CSE 312

Foundations of Computing II

Lecture 14: Continuous RV



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Slide Credit: Based on Stefano Tessaro's slides for 312 19au incorporating ideas from Alex Tsun's and Anna Karlin's slides for 312 20su and 20au

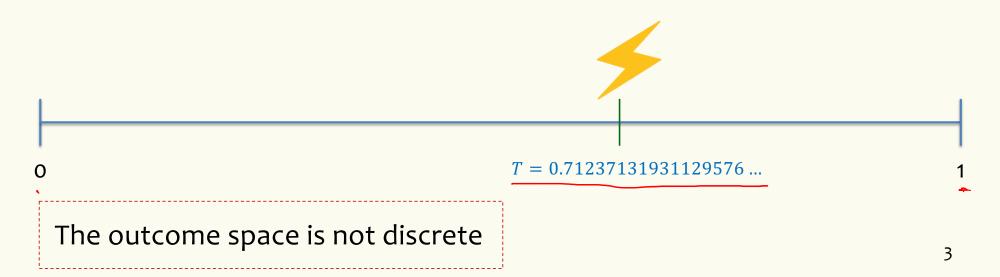
Agenda

- Continuous Random Variables
- Probability Density Function
- Cumulative Distribution Function

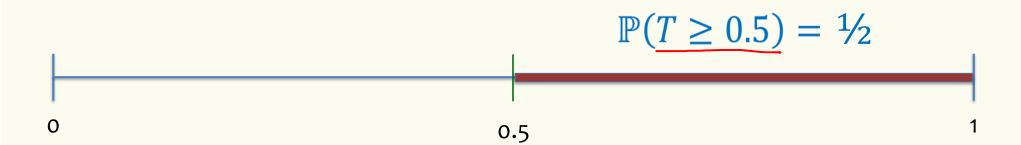
Often we want to model experiments where the outcome is <u>not</u> discrete.

Example – Lightning Strike

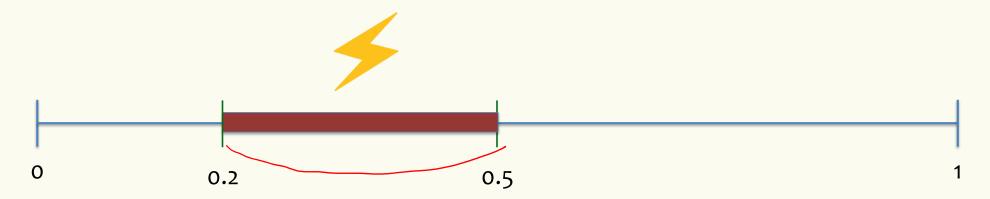
- T = time of lightning strike
- Every time within [0,1] is equally likely
 - Time measured with infinitesimal precision.



- T = time of lightning strike
- Every point in time within [0,1] is equally likely

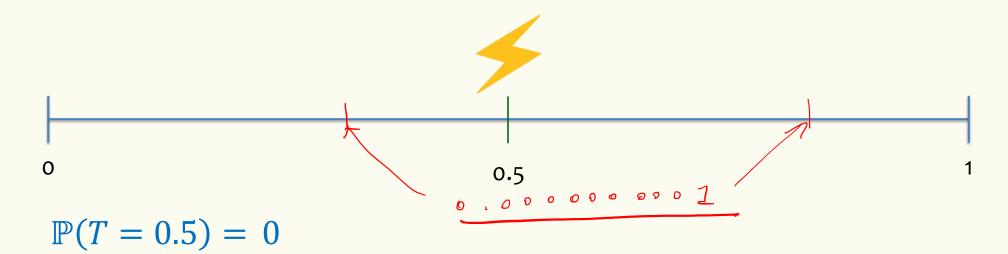


- T = time of lightning strike
- Every point in time within [0,1] is equally likely



$$\mathbb{P}(0.2 \le T \le 0.5) = 0.5 - 0.2 = 0.3$$

- T = time of lightning strike
- Every point in time within [0,1] is equally likely



Bottom line

- This gives rise to a different type of random variable
- $\mathbb{P}(T = x) = 0 \text{ for all } x \in [0,1]$
- Yet, somehow we want

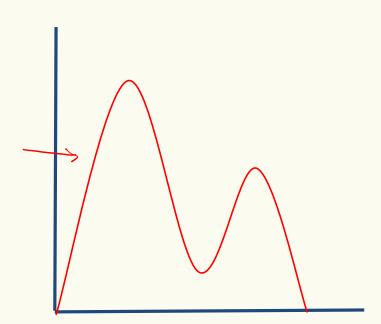
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(-\mathbb{P}(T \in [0,1]) = 1
-\mathbb{P}(T \in [a,b]) = b - a
-\dots
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• How do we model the behavior of T?

