### CSE 312 Foundations of Computing II

#### Lecture 19: Maximum Likelihood Estimation



#### **Aleks Jovcic**

Slide Credit: Based on Stefano Tessaro's slides for 312 19au incorporating ideas from Anna Karlin, Alex Tsun, Rachel Lin, Hunter Schafer & myself ©

#### **Final Pset**

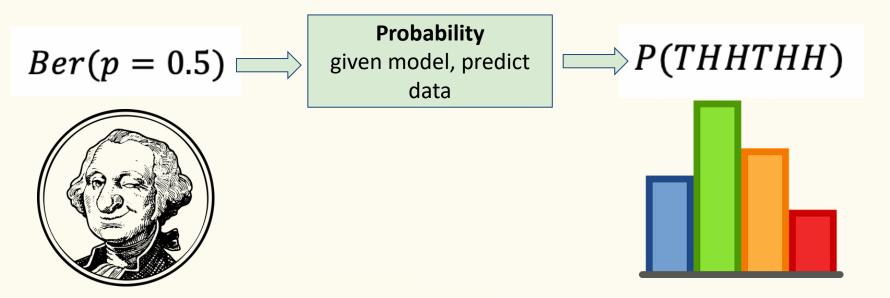
- Slightly longer, slightly harder, less time to work
- Released Tuesday, August 16<sup>th</sup> at 11:59pm PST
- Due Friday, August 19<sup>th</sup> at 11:59pm PST
  - No late days can be spent!
  - If something comes up, please let me know as soon as possible
- Individual, but working and studying together is encouraged
  - No office hours during this time
  - Prepare and go to office hours ahead of time
  - There will be a form to find classmates to work with as needed
  - Remember that you do not need to typeset if that will take up too much time
- TA-led review session, in-class review, TBA

#### Agenda

- Idea: Estimation
- Maximum Likelihood Estimation
- MLE with continuous random variables
- General Steps

