

## Maximum Likelihood Estimators

### Parameter estimation

Given independent samples  $x_1, x_2, \dots, x_n$  from a distribution  $f(x|\theta)$  (either a PMF or PDF), where  $\theta$  is a list of the distribution's parameters, estimate  $\theta$ .

Ex: Given samples HHTHT from independent flips of a coin, estimate  $\theta = P(\text{heads})$ .

$P(x|\theta)$ : Probability of event  $x$  given parameters  $\theta$ .  
Viewed as a function of  $x$  ( $\theta$  fixed), it's a probability.  
Viewed as a function of  $\theta$  ( $x$  fixed), it's called a likelihood and often written  $L(x|\theta)$ .

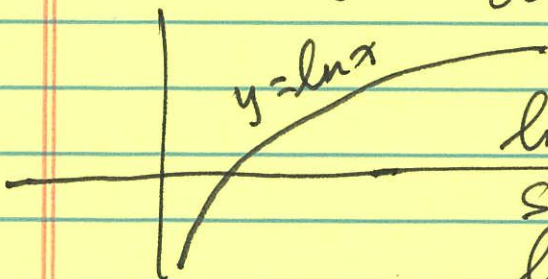
Maximum likelihood estimation:

$$\text{What } \theta \text{ maximizes } L(x_1, x_2, \dots, x_n | \theta) \\ = \prod_{i=1}^n f(x_i | \theta)$$

where  $x_1, x_2, \dots, x_n$  are ind. samples from  $f(x|\theta)$ ?

$$\text{Approach: } \frac{\partial}{\partial \theta} L(x_1, \dots, x_n | \theta) = 0$$

$$\text{Or more often: } \frac{\partial}{\partial \theta} \ln L(x_1, \dots, x_n | \theta) = 0$$



$\ln$  monotonically increasing  
so any  $\theta$  that maximizes  $\ln L$  also maximizes  $L$ .

Taking  $\ln$  turns  $\prod$  into  $\sum$ , and so taking the derivative is much simpler.

First application:

$\theta = P(\text{head})$ ,  $n$  ind. flips  $x_1, x_2, \dots, x_n$  yielding  $n_0$  tails and  $n_1$  heads,  $n_0 + n_1 = n$ .

$$L(x_1, x_2, \dots, x_n | \theta) = (1-\theta)^{n_0} \theta^{n_1}$$

$$\ln L(x_1, x_2, \dots, x_n | \theta) = n_0 \ln(1-\theta) + n_1 \ln(\theta)$$

$$\frac{\partial}{\partial \theta} \ln L(x_1, x_2, \dots, x_n | \theta) = \frac{-n_0}{1-\theta} + \frac{n_1}{\theta} = 0$$

Let  $\hat{\theta}$  be the solution to this.

$$\frac{n_1}{\hat{\theta}} = \frac{n_0}{1-\hat{\theta}}$$

$$n_1(1-\hat{\theta}) = n_0 \hat{\theta}$$

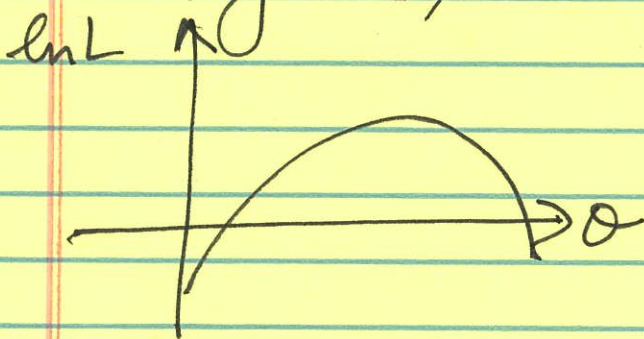
$$n_1 = n_0 \hat{\theta} + n_1 \hat{\theta}$$

$$\hat{\theta} = \frac{n_1}{n_0 + n_1} = \frac{n_1}{n}$$

Is  $\theta = \hat{\theta}$  a maximum?

$$\frac{\partial^2}{\partial \theta^2} \ln L(x_1, x_2, \dots, x_n | \theta) = -\frac{n_0}{(1-\theta)^2} - \frac{n_1}{\theta^2} < 0$$

so  $\ln L(x_1, \dots, x_n | \theta)$  is concave downward everywhere, and  $\theta = \hat{\theta}$  is a maximum.



Second application:

$x_i \sim N(\mu, \sigma^2)$ , where  $\mu, \sigma^2$  are unknown.

Let  $\theta_1 = \mu$  and  $\theta_2 = \sigma^2$ .

$$L(x_1, x_2, \dots, x_n | \theta_1, \theta_2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\theta_2}} e^{-\frac{(x_i - \theta_1)^2}{2\theta_2}}$$

$$\ln L(x_1, x_2, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n \left( -\frac{1}{2} \ln(2\pi\theta_2) - \frac{(x_i - \theta_1)^2}{2\theta_2} \right)$$

$$\frac{\partial}{\partial \theta_1} \ln L(x_1, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n \frac{2(x_i - \theta_1)}{2\theta_2} = 0$$

$$\sum_{i=1}^n (x_i - \hat{\theta}_1) = 0$$

$$\sum_{i=1}^n x_i - n\hat{\theta}_1 = 0$$

$$\hat{\theta}_1 = \frac{1}{n} \sum_{i=1}^n x_i$$

So sample mean is MLE for population mean  $\mu$ .

$$\frac{\partial^2}{\partial \theta_1^2} \ln L(x_1, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n -\frac{1}{\theta_2} = -\frac{n}{\theta_2} < 0,$$

so  $\ln L$  is concave downward everywhere as a function of  $\theta_1$ , so  $\theta_1 = \hat{\theta}_1$  is a max.

~~Next~~ Next up:

$$\frac{\partial}{\partial \theta_2} \ln L(x_1, \dots, x_n | \theta_1, \theta_2)$$