## CSE 312

## Foundations of Computing II

## Lecture 17: The Normal Distribution

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Slide Credit: Based on Stefano Tessaro's slides for 312 19au incorporating ideas from Anna Karlin, Alex Tsun, Rachel Lin, Hunter Schafer \& myself ©

## Agenda

- Law of Total Expectation (LTE) Practice
- Chebyshev's Inequality
- The Normal Distribution
- Practice with Normals

Example: Flipping Coins LTE: $E[X]=\sum_{j=1}^{n} E[X \mid A ;)$ $y=2$
suppose wanted g analyze flipping a random number of coins. Suppose someone $\mathrm{A}_{i}$ gave us $Y \sim$ Poi ( 5 ) fair coins and
heads $X$ from flipping those coins.

$$
\begin{aligned}
& E[X]=0.5 ? ? \\
& \begin{array}{l}
E[X \mid Y=2]=0.52 \\
\sum_{y=0}^{\infty} 0.5 y \cdot e^{-5} \cdot \frac{5^{\varphi}}{y!}
\end{array}
\end{aligned}
$$

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## Using variance

- If we know more about the random variable, e.g. its variance, we can get a better bound!

Markov's inequality

## Chebyshev's Inequality

Theorem. Let $X$ be a random variable. Then, for any $t>0$,

$$
\mathbb{P}(|X-\mathbb{E}(X)| \geq t) \leq \frac{\operatorname{Var}(X)}{t^{2}}
$$

Proof: Define $Z=X-\mathbb{E}(X)$

$$
E(X)=\left[(X]^{7}\right. \text { Definition of Variance }
$$ 2

$$
\begin{gathered}
\mathbb{P}(|Z| \geq t)=\mathbb{P}\left(Z^{2} \geq t^{2}\right) \leq \frac{\mathbb{E}\left(Z^{2}\right)}{t^{2}}=\frac{t^{2}}{\operatorname{Var}(X)} \\
|Z| \geq t \text { iff } Z^{2} \geq t^{2} \quad \text { Markov's inequality }\left(Z^{2} \geq 0\right)
\end{gathered}
$$

## Example - Binomial Random Variable

Chebychev's Inequality $\mathbb{P}(|X-\mathbb{E}(X)| \geq t) \leq \frac{\operatorname{Var}(X)}{t^{2}}$.

Let $X$ be Binomial RV with parameters. $n, p=0.5$

$$
\mathbb{E}(X)=\frac{n}{2}
$$

$$
\ln \operatorname{lar}(x)=\frac{3 n}{n}-\frac{n}{4}
$$

What is the probability that $X \geq \frac{3 n}{4}$ ?
Chebychev's inequality: $\mathbb{P}\left(x \geq \frac{3 n}{4}\right)=\mathbb{S}\left(x-\frac{n}{2} \geq \frac{n}{4}\right) \stackrel{\left(\frac{n}{n}\right)^{2}}{\leftrightarrows}$

Markov's inequality: $\mathbb{P}\left(X \geq \frac{3 n}{4}\right) \leq \frac{4}{3 n} \cdot \frac{n}{2}=\frac{2}{3}$

$$
\mathbb{P}\left(\frac{\left|X-\frac{n}{2}\right| \geq \frac{n}{4}}{7}\right)
$$

## Tail Bounds

Useful for approximations of complex systems. How good the approximation is depends on the actual distribution and the context you are using it in.

- Usually loose upper-bounds are okay when designing for worstcase

Generally, the more you know about your random variable the better tail bounds you can get.

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## Review - Continuous RVs

Probability Density Function (PDF).
$f: \mathbb{R} \rightarrow \mathbb{R}$ s.t.

- $f(x) \geq 0$ for all $x \in \mathbb{R}$


Cumulative Density Function (CDF).

$$
F(y)=\int_{-\infty}^{y} f(x) \mathrm{d} x
$$

- $\int_{-\infty}^{+\infty} f(x) \mathrm{d} x=1$


Density $\neq$ Probability !
$=X \leq b)=F_{F(\nu)}(b)-F_{X}(c)$

$$
F(y)=\mathbb{P}(X \leq y)
$$

## Review - Continuous RVs



$$
\mathbb{P}(X \in[a, b])=\int_{a}^{b} f_{X}(x) \mathrm{d} x=F_{X}(b)-F_{X}(a)
$$

