



ANNA KARLIN  
MOST SLIDES BY ALEX TSUN

# POISSON RV EXAMPLE



Suppose Lookbook gets on average 120 new users per hour, and Quickgram gets 180 new users per hour, independently. What is the probability that, combined, less than 2 users sign up in the next minute?

Convert  $\lambda$ 's to the same unit of interest. For us, it's a minute.

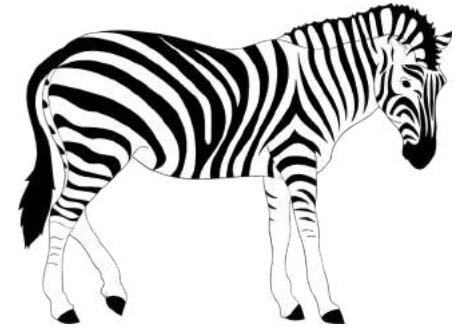
$$X \sim \text{Poi}(2 \text{ users/min})$$

$$Y \sim \text{Poi}(3 \text{ users/min})$$

$$Z = X + Y \sim \text{Poi}(2 + 3) = \text{Poi}(5)$$

$$P(Z < 2) = p_Z(0) + p_Z(1) = e^{-5} \frac{5^0}{0!} + e^{-5} \frac{5^1}{1!} = 6e^{-5} \approx 0.04$$

# THE ZOO OF DISCRETE RV'S



- THE BERNOULLI RV
- THE BINOMIAL RV
- THE GEOMETRIC RV
- THE UNIFORM RV
- THE POISSON RV

- THE NEGATIVE BINOMIAL RV
- THE HYPERGEOMETRIC RV

## random variables

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### Important Examples:

**Uniform(a,b):**  $P(X = i) = \frac{1}{b - a + 1}$       $\mu = \frac{a + b}{2}, \sigma^2 = \frac{(b - a)(b - a + 2)}{12}$

**Bernoulli(p):**  $P(X = 1) = p, P(X = 0) = 1 - p$       $\mu = p, \sigma^2 = p(1 - p)$

**Binomial(n,p)**  $P(X = i) = \binom{n}{i} p^i (1 - p)^{n - i}$       $\mu = np, \sigma^2 = np(1 - p)$

**Poisson( $\lambda$ ):**      $P(X = i) = e^{-\lambda} \frac{\lambda^i}{i!}$       $\mu = \lambda, \sigma^2 = \lambda$

**Bin(n,p)  $\approx$  Poi( $\lambda$ )** where  $\lambda = np$  fixed,  $n \rightarrow \infty$  (and so  $p = \lambda/n \rightarrow 0$ )

**Geometric(p)**      $P(X = k) = (1 - p)^{k-1} p$       $\mu = 1/p, \sigma^2 = (1 - p)/p^2$

# PROBABILITY

## 4.1 CONTINUOUS RANDOM VARIABLES BASICS

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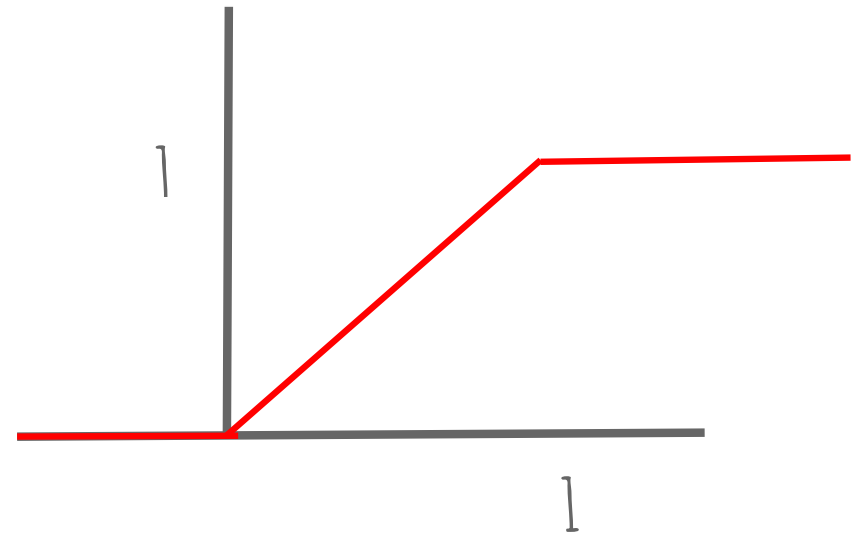
# AGENDA

- PROBABILITY DENSITY FUNCTIONS (PDFS)
- CUMULATIVE DISTRIBUTION FUNCTIONS (CDFs)
- FROM DISCRETE TO CONTINUOUS



# CDF INTUITION

$$F_X(w) = P(X \leq w)$$

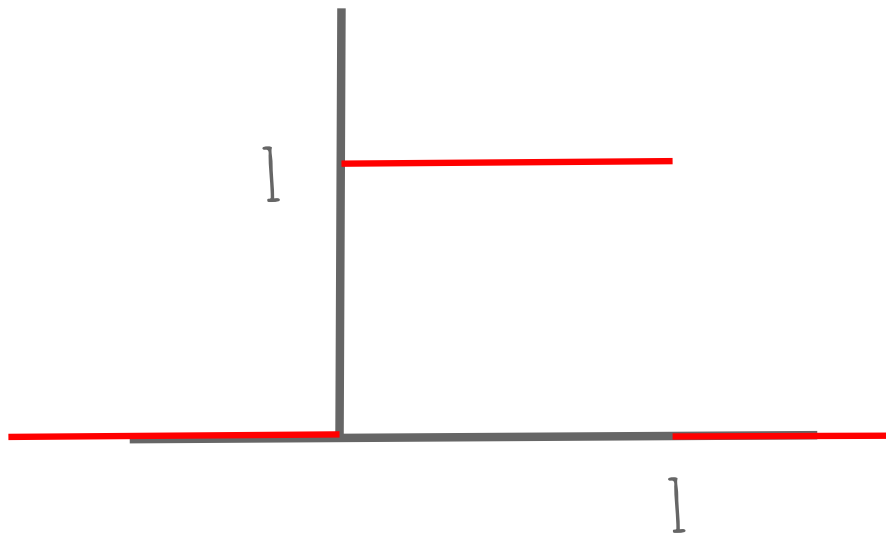


$$F_X(w) = \begin{cases} 0, & w < 0 \\ w, & 0 \leq w \leq 1 \\ 1, & w > 1 \end{cases}$$



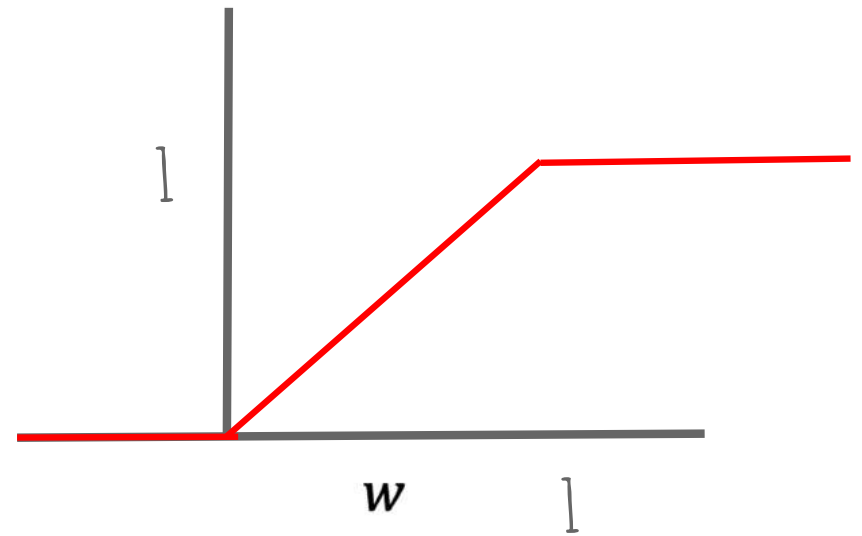
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$f_X(v)$



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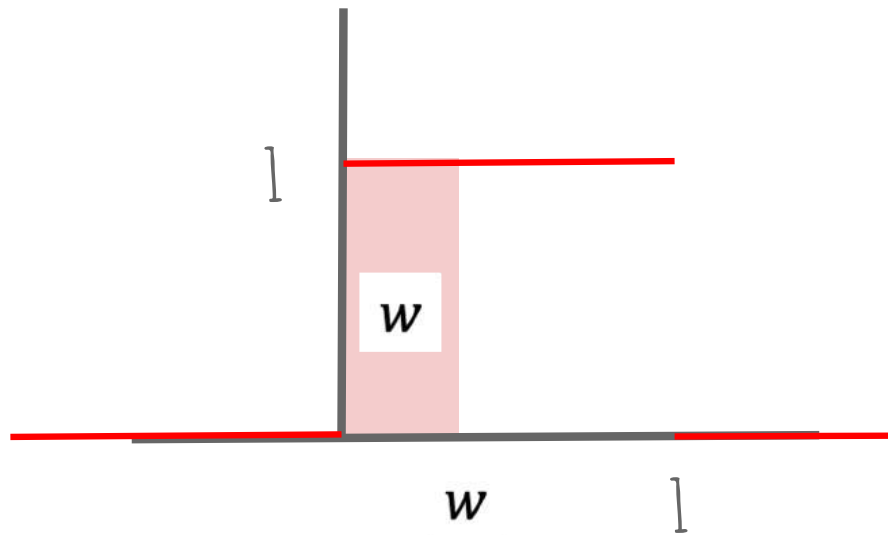


