

CSE 312

Foundations of Computing II

Lecture 8: Linearity of Expectation

Last Class:

- Random Variables
- Probability Mass Function (PMF)
- Cumulative Distribution Function (CDF)
- Expectation

Today:

- More Expectation Examples
- Linearity of Expectation
- Indicator Random Variables

Kandinsky

[This Photo](#) by Unknown Author is licensed under [CC BY-SA](#)



Review Random Variables

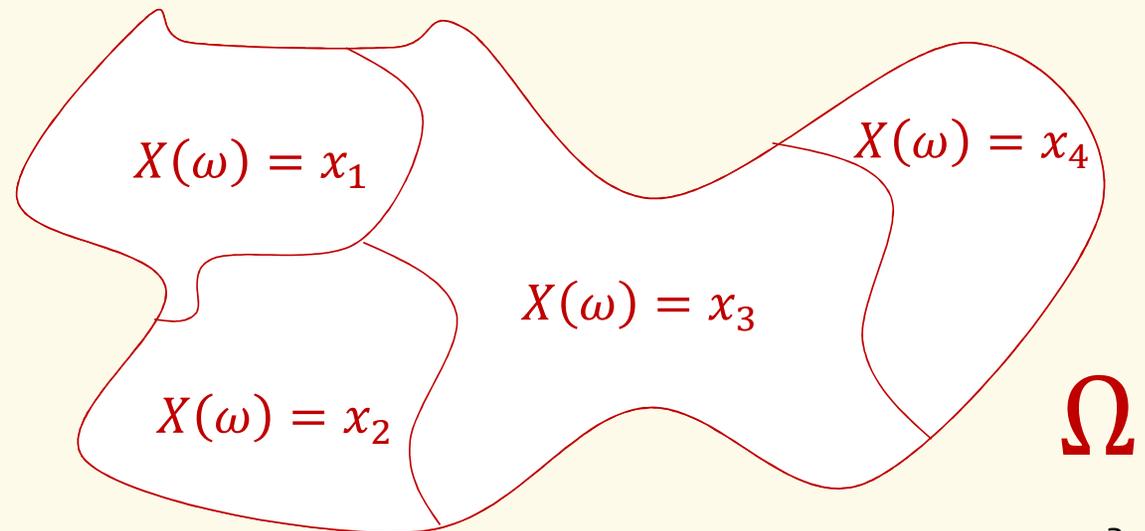
Definition. A **random variable (RV)** for a probability space (Ω, P) is a function $X: \Omega \rightarrow \mathbb{R}$.

The set of values that X can take on is its *range/support*: $X(\Omega)$ or Ω_X

$$\{X = x_i\} = \{\omega \in \Omega \mid X(\omega) = x_i\}$$

Random variables **partition**
the sample space.

$$\sum_{x \in X(\Omega)} P(X = x) = 1$$



Review PMF and CDF

Definitions:

For a RV $X: \Omega \rightarrow \mathbb{R}$, the **probability mass function (pmf)** of X specifies, for any real number x , the probability that $X = x$

$$p_X(x) = P(X = x) = P(\{\omega \in \Omega \mid X(\omega) = x\})$$

$$\sum_{x \in \Omega_X} p_X(x) = 1$$

For a RV $X: \Omega \rightarrow \mathbb{R}$, the **cumulative distribution function (cdf)** of X specifies, for any real number x , the probability that $X \leq x$

$$F_X(x) = P(X \leq x)$$

Review Expected Value of a Random Variable

Definition. Given a discrete RV $X: \Omega \rightarrow \mathbb{R}$, the **expectation** or **expected value** or **mean** of X is

