Section 8
CSE312 23Wi

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Administrative

- Pset 6 due yesterday
- Pset 7 due **Wednesday March 1**

Review: Joint Distributions

1) Multivariate: Discrete to Continuous:

	Discrete	Continuous
Joint PMF/PDF	$p_{X,Y}(x,y) = \mathbb{P}(X = x, Y = y)$	$f_{X,Y}(x,y) \neq \mathbb{P}(X = x, Y = y)$
Joint range/support $\Omega_{X,Y}$	$\{(x,y) \in \Omega_X \times \Omega_Y : p_{X,Y}(x,y) > 0\}$	$\{(x,y) \in \Omega_X \times \Omega_Y : f_{X,Y}(x,y) > 0\}$
Joint CDF	$F_{X,Y}(x,y) = \sum_{t \leqslant x, s \leqslant y} p_{X,Y}(t,s)$	$F_{X,Y}(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f_{X,Y}(t,s)\, ds dt$
Normalization	$\sum_{x,y} p_{X,Y}(x,y) = 1$	$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y)\, dx dy = 1$
Marginal PMF/PDF	$p_X(x) = \sum_y p_{X,Y}(x,y)$	$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy$
Expectation	$\mathbb{E}[g(X,Y)] = \sum_{x,y} g(x,y) p_{X,Y}(x,y)$	$\mathbb{E}[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{X,Y}(x,y) dx dy$
Independence must have	$\forall x, y, p_{X,Y}(x,y) = p_X(x) p_Y(y)$ $\Omega_{X,Y} = \Omega_X \times \Omega_Y$	$\forall x, y, f_{X,Y}(x,y) = f_X(x) f_Y(y)$ $\Omega_{X,Y} = \Omega_X \times \Omega_Y$
Conditional PMF/PDF	$p_{X\|Y}(x\|y) = \frac{p_{X,Y}(x,y)}{p_Y(y)}$	$f_{X\|Y}(x\|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$
Conditional Expectation	$\mathbb{E}[X\|Y = y] = \sum_x x \cdot p_{X\|Y}(x\|y)$	$\mathbb{E}[X\|Y = y] = \int_{-\infty}^{\infty} x f_{X\|Y}(x\|y) dx$

Review: Central Limit Theorem

3) Central Limit Theorem (CLT): Let $X_1, \ldots, X_n$ be iid random variables with $\mathbb{E}[X_i] = \mu$ and $Var(X_i) = \sigma^2$. Let $X = \sum_{i=1}^{n} X_i$, which has $\mathbb{E}[X] = n\mu$ and $Var(X) = n\sigma^2$. Let $\overline{X} = \frac{1}{n}\sum_{i=1}^{n} X_i$, which has $\mathbb{E}[\overline{X}] = \mu$ and $Var(\overline{X}) = \frac{\sigma^2}{n}$. $\overline{X}$ is called the *sample mean*. Then, as $n \to \infty$, $\overline{X}$ approaches the normal distribution $\mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$. Standardizing, this is equivalent to $Y = \frac{\overline{X}-\mu}{\sigma/\sqrt{n}}$ approaching $\mathcal{N}(0,1)$. Similarly, as $n \to \infty$, X approaches $\mathcal{N}(n\mu, n\sigma^2)$ and $Y' = \frac{X-n\mu}{\sigma\sqrt{n}}$ approaches $\mathcal{N}(0,1)$.

It is no surprise that $\overline{X}$ has mean μ and variance σ^2/n – this can be done with simple calculations. The importance of the CLT is that, for large n, regardless of what distribution X_i comes from, $\overline{X}$ is *approximately normally distributed with mean μ and variance σ^2/n*. Don't forget the continuity correction, only when $X_1, \ldots, X_n$ are discrete random variables.

Here is the **Standard normal table.**

Problem 1: Joint PMFs

Suppose X and Y have the following joint PMF:

X/Y	1	2	3
0	0	0.2	0.1
1	0.3	0	0.4

a) Identify the range of X (Ω_X), the range of Y (Ω_Y), and their joint range ($\Omega_{X,Y}$).

b) Find the marginal PMF for X, $p_X(x)$ for $x \in \Omega_X$.

c) Find the marginal PMF for Y, $p_Y(y)$ for $y \in \Omega_Y$.

d) Are X and Y independent? Why or why not?

e) Find $\mathbb{E}[X^3Y]$.

Problem 1 Solution: a)

a) Identify the range of X (Ω_X), the range of Y (Ω_Y), and their joint range ($\Omega_{X,Y}$).

$\Omega_X = \{0, 1\}$, $\Omega_Y = \{1, 2, 3\}$, and $\Omega_{X,Y} = \{(0, 2), (0, 3), (1, 1), (1, 3)\}$

Problem 1 Solution: b)

b) Find the marginal PMF for X, $p_X(x)$ for $x \in \Omega_X$.

$$p_X(0) = \sum_y p_{X,Y}(0,y) = 0 + 0.2 + 0.1 = 0.3$$

$$p_X(1) = 1 - p_X(0) = 0.7$$

Problem 1 Solution: c)

c) Find the marginal PMF for Y, $p_Y(y)$ for $y \in \Omega_Y$.

$$p_Y(1) = \sum_x p_{X,Y}(x,1) = 0 + 0.3 = 0.3$$

$$p_Y(2) = \sum_x p_{X,Y}(x,2) = 0.2 + 0 = 0.2$$

$$p_Y(3) = \sum_x p_{X,Y}(x,3) = 0.1 + 0.4 = 0.5$$

Problem 1 Solution: d)

d) Are X and Y independent? Why or why not?