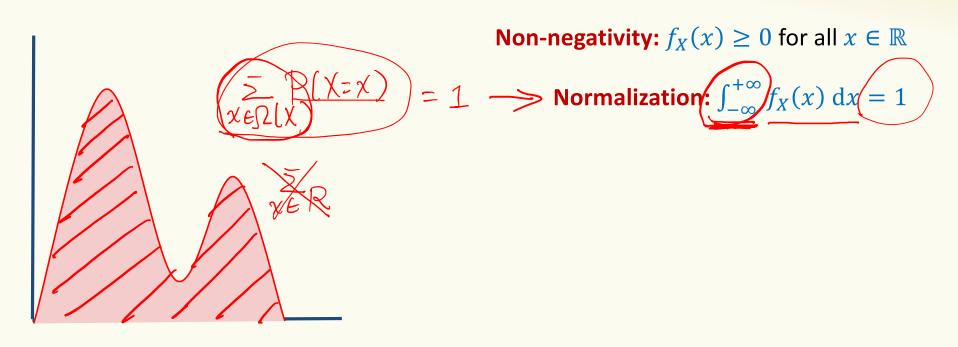
Definition. A continuous random variable X is defined by a probability density function (PDF) $f_X: \mathbb{R} \to \mathbb{R}$, such that

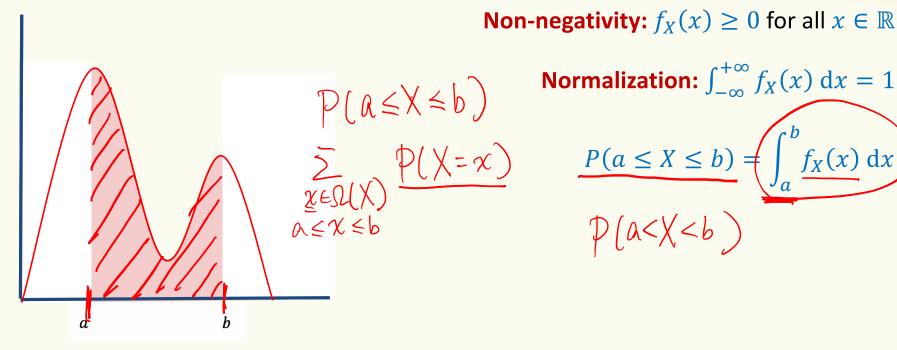
MASS

fx: x -> P(x=x)



Non-negativity: $f_X(x) \ge 0$ for all $x \in \mathbb{R}$



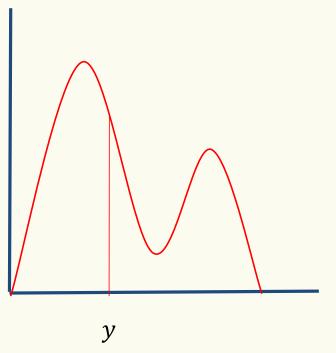


$$P(\alpha \leq X \leq b)$$
Normalization:
$$\int_{-\infty}^{+\infty} f_X(x) dx = 1$$

$$P(a \leq X \leq b) = \int_{a}^{b} f_X(x) dx$$

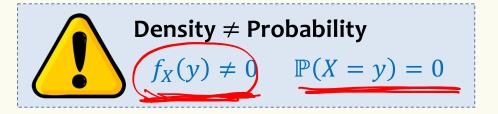
$$P(\alpha \leq X \leq b) = \int_{a}^{b} f_X(x) dx$$