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot P(\omega)$$

or equivalently

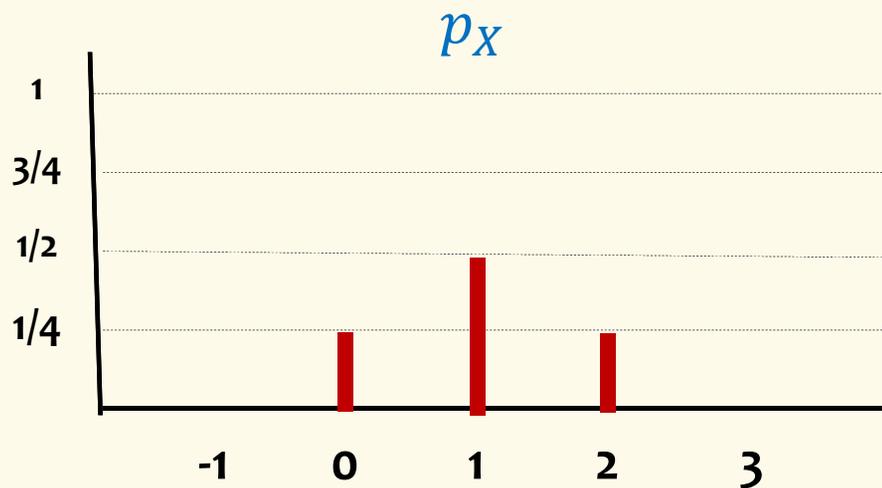
$$\mathbb{E}[X] = \sum_{x \in X(\Omega)} x \cdot P(X = x) = \sum_{x \in \Omega_X} x \cdot p_X(x)$$

Intuition: “Weighted average” of the possible outcomes (weighted by probability)

Expectation

Example. Two fair coin flips
 $\Omega = \{TT, HT, TH, HH\}$
 $X =$ number of heads

What is $\mathbb{E}[X]$?



$$\begin{aligned}\mathbb{E}[X] &= 0 \cdot p_X(0) + 1 \cdot p_X(1) + 2 \cdot p_X(2) \\ &= 0 \cdot \frac{1}{4} + 1 \cdot \frac{1}{2} + 2 \cdot \frac{1}{4} = \frac{1}{2} + \frac{1}{2} = 1\end{aligned}$$

0

6

Another Interpretation

“If X is how much you win playing the game in one round. How much would you expect to win, on average, per game, when repeatedly playing?”

Answer: $E[X]$

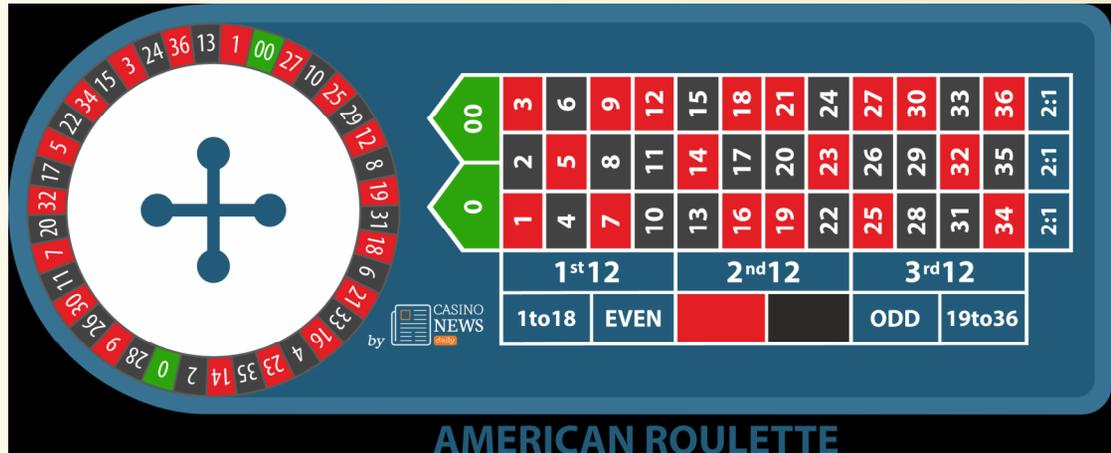
Roulette (USA)

Ω :

Numbers 1-36

- 18 Red
- 18 Black

Green 0 and 00



RVs for gains from some bets:

Note 0 and 00 are not EVEN

RV RED: If Red number turns up +1, if Black number, 0, or 00 turns up -1

$$\mathbb{E}[\text{RED}] = (+1) \cdot \frac{18}{38} + (-1) \cdot \frac{20}{38} = -\frac{2}{38} \approx -5.26\%$$