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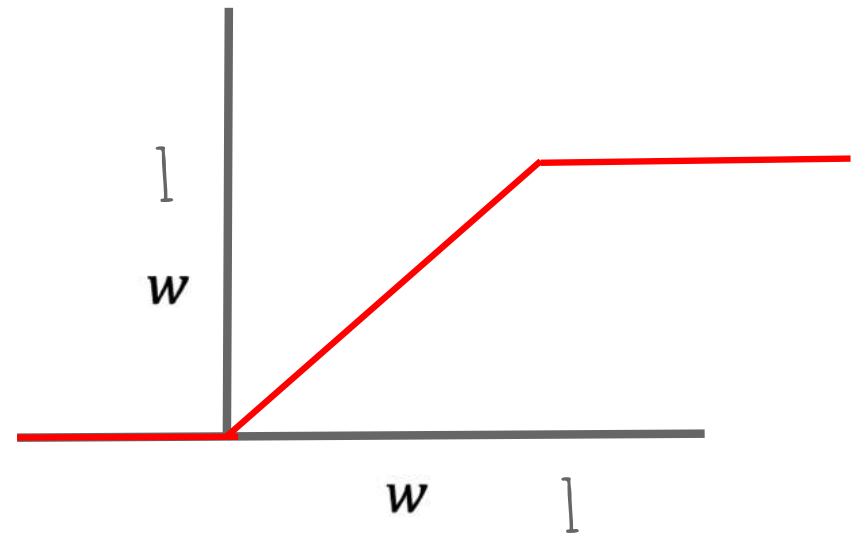
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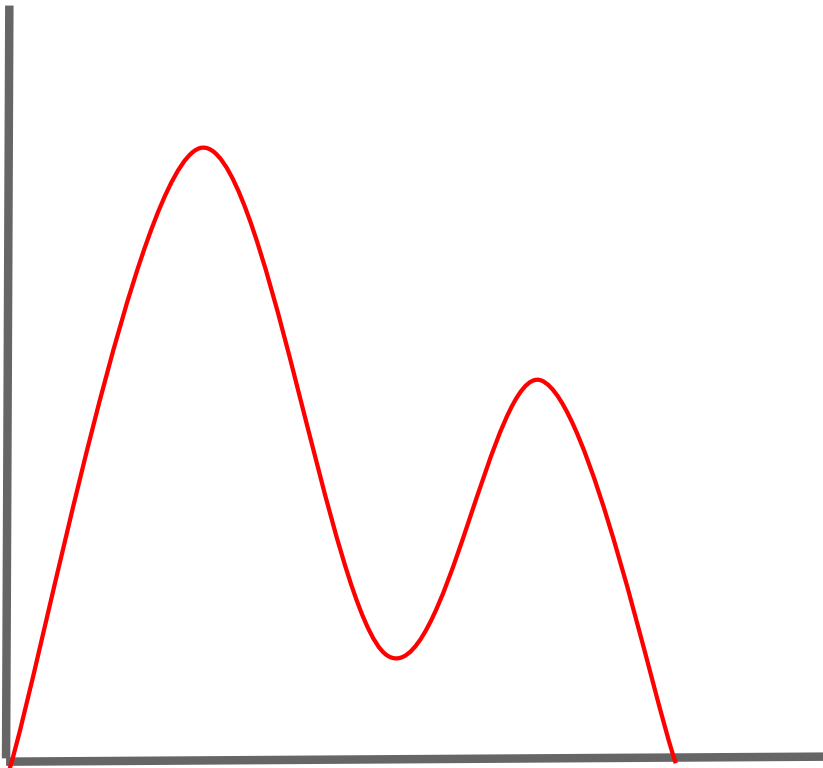
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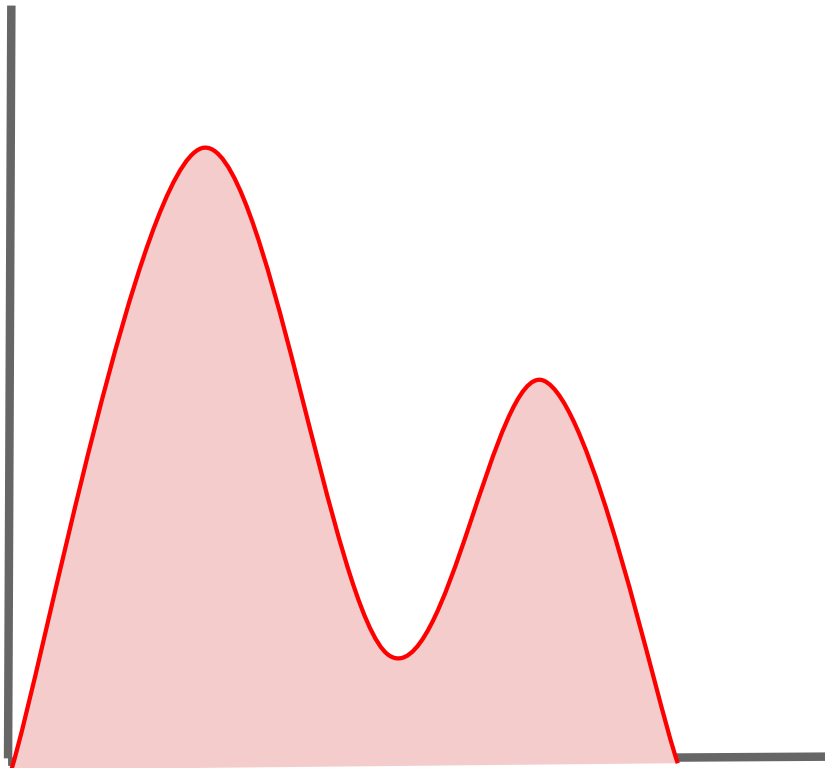
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# PDF INTUITION

$$f_X(z) \geq 0 \text{ for all } z \in \mathbb{R}$$



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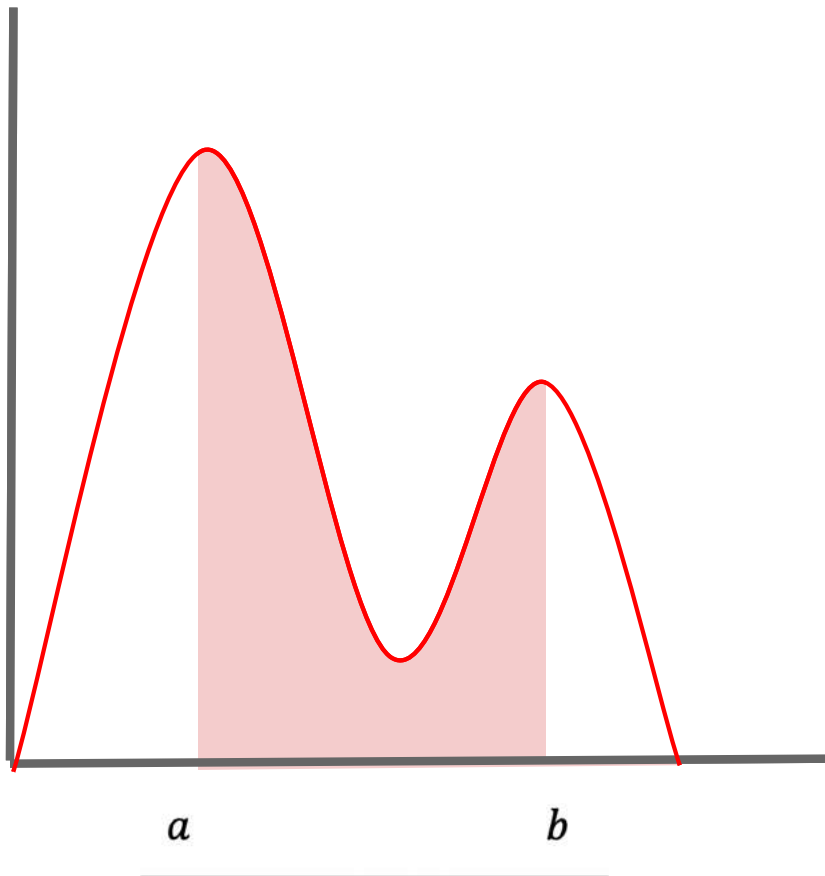


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$$\int_{-\infty}^{\infty} f_X(t) dt = 1$$



