CSE 312

Foundations of Computing II

Lecture 7: Chain Rule and Independence



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Slide Credit: Based on Stefano Tessaro's slides for 312 19au incorporating ideas from Alex Tsun, Rachel Lin, Hunter Schafer & myself ©

Announcements

- No concept check today!
- Section tomorrow is **important** with new content that you will need on pset 3, problem 7. Bring your laptops.
- I have to be out of town (and will be largely unreachable) Thursday-Saturday – Aleks will give Friday's lecture!
- Quiz 1 out later next week. Will cover material from the first two problem sets.

Friday 10/8: Bayes Theorem with Law of Total Probability

Bayes Theorem with LTP: Let $E_1, E_2, ..., E_n$ be a partition of the sample space, and F and event. Then,

$$P(E_1|F) = \frac{P(F|E_1)P(E_1)}{P(F)} = \frac{P(F|E_1)P(E_1)}{\sum_{i=1}^{n} P(F|E_i)P(E_i)}$$

Simple Partition: In particular, if E is an event with non-zero probability, then

$$P(E|F) = \frac{P(F|E)P(E)}{P(F|E)P(E) + P(F|E^C)P(E^C)}$$

Monday 10/10: Chain Rule



$$\mathbb{P}(\mathcal{B}|\mathcal{A}) = \frac{\mathbb{P}(\mathcal{A} \cap \mathcal{B})}{\mathbb{P}(\mathcal{A})} \qquad \qquad \mathbb{P}(\mathcal{A})\mathbb{P}(\mathcal{B}|\mathcal{A}) = \mathbb{P}(\mathcal{A} \cap \mathcal{B})$$

$$\mathbb{P}(\mathcal{A})\mathbb{P}(\mathcal{B}|\mathcal{A}) = \mathbb{P}(\mathcal{A} \cap \mathcal{B})$$

Theorem. (Chain Rule) For events $\mathcal{A}_1, \mathcal{A}_2, \dots, \mathcal{A}_n$,

$$\mathbb{P}(\mathcal{A}_1 \cap \cdots \cap \mathcal{A}_n) = \mathbb{P}(\mathcal{A}_1) \cdot \mathbb{P}(\mathcal{A}_2 | \mathcal{A}_1) \cdot \mathbb{P}(\mathcal{A}_3 | \mathcal{A}_1 \cap \mathcal{A}_2)$$

$$\cdots \mathbb{P}(\mathcal{A}_n | \mathcal{A}_1 \cap \mathcal{A}_2 \cap \cdots \cap \mathcal{A}_{n-1})$$

An easy way to remember: We have n tasks and we can do them sequentially, conditioning on the outcome of previous tasks

Monday: Independence

Definition. Two events \mathcal{A} and \mathcal{B} are (statistically) **independent** if

$$\mathbb{P}(\mathcal{A} \cap \mathcal{B}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}).$$

Alternatively,

- If $\mathbb{P}(A) \neq 0$, equivalent to $\mathbb{P}(B|A) = \mathbb{P}(B)$
- If $\mathbb{P}(\mathcal{B}) \neq 0$, equivalent to $\mathbb{P}(\mathcal{A}|\mathcal{B}) = \mathbb{P}(\mathcal{A})$

"The probability that \mathcal{B} occurs after observing \mathcal{A} " -- Posterior = "The probability that \mathcal{B} occurs" -- Prior

Agenda

- A Sequential Process Defined Using Independence
- Independence As An Assumption
- Sometimes Independence Occurs for Nonobvious Reasons
- Conditional Independence
- Correlation vs Causation
- Information Cascades

Example - Throwing A Die Repeatedly

Alice and Bob are playing the following game.

A 6-sided die is thrown, and each time it's thrown, regardless of the history, it is equally likely to show any of the six numbers

If it shows 1, $2 \rightarrow Alice wins$.

If it shows $3 \rightarrow Bob$ wins.

Otherwise, play another round

What is Pr(Alice wins on 1st round) = 3

Pr(Alice wins on i^{th} round) = ?

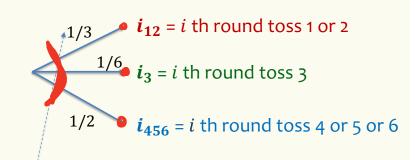
Pr(Alice wins) = ?