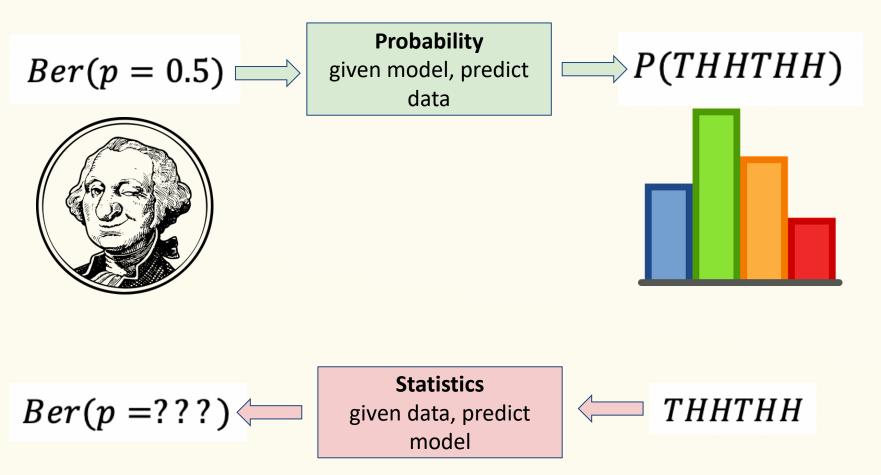
#### **Probability vs statistics**



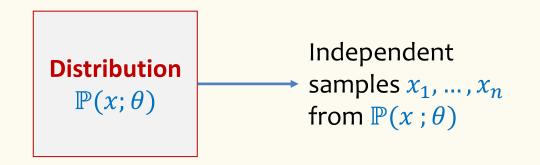


#### **Probability vs statistics**





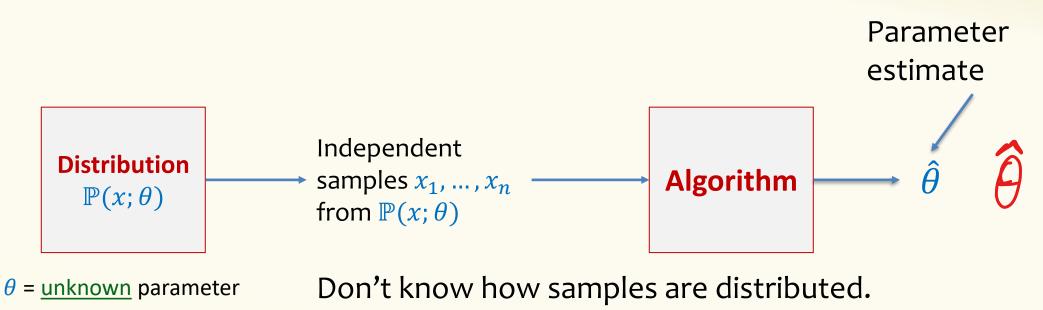
#### **Probability: Viewpoint up to Now**



 $\theta = \underline{known}$  parameter

 $\theta$  tells us how samples are distributed.  $\mathbb{P}(x; \theta)$  viewed as a function of x (fixed  $\theta$ )

#### **Statistics: Parameter Estimation – Workflow**



#### Example

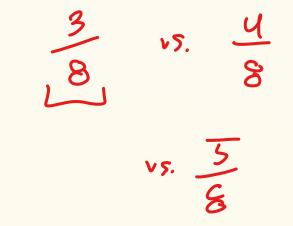
#### NBe (0)

Suppose we have a mystery coin with some probability if of coming up heads. We flip the coin 8 times, independent of other flips and see the following sequence. of flips

#### TTHTHTTH

Given this data,	, what would you	estimate <i>p</i> is?
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Ро	ll:				
а.	1/2				
<i>b.</i>	5/8				
С.	3/8				
<i>d.</i>	1/4				



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#### Likelihood

Say we see outcome HHTHH.

You tell me your best guess about the value of the unknown parameter  $\theta$  (aka p) is 4/5. Is there some way that you can argue "objectively" that this is the best estimate?

#### Likelihood

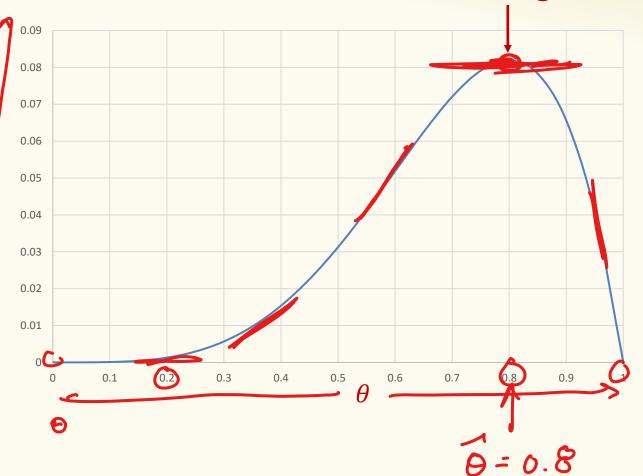


#### **Max Prob of seeing HHTHH**

Say we see outcome HHTHH.

You tell me your best guess about the value of the unknown parameter  $\theta$  (aka p) is 4/5. Is there some way that you can argue "objectively" that this is the best estimate?

$$\mathcal{L}(HHTHH \mid \theta) = \theta^{4}(1-\theta)$$



# Likelihood of Different Observations (Discrete case) $\int_{L} \int_{L} \int_{L$

 $= \mathcal{P}(X, 10) \cdot \mathcal{P}(X, i0) \cdot \dots \cdot \mathcal{P}(X, i0)$ 

(Discrete case)

**Definition.** The **likelihood** of independent observations  $x_1, \dots, x_n$  is  $\mathcal{L}(x_1, \dots, x_n | \theta) = \prod_{i=1}^n \mathbb{P}(x_i; \theta)$ 

**Maximum Likelihood Estimation (MLE).** Given data  $x_1, ..., x_n$ , find  $\hat{\theta}$  ("the MLE") of model such that  $L(x_1, ..., x_n | \hat{\theta})$  is maximized!  $\hat{\theta} = \operatorname*{argmax}_{\theta} \mathcal{L}(x_1, ..., x_n | \theta)$ 

#### Likelihood vs. Probability

A **probability function**  $Pr(x; \theta)$  is a function with input being an event x for some fixed probability model (w/ param  $\theta$ ).