# PDF INTUITION



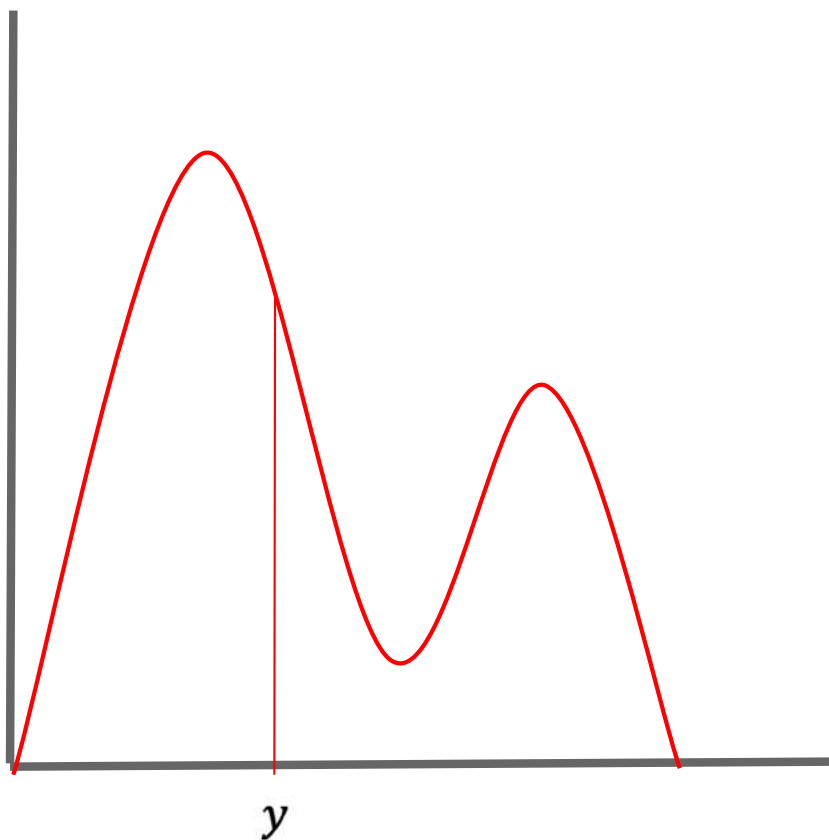
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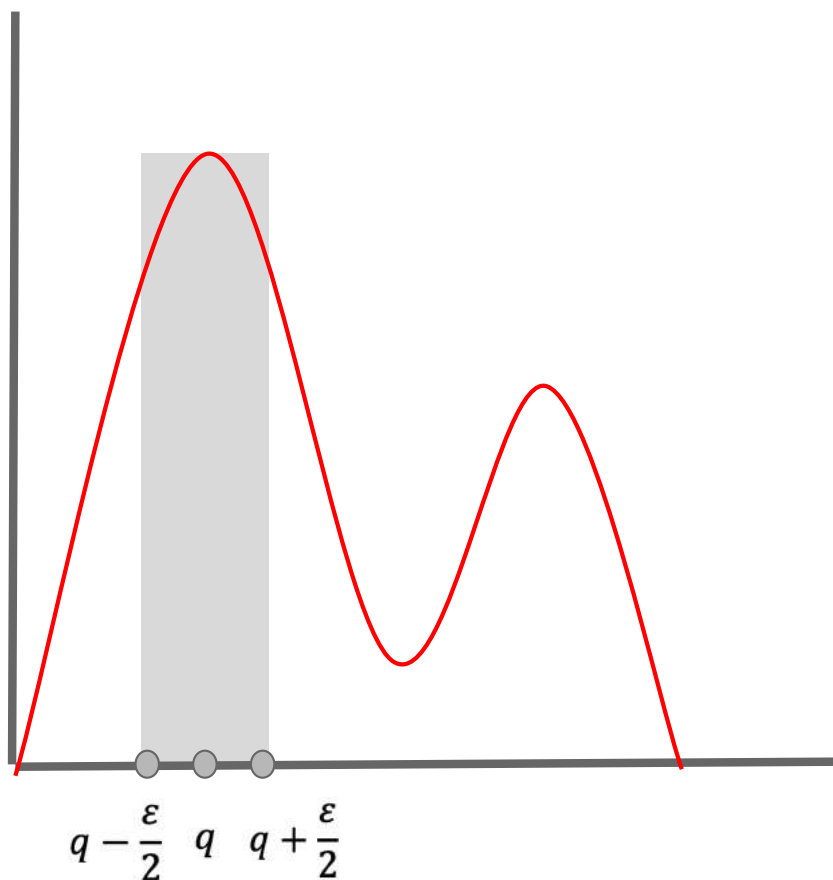
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$$P(a \leq X \leq b) = \int_a^b f_X(w) dw$$

$$P(X = y) = P(y \leq X \leq y) = \int_y^y f_X(w) dw = 0$$



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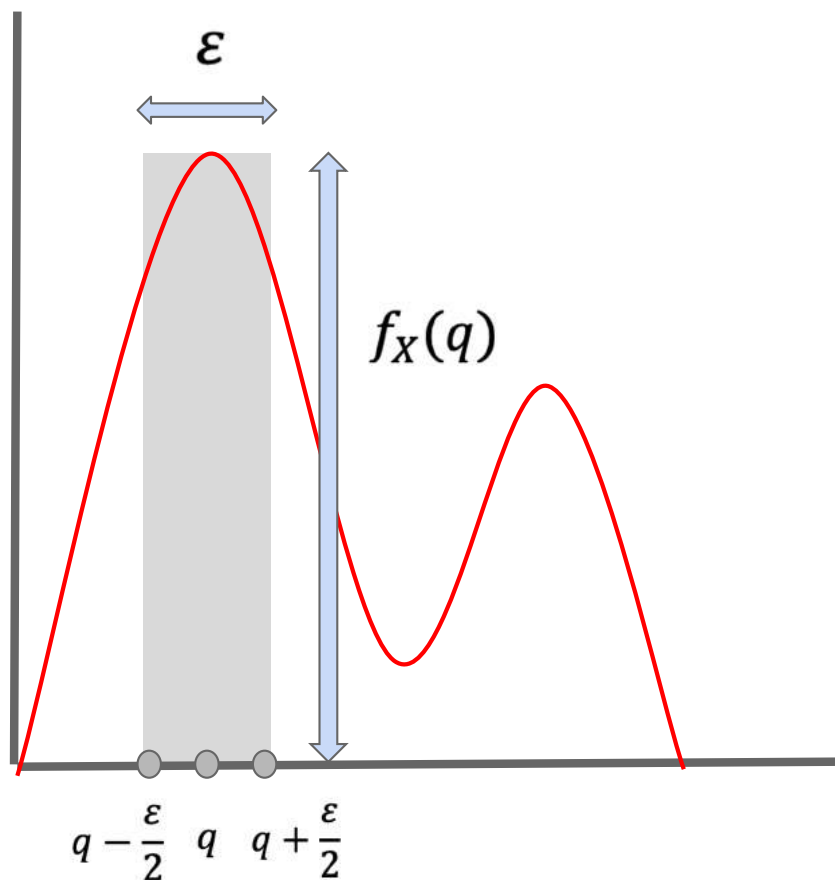
$$P(a \leq X \leq b) = \int_a^b f_X(w) dw$$

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$$P(X \approx q) \approx P\left(q - \frac{\epsilon}{2} \leq X \leq q + \frac{\epsilon}{2}\right)$$



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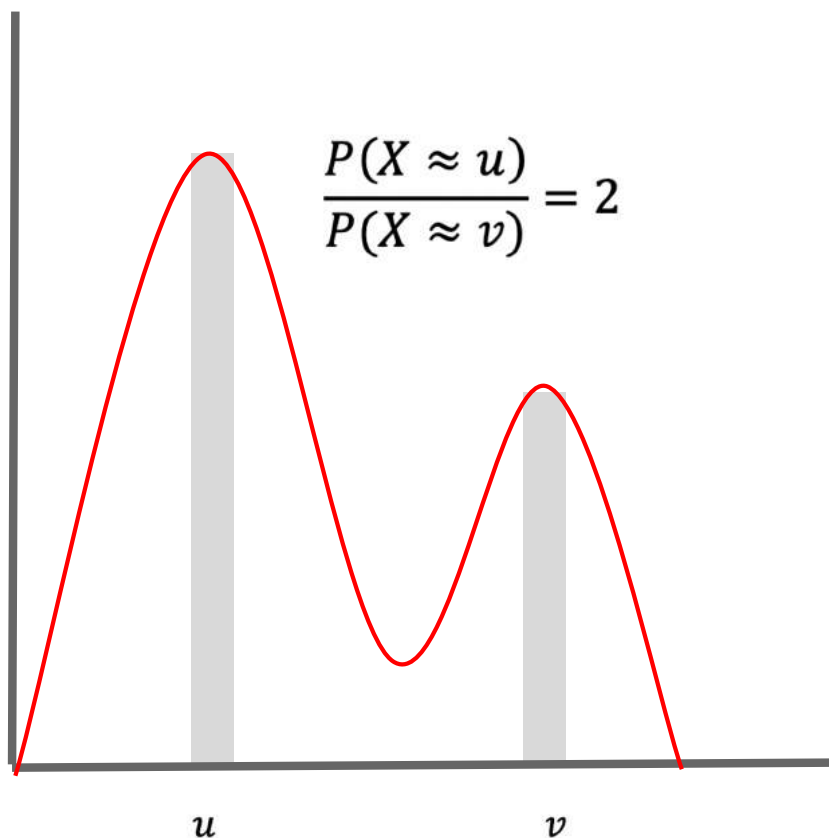
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$$P(a \leq X \leq b) = \int_a^b f_X(w) dw$$

$$P(X = y) = P(y \leq X \leq y) = \int_y^y f_X(w) dw = 0$$

$$P(X \approx q) \approx P\left(q - \frac{\varepsilon}{2} \leq X \leq q + \frac{\varepsilon}{2}\right) \approx \varepsilon f_X(q)$$

$$\frac{P(X \approx u)}{P(X \approx v)} = \frac{\varepsilon f_X(u)}{\varepsilon f_X(v)} = \frac{f_X(u)}{f_X(v)}$$