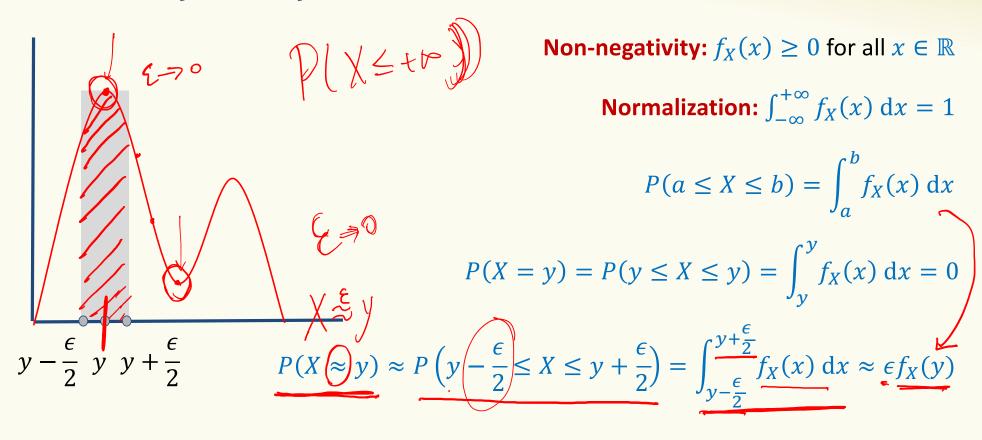
$$P(\alpha \leq X \leq b) = \int_{a}^{b} f_X(x) dx$$

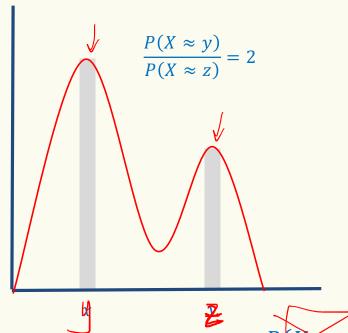


Non-negativity: $f_X(x) \ge 0$ for all $x \in \mathbb{R}$

$$P(X = y) = P(y \le X \le y) = \int_{y}^{b} f_{X}(x) dx$$







Non-negativity: $f_X(x) \ge 0$ for all $x \in \mathbb{R}$

$$P(a \le X \le b) = \int_{a}^{b} f_X(x) \, \mathrm{d}x$$

$$P(X = y) = P(y \le X \le y) = \int_{y}^{y} f_X(x) dx = 0$$

$$P(x = y) \approx P(y - \frac{\epsilon}{2} \le X \le y + \frac{\epsilon}{2}) = \int_{y - \frac{\epsilon}{2}}^{y + \frac{\epsilon}{2}} f_X(x) dx \approx \epsilon f_X(y)$$

$$P(X \approx y) \approx \frac{\langle f_X(y) \rangle}{\langle f_X(z) \rangle} = \frac{f_X(y)}{f_X(z)}$$

Definition. A continuous random variable X is defined by a **probability density function** (PDF) $f_X: \mathbb{R} \to \mathbb{R}$, such that

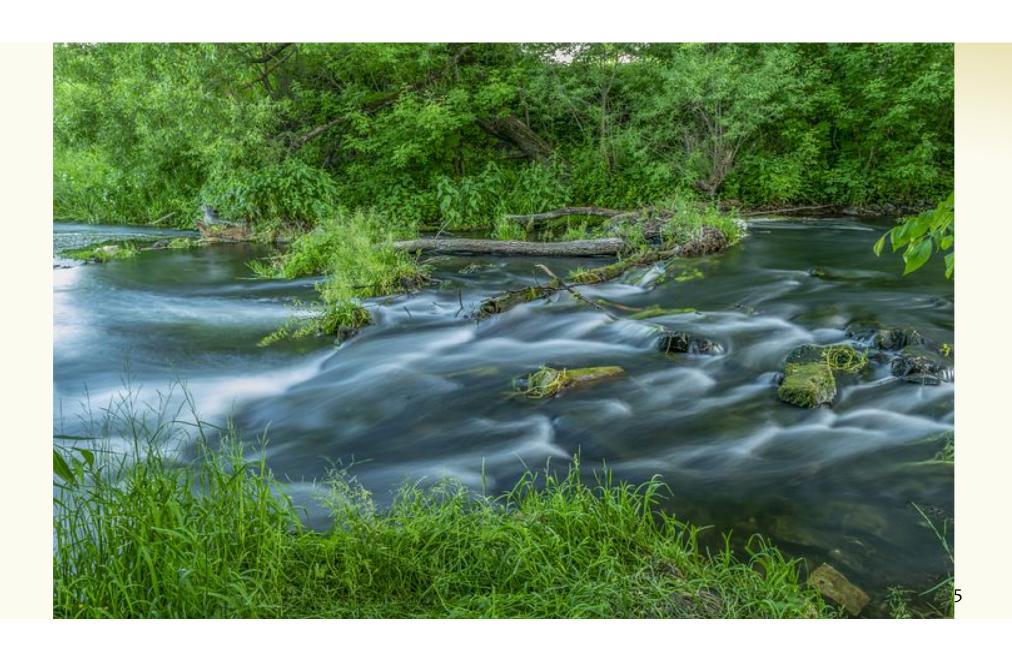
Non-negativity: $f_X(x) \ge 0$ for all $x \in \mathbb{R}$