Sequential Process – defined in terms of independence

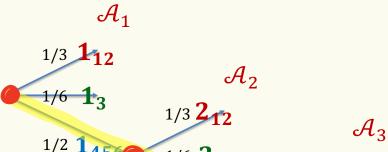
A 6-sided die is thrown, and each time it's thrown, regardless of the history, it is equally likely to show any of the six numbers

Local Rules: In each round

- If it shows 1,2 → Alice wins
- If it shows 3 → Bob wins
- Else, play another round



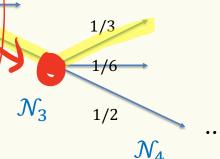
Pr (Alice wins on i -th round | nobody won in rounds 1.. i-1) = 1/3



Local Rules: In each round

- If it shows 1,2 → Alice wins
- If it shows 3 → Bob wins
- Else, play another round

- \mathcal{A}_i = Alice wins in round i
- \mathcal{N}_i = nobody wins in rounds 1..i



1/3

1/6

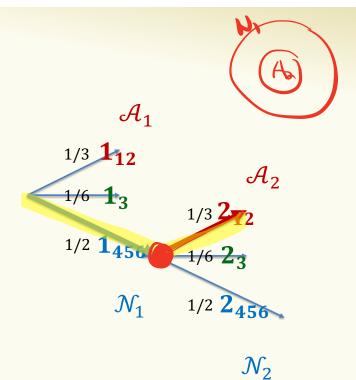
Events:

- \mathcal{A}_i = Alice wins in round i
- \mathcal{N}_i = nobody wins in rounds 1..i

$$\mathbb{P}(A_2) = Pr(u, nA_b)$$

$$= Pr(u,) Pr(A_b, u,)$$

$$= \frac{1}{3}$$



2nd roll indep of 1st roll

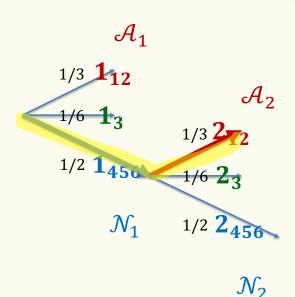
Events:

- \mathcal{A}_i = Alice wins in round i
- \mathcal{N}_i = nobody wins in rounds 1..i

$$\mathbb{P}(\mathcal{A}_2) = \mathcal{P}(\mathcal{N}_1 \cap \mathcal{A}_2)$$

$$= \mathcal{P}(\mathcal{N}_1) \times \mathcal{P}(\mathcal{A}_2 | \mathcal{N}_1)$$

$$= \frac{1}{2} \times \frac{1}{3} = \frac{1}{6}$$



The event \mathcal{A}_2 implies \mathcal{N}_1 , and this means that $\mathcal{A}_2 \cap \mathcal{N}_1 = \mathcal{A}_2$

2nd roll indep of 1st roll

 $\mathbb{P}(\underline{\mathcal{A}_i}) = \mathcal{P}(\mathcal{N}_1 \cap \mathcal{N}_2 \cap \dots \cap \mathcal{N}_{i-1} \cap \mathcal{A}_i)$

Events:

- \mathcal{A}_i = Alice wins in round i
- \mathcal{N}_i = nobody wins in round $\mathbf{\xi}$

 $= \underbrace{\mathcal{P}(\mathcal{N}_{1})} \times \underbrace{\mathcal{P}(\mathcal{N}_{2}|\mathcal{N}_{1})} \times \underbrace{\mathcal{P}(\mathcal{N}_{3}|\mathcal{N}_{1}\cap\mathcal{N}_{2})} \times \underbrace{\mathcal{P}(\mathcal{N}_{i-1}|\mathcal{N}_{1}\cap\mathcal{N}_{2}\cap\cdots\cap\mathcal{N}_{i-1})} \times \underbrace{\mathcal{P}(\mathcal{A}_{i}|\mathcal{N}_{1}\cap\mathcal{N}_{2}\cap\cdots\cap\mathcal{N}_{i-1})}_{i-1}$

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$$\mathcal{A}_i$$
 = Alice wins in round i $\mathbb{P}(\mathcal{A}_i) = \left(\frac{1}{2}\right)^{l-1} \times \frac{1}{3}$

What is the probability that Alice wins?