## Exponential Distribution

Definition. An exponential random variable $X$ with parameter $\lambda \geq 0$ is follows the exponential density

$$
f_{X}(x)=\left\{\begin{array}{cc}
\lambda e^{-\lambda x} & x \geq 0 \\
0 & x<0
\end{array}\right.
$$

We write $X \sim \operatorname{Exp}(\lambda)$ and say $X$ that follows the exponential distribution.

$$
\begin{aligned}
& \text { CDF: For } y \geq 0, \\
& F_{X}(y)=1-e^{-\lambda y}
\end{aligned}
$$



## The Normal Distribution

Definition. A Gaussian (or normal) random variable with parameters $\mu \in \mathbb{R}$ and $\sigma \geq 0$ has density

$$
f_{X}(x)=\frac{1}{\sqrt{2 \pi} \sigma} e^{-\frac{(x-\mu)^{2}}{2 \sigma^{2}}}
$$

(We say that $X$ follows the Normal Distribution, and write $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$ )

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$$
f_{X}(x)=\frac{1}{\sqrt{2 \pi} \sigma} e^{-\frac{(x-\mu)^{2}}{2 \sigma^{2}}}
$$

(We say that $X$ follows the Normal Distribution, and write $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$ )
Fact. If $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, then $\mathbb{E}(X)=\mu$, and $\operatorname{Var}(X)=\sigma^{2}$
Expectation follows from density being symmetric around $\mu, f_{X}(\mu-x)=f_{X}(\mu+x)$

The Normal Distribution
Aka a "Bell Curve" (imprecise name)


## Shifting and Scaling - turning one normal dist into another

Fact. If $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, then $Y=a X+b \sim \mathcal{N}\left(a \mu+b, a^{2} \sigma^{2}\right)$


Note: $\frac{X-\mu}{\sigma} \sim \underline{\mathcal{N}(0,1)}$

## CDF of normal distribution

Fact. If $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, then $Y=a X+b \sim \mathcal{N}\left(a \mu+b, a^{2} \sigma^{2}\right)$

Standard (unit) normal $\mathbb{Z} \mathcal{N}(0,1)$
CDF. $\Phi(z)=\mathbb{P}(Z \leq z)=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{Z} e^{-x^{2} / 2} \mathrm{~d} x$ for $Z \sim \mathcal{N}(0,1)$
Note: $\Phi(z)$ has no closed form - generally given via tables