# PROBABILITY DENSITY FUNCTIONS (PDFS)

**Probability Density Function (PDF):** Let  $X$  be a continuous rv (one whose range is typically an interval or union of intervals). The probability density function (PDF) of  $X$  is the function  $f_X: \mathbb{R} \rightarrow \mathbb{R}$  such that

- $f_X(z) \geq 0$  for all  $z \in \mathbb{R}$ .
- $\int_{-\infty}^{\infty} f_X(t) dt = 1$ .
- $P(a \leq X \leq b) = \int_a^b f_X(w) dw$ .
- $P(X = y) = 0$  for any  $y \in \mathbb{R}$ .
- The probability that  $X$  is close to  $q$  is proportional to  $f_X(q)$ :  $P(X \approx q) \approx P\left(q - \frac{\varepsilon}{2} \leq X \leq q + \frac{\varepsilon}{2}\right) \approx \varepsilon f_X(q)$ .
- Ratios of probabilities of being near points are maintained:  $\frac{P(X \approx u)}{P(X \approx v)} =$

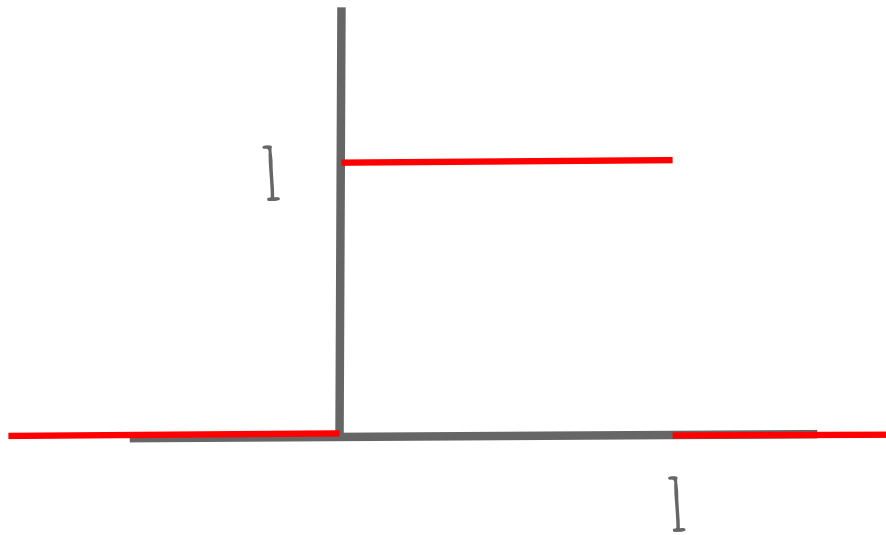
$$\frac{\varepsilon f_X(u)}{\varepsilon f_X(v)} = \frac{f_X(u)}{f_X(v)}$$

# RANDOM PICTURE



# CDF INTUITION

$f_X(v)$

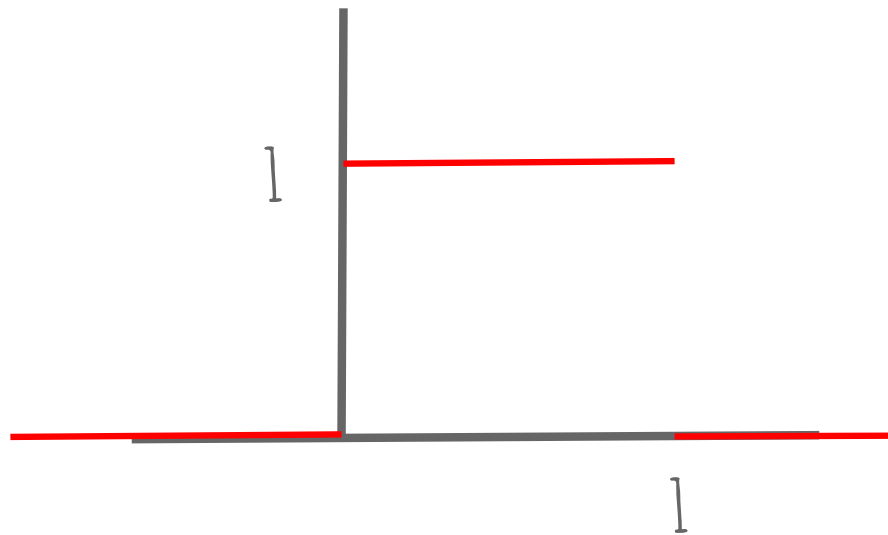


$$f_X(v) = \begin{cases} 1, & 0 \leq v \leq 1 \\ 0, & \text{otherwise} \end{cases}$$



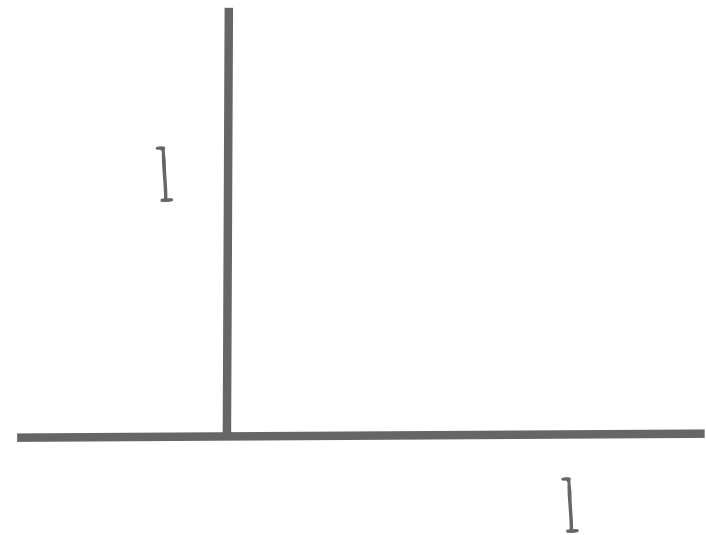
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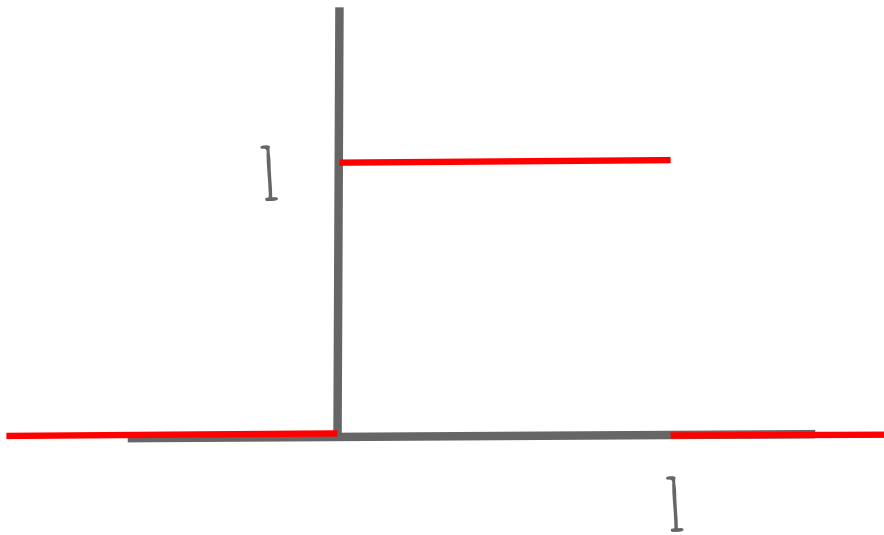


$$F_X(w) = \begin{cases} 0, & w < 0 \\ w, & 0 \leq w \leq 1 \\ 1, & w > 1 \end{cases}$$



