

Problem 3 on pset 3:
parts d) & e) extra credit.

3.2 MORE ON EXPECTATION

3.3 VARIANCE AND STANDARD DEVIATION

ANNA KARLIN

MOST SLIDES BY ALEX TSUN

AGENDA

- LINEARITY OF EXPECTATION (LOE) *recap*
- LAW OF THE UNCONSCIOUS STATISTICIAN (LOTUS)
- VARIANCE
- INDEPENDENCE OF RANDOM VARIABLES
- PROPERTIES OF VARIANCE

LINEARITY OF EXPECTATION (LOE)

Linearity of Expectation: Let Ω be the sample space of an experiment, $X, Y: \Omega \rightarrow \mathbb{R}$ be (possibly "dependent") random variables both defined on Ω , and $a, b, c \in \mathbb{R}$ be scalars. Then,

$$E[X + Y] = E[X] + E[Y]$$

and

$$E[aX + b] = aE[X] + b$$

Combining them gives,

$$E[aX + bY + c] = aE[X] + bE[Y] + c$$

COROLLARY: LINEARITY FOR SUM OF LOTS OF R.V.S

$$E(X_1 + X_2 + \dots + X_n) = E(X_1) + E(X_2) + \dots + E(X_n)$$

Proof by induction!

INDICATOR RANDOM VARIABLE

- For any event A , can define the indicator random variable for A



$$X_A = \begin{cases} 1 \\ 0 \end{cases}$$

$w \in A$
o.w.

COMPUTING COMPLICATED EXPECTATIONS

- Often boils down to finding the right way to decompose the random variable into simple random variables (often indicator random variables) and then applying linearity of expectation.

LINEARITY IS SPECIAL!

- In general $E(g(X)) \neq g(E(X))$

$$X = \begin{cases} 1 & \text{with prob } 1/2 \\ -1 & \text{with prob } 1/2 \end{cases}$$

$$E(X) = 1 \cdot \frac{1}{2} + (-1) \cdot \frac{1}{2} = 0$$

$$E(X^2) = 1 \cdot 1 = 1$$

$$E(X^2) \neq [E(X)]^2$$

$$Y = X^2$$

$Y = 1$ with prob 1

~~$$\begin{aligned} E[X \cdot Y] &= E[X] \cdot E[Y] \\ E[X^2] &= E[X]^2 \\ E[X/Y] &= E[X] / E[Y] \\ E[\sinh(X)] &= \sinh(E[X]) \\ &\vdots \end{aligned}$$~~

	$E(X)$	$E(X^2)$
a)	0	0
b)	0	1
c)	1	1
d)	-1	1

$$E(X^2) = (1)^2 \cdot \frac{1}{2} + (-1)^2 \cdot \frac{1}{2}$$

$$g(x) = x^2$$

LINEARITY IS SPECIAL!

- In general $E(g(X)) \neq g(E(X))$

- How DO we compute $E(g(X))$?

$$E(Y)$$

proof of
 $P_Y(y)$

$$E(Y) = \sum_{y \in \mathcal{U}_Y} y P_Y(y)$$

HOMWORKS OF 3 STUDENTS RETURNED RANDOMLY

- Each permutation equally likely
- X : # people who get their own homework
- What is $E(X^2 \bmod 2)$?

Prob	Outcome w	$X(w)$
1/6	1 2 3	3
1/6	1 3 2	1
1/6	2 1 3	1
1/6	2 3 1	0
1/6	3 1 2	0
1/6	3 2 1	1

$Y(w)$

$3^2 \bmod 2 = 1$

1

1

0

0

0

1

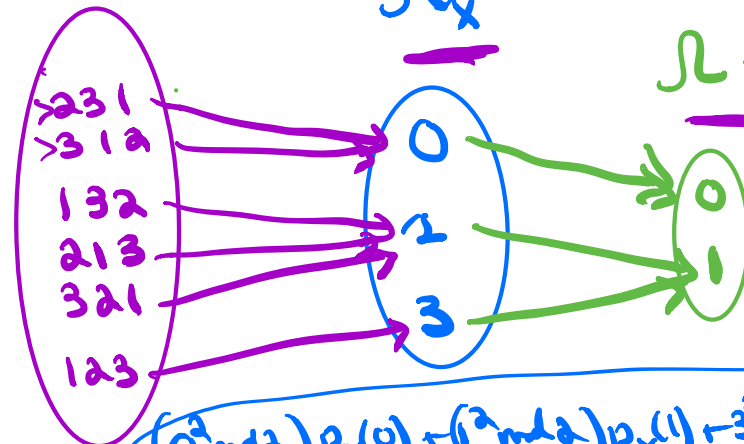
$$Y = X^2 \bmod 2.$$

$$g(x) = x^2 \bmod 2.$$

Ω

Ω_X

Ω_Y



$$P_X(x) = \begin{cases} \frac{1}{3} & x=0 \\ \frac{2}{6} & x=1 \\ \frac{1}{6} & x=3 \end{cases}$$