 $\sum_{x} \Pr(x; \theta) = 1$ 

A likelihood function  $\mathcal{L}(x | \theta)$  is a function with input being  $\theta$  (the param of the prob. Model) for some fixed dataset x.

These notions are very closely connected, but answer different questions. We are trying to find the  $\theta$  that maximizes likelihood, thus we are looking for the **maximum likelihood estimator**.

#### **Example – Coin Flips**

Observe: Coin-flip outcomes  $x_1, ..., x_n$ , with  $n_H$  heads,  $n_T$  tails  $- \text{l.e.}, n_H + n_T = n$  **Goal:** estimate  $\theta$  = prob. heads.  $L(x_1, ..., x_n | \theta) = \theta^{n_H} (1 - \theta)^{n_T}$  $= \int_{i=1}^{n_H} \mathbb{P}(x_i; \theta) = \int_{i=1}^{n_H} \mathbb{P}(x_i; \theta)$ 

While it is not difficult to compute this derivative, we make our lives easier by observing that we are always taking a derivative of a product....

#### Log-Likelihood

We can save some work if we work with the **log-likelihood** instead of the likelihood directly.

a > b (n(a) > b (b)

**Definition.** The **log-likelihood** of independent observations  $x_1, \dots, x_n$  is  $\mathcal{LL}(x_1, \dots, x_n | \theta) = \ln \mathcal{L}(x_1, \dots, x_n | \theta)$  $= \ln \prod_{i=1}^n \mathbb{P}(x_i; \theta) = \sum_{i=1}^n \ln \mathbb{P}(x_i; \theta)$ 

Useful log properties

$$\log(ab) = \log(a) + \log(b) \\ \log(a/b) = \log(a) - \log(b) \\ \log(a^b) = blog(a)$$

**Example – Coin Flips**  $l_{n}(a') = 5l_{n}(a)$ KILT LIK ln: loge Observe: Coin-flip outcomes  $x_1, \dots, x_n$ , with  $n_H$  heads,  $n_T$  tails  $-1.e., n_H + n_T = n$ **Goal:** estimate  $\theta$  = prob. heads.  $\frac{n_{+1}}{\tilde{\Theta}} - \frac{n_{-1}}{1 - \tilde{\Theta}} = 0$  $\mathcal{L}(x_1,\ldots,x_n|\theta) = \underline{\theta}^{n_H}(1-\theta)^{n_T}$  $\frac{r_{H}}{2} = \hat{\Theta}$  $\ln \mathcal{L}(x_1, \dots, x_n | \theta) = \mathcal{L}(\theta'') + \mathcal{L}(((-\theta))')$  $\frac{n_{r'}}{n_{r'}} \sim \frac{n_{r'}}{n_{r'}}$ = n, ln (0) + n- ln (1.0)  $n_{\rm FI}(1-\hat{\Theta}) = n_{\rm T}\hat{\Theta}$  $\frac{\partial}{\partial \theta} \mathcal{L}((X_{1},..,Y_{n}|\theta)) = \frac{n_{H}}{\theta} - \frac{n_{T}}{1-\theta}$  $\Lambda_{FI} - \Lambda_T \tilde{\Theta}$  $U^{+1} = (V^{\perp} \rightarrow V^{\perp}) \stackrel{\sim}{\Theta}$ 

#### **Example – Coin Flips**

Observe: Coin-flip outcomes  $x_1, ..., x_n$ , with  $n_H$  heads,  $n_T$  tails - I.e.,  $n_H + n_T = n$ Goal: estimate  $\theta$  = prob. heads.

$$\mathcal{L}(x_1,\ldots,x_n|\theta) = \theta^{n_H}(1-\theta)^{n_T}$$

$$\ln \mathcal{L}(x_1, \dots, x_n | \theta) = n_H \ln \theta + n_T \ln(1 - \theta)$$
  