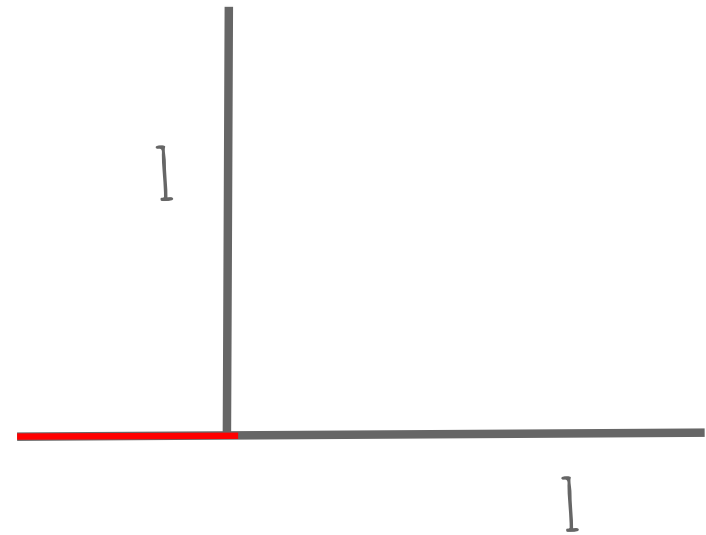
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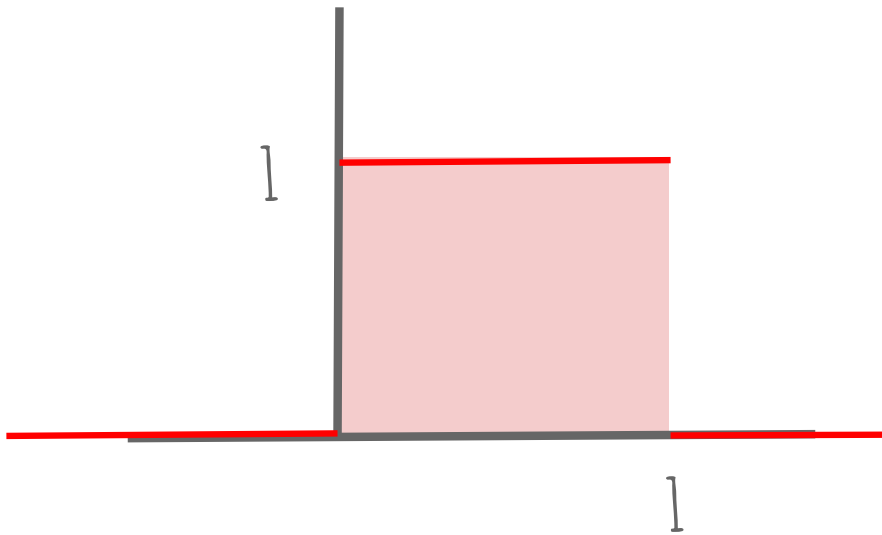


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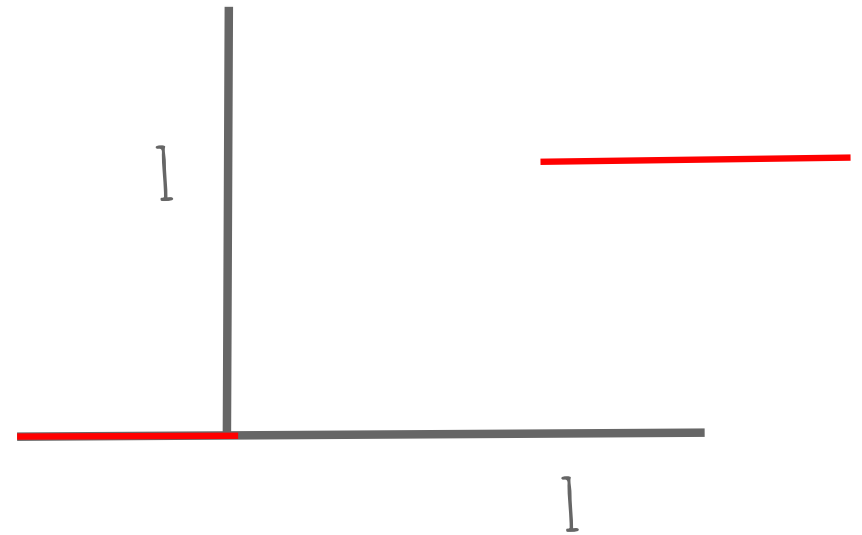
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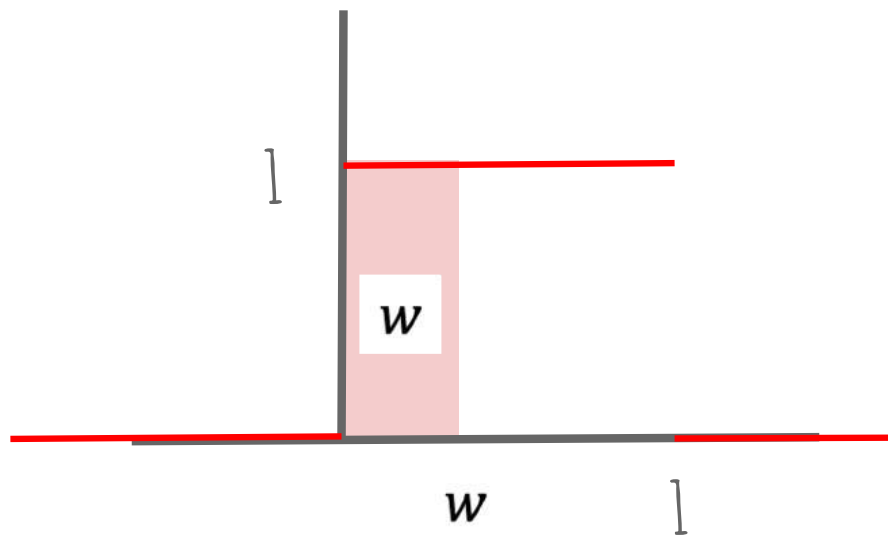
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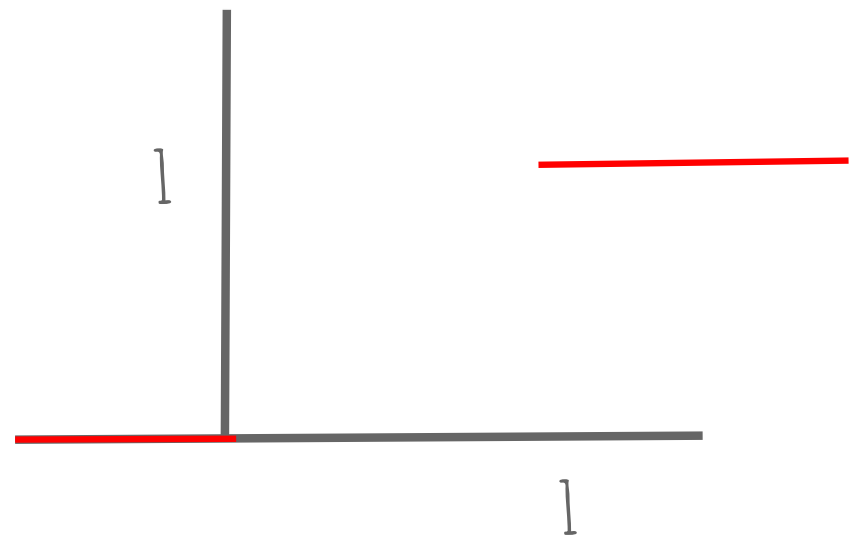
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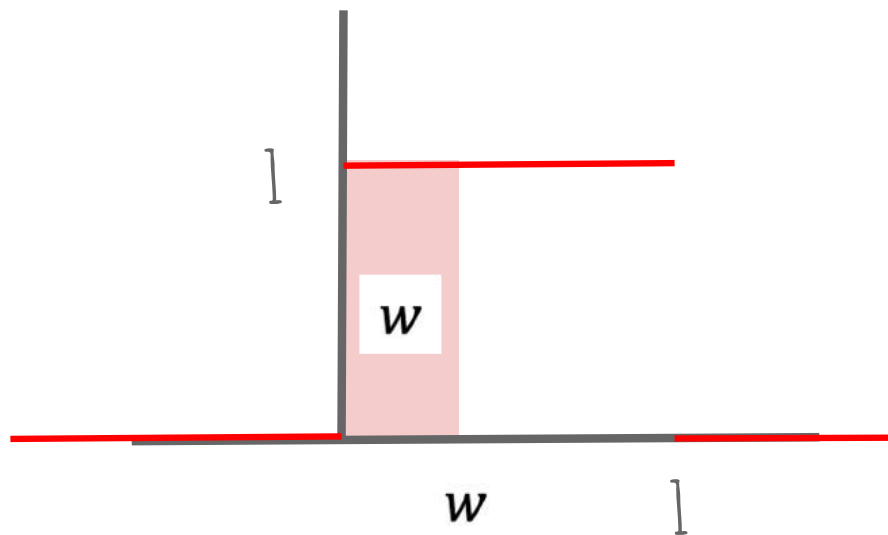
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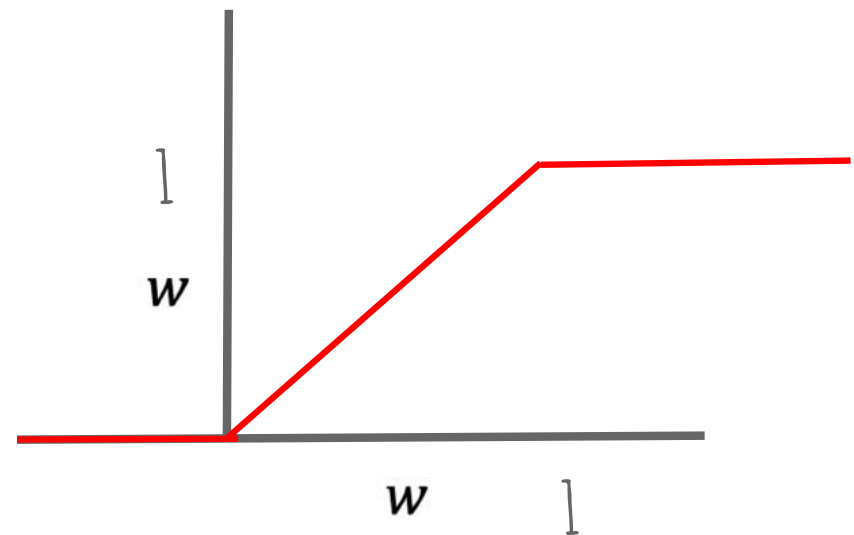
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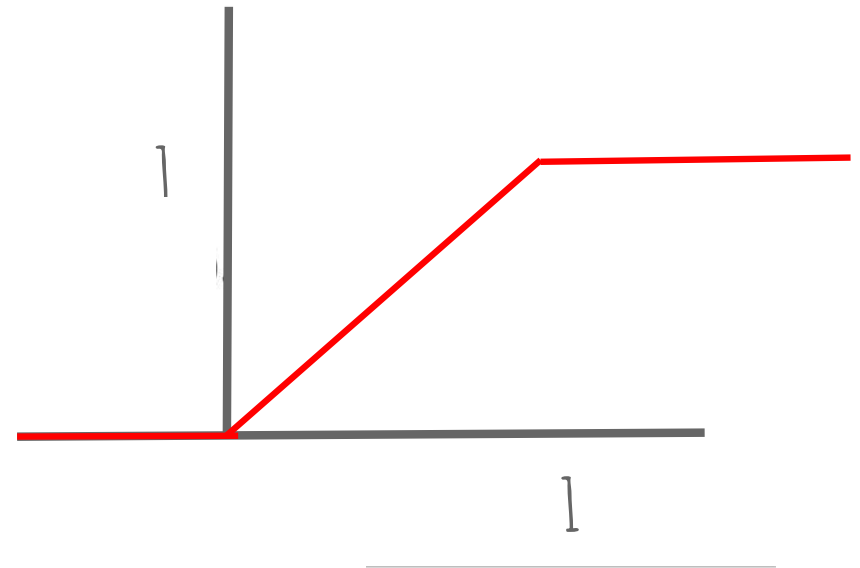