$$P_Y(y) = \begin{cases} \frac{1}{3} & y=0 \\ \frac{2}{3} & y=1 \end{cases}$$

$$Y = X^2 \bmod 2.$$

$$(0^2 \bmod 2) p_X(0) + (1^2 \bmod 2) p_X(1) + (3^2 \bmod 2) p_X(3)$$

$$E(Y) = \sum_{\omega \in \Omega} g(X(\omega)) P(\omega)$$

$$= \sum_{x \in \Omega_X} g(x) p_X(x) = \sum_{y \in \Omega_Y} y p_Y(y)$$

$$Y = g(X)$$

LAW OF THE UNCONSCIOUS STATISTICIAN (LOTUS)

Law of the Unconscious Statistician (LOTUS): Let X be a discrete random variable with range Ω_X and $g: D \rightarrow \mathbb{R}$ be a function defined at least over Ω_X ($\Omega_X \subseteq D$). Then,

$$E[g(X)] = \sum_{b \in \Omega_X} g(b) p_X(b)$$

Note that in general, $E[g(X)] \neq g(E[X])$. For example, $E[X^2] \neq (E[X])^2$, or $E[\log(X)] \neq \log(E[X])$.

$$E[X^2] = \sum_{k \in \Omega_X} k^2 p_X(k)$$



VARIANCE (INTUITION)



WHICH GAME WOULD YOU RATHER PLAY? WE FLIP A FAIR COIN.

GAME 1:

- IF HEADS, YOU PAY ME \$1.
- IF TAILS, I PAY YOU \$1.

$$1 \cdot \frac{1}{2} + (-1) \cdot \frac{1}{2}$$

GAME 2:

- IF HEADS, YOU PAY ME \$1000.
- IF TAILS, I PAY YOU \$1000.

$$1000 \cdot \frac{1}{2} + (-1000) \cdot \frac{1}{2}$$

my
E(payoff)

	Game 1	Game 2
a)	0	0
b)	0	1000
c)	1	0
d)	1	1000

VARIANCE (INTUITION)



HOW FAR IS A RANDOM VARIABLE FROM ITS MEAN, ON AVERAGE?

$$X - E[X]$$

$$E(X - E(X)) = E(X) - \underbrace{E(E(X))}_{\text{number.}} = 0$$

$= E(X)$

VARIANCE (INTUITION)

HOW FAR IS A RANDOM VARIABLE FROM ITS MEAN, ON AVERAGE?

$$|X - E[X]|$$



VARIANCE (INTUITION)



HOW FAR IS A RANDOM VARIABLE FROM ITS MEAN, ON AVERAGE?

$$E[|X - E[X]|] \neq \sigma_x$$

$$E[(X - E[X])^2]$$

VARIANCE AND STANDARD DEVIATION (SD)

Variance: The variance of a random variable X is

$$E(aX) = aE(X)$$

$$\text{Var}(X) = E[(X - E[X])^2]$$

The variance is always nonnegative since we take an expectation of a nonnegative random variable $(X - E[X])^2$. We can also show that for any scalars $a, b \in \mathbb{R}$,

$$\text{Var}(aX + b) = a^2 \text{Var}(X)$$

$$\begin{aligned} E[(X - E(X))^2] &= E[X^2 - 2\mu X + \mu^2] \\ &= E[X^2] + E[-2\mu X] + E[\mu^2] \\ &\quad \quad \quad -2\mu E(X) \quad \quad \quad \mu^2 \\ &\quad \quad \quad -2\mu^2 \\ &= E[X^2] - \mu^2 \end{aligned}$$

VARIANCE AND STANDARD DEVIATION (SD)

Variance: The variance of a random variable X is

$$\text{Var}(X) = E[(X - E[X])^2] = E[X^2] - (E[X])^2$$

MORE USEFUL

The variance is always nonnegative since we take an expectation of a nonnegative random variable $(X - E[X])^2$. We can also show that for any scalars $a, b \in \mathbb{R}$,

$$\text{Var}(aX + b) = a^2 \text{Var}(X)$$

VARIANCE AND STANDARD DEVIATION (SD)

Variance: The variance of a random variable X is

MORE USEFUL

$$\text{Var}(X) = E[(X - E[X])^2] = E[X^2] - (E[X])^2$$

The variance is always nonnegative since we take an expectation of a nonnegative random variable $(X - E[X])^2$. We can also show that for any scalars $a, b \in \mathbb{R}$,

$$\text{Var}(aX + b) = a^2 \text{Var}(X)$$

Standard Deviation (SD): The standard deviation of a random variable X is

$$\sigma_X = \sqrt{\text{Var}(X)}$$

We want this because the units of variance are squared in terms of the original variable X , and this "undo's" our squaring, returning the units to the same as X .