RV 1st12: If number 1-12 turns up +2, if number 13-36, 0, or 00 turns up -1

$$\mathbb{E}[1^{\text{st}}12] = (+2) \cdot \frac{12}{38} + (-1) \cdot \frac{26}{38} = -\frac{2}{38} \approx -5.26\%$$

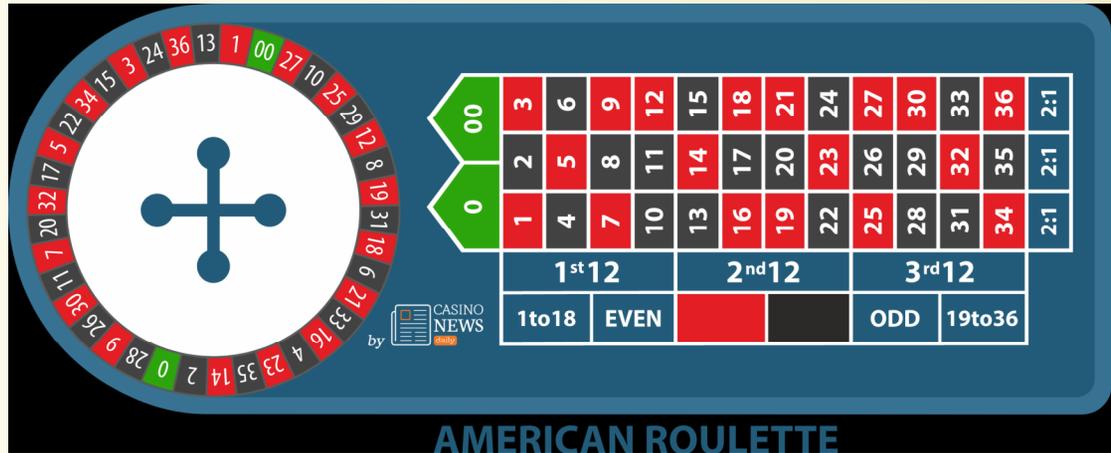
Roulette (USA)

Ω :

Numbers 1-36

- 18 Red
- 18 Black

Green 0 and 00



An even worse bet:

Note 0 and 00 are not EVEN

RV BASKET: If 0, 00, 1, 2, or 3 turns up +6 otherwise -1

$$\mathbb{E}[\text{BASKET}] = (+6) \cdot \frac{5}{38} + (-1) \cdot \frac{33}{38} = -\frac{3}{38} \approx -7.89\%$$

Example: Returning Homeworks

- Class with 3 students, randomly hand back homeworks. All permutations equally likely.
- Let X be the number of students who get their own HW

$\Pr(\omega)$	ω	$X(\omega)$
1/6	1, 2, 3	3
1/6	1, 3, 2	1
1/6	2, 1, 3	1
1/6	2, 3, 1	0
1/6	3, 1, 2	0
1/6	3, 2, 1	1

$$\begin{aligned}\mathbb{E}[X] &= 3 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} + 0 \cdot \frac{1}{6} + 0 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} \\ &= 6 \cdot \frac{1}{6} = 1\end{aligned}$$

Example – Flipping a biased coin until you see heads

- Biased coin:

$$P(H) = q > 0$$

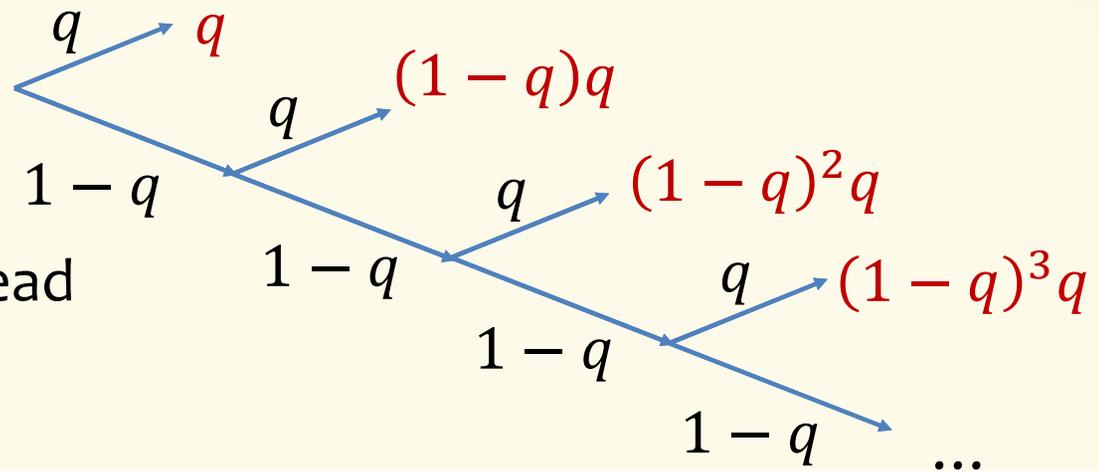
$$P(T) = 1 - q$$

- $Z = \#$ of coin flips until first head

$$P(Z = i) = q (1 - q)^{i-1}$$

$$\mathbb{E}[Z] = \sum_{i=1}^{\infty} i \cdot P(Z = i) = \sum_{i=1}^{\infty} i \cdot q(1 - q)^{i-1}$$

Converges, so $\mathbb{E}[Z]$ is finite



Can calculate this directly but...

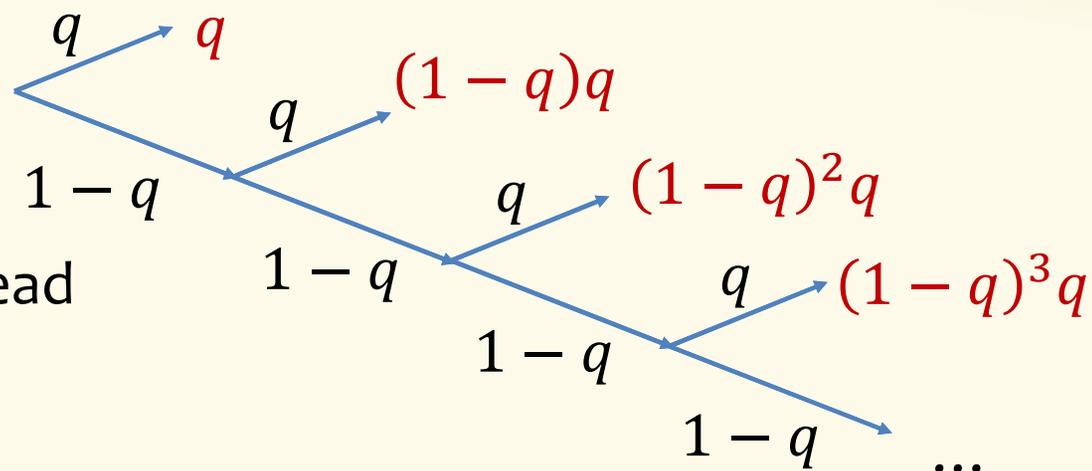
Example – Flipping a biased coin until you see heads

- Biased coin:

$$P(H) = q > 0$$

$$P(T) = 1 - q$$

- $Z = \#$ of coin flips until first head



Another view: If you get heads first try you get $Z = 1$;
If you get tails you have used one try and have the same experiment left

$$\mathbb{E}[Z] = q + (1 - q)(1 + \mathbb{E}[Z])$$

Solving gives $q \cdot \mathbb{E}[Z] = q + (1 - q) = 1$ Implies $\mathbb{E}[Z] = 1/q$

Expected Value of $X = \#$ of heads

Each coin shows up heads half the time.