No, since a necessary condition is that $\Omega_{X,Y} = \Omega_X \times \Omega_Y$.

Problem 1 Solution: e)

e) Find $\mathbb{E}[X^3Y]$.

Note that $X^3 = X$ since X takes values in $\{0, 1\}$.

$$\mathbb{E}[X^3Y] = \mathbb{E}[XY] = \sum_{(x,y)\in\Omega_{X,Y}} xyp_{X,Y}(x,y) = 1 \cdot 1 \cdot 0.3 + 1 \cdot 3 \cdot 0.4 = 1.5$$

Problem 9: Confidence Intervals

Suppose that $X_1, \ldots, X_n$ are i.i.d. samples from a normal distribution with unknown mean μ and variance 36. How big does n need to be so that μ is in

$$[\overline{X} - 0.11, \overline{X} + 0.11]$$

with probability at least 0.97?

Recall that

$$\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_i.$$

You may use the fact that $\Phi^{-1}(0.985) = 2.17$.

Problem 9 Solution

Our goal is to find n such that μ lies within 0.11 of $\bar{X}$ 97% of the time. This is equivalent to finding n such that the probability that μ lies outside the range is less than 3%.

$$\mathbb{P}(|\bar{X} - \mu| > 0.11) \leqslant 0.03$$

Problem 9 Solution

Let us define $Z = \frac{\bar{X}-\mu}{\sigma}$. We can solve for σ by using the Properties of Variance. Since

$$\bar{X} = \frac{1}{n}\sum_{i=1}^{n} X_i$$

we can say that

$$\text{Var}(\bar{X}) = \text{Var}(\frac{1}{n}\sum_{i=1}^{n} X_i)$$

Using the Properties of Variance and the fact that X_i's are i.i.d., $\text{Var}(\bar{X}) = \frac{1}{n^2} \cdot n \cdot 36 = \frac{36}{n}$, so $\sigma = \frac{6}{\sqrt{n}}$.

Problem 9 Solution

$$\mathbb{P}(|\bar{X} - \mu| > 0.11) \leqslant 0.03$$

$$\mathbb{P}(|Z| \cdot \sigma > 0.11) \leqslant 0.03 \quad \text{[Definition of } Z\text{]}$$

Problem 9 Solution

$$\mathbb{P}(|\bar{X} - \mu| > 0.11) \leqslant 0.03$$

$$\mathbb{P}(|Z| \cdot \sigma > 0.11) \leqslant 0.03 \qquad \text{[Definition of } Z\text{]}$$

$$\mathbb{P}\left(|Z| > \frac{0.11}{6}\sqrt{n}\right) \leqslant 0.03$$

$$\mathbb{P}\left(Z < -\frac{0.11}{6}\sqrt{n}\right) \leqslant 0.015 \qquad \text{[Symmetry of Normal Dist.]}$$

$$\Phi\left(-\frac{0.11}{6}\sqrt{n}\right) \leqslant 0.015 \qquad \text{[CDF of Standard Norm.]}$$

Problem 9 Solution

$$
\begin{aligned}
-\frac{0.11}{6}\sqrt{n} &\leqslant -\Phi^{-1}(0.985) \\
\sqrt{n} &\geqslant \frac{6 \cdot \Phi^{-1}(0.985)}{0.11} \\
n &\geqslant \left(\frac{6 \cdot \Phi^{-1}(0.985)}{0.11}\right)^2 \\
&\approx 14009.95
\end{aligned}
$$

Therefore the final answer should be 14010.

Problem 6: Continuous Joint Density

The joint density of X and Y is given by

$$f_{X,Y}(x,y) = \begin{cases} xe^{-(x+y)} & x > 0, y > 0 \\ 0 & \text{otherwise.} \end{cases}$$

and the joint density of W and V is given by

$$f_{W,V}(w,v) = \begin{cases} 2 & 0 < w < v, 0 < v < 1 \\ 0 & \text{otherwise.} \end{cases}$$

Are X and Y independent? Are W and V independent?

Problem 6 Solution

For two random variables X, Y to be independent, we must have $f_{X,Y}(x,y) = f_X(x) f_Y(y)$ for all $x \in \Omega_X,\ y \in \Omega_Y$. Let's start with X and Y by finding their marginal PDFs. By definition, and using the fact that the joint PDF is 0 outside of $y > 0$, we get:

$$f_X(x) = \int_0^\infty x e^{-(x+y)} dy = e^{-x} x$$

Problem 6 Solution

We do the same to get the PDF of Y, again over the range $x > 0$:

$$f_Y(y) = \int_0^\infty xe^{-(x+y)}dx = e^{-y}$$

Since $e^{-x}x \cdot e^{-y} = xe^{-x-y} = xe^{-(x+y)}$ for all $x, y > 0$, X and Y are independent.

Problem 6 Solution

We can see that W and V are not independent simply by observing that $\Omega_W = (0,1)$ and $\Omega_V = (0,1)$, but $\Omega_{W,V}$ is not equal to their Cartesian product. Specifically, looking at their range of $f_{W,V}(w,v)$. Graphing it with w as the "x-axis" and v as the "y-axis", we see that :