$$P(a \le X \le b) = \int_{a}^{b} f_X(x) \, \mathrm{d}x$$

$$P(X = y) = P(y \le X \le y) = \int_{y}^{y} f_X(x) dx = 0$$

$$P(X \approx y) \approx P\left(y - \frac{\epsilon}{2} \le X \le y + \frac{\epsilon}{2}\right) = \int_{y - \frac{\epsilon}{2}}^{y + \frac{\epsilon}{2}} f_X(x) \, \mathrm{d}x \approx \epsilon f_X(y)$$

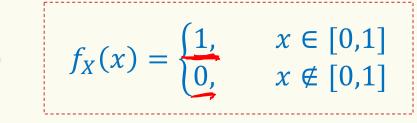
$$\frac{P(X \approx y)}{P(X \approx z)} \approx \frac{\epsilon f_X(y)}{\epsilon f_X(z)} = \frac{f_X(y)}{f_X(z)}$$



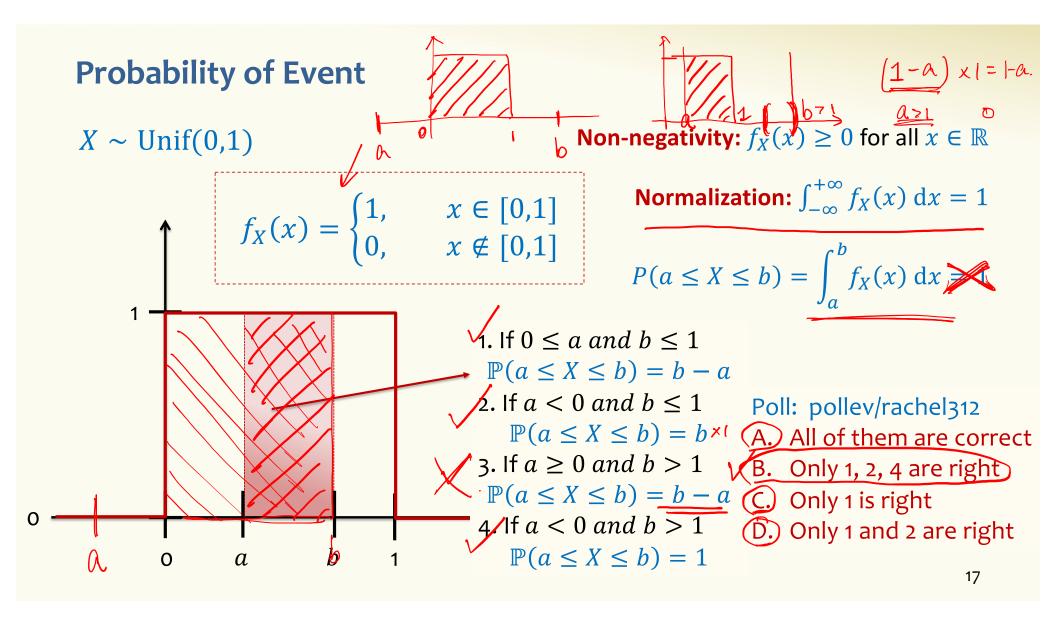
PDF of Uniform RV