$$
P(z)=P(Z \leq 2)
$$

Table of $\Phi(z)$ CDF of

## Standard Normal Distn

## Make sure to use the one linked on the site!



| $\Phi \text { Table: } \mathbb{P}(Z \leq z) \text { when } Z \sim \mathcal{N}(0,1)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $z$ | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| 0.0 | 0.5 | 0.50399 | 0.50798 | 0.51197 | 0.51595 | 0.51994 | 0.52392 | 0.5279 | 0.53188 | 0.53586 |
| 0.1 | 0.53983 | 0.5438 | 0.54776 | 0.55172 | 0.55567 | 0.55962 | 0.56356 | 0.56749 | 0.57142 | 0.57535 |
| 0.2 | 0.57926 | 0.58317 | 0.58706 | 0.59095 | 0.59483 | 0.59871 | 0.60257 | 0.60642 | 0.61026 | 0.61409 |
| 0.3 | 0.6170 | . 62172 | 069552 | 0.6293 | 0.63307 | 0.63683 | 0.64058 | 0.64431 | 0.64803 | 0.65173 |
|  | 0.65542 | . 65 | 0.66276) | D. 6664 | 0.67003 | 0.67364 | 0.67724 | 0.68082 | 0.68439 | 0.68793 |
| 0.5 | 0.60146 | 0.69497 | cooer | 0.70194 | 0.7054 | 0.70884 | 0.71226 | 0.71566 | 0.71904 | 0.7224 |
| 0.6 | 0.72575 | 0.72907 | 0.73237 | 0.73565 | 0.73891 | 0.74215 | 0.74537 | 0.74857 | 0.75175 | 0.7549 |
| 0.7 | 0.75804 | 0.76115 | 0.76424 | 0.7673 | 0.77035 | 0.77337 | 0.77637 | 0.77935 | 0.7823 | 0.78524 |
| 0.8 | 0.78814 | 0.79103 | 0.79389 | 0.79673 | 0.79955 | 0.80234 | 0.80511 | 0.80785 | 0.81057 | 0.81327 |
| 0.9 | 0.81594 | 0.81859 | 0.82121 | 0.82381 | 0.82639 | 0.82894 | 0.83147 | 0.83398 | 0.83646 | 0.83891 |
| 1.0 | 0.84134 | 0.84375 | 0.84614 | 0.84849 | 0.85083 | 0.85314 | 0.85543 | 0.85769 | 0.85993 | 0.86214 |
| 1.1 | 0.86433 | 0.8665 | 0.86864 | 0.87076 | 0.87286 | 0.87493 | 0.87698 | 0.879 | 0.881 | 0.88298 |
| 1.2 | 0.88493 | 0.88686 | 0.88877 | 0.89065 | 0.89251 | 0.89435 | 0.89617 | 0.89796 | 0.89973 | 0.90147 |
| 1.3 | 0.9032 | 0.9049 | 0.90658 | 0.90824 | 0.90988 | 0.91149 | 0.91309 | 0.91466 | 0.91621 | 0.91774 |
| 1.4 | 0.91924 | 0.92073 | 0.9222 | 0.92364 | 0.92507 | 0.92647 | 0.92785 | 0.92922 | 0.93056 | 0.93189 |
| 1.5 | 0.93319 | 0.93448 | 0.93574 | 0.93699 | 0.93822 | 0.93943 | 0.94062 | 0.94179 | 0.94295 | 0.94408 |
| 1.6 | 0.9452 | 0.9463 | 0.94738 | 0.94845 | 0.9495 | 0.95053 | 0.95154 | 0.95254 | 0.95352 | 0.95449 |
| 1.7 | 0.95543 | 0.95637 | 0.95728 | 0.95818 | 0.95907 | 0.95994 | 0.9608 | 0.96164 | 0.96246 | 0.96327 |
| 1.8 | 0.96407 | 0.96485 | 0.96562 | 0.96638 | 0.96712 | 0.96784 | 0.96856 | 0.96926 | 0.96995 | 0.97062 |
| 1.9 | 0.97128 | 0.97193 | 0.97257 | 0.9732 | 0.97381 | 0.97441 | 0.975 | 0.97558 | 0.97615 | 0.9767 |
| 20 | 0.97725 | 0.97778 | 0.97831 | 0.97882 | 0.97932 | 0.97982 | 0.9803 | 0.98077 | 0.98124 | 0.98169 |
| 21 | 0.88214 | 0.98257 | 0.983 | 0.98341 | 0.98382 | 0.98422 | 0.98461 | 0.985 | 0.98537 | 0.98574 |
| 2.2 | 0.9861 | 0.98645 | 0.98679 | 0.98713 | 0.98745 | 0.98778 | 0.98809 | 0.9884 | 0.9887 | 0.98899 |
| 2.8 | 0.98928 | 0.98956 | 0.98983 | 0.9901 | 0.99036 | 0.99061 | 0.99086 | 0.99111 | 0.99134 | 0.99158 |
| 2.4 | 0.9918 | 0.99202 | 0.99224 | 0.99245 | 0.99266 | 0.99286 | 0.99305 | 0.99324 | 0.99343 | 0.99361 |
| 2.5 | 0.99379 | 0.99396 | 0.99413 | 0.9943 | 0.99446 | 0.99461 | 0.99477 | 0.99492 | 0.99506 | 0.9952 |
| 2.6 | 0.99534 | 0.99547 | 0.9956 | 0.99573 | 0.99585 | 0.99598 | 0.99609 | 0.99621 | 0.99632 | 0.99643 |
| 2.7 | 0.99653 | 0.99664 | 0.99674 | 0.99683 | 0.99693 | 0.99702 | 0.99711 | 0.9972 | 0.99728 | 0.99736 |
| 2.8 | 0.99744 | 0.99752 | 0.9976 | 0.99767 | 0.99774 | 0.99781 | 0.99788 | 0.99795 | 0.99801 | 0.99807 |
| 2.9 | 0.99813 | 0.99819 | 0.99825 | 0.99831 | 0.99836 | 0.99841 | 0.99846 | 0.99851 | 0.99856 | 0.99861 |
| 3.0 | 0.99865 | 0.99869 | 0.99874 | 0.99878 | 0.99882 | 0.99886 | 0.99889 | 0.99893 | 0.99896 | 0.999 |

$$
\begin{aligned}
& \begin{array}{lll}
\text { Example } & \mu=0.4 \\
\text { Let } X \sim N(0.4,4) .
\end{array} \sigma^{2}=4 \quad \frac{X-\mu}{\sigma} \\
& \mathbb{P}(x \leq 1.2)=\mathbb{P}\left(\frac{x-0.4}{2} \leq \frac{1.2-0.4}{2}\right) \\
& =\pi(z \leq 0.4) \\
& =\phi(0.4)=0.66542
\end{aligned}
$$