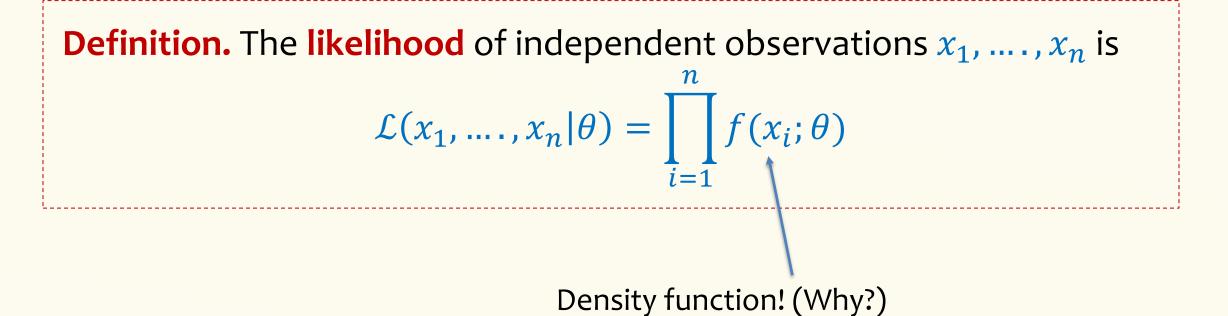
$$\frac{\partial}{\partial \theta} \ln \mathcal{L}(x_1, \dots, x_n | \theta) = n_H \cdot \frac{1}{\theta} - n_T \cdot \frac{1}{1 - \theta}$$
  
Solve  $n_H \cdot \frac{1}{\theta} - n_T \cdot \frac{1}{1 - \theta} = 0$  -----

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#### **The Continuous Case**

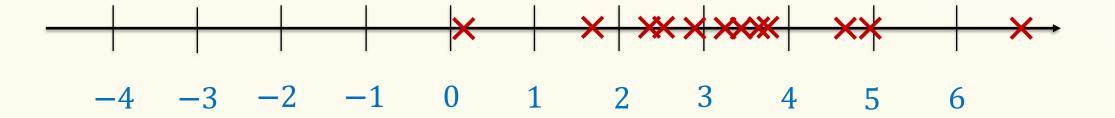
## Given *n* samples $x_1, ..., x_n$ from a Gaussian $\mathcal{N}(\mu, \sigma^2)$ , estimate $\theta = (\mu, \sigma^2)$



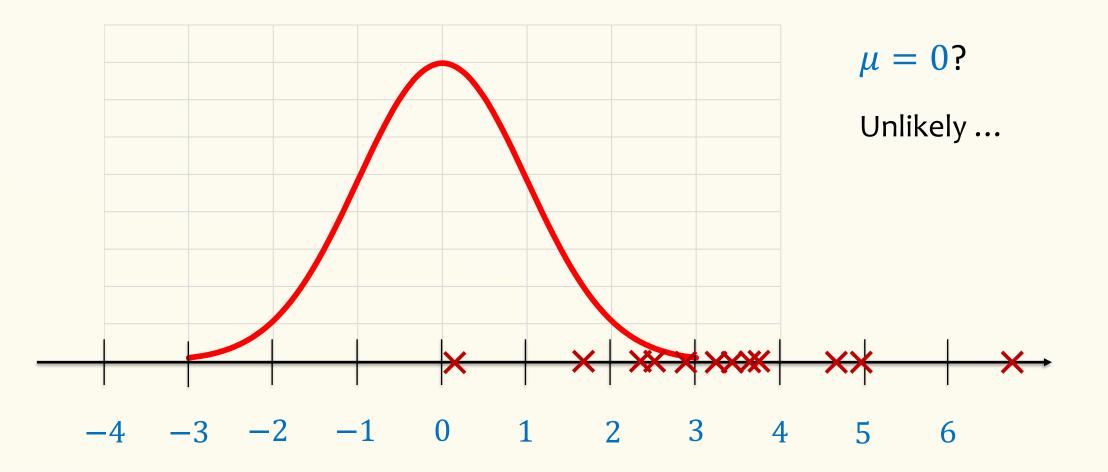
#### Why density?