$$X \sim \text{Unif}(0,1)$$

✓ Non-negativity: $f_X(x) \ge 0$ for all $x \in \mathbb{R}$



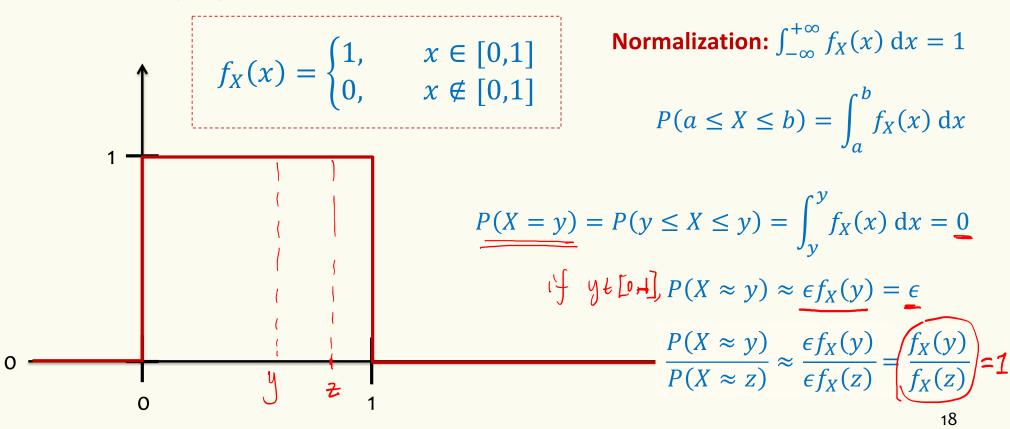
$$\int_{-\infty}^{+\infty} f_X(x) \, \mathrm{d}x = \int_0^1 f_X(x) \, \mathrm{d}x = \underbrace{1 \cdot 1}_{16} = 1$$



Probability of Event

 $X \sim \text{Unif}(0,1)$

Non-negativity: $f_X(x) \ge 0$ for all $x \in \mathbb{R}$

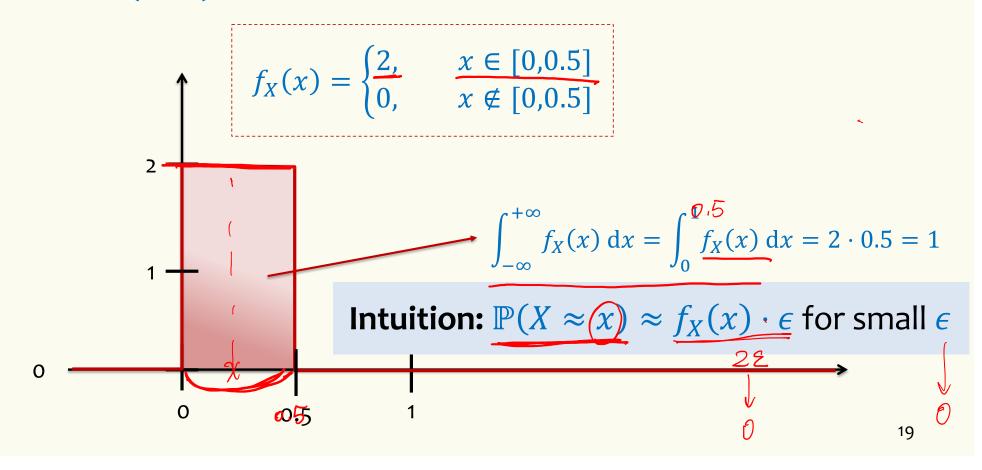


PDF of Uniform RV

Density ≠ **Probability**

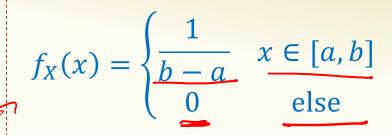
 $f_X(x) \gg 1$ is possible!