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$$F_X(t) = P(X \leq t) = \int_{-\infty}^t f_X(w) dw \text{ for all } t \in \mathbb{R}.$$

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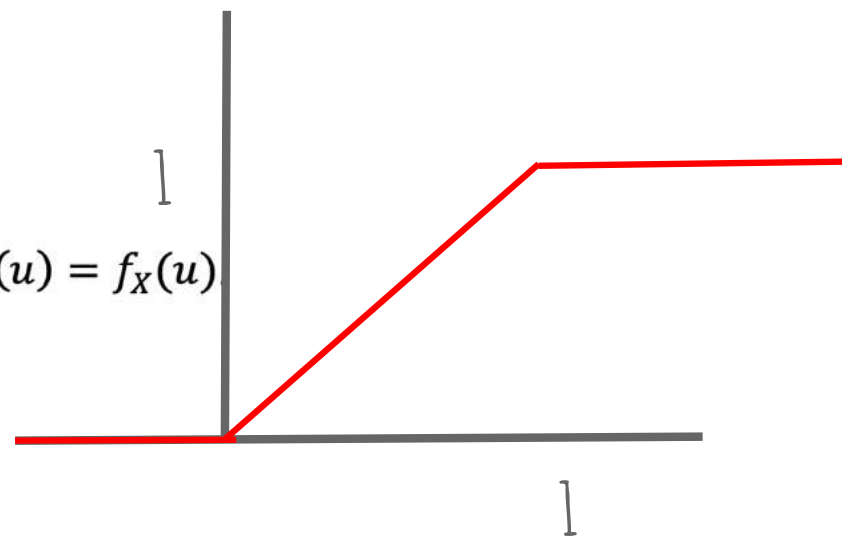
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Hence, by the Fundamental Theorem of Calculus,  $\frac{d}{du} F_X(u) = f_X(u)$



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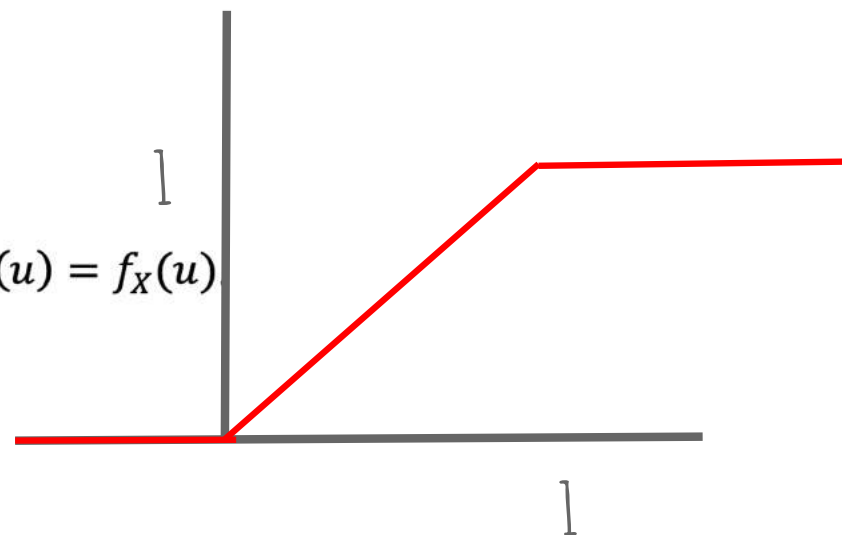


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$$P(a \leq X \leq b) = F_X(b) - F_X(a).$$



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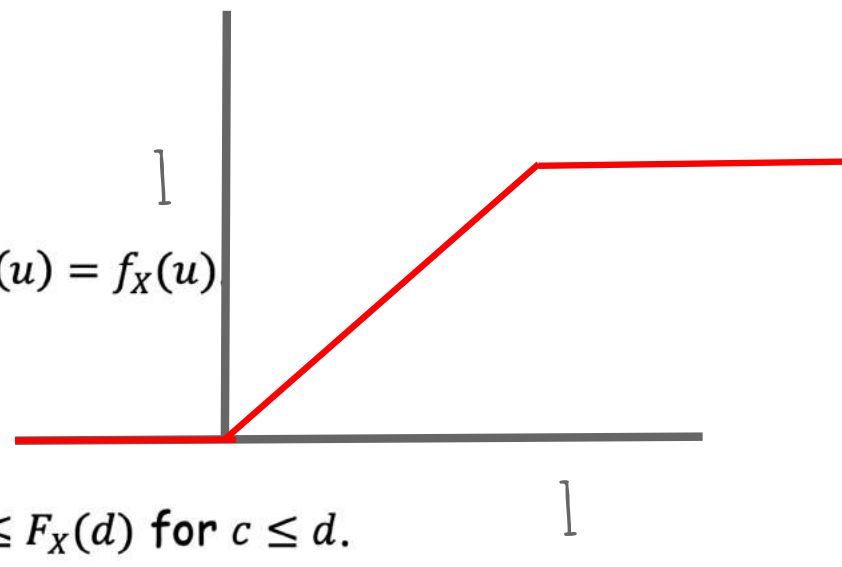
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$F_X$  is monotone increasing, since  $f_X \geq 0$ . That is,  $F_X(c) \leq F_X(d)$  for  $c \leq d$ .



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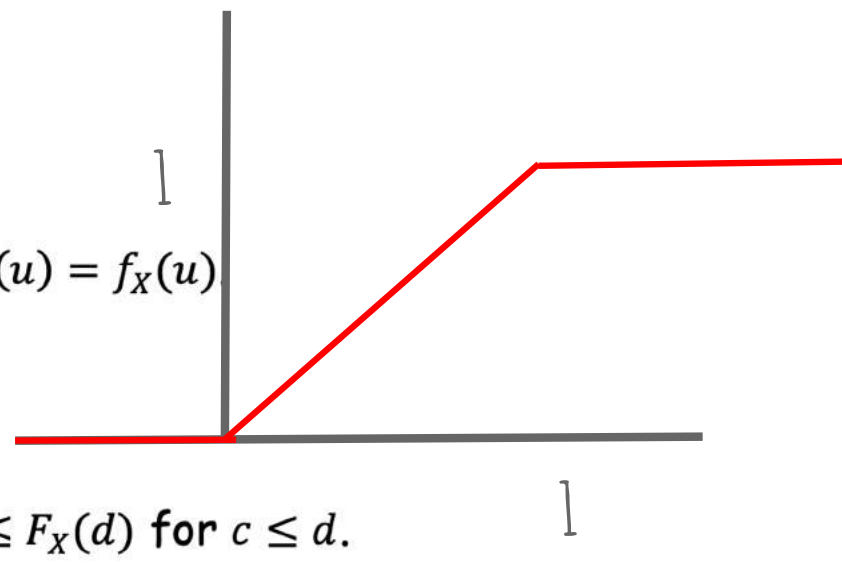
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$$\lim_{v \rightarrow -\infty} F_X(v) = P(X \leq -\infty) = 0.$$