Two fair coins



Glued coins



Attached coins



$$P(HT) = P(TH) = 0.25$$

$$P(HH) = P(TT) = 0.25$$

$$P(HT) = P(TH) = 0.5$$

$$P(HH) = P(TT) = 0$$

$$P(HH) = P(TT) = 0.4$$

$$P(HT) = P(TH) = 0.1$$

$$\mathbb{E}(X) = 1 \cdot \frac{1}{2} + 2 \cdot \frac{1}{4} = 1$$

$$\mathbb{E}(X) = 1 \cdot 1 = 1$$

$$\mathbb{E}(X) = 1 \cdot 0.2 + 2 \cdot 0.4 = 1$$

Linearity of Expectation

Theorem. For **any** two random variables X and Y
(X, Y do not need to be independent)

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y].$$

Or, more generally: For any random variables X_1, \dots, X_n ,

$$\mathbb{E}[X_1 + \dots + X_n] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n].$$

Because: $\mathbb{E}[X_1 + \dots + X_n] = \mathbb{E}[(X_1 + \dots + X_{n-1}) + X_n]$
 $= \mathbb{E}[X_1 + \dots + X_{n-1}] + \mathbb{E}[X_n] = \dots$

Linearity of Expectation – Proof

Theorem. For **any** two random variables X and Y
(X, Y do not need to be independent)

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y].$$

$$\begin{aligned}\mathbb{E}[X + Y] &= \sum_{\omega} P(\omega)(X(\omega) + Y(\omega)) \\ &= \sum_{\omega} P(\omega)X(\omega) + \sum_{\omega} P(\omega)Y(\omega) \\ &= \mathbb{E}[X] + \mathbb{E}[Y]\end{aligned}$$

Example – Coin Tosses

We flip n coins, each one heads with probability p
 Z is the number of heads, what is $\mathbb{E}(Z)$?

Example – Coin Tosses – The brute force method

We flip n coins, each one heads with probability p ,

Z is the number of heads, what is $\mathbb{E}[Z]$?

$$\begin{aligned}\mathbb{E}[Z] &= \sum_{k=0}^n k \cdot P(Z = k) = \sum_{k=0}^n k \cdot \binom{n}{k} p^k (1-p)^{n-k} \\ &= \sum_{k=0}^n k \cdot \frac{n!}{k! (n-k)!} p^k (1-p)^{n-k} = \sum_{k=1}^n \frac{n!}{(k-1)! (n-k)!} p^k (1-p)^{n-k} \\ &= np \sum_{k=1}^n \frac{(n-1)!}{(k-1)! (n-k)!} p^{k-1} (1-p)^{n-k} \\ &= np \sum_{k=0}^{n-1} \frac{(n-1)!}{k! (n-1-k)!} p^k (1-p)^{(n-1)-k} \\ &= np \sum_{k=0}^{n-1} \binom{n-1}{k} p^k (1-p)^{(n-1)-k} = np(p + (1-p))^{n-1} = np \cdot 1 = np\end{aligned}$$



This Photo by Unknown Author is licensed under [CC BY-NC](https://creativecommons.org/licenses/by-nc/4.0/)

Can we solve it more elegantly, please?

Computing complicated expectations

Often boils down to the following three steps:

- Decompose: Finding the right way to decompose the random variable into sum of simple random variables

$$X = X_1 + \cdots + X_n$$

- LOE: Apply linearity of expectation.

$$\mathbb{E}[X] = \mathbb{E}[X_1] + \cdots + \mathbb{E}[X_n].$$

- Conquer: Compute the expectation of each X_i

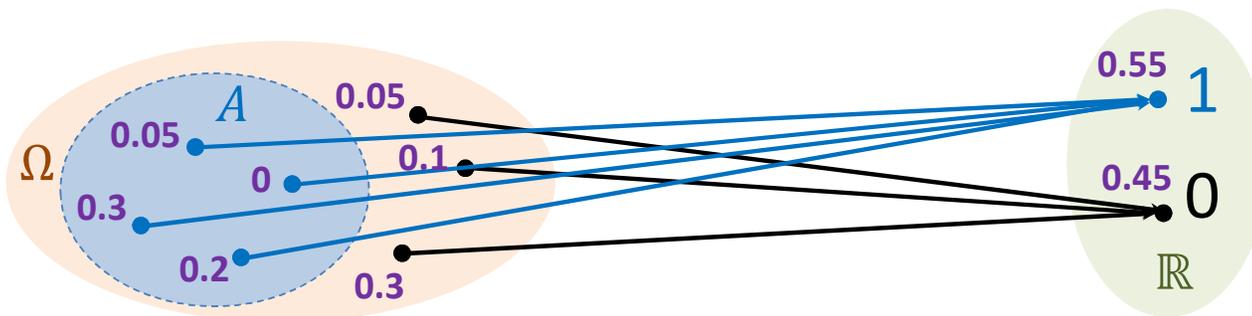
Often, X_i are **indicator** (0/1) random variables.

