

Markov Chains and PageRank

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Suppose $\hat{\Theta}(x_1, \dots, x_n)$ is estimator for
param Θ

MLE estimators for

μ, σ^2 of normal

$$\hat{\Theta}_1 = \frac{1}{n} \sum_{i=1}^n x_i$$

for μ unbiased

$$\hat{\Theta}_2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\Theta}_1)^2$$

for σ^2 biased.

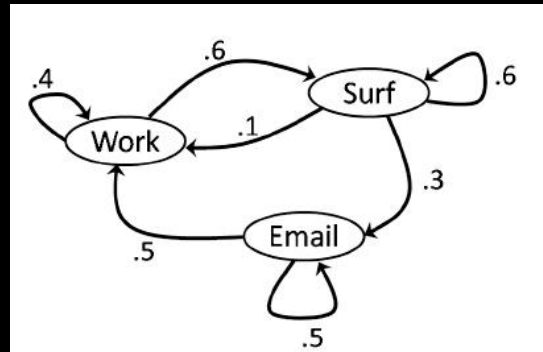
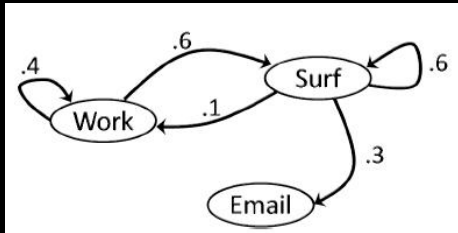
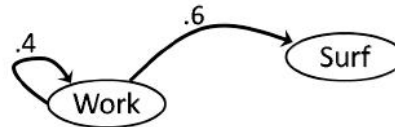
$\hat{\Theta}(x_1, \dots, x_n)$ unbiased if

$$E(\hat{\Theta}(x_1, \dots, x_n)) = \Theta$$

(true param
value)

A first Markov chain

Work



My daily life in a nutshell!

Finite Markov Chain

- set of states $S = \{1, 2, \dots, n\}$
- X_t : state at time t
- transition probabilities

$$P_{ij} = \Pr(X_{t+1} = j \mid X_t = i)$$

Collect these transition probs
into n by n matrix

P where $(P)_{ij}$ entry of $P = P_{ij}$

sum of entries in i th row

$$\sum_{j=1}^n P_{ij} = 1 \quad \forall i$$



$$S = \{W, S, E\}$$

	W	S	E
W	.4	.6	0
S	.1	.6	.3
E	.5	0	.5

$$P_{SE} = 0.3$$

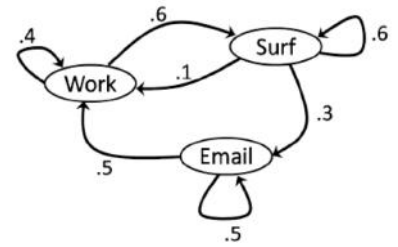
$$\begin{array}{c}
 W \quad S \quad E \\
 W \begin{pmatrix} .4 & .6 & 0 \\ .1 & .6 & .3 \\ .5 & 0 & .5 \end{pmatrix} \\
 S \\
 E
 \end{array}$$

$$P_{SE} = 0.3$$

t	q_w^+	q_s^+	q_E^+
0	1	0	0
1	0.4	0.6	0
2			

$$0.4 \cdot 0.4 + 0.6 \cdot 0.1 + 0 \cdot 0.5$$

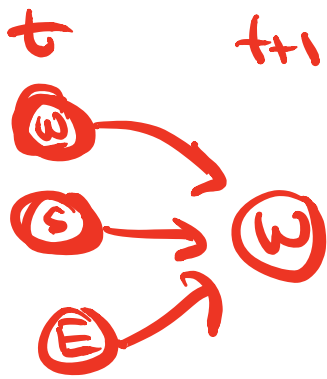
$$0.4 \cdot 0.6 + 0.6 \cdot 0.6 + 0 \cdot 0$$



$$q_w^+ = \Pr(X_t = w | X_0 = w)$$

$$q_s^+ = \Pr(X_t = s | X_0 = w)$$

$$q_E^+ = \Pr(X_t = E | X_0 = w)$$



0.4 0.6 0

$$(q_w^+, q_s^+, q_e^+)$$

$$\begin{matrix} W \\ S \\ E \end{matrix} \begin{pmatrix} .4 & .6 & 0 \\ .1 & .6 & .3 \\ .5 & 0 & .5 \end{pmatrix} = (q_w^{t+1}, q_s^{t+1}, q_e^{t+1})$$