- Density ≠ probability, but:
  - For maximizing likelihood, we really only care about relative likelihoods, and density captures that
  - has desired property that likelihood increases with better fit to the model

*n* samples  $x_1, ..., x_n \in \mathbb{R}$  from Gaussian  $\mathcal{N}(\mu, 1)$ . <u>Most likely</u>  $\mu$ ? [i.e., we are given the <u>promise</u> that the variance is one]

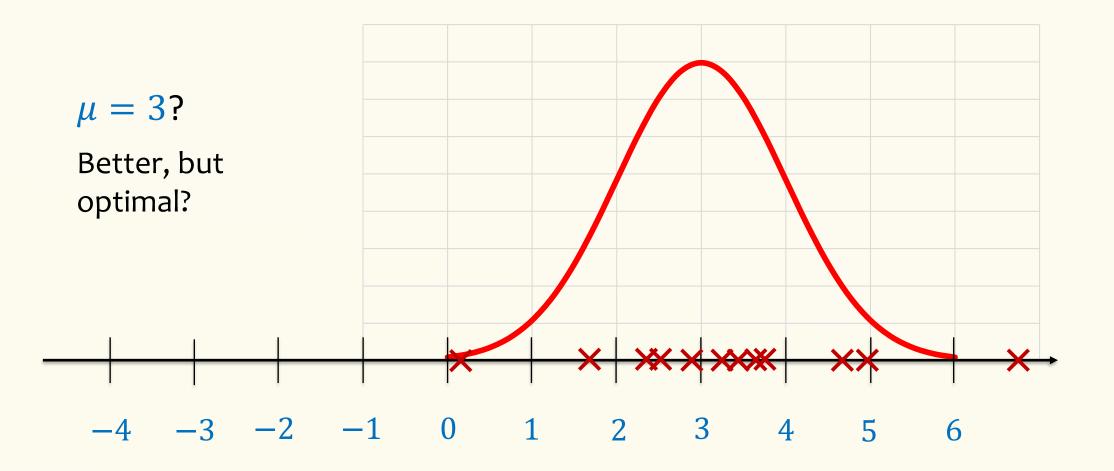


*n* samples  $x_1, \ldots, x_n \in \mathbb{R}$  from Gaussian  $\mathcal{N}(\mu, 1)$ . Most likely  $\mu$ ?



23

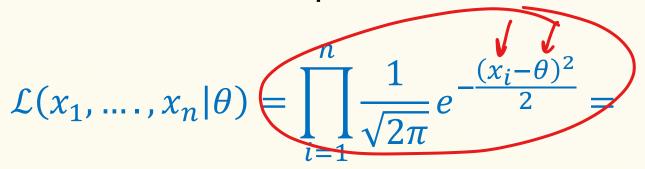
*n* samples  $x_1, \ldots, x_n \in \mathbb{R}$  from Gaussian  $\mathcal{N}(\mu, 1)$ . Most likely  $\mu$ ?



24

Normal outcomes  $x_1, \dots, x_n$ , known variance  $\sigma^2 = 1$ 

**Goal:** estimate  $\theta$  expectation



$$log(ab) = log(a) + log(b)$$
  

$$log(a/b) = log(a) - log(b)$$
  

$$log(ab) = blog(a) 25$$

Normal outcomes  $x_1, ..., x_n$ , known variance  $\sigma^2 = 1$ 

**Goal:** estimate  $\theta$  expectation

$$\mathcal{L}(x_1, \dots, x_n | \theta) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} e^{-\frac{(x_i - \theta)^2}{2}} = \left(\frac{1}{\sqrt{2\pi}}\right)^n \prod_{i=1}^n e^{-\frac{(x_i - \theta)^2}{2}}$$
$$\ln \mathcal{L}(x_1, \dots, x_n | \theta) = -n \frac{\ln 2\pi}{2} - \sum_{i=1}^n \frac{(x_i - \theta)^2}{2}$$

**Goal:** estimate  $\theta$  = expectation

Normal outcomes  $x_1, ..., x_n$ , known variance  $\sigma^2 = 1$ 

$$\ln \mathcal{L}(x_1, \dots, x_n | \theta) = -n \frac{\ln 2\pi}{2} - \sum_{i=1}^n \frac{(x_i - \theta)^2}{2}$$