$$P(A, \cup A_{\sigma} \cup A_{3} \cup \cdots) = \sum_{i=1}^{\infty} P(A_{i}) = \sum_{i=1}^{\infty} (\frac{1}{2})^{i-1} \frac{1}{3}$$



$$\mathcal{A}_i$$
 = Alice wins in round i $\mathbb{P}(\mathcal{A}_i) = \left(\frac{1}{2}\right)^{i-1} \times \frac{1}{3}$



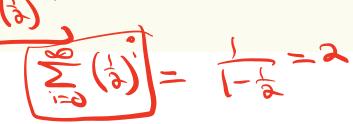
What is the probability that Alice wins?

$$\mathbb{P}(\mathcal{A}_1 \cup \mathcal{A}_2 \cup \cdots) = \sum_{i=1}^{\infty} \mathbb{P}(\mathcal{A}_i) \qquad \text{All } c$$

All \mathcal{A}_i 's are disjoint.

$$\sum_{i=1}^{\infty} \left(\frac{1}{2}\right)^{i-1} \times \frac{1}{3} = \frac{1}{3} \times 2 = \frac{2}{3}$$

Fact. If
$$|x| < 1$$
, then $\sum_{i=0}^{\infty} x^i = \frac{1}{1-x}$.





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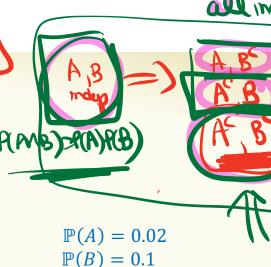


Independence as an assumption

- People often assume it without justification.
- Example: A sky diver has two chutes

A: event that the main chute doesn't open

B: event that the backup doesn't open



$$(2r(AUB) = 1 - P(A^c \cap B^c)$$

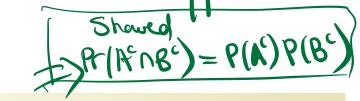
$$= (-P(A^c) P(B^c))$$

$$= 1 - (P_{-}(A) + P(B)) - P(A \cap B)$$

$$= 1 - P_{-}(A) - P(B) + P(A) + P(B)$$

$$= (1 - P_{-}(A)) - P(B) + P(A) + P(B)$$

$$= (1 - P_{-}(A)) - P(B) + P(B)$$



Independence as an assumption

- People often assume it without justification.
- Example: A sky diver has two chutes

A: event that the main chute doesn't open

B: event that the backup doesn't open

$$\mathbb{P}(A) = 0.02$$

 $\mathbb{P}(B) = 0.1$

What is the chance that at least one opens assuming independence?

Assuming independence doesn't justify the assumption! Both chutes could fail because of the same rare event e.g., freezing rain.

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Independence – Another Look

Definition. Two events \mathcal{A} and \mathcal{B} are (statistically) **independent** if

$$\mathbb{P}(\mathcal{A} \cap \mathcal{B}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}).$$

"Equivalently." $\mathbb{P}(A|B) = \mathbb{P}(A)$.

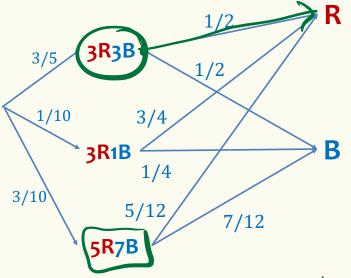
Events generated independently

their probabilities satisfy independence



This can be counterintuitive!

Sequential Process



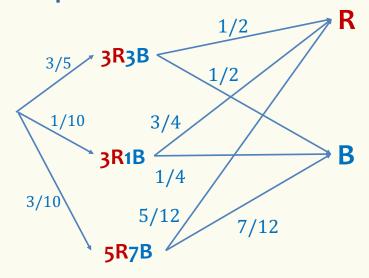
Setting: An urn contains:

- 3 red and 3 blue balls w/ probability 3/5
- 3 red and 1 blue balls w/ probability 1/10
- 5 red and 7 blue balls w/ probability 3/10 We draw a ball at random from the urn.

Are R and 3R3B independent?

$$P(R) = P(R|3R3B)P(3R3B) + P(R|3R1B)P(3R1B)$$
+ $P(R|5R7B)P(5R7B)$

Sequential Process



Are R and 3R3B independent?