Indicator random variables

For any event A , can define the **indicator** random variable X_A for A

$$X_A = \begin{cases} 1 & \text{if event } A \text{ occurs} \\ 0 & \text{if event } A \text{ does not occur} \end{cases}$$

$$\begin{aligned} P(X_A = 1) &= P(A) \\ P(X_A = 0) &= 1 - P(A) \end{aligned}$$



Example – Coin Tosses

We flip n coins, each one heads with probability p
 Z is the number of heads, what is $\mathbb{E}[Z]$?

$$- X_i = \begin{cases} 1, & i^{\text{th}} \text{ coin flip is heads} \\ 0, & i^{\text{th}} \text{ coin flip is tails.} \end{cases}$$

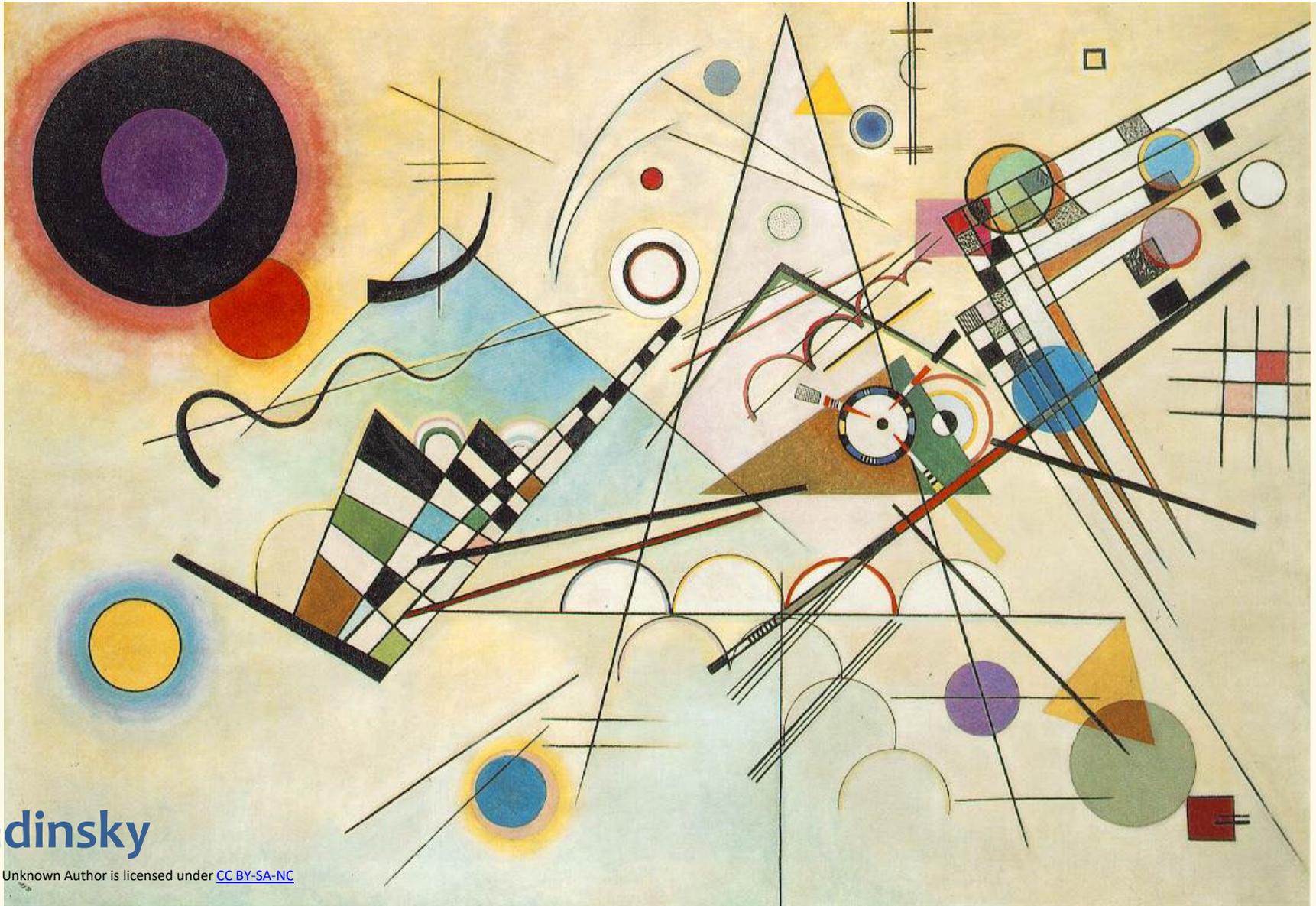
$$\text{Fact. } Z = X_1 + \dots + X_n$$

Linearity of Expectation:

$$\mathbb{E}[Z] = \mathbb{E}[X_1 + \dots + X_n] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n] = n \cdot p$$

$$\begin{aligned} P(X_i = 1) &= p \\ P(X_i = 0) &= 1 - p \end{aligned}$$

$$\mathbb{E}[X_i] = p \cdot 1 + (1 - p) \cdot 0 = p$$



Kandinsky

[This Photo](#) by Unknown Author is licensed under [CC BY-SA-NC](#)

Example: Returning Homeworks

- Class with n students, randomly hand back homeworks. All permutations equally likely.
- Let X be the number of students who get their own HW

What is $\mathbb{E}[X]$? Use linearity of expectation!

Decompose: What is X_i ?

$\Pr(\omega)$	ω	$X(\omega)$
1/6	1, 2, 3	3
1/6	1, 3, 2	1
1/6	2, 1, 3	1
1/6	2, 3, 1	0
1/6	3, 1, 2	0
1/6	3, 2, 1	1

$X_i = 1$ iff i^{th} student gets own HW back

LOE: $\mathbb{E}[X] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n]$

Conquer: What is $\mathbb{E}[X_i]$? A. $\frac{1}{n}$ B. $\frac{1}{n-1}$ C. $\frac{1}{2}$