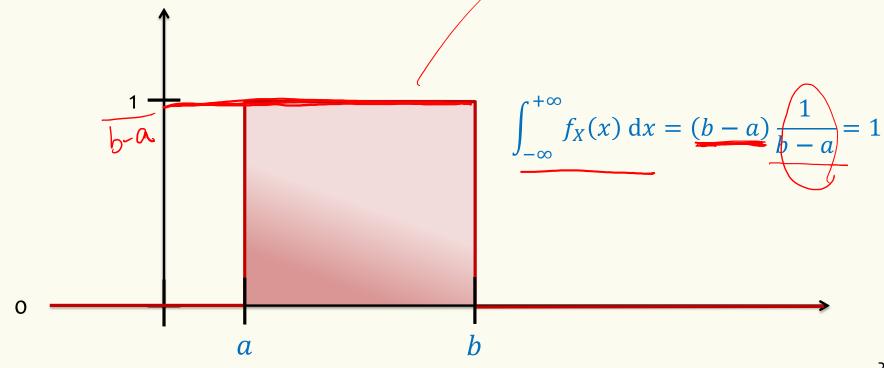
$$X \sim \text{Unif}(0,0.5)$$



Uniform Distribution

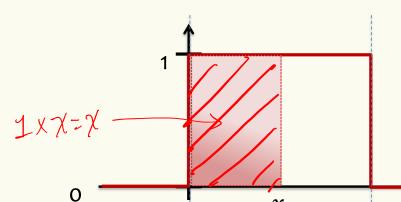
$$X \sim \text{Unif}(a, b)$$







Probability Density Function



$$f_T(x) = \begin{cases} 1, & x \in [0,1] \\ 0, & x \notin [0,1] \end{cases}$$

Cumulative Distribution Function

$$\frac{1}{x}$$

$$F_T(x) = P(T \le x) = \begin{cases} 0 & x \le 0 \\ ? & 0 \le x \le 1 \\ 1 & 1 \le x \end{cases}$$

Cumulative Distribution Function



$$F_X(a) = \mathbb{P}(X \le a) = \int_{-\infty}^a f_X(x) \, \mathrm{d}x$$

$$o \le \left(\int_{c}^{d} f(x) dx\right) + \int_{-\infty}^{c} f(x) dx = \int_{-\infty}^{d} f(x) dx$$
By the fundamental theorem of Calculus $f_X(x) = \frac{d}{dx} F(x)$

Therefore:
$$\mathbb{P}(X \in [a,b]) = F(b) - F(a) = \mathbb{P}(X \leq a)$$

 F_X is monotone increasing, since $f_X(x) \geq 0$. That is $F_X(c) \leq F_X(d)$ for $c \leq d$

$$\lim_{a \to -\infty} F_X(a) = P(X \le -\infty) = 0 \quad \lim_{a \to +\infty} F_X(a) = P(X \le +\infty) = 1$$

From Discrete to Continuous

 $\mathcal{P}(a \leq X \leq b)$

| | Discrete | Continuous |
|---------------|---|---|
| PMF/PDF | $p_X(x) = P(X = x)$ | $f_X(x) \neq P(X = x) = 0$ |
| CDF | $F_X(x) = \sum_{t \leq x} p_X(t)$ | $F_X(x) = \int_{-\infty}^x f_X(t) dt$ |
| Normalization | $\sum_{x} p_X(x) = 1$ | $\int_{-\infty}^{\infty} f_X(x) dx = 1$ |
| Expectation | $\mathbb{E}[g(X)] = \sum_{x} g(x) p_X(x)$ | $\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx$ |

Expectation of a Continuous RV

Definition. The **expected value** of a continuous RV *X* is defined as

$$\mathbb{E}(X) = \int_{-\infty}^{+\infty} f_X(x) \cdot x \, \mathrm{d}x$$

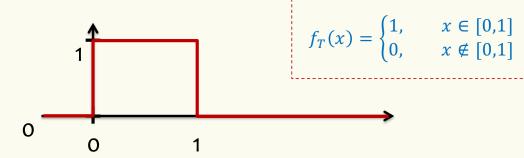
Fact.
$$\mathbb{E}(aX + bY + c) = a\mathbb{E}(X) + b\mathbb{E}(Y) + c$$