$$q_w^{t+1} = q_w^+ P_{ww} + q_s^+ P_{sw} + q_e^+ P_{ew}$$

$S = \{1, 2, \dots, n\}$ P TPM

$$P_{ij} = \Pr(X_{t+1} = j | X_t = i)$$

$$q^+ = (q_1^+, \dots, q_n^+)$$

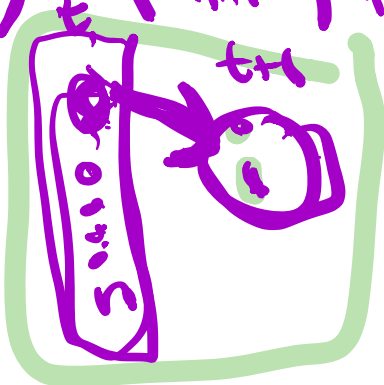
$$q_i^+ = \Pr(X_+ = i | X_0 = 1)$$

$$q^{t+1} = q^+ P$$

$$q^0 = (1, 0, 0, \dots, 0)$$

$$\equiv q_i^{t+1} = \sum_{j=1}^n q_j^+ P_{ji} \quad \forall i$$

$$LTP = \sum_{j=1}^n \Pr(X_t=j | X_0=1) \Pr(X_{t+1}=i | X_t=j)$$



$$\begin{aligned}
 q^{(3)} &= q^{(2)} P = q^{(1)} P^2 = q^{(0)} P^3 \\
 q^{(2)} &= q^{(1)} P = q^{(0)} P^2 \\
 q^{(1)} &= q^{(0)} P = q^{(0)} P
 \end{aligned}$$

$$q^{(t)} = \begin{pmatrix} 1 & 0 & 0 & \dots & 0 \end{pmatrix} P^t$$

$$(q_1^+, q_2^+, \dots, q_n^+) = \begin{pmatrix} 1 & 0 & 0 & \dots & 0 \end{pmatrix} \begin{pmatrix} p_{11} & p_{12} & \dots & p_{1n} \end{pmatrix}^t$$

$$P^2 = \begin{pmatrix} p_{11} & p_{12} \\ p_{21} & p_{22} \end{pmatrix} \begin{pmatrix} p_{11} & p_{12} \\ p_{21} & p_{22} \end{pmatrix}$$

$$P_{ij}^2 = \sum_k p_{ik} p_{kj} = \sum_k P(X_{t+2}=j | X_{t+1}=k) \cdot \Pr(X_{t+1}=k | X_t=i)$$

\swarrow \searrow
 i \dots j

$$= \Pr(X_{t+2}=j | X_t=i)$$

$$P_{ij}^t$$

$$= \Pr(X_t=j | X_0=i)$$

$$= \Pr(X_{t+c}=j | X_c=i)_{\forall c}$$

Prob that
in t steps
starting at
 i I end up
at j

$$P^2 = \begin{matrix} & W & S & E \\ \begin{matrix} W \\ S \\ E \end{matrix} & \begin{pmatrix} .22 & .6 & .18 \\ .25 & .42 & .33 \\ .45 & .3 & .25 \end{pmatrix} \end{matrix}$$

$$P^3 = \begin{matrix} & W & S & E \\ \begin{matrix} W \\ S \\ E \end{matrix} & \begin{pmatrix} .238 & .492 & .270 \\ .307 & .402 & .291 \\ .335 & .450 & .215 \end{pmatrix} \end{matrix}$$

$$P^{10} \approx \begin{matrix} & W & S & E \\ \begin{matrix} W \\ S \\ E \end{matrix} & \begin{pmatrix} .2940 & .4413 & .2648 \\ .2942 & .4411 & .2648 \\ .2942 & .4413 & .2648 \end{pmatrix} \end{matrix}$$

$$Pr(X_{10} = E | X_0 = S)$$

$$P^{30} \approx \begin{matrix} & W & S & E \\ \begin{matrix} W \\ S \\ E \end{matrix} & \begin{pmatrix} .29411764705 & .44117647059 & .26470588235 \\ .29411764706 & .44117647058 & .26470588235 \\ .29411764706 & .44117647059 & .26470588235 \end{pmatrix} \end{matrix}$$