Poll: pollev.com/paulbeame028

Pairs with the same birthday

- In a class of m students, on average how many pairs of people have the same birthday (assuming 365 equally likely birthdays)?

Decompose: Indicator events involve **pairs** of students (i, j) for $i \neq j$
 $X_{ij} = 1$ iff students i and j have the same birthday

LOE: $\binom{m}{2}$ indicator variables X_{ij}

Conquer: $\mathbb{E}[X_{ij}] = \frac{1}{365}$ so total expectation is $\frac{\binom{m}{2}}{365} = \frac{m(m-1)}{730}$ pairs

Linearity of Expectation – Even stronger

Theorem. For any random variables X_1, \dots, X_n , and real numbers $a_1, \dots, a_n \in \mathbb{R}$,

$$\mathbb{E}[a_1X_1 + \dots + a_nX_n] = a_1\mathbb{E}[X_1] + \dots + a_n\mathbb{E}[X_n].$$

Very important: In general, we do not have $\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$

Linearity is special!

In general $\mathbb{E}[g(X)] \neq g(\mathbb{E}(X))$

E.g., $X = \begin{cases} +1 & \text{with prob } 1/2 \\ -1 & \text{with prob } 1/2 \end{cases}$

Then: $\mathbb{E}[X^2] \neq \mathbb{E}[X]^2$

How DO we compute $\mathbb{E}[g(X)]$?

Expected Value of $g(X)$

Definition. Given a discrete RV $X: \Omega \rightarrow \mathbb{R}$, the **expectation** or **expected value** or **mean** of $g(X)$ is

$$\mathbb{E}[g(X)] = \sum_{\omega \in \Omega} g(X(\omega)) \cdot P(\omega)$$

or equivalently

$$\mathbb{E}[g(X)] = \sum_{x \in X(\Omega)} g(x) \cdot P(X = x) = \sum_{x \in \Omega_X} g(x) \cdot p_X(x)$$



Kandinsky

This Photo by Unknown Author is licensed under [CC BY-SA](#)