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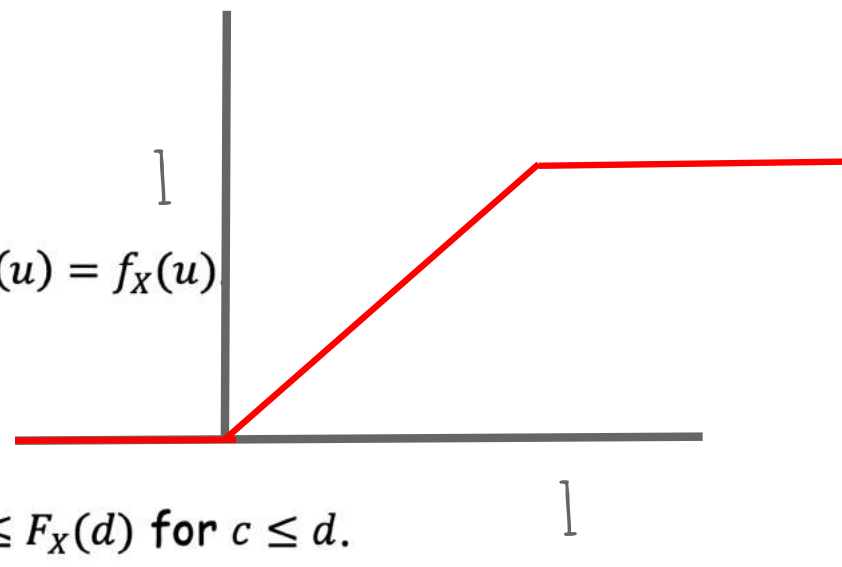
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$$\lim_{v \rightarrow +\infty} F_X(v) = P(X \leq +\infty) = 1.$$



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# CUMULATIVE DISTRIBUTION FUNCTIONS (CDFs)

**Cumulative Distribution Function (CDF):** Let  $X$  be a continuous rv (one whose range is typically an interval or union of intervals). The cumulative distribution function (CDF) of  $X$  is the function  $F_X: \mathbb{R} \rightarrow \mathbb{R}$  such that

- $F_X(t) = P(X \leq t) = \int_{-\infty}^t f_X(w)dw$  for all  $t \in \mathbb{R}$ .
- Hence, by the Fundamental Theorem of Calculus,  $\frac{d}{du}F_X(u) = f_X(u)$ .
- $P(a \leq X \leq b) = F_X(b) - F_X(a)$ .
- $F_X$  is monotone increasing, since  $f_X \geq 0$ . That is,  $F_X(c) \leq F_X(d)$  for  $c \leq d$ .
- $\lim_{v \rightarrow -\infty} F_X(v) = P(X \leq -\infty) = 0$ .
- $\lim_{v \rightarrow +\infty} F_X(v) = P(X \leq +\infty) = 1$ .





# FROM DISCRETE TO CONTINUOUS

	<b>Discrete</b>	<b>Continuous</b>
<b>PMF/PDF</b>	$p_X(x) = P(X = x)$	$f_X(x) \neq P(X = x) = 0$
<b>CDF</b>	$F_X(x) = \sum_{t \leq x} p_X(t)$	$F_X(x) = \int_{-\infty}^x f_X(t) dt$
<b>Normalization</b>	$\sum_x p_X(x) = 1$	$\int_{-\infty}^{\infty} f_X(x) dx = 1$
<b>Expectation</b>	$\mathbb{E}[g(X)] = \sum_x g(x)p_X(x)$	$\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} g(x)f_X(x) dx$



# PROBABILITY

## 4.2 ZOO OF CONTINUOUS RVS



# AGENDA

- THE (CONTINUOUS) UNIFORM RV
- THE EXPONENTIAL RV
- MEMORYLESSNESS

# THE (CONTINUOUS) UNIFORM RV

**Uniform (Continuous) RV:**  $X \sim \text{Unif}(a, b)$  where  $a < b$  are real numbers, if and only if  $X$  has the following pdf:

$$f_X(x) = \begin{cases} \frac{1}{b-a}, & x \in [a, b] \\ 0, & \text{otherwise} \end{cases}$$

$X$  is equally likely to take on any value in  $[a, b]$ .

# THE UNIFORM (CONTINUOUS) RV

**Uniform (Continuous) RV:**  $X \sim \text{Unif}(a, b)$  where  $a < b$  are real numbers, if and only if  $X$  has the following pdf:

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$X$  is equally likely to take on any value in  $[a, b]$ .

$$E[X] = \frac{a+b}{2} \quad \text{Var}(X) = \frac{(b-a)^2}{12}$$

The cdf is

$$F_X(x) = \begin{cases} 0, & x < a \\ \frac{x-a}{b-a}, & a \leq x \leq b \\ 1, & x > b \end{cases}$$

# THE EXPONENTIAL PDF/CDF



Recall the Poisson Process with parameter  $\lambda > 0$  has events happening at average rate of  $\lambda$  per unit of time forever. The exponential RV measures the time until the first occurrence of an event, so is a continuous RV with range  $[0, \infty)$  (unlike the Poisson RV, which counts the number of occurrences in a unit of time, with range  $\{0, 1, 2, \dots\}$ .)



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Let  $Y \sim \text{Exp}(\lambda)$  be the time until the first event. We'll compute  $F_Y(t)$  and  $f_Y(t)$ .

Let  $X(t) \sim \text{Poi}(\lambda t)$  be the # of events in the first  $t$  units of time, for  $t \geq 0$ .

# THE EXPONENTIAL PDF/CDF



Recall the Poisson Process with parameter  $\lambda > 0$  has events happening at average rate of  $\lambda$  per unit of time forever. The exponential RV measures the time until the first occurrence of an event, so is a continuous RV with range  $[0, \infty)$  (unlike the Poisson RV, which counts the number of occurrences in a unit of time, with range  $\{0, 1, 2, \dots\}$ .)

Let  $Y \sim \text{Exp}(\lambda)$  be the time until the first event. We'll compute  $F_Y(t)$  and  $f_Y(t)$ .

Let  $X(t) \sim \text{Poi}(\lambda t)$  be the # of events in the first  $t$  units of time, for  $t \geq 0$ .

$$P(Y > t) = P(\text{no events in first } t \text{ units}) = P(X(t) = 0) = e^{-\lambda t} \frac{(\lambda t)^0}{0!} = e^{-\lambda t}$$

$$F_Y(t) = P(Y \leq t) = 1 - P(Y > t) = 1 - e^{-\lambda t}$$

$$f_Y(t) = \frac{d}{dt} F_Y(t) = \lambda e^{-\lambda t}$$

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$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^{\infty} x \cdot \lambda e^{-\lambda x} dx = \frac{1}{\lambda}$$

$$E[X^2] = \int_{-\infty}^{\infty} x^2 f_X(x) dx = \int_0^{\infty} x^2 \cdot \lambda e^{-\lambda x} dx = \frac{2}{\lambda^2}$$

$$\text{Var}(X) = E[X^2] - E[X]^2 = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$$

# THE EXPONENTIAL RV

**Exponential RV**:  $X \sim \text{Exp}(\lambda)$ , if and only if  $X$  has the following pdf:

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

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$$E[X] = \frac{1}{\lambda} \quad \text{Var}(X) = \frac{1}{\lambda^2}$$

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# RANDOM PICTURE



# MEMORYLESSNESS (INTUITION)

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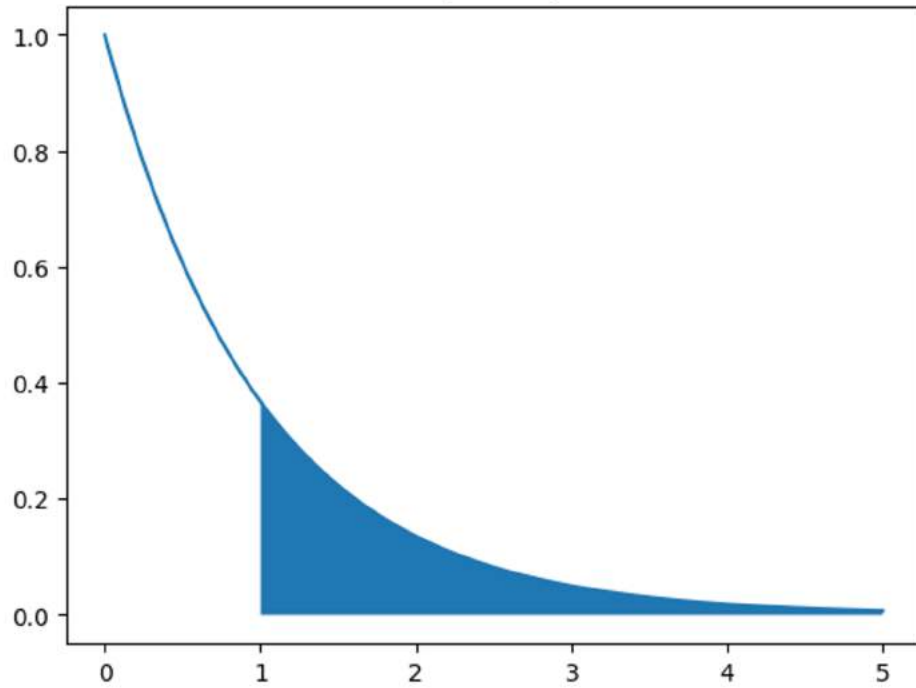
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The only memoryless RVs are the **Geometric** (discrete) and **Exponential** (continuous)!