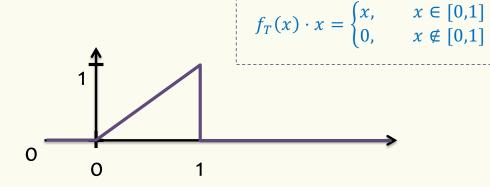
Definition. The variance of a continuous RV X is defined as

$$Var(X) = \int_{-\infty}^{+\infty} f_X(x) \cdot (x - \mathbb{E}(X))^2 dx = \mathbb{E}(X^2) - \mathbb{E}(X)^2$$

Expectation of a Continuous RV

Example. $T \sim \text{Unif}(0,1)$





Definition.

$$\mathbb{E}(X) = \int_{-\infty}^{+\infty} f_X(x) \cdot x \, \mathrm{d}x$$

$$\mathbb{E}(T) = \frac{1}{2}1^2 = \frac{1}{2}$$

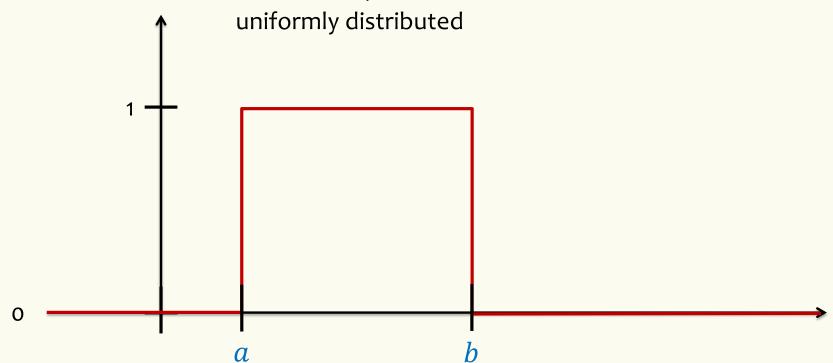
Area of triangle

Uniform Distribution

 $X \sim \text{Unif}(a, b)$

We also say that *X* follows the uniform distribution / is uniformly distribute

$$f_X(x) = \begin{cases} \frac{1}{b-a} & x \in [a,b] \\ 0 & \text{else} \end{cases}$$



Uniform Density – Expectation

$$X \sim \text{Unif}(a, b)$$

$$f_X(x) = \begin{cases} \frac{1}{b-a} & x \in [a,b] \\ 0 & \text{else} \end{cases}$$

$$\mathbb{E}(X) = \int_{-\infty}^{+\infty} f_X(x) \cdot x \, dx$$

$$= \frac{1}{b-a} \int_a^b x \, dx = \frac{1}{b-a} \left(\frac{x^2}{2}\right) \Big|_a^b = \frac{1}{b-a} \left(\frac{b^2 - a^2}{2}\right)$$

$$= \frac{(b-a)(a+b)}{2(b-a)} = \frac{a+b}{2}$$

Uniform Density – Variance

$$X \sim \text{Unif}(a, b)$$

$$\mathbb{E}(X^2) = \int_{-\infty}^{+\infty} f_X(x) \cdot x^2 \, \mathrm{d}x$$

$$= \frac{1}{b-a} \int_{a}^{b} x^{2} dx = \frac{1}{b-a} \left(\frac{x^{3}}{3}\right) \Big|_{a}^{b} = \frac{b^{3}-a^{3}}{3(b-a)}$$
$$= \frac{(b-a)(b^{2}+ab+a^{2})}{3(b-a)} = \frac{b^{2}+ab+a^{2}}{3}$$

 $f_X(x) = \begin{cases} \frac{1}{b-a} & x \in [a,b] \\ 0 & \text{else} \end{cases}$

Uniform Density – Variance

$$\mathbb{E}(X^2) = \frac{b^2 + ab + a^2}{3} \qquad \mathbb{E}(X) = \frac{a+b}{2}$$

$$X \sim \text{Unif}(a, b)$$

$$Var(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2$$

$$= \frac{b^2 + ab + a^2}{3} - \frac{a^2 + 2ab + b^2}{4}$$

$$= \frac{4b^2 + 4ab + 4a^2}{12} - \frac{3a^2 + 6ab + 3b^2}{12}$$

$$= \frac{b^2 - 2ab + a^2}{12} = \frac{(b - a)^2}{12}$$