**Goal:** estimate  $\theta$  = expectation

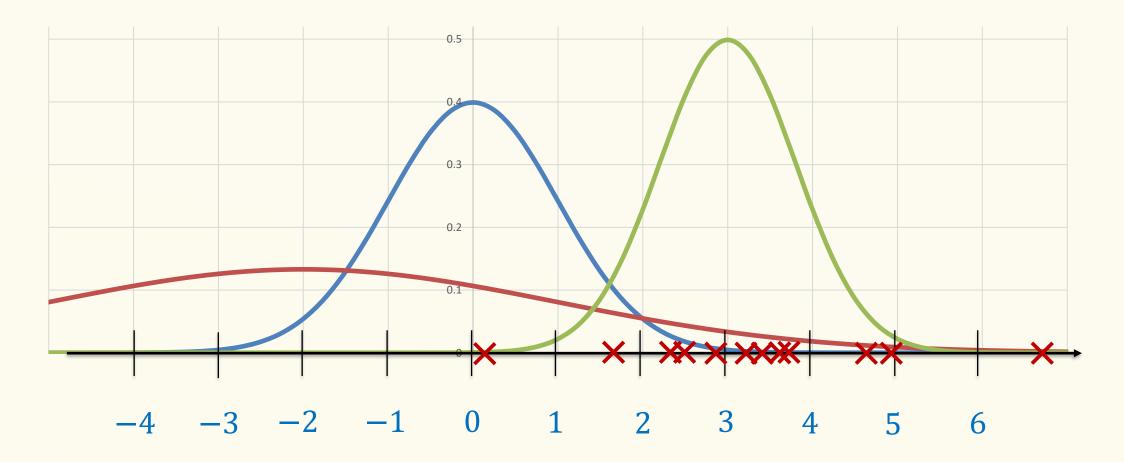
Normal outcomes  $x_1, ..., x_n$ , known variance  $\sigma^2 = 1$ 

$$\ln \mathcal{L}(x_1, \dots, x_n | \theta) = -n \frac{\ln 2\pi}{2} - \sum_{i=1}^n \frac{(x_i - \theta)^2}{2}$$
  
lote:  $\frac{\partial}{\partial \theta} \frac{(x_i - \theta)^2}{2} = \frac{1}{2} \cdot 2 \cdot (x_i - \theta) \cdot (-1) = \theta - x_i$   
 $\frac{\partial}{\partial \theta} \ln \mathcal{L}(x_1, \dots, x_n | \theta) = \sum_{i=1}^n (x_i - \theta) = \sum_{i=1}^n x_i - n\theta = 0$ 

$\hat{ heta} =$	$\sum_{i}^{n} x_{i}$
	$\overline{n}$

In other words, MLE is the sample mean of the data.

**Next steps:** *n* samples  $x_1, ..., x_n \in \mathbb{R}$  from Gaussian  $\mathcal{N}(\mu, \sigma^2)$ . Most likely  $\mu$  and  $\sigma^2$ ?



29

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#### **General Recipe**

1. Input Given *n* iid samples  $x_1, ..., x_n$  from parametric model with parameters  $\theta$ .

- 2. Likelihood Define your likelihood  $\mathcal{L}(x_1, \dots, x_n | \theta)$ .
  - For discrete  $\mathcal{L}(x_1, \dots, x_n | \theta) = \prod_{i=1}^n \Pr(x_i; \theta)$
  - For continuous  $\mathcal{L}(x_1, \dots, x_n | \theta) = \prod_{i=1}^n f(x_i; \theta)$
- 3. Log Compute  $\ln \mathcal{L}(x_1, \dots, x_n | \theta)$
- 4. Differentiate Compute  $\frac{\partial}{\partial \theta} \ln \mathcal{L}(x_1, \dots, x_n | \theta)$
- 5. Solve for  $\hat{\theta}$  by setting derivative to 0 and solving for max.

Generally, you need to do a second derivative test to verify it is a maximum, but we won't ask you to do that in CSE 312.

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