^2}{n}$. \bar{X} is called the *sample mean*. Then, as $n \rightarrow \infty$, \bar{X} approaches the normal distribution $\mathcal{N}(\mu, \frac{\sigma^2}{n})$. Standardizing, this is equivalent to $Y = \frac{\bar{X}-\mu}{\sigma/\sqrt{n}}$ approaching $\mathcal{N}(0, 1)$. Similarly, as $n \rightarrow \infty$, X approaches $\mathcal{N}(n\mu, n\sigma^2)$ and $Y' = \frac{X-n\mu}{\sigma\sqrt{n}}$ approaches $\mathcal{N}(0, 1)$.

It is no surprise that \bar{X} has mean μ and variance σ^2/n – this can be done with simple calculations. The importance of the CLT is that, for large n , regardless of what distribution X_i comes from, \bar{X} is *approximately normally distributed with mean μ and variance σ^2/n* . Don't forget the continuity correction, only when X_1, \dots, X_n are discrete random variables.

Here is the **Standard normal table**.

4) Uniform: $X \sim \text{Uniform}(a, b)$ iff X has the following probability density function:

$$f_X(x) = \begin{cases} \frac{1}{b-a} & \text{if } x \in [a, b] \\ 0 & \text{otherwise} \end{cases}$$

$\mathbb{E}[X] = \frac{a+b}{2}$ and $\text{Var}(X) = \frac{(b-a)^2}{12}$. This represents each real number from $[a, b]$ to be equally likely.

5) Exponential: $X \sim \text{Exponential}(\lambda)$ iff X has the following probability density function:

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

$\mathbb{E}[X] = \frac{1}{\lambda}$ and $\text{Var}(X) = \frac{1}{\lambda^2}$. $F_X(x) = 1 - e^{-\lambda x}$ for $x \geq 0$. The exponential random variable is the continuous analog of the geometric random variable: it represents the waiting time to the next event, where $\lambda > 0$ is the average number of events per unit time. Note that the exponential measures how much time passes until the next event (any real number, continuous), whereas the Poisson measures how many events occur in a unit of time (nonnegative integer, discrete). The exponential random variable X is memoryless:

$$\text{for any } s, t \geq 0, \mathbb{P}(X > s + t \mid X > s) = \mathbb{P}(X > t)$$

The geometric random variable also has this property.

Task 1 – Joint PMF's

Suppose X and Y have the following joint PMF:

X/Y	1	2	3
0	0	0.2	0.1
1	0.3	0	0.4

- Identify the range of X (Ω_X), the range of Y (Ω_Y), and their joint range ($\Omega_{X,Y}$).
- Find the marginal PMF for X , $p_X(x)$ for $x \in \Omega_X$.
- Find the marginal PMF for Y , $p_Y(y)$ for $y \in \Omega_Y$.
- Are X and Y independent? Why or why not?
- Find $\mathbb{E}[X^3Y]$.

Task 2 – Do You “Urn” to Learn More About Probability?

Suppose that 3 balls are chosen without replacement from an urn consisting of 5 white and 8 red balls. Let $X_i = 1$ if the i -th ball selected is white and let it be equal to 0 otherwise. Give the joint probability mass function of

- X_1, X_2
- X_1, X_2, X_3

Task 3 – Trinomial Distribution

A generalization of the Binomial model is when there is a sequence of n independent trials, but with three outcomes, where $\mathbb{P}(\text{outcome } i) = p_i$ for $i = 1, 2, 3$ and of course $p_1 + p_2 + p_3 = 1$. Let X_i be the number of times outcome i occurred for $i = 1, 2, 3$, where $X_1 + X_2 + X_3 = n$. Find the joint PMF $p_{X_1, X_2, X_3}(x_1, x_2, x_3)$ and specify its value for all $x_1, x_2, x_3 \in \mathbb{R}$.

Are X_1 and X_2 independent?

Task 4 – Successes

Consider a sequence of independent Bernoulli trials, each of which is a success with probability p . Let X_1 be the number of failures preceding the first success, and let X_2 be the number of failures between the first 2 successes. Find the joint pmf of X_1 and X_2 . Write an expression for $E[\sqrt{X_1 X_2}]$. You can leave your answer in the form of a sum.

Task 5 – Who fails first?

Here's a question that commonly comes up in industry, but isn't immediately obvious. You have a disk with probability p_1 of failing each day. You have a CPU which independently has probability p_2 of failing each day. What is the probability that your disk fails *before* your CPU?

- Compute the probability by summing over the relevant part of the probability space.
- Try to provide an intuitive reason for the answer.
- Recompute the probability using the law of total probability, conditioning on the value of X_1 .

Task 6 – Continuous joint density

The joint density of X and Y is given by

$$f_{X,Y}(x,y) = \begin{cases} xe^{-(x+y)} & x > 0, y > 0 \\ 0 & \text{otherwise.} \end{cases}$$

and the joint density of W and V is given by

$$f_{W,V}(w,v) = \begin{cases} 2 & 0 < w < v, 0 < v < 1 \\ 0 & \text{otherwise.} \end{cases}$$

Are X and Y independent? Are W and V independent?

Task 7 – Law of Total Probability Review

- (Discrete version) Suppose we flip a coin with probability U of heads, where U is equally likely to be one of $\Omega_U = \{0, \frac{1}{n}, \frac{2}{n}, \dots, 1\}$ (notice this set has size $n+1$). Let H be the event that the coin comes up heads. What is $\mathbb{P}(H)$?
- Now suppose $U \sim \text{Uniform}(0,1)$ has the *continuous* uniform distribution over the interval $[0,1]$. What is $\mathbb{P}(H)$?

Use the continuous version of the law of total probability: suppose E is an event, and X is a continuous random variable with density function $f_X(x)$. Then

$$\mathbb{P}(E) = \int_{-\infty}^{\infty} \mathbb{P}(E | X = x) f_X(x) dx$$

Task 8 – Normal Approximation of a Sum

Imagine that we are trying to transmit a signal. During the transmission, there are a hundred sources independently making low noise. Each source produces an amount of noise that is Uniformly distributed between $a = -1$ and $b = 1$. If the total amount of noise is greater than 10 or less than -10 , then it corrupts the signal. However, if the absolute value of the total amount of noise is under 10, then it is not a problem. What is the approximate probability that the absolute value of the total amount of noise from the 100 signals is less than 10?

Task 9 – Confidence Intervals

Suppose that X_1, \dots, X_n are i.i.d. samples from a normal distribution with unknown mean μ and variance 36. How big does n need to be so that μ is in

$$[\bar{X} - 0.11, \bar{X} + 0.11]$$

with probability at least 0.97?

Recall that

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i.$$

You may use the fact that $\Phi^{-1}(0.985) = 2.17$.