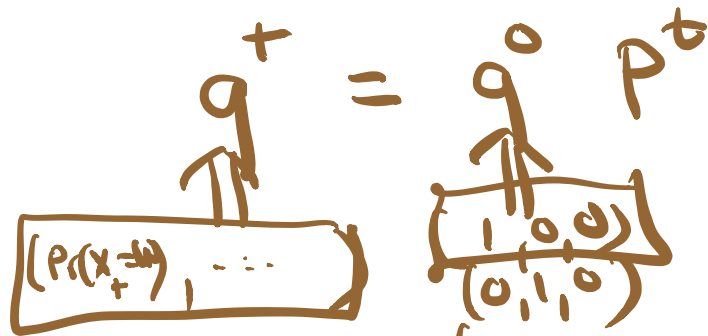
$$P^{60} \approx \begin{matrix} & W & S & E \\ \begin{matrix} W \\ S \\ E \end{matrix} & \begin{pmatrix} .294117647058823 & .441176470588235 & .264705882352941 \\ .294117647068823 & .441176470588235 & .264705882352941 \\ .294117647068823 & .441176470588235 & .264705882352941 \end{pmatrix} \end{matrix}$$

$$Pr(X_{60} = E | X_0 = S)$$

$$Pr(X_{60} = E | X_0 = W)$$

$$Pr(X_{60} = E | X_0 = E)$$

converging to fixed #'s
in each column



$$q^{t+1} = q^t P$$



$$q^{t+1} = q^t = \pi = (\pi_1, \dots, \pi_n)$$

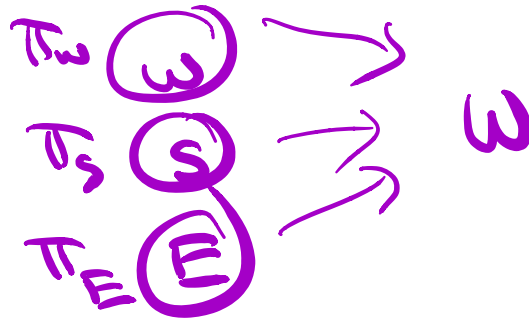
once we reach fixed distribution over states.

$$\boxed{\pi = \pi P}$$

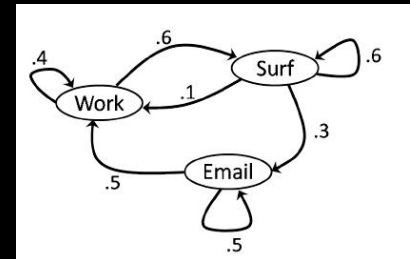
$$\pi_w = \pi_w P_{ww} + \pi_s P_{sw} + \pi_E P_{EW}$$

$$\pi_s = \pi_w P_{ws} + \pi_s P_{ss} + \pi_E P_{ES}$$

$$\pi_E = \pi_w P_{wE} + \pi_s P_{sE} + \pi_E P_{EE}$$



$$\pi[S] = \frac{15}{34}, \pi[W] = \frac{10}{34}, \pi[E] = \frac{9}{34}.$$



$$\begin{cases} \pi[W] & = & .4\pi[W] + .1\pi[S] + .5\pi[E] \\ \pi[S] & = & .6\pi[W] + .6\pi[S] + 0\pi[E] \\ \pi[E] & = & 0\pi[W] + .3\pi[S] + .5\pi[E] \end{cases}$$

$$\pi[W] + \pi[S] + \pi[E] = 1$$

Fundamental Thm of MCs.

$$\lim_{t \rightarrow \infty} P_{ij}^t = \pi_j \quad \forall i$$

$$\Pr(X_t = j | X_0 = i)$$

stationary distn

(fine print)

Assumes $q_0 = (1, 0, \dots, 0)$

$$q_i^t = \Pr(X_t = i | X_0 = 1)$$

$$P_{ij}^t = \Pr(X_t = j | X_0 = i)$$

$$q^{t+1} = q^t P$$

$$= q^0 P^t$$