# MEMORYLESSNESS (INTUITION)



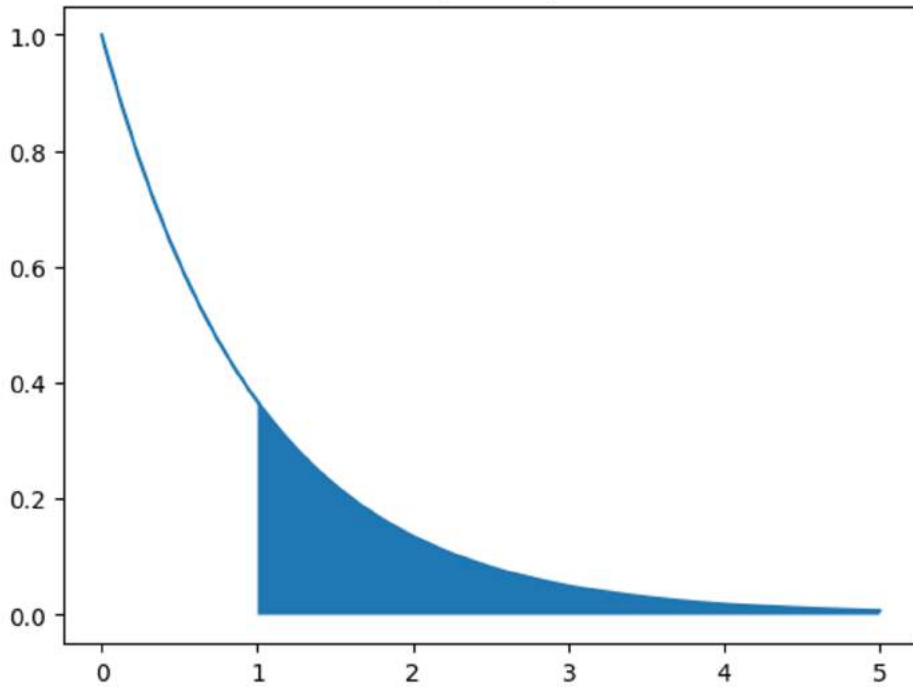
$P(X > 1)$



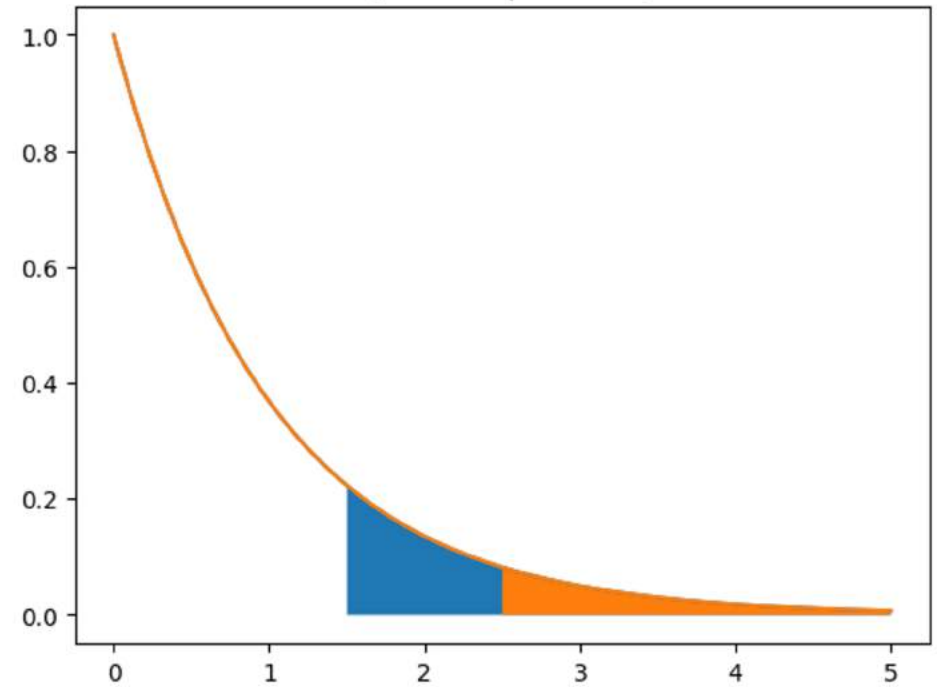
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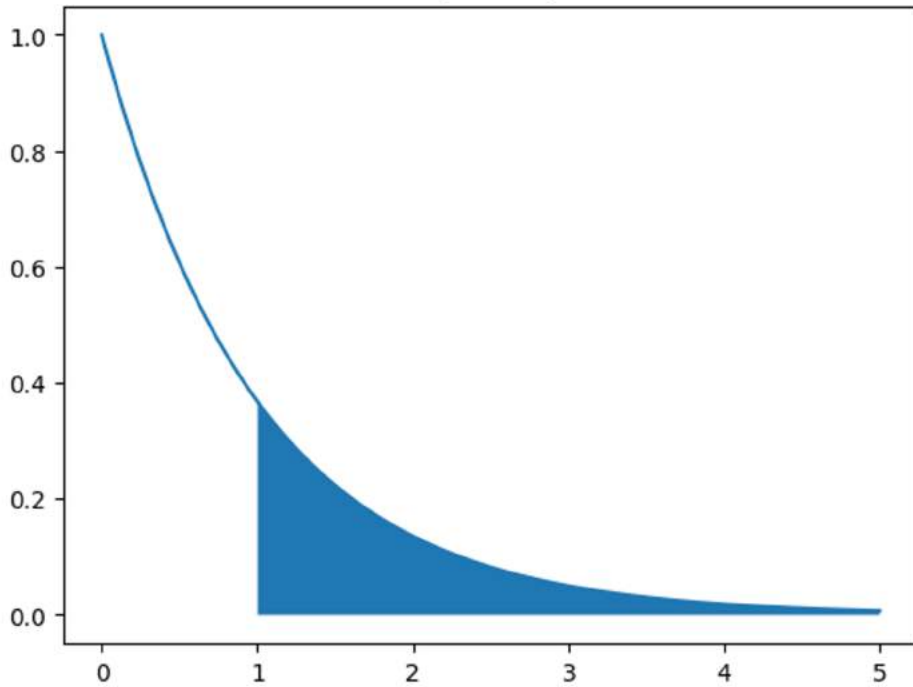
$P(X > 2.5 | X > 1.5)$



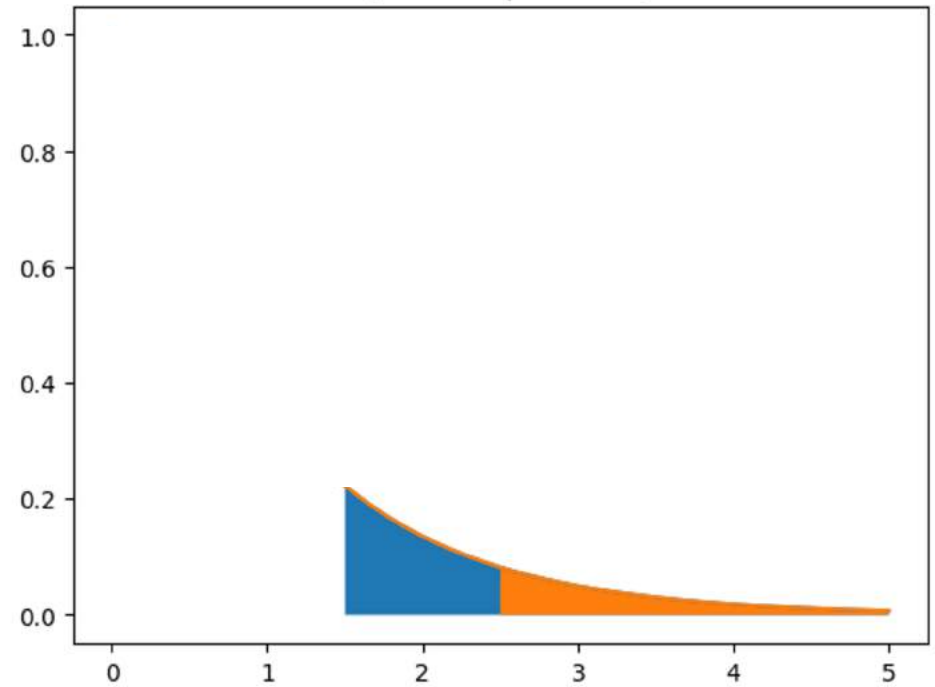
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# MEMORYLESSNESS OF EXPONENTIAL (PROOF)



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$$\begin{aligned} P(X > s + t | X > s) &= \frac{P(X > s | X > s + t)P(X > s + t)}{P(X > s)} \\ &= \frac{P(X > s + t)}{P(X > s)} \\ &= \frac{e^{-\lambda(s+t)}}{e^{-\lambda s}} \\ &= e^{-\lambda t} \\ &= P(X > t) \end{aligned}$$



# THE GAMMA RV

**Gamma RV**:  $X \sim \text{Gamma}(r, \lambda)$  if and only if  $X$  has the following pdf:

$$f_X(x) = \begin{cases} \frac{\lambda^r}{(r-1)!} x^{r-1} e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

$X$  is the sum of  $r$  independent  $\text{Exp}(\lambda)$  random variables.