## Example

Let $X \sim \mathcal{N}\left(0.4,4=2^{2}\right)$.

$$
\begin{aligned}
& \mathbb{P}(X \leq 1.2)=\mathbb{P}\left(\frac{X-0.4}{2} \leq \frac{1.2-0.4}{2}\right)
\end{aligned}
$$

$$
\begin{aligned}
& \text { Example } \frac{5-3}{\sqrt{16}} \frac{2-3}{4} \\
& \operatorname{Let} X \sim \mathcal{N}(3,16) . \\
& \mathbb{P}(2<x<5)-\mathbb{P}\left(-\frac{1}{4} \leq \frac{x-3}{4} \leq \frac{1}{2}\right) \\
& \mathbb{P}(2 \geq 0.2 \overline{5}) \mathbb{P}\left(-\frac{1}{4} \leq 2 \leq-\frac{1}{2}\right) \quad \mathbb{P}(2 \leq-0.75) \\
& 1-\mathbb{P}(2 \leq 0.75) \phi(0.5)-\Phi(-0.25) \\
& \rightarrow(0.5)-(1-\phi(0.75))
\end{aligned}
$$

## Example

Let $X \sim \mathcal{N}(3,16)$.

$$
\begin{aligned}
\mathbb{P}(2<X<5) & =\mathbb{P}\left(\frac{2-3}{4}<\frac{X-3}{4}<\frac{5-3}{4}\right) \\
& =\mathbb{P}\left(-\frac{1}{4}<Z<\frac{1}{2}\right) \\
& =\Phi\left(\frac{1}{2}\right)-\Phi\left(-\frac{1}{4}\right) \\
& =\Phi\left(\frac{1}{2}\right)-\left(1-\Phi\left(\frac{1}{4}\right)\right) \approx 0.29017
\end{aligned}
$$

## Example - Off by Standard Deviations

Let $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$.
$\mathbb{P}(|X-\mu|<k \sigma)=$

## Example - Off by Standard Deviations

Let $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$.

$$
\begin{aligned}
\mathbb{P}(|X-\mu|<k \sigma) & =\mathbb{P}\left(\frac{|X-\mu|}{\sigma}<k\right)= \\
& =\mathbb{P}\left(-k<\frac{X-\mu}{\sigma}<k\right)=\Phi(k)-\Phi(-k)
\end{aligned}
$$

e.g. $k=1: 68 \%, k=2: 95 \%, k=3: 99 \%$

Summary of procedure for doing calculations with normal r.v.
If $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, then $\frac{X-\mu}{\sigma} \sim \mathcal{N}(0,1)$

Therefore,

$$
F_{X}(z)=\mathbb{P}(X \leq z)=\mathbb{P}\left(\frac{X-\mu}{\sigma} \leq \frac{z-\mu}{\sigma}\right)=\Phi\left(\frac{z-\mu}{\sigma}\right)
$$

## CDF of normal distribution

Fact. If $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, then $Y=a X+b \sim \mathcal{N}\left(a \mu+b, a^{2} \sigma^{2}\right)$

Standard (unit) normal $Z \sim \mathcal{N}(0,1)$
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If $X \sim \mathcal{N}\left(\mu, \sigma^{2}\right)$, then $F_{X}(z)=\mathbb{P}(X \leq z)=\mathbb{P}\left(\frac{X-\mu}{\sigma} \leq \frac{z-\mu}{\sigma}\right)=\Phi\left(\frac{z-\mu}{\sigma}\right)$

## Closure of the normal -- under addition

Fact. If $X \sim \mathcal{N}\left(\mu_{X}, \sigma_{X}^{2}\right), \mathrm{Y} \sim \mathcal{N}\left(\mu_{Y}, \sigma_{Y}^{2}\right)$ (both independent normal RV) then $\mathrm{a} X+b Y+c \sim \mathcal{N}\left(a \mu_{X}+b \mu_{Y}+c, a^{2} \sigma_{X}^{2}+b^{2} \sigma_{Y}^{2}\right)$

Note: The special thing is that the sum of normal RVs is still a normal RV.

The values of the expectation and variance is not surprising.

- Linearity of expectation (always true)
- When $X$ and $Y$ are independent, $\operatorname{Var}(a X+b Y)=a^{2} \operatorname{Var}(X)+b^{2} \operatorname{Var}(Y)$