Setting: An urn contains:

- 3 red and 3 blue balls w/ probability 3/5
- 3 red and 1 blue balls w/ probability 1/10
- 5 **red** and 7 **blue** balls w/ probability 3/10 We draw a ball at random from the urn.

$$\mathbb{P}(\mathbf{R}) = \frac{3}{5} \times \frac{1}{2} + \frac{1}{10} \times \frac{3}{4} + \frac{3}{10} \times \frac{5}{12} = \frac{1}{2}$$

$$\mathbb{P}(3R3B) \times \mathbb{P}(R \mid 3R3B)$$

Independent! $\mathbb{P}(R) = \mathbb{P}(R \mid 3R3B)$

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Conditional Independence

Definition. Two events \mathcal{A} and \mathcal{B} are **independent** conditioned on $\underline{\mathcal{C}}$ if $\mathbb{P}(\mathcal{C}) \neq 0$ and $\mathbb{P}(\mathcal{A} \cap \mathcal{B} \mid \mathcal{C}) = \mathbb{P}(\mathcal{A} \mid \mathcal{C}) \cdot \mathbb{P}(\mathcal{B} \mid \mathcal{C})$.



Plain Independence. Two events \mathcal{A} and \mathcal{B} are independent if

$$\mathbb{P}(\mathcal{A} \cap \mathcal{B}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}).$$

Equivalence:

- If $\mathbb{P}(A) \neq 0$, equivalent to $\mathbb{P}(B|A) = \mathbb{P}(B)$
- If $\mathbb{P}(\mathcal{B}) \neq 0$, equivalent to $\mathbb{P}(\mathcal{A}|\mathcal{B}) = \mathbb{P}(\mathcal{A})$

Conditional Independence

Definition. Two events \mathcal{A} and \mathcal{B} are **independent** conditioned on \mathcal{C} if $\mathbb{P}(\mathcal{C}) \neq 0$ and $\mathbb{P}(\mathcal{A} \cap \mathcal{B} \mid \mathcal{C}) = \mathbb{P}(\mathcal{A} \mid \mathcal{C}) \cdot \mathbb{P}(\mathcal{B} \mid \mathcal{C})$.

Equivalence:

- If $\mathbb{P}(A \cap C) \neq 0$, equivalent to $\mathbb{P}(B|A \cap C) = \mathbb{P}(B|C)$
- If $\mathbb{P}(\mathcal{B} \cap \mathcal{C}) \neq 0$, equivalent to $\mathbb{P}(\mathcal{A} | \mathcal{B} \cap \mathcal{C}) = \mathbb{P}(\mathcal{A} | \mathcal{C})$

Example - More coin tossing

Suppose there is a coin C1 with Pr(Head) = 0.3 and a coin C2 with Pr(Head) = 0.9. We pick one randomly with equal probability and flip that coin twice independently. What is the probability we get all heads?

$$Pr(HH) = Pr(HH \mid C1) Pr(C1) + Pr(HH \mid C2) Pr(C2)$$

$$= P(H \mid C1) P(H \mid C_1) P(C1) + P(HH \mid C2) Pr(C2)$$

$$= P(H \mid C1) P(H \mid C_1) P(C1) + P(HH \mid C2) Pr(C2)$$

$$= P(H \mid C1) P(H \mid C_1) P(C1) + P(HH \mid C2) Pr(C2)$$

$$= P(H \mid C1) P(H \mid C_1) P(C1) + P(HH \mid C2) Pr(C2)$$

$$= P(H \mid C1) P(H \mid C_1) P(C1) + P(HH \mid C2) Pr(C2)$$

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$$= P(H \mid C1) P(H \mid C_1) P(C1) + P(HH \mid C2) Pr(C2)$$

Suppose there is a coin C1 with Pr(Head) = 0.3 and a coin C2 with Pr(Head) = 0.9. We pick one randomly with equal probability and flip that coin 2 times independently. What is the probability we get all heads?

$$Pr(HH) = Pr(HH \mid C1) Pr(C1) + Pr(HH \mid C2) Pr(C2)$$
 LTP

= $Pr(H \mid C2)^2 Pr(C1) + Pr(H \mid C2)^2 Pr(C2)$ Conditional Independence

$$= 0.3^2 \cdot 0.5 + 0.9^2 \cdot 0.5 = 0.45$$

$$Pr(H) = Pr(H \mid C1) Pr(C1) + Pr(H \mid C2) Pr(C2) = 0.6$$