VARIANCE (PROPERTY)

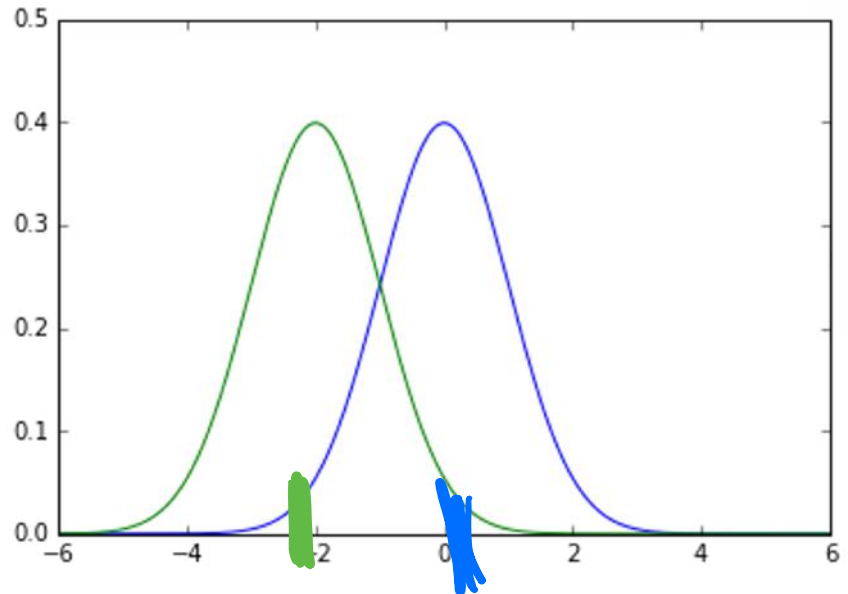
$\text{Var}(aX)$

//

$$\text{Var}(aX + b) = a^2 \text{Var}(X)$$



$$\text{Var}(X + b) = \text{Var}(X)$$



VARIANCE (PROPERTY)

$$\text{Var}(aX + b) = a^2 \text{Var}(X)$$

$$\begin{aligned} \text{Var}(aX) &= E[(aX)^2] - (E[aX])^2 = E[a^2X^2] - (aE[X])^2 \\ &= a^2E[X^2] - a^2(E[X])^2 = a^2(E[X^2] - E[X]^2) = a^2 \text{Var}(X) \end{aligned}$$

VARIANCE (EXAMPLE)



Let X be the outcome of a fair 6-sided die roll. What is $Var(X)$?

$$Var(X) = E[X^2] - E[X]^2$$

$$E[X] = 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6} + \dots + 6 \cdot \frac{1}{6} = 3.5$$

$$E[X^2] = 1^2 \cdot \frac{1}{6} + 2^2 \cdot \frac{1}{6} + 3^2 \cdot \frac{1}{6} + \dots + 6^2 \cdot \frac{1}{6} = \frac{91}{6}$$

$$Var(X) = E[X^2] - E[X]^2 = \frac{91}{6} - (3.5)^2 = \frac{35}{12}$$

LOTUS

VARIANCE



WHICH GAME WOULD YOU RATHER PLAY? WE FLIP A FAIR COIN.

GAME 1:

- IF HEADS, YOU PAY ME \$1.
- IF TAILS, I PAY YOU \$1.

GAME 2:

- IF HEADS, YOU PAY ME \$1000.
- IF TAILS, I PAY YOU \$1000.

$$\begin{aligned} \text{Var}(X) &= E(X^2) - [E(X)]^2 \\ &= E(X^2) - 0^2 \\ &= E(X^2) \end{aligned}$$

$$1^2 \cdot \frac{1}{2} + (-1)^2 \cdot \frac{1}{2} = 1$$

$X = \text{payoff in game}$

$\text{Var}(X)$

	<u>Game 1</u>	<u>Game 2</u>
a)	0	0
b)	1	1000
c)	1	1,000,000

$$(1000)^2 \cdot \frac{1}{2} + (-1000)^2 \cdot \frac{1}{2}$$

Game 1

$$\sigma_x = 1$$

$$\sigma_x = \sqrt{\text{Var}(X)}$$

Game 2

$$\sigma_x = \underline{\underline{1000}}$$

IN GENERAL, $\text{VAR}(X+Y) \neq \text{VAR}(X) + \text{VAR}(Y)$

Example 1:

- $X = \pm 1$ each with prob $\frac{1}{2}$
- $E(X) = 0$ $\text{Var}(X) = 1$
- How about $Y = -X$

$$\text{Var}(X) + \text{Var}(Y) = 2$$

$$\text{Var}(X+Y) = \text{Var}(0) = 0$$

$$E(Y) = 0 \quad \text{Var}(Y) = 1$$

Example 2: What is $\text{Var}(X+X)$?

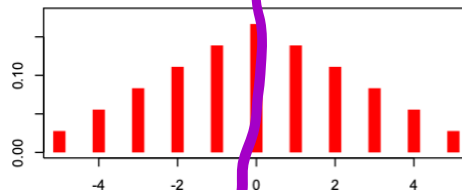
$$\underline{\underline{\neq 2 \text{Var}(X)}}$$

$$\text{Var}(2X) = 2^2 \text{Var}(X)$$

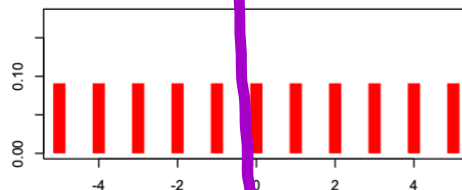
$$\text{Var}(aX) = a^2 \text{Var}(X)$$

VARIANCE IN PICTURES

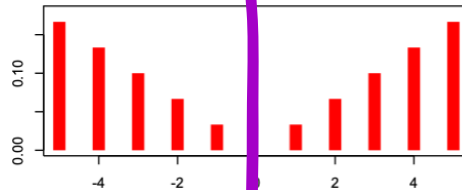
$$\sigma^2 = 5.83$$



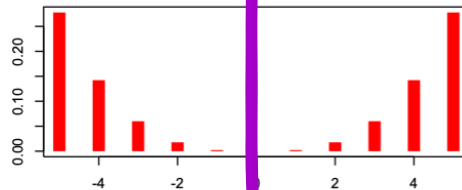
$$\sigma^2 = 10$$



$$\sigma^2 = 15$$



$$\sigma^2 = 19.7$$



RANDOM PICTURE



RANDOM VARIABLES AND INDEPENDENCE

Random variable X and event E are independent if the event E is independent of the event $\{X=x\}$ (for any fixed x), i.e.

$$\forall x \quad \underline{P(X = x \text{ and } E)} = \underline{P(X=x) \cdot P(E)}$$

Two random variables X and Y are independent if the events $\{X=x\}$ and $\{Y=y\}$ are independent for any fixed x, y, i.e.

$$\forall x, y \quad P(X = x \text{ and } Y=y) = P(X=x) \cdot P(Y=y)$$

$$P(X=x | Y=y) = P(X=x)$$

Intuition as before: knowing X doesn't help you guess Y or E and vice versa.

EXAMPLE

Random variable X and event E are independent if the event E is independent of the event $\{X=x\}$ (for any fixed x), i.e.

$$\forall x \ P(X = x \text{ and } E) = P(X=x) \cdot P(E)$$

Example: Let X be number of heads in n independent coin flips. Let E be the event that the number of heads is even.

$$Pr(X=k) = \binom{n}{k} \frac{1}{2^n}$$

X & E indep?

E happen

$$Pr(X=5 | E) \stackrel{?}{=} P(X=5)$$

0 = 0

- a) yes
- b) no

EXAMPLE

Two random variables X and Y are independent if the events $\{X=x\}$ and $\{Y=y\}$ are independent (for any fixed x, y), i.e.
 $\forall x, y P(X = x \text{ and } Y=y) = P(X=x) \cdot P(Y=y)$

Example: Let X be number of heads in first n of $2n$ independent coin flips, Y be number in the last n flips, and let Z be the total.

$n > 2$

$$Pr(X=n | Z=2) = 0$$

$$\neq P(X=n)$$

	$X \& Y$ indep.	$X \& Z$ indep.
a)	no	no
b)	no	yes
c)	yes	no
d)	yes	yes

EXAMPLE CONTINUED

Example: Let X be number of heads in first n of $2n$ independent coin flips, Y be number in the last n flips, and let Z be the total.

$$P(X = j) = \binom{n}{j} 2^{-n}$$

$$P(Y = k) = \binom{n}{k} 2^{-n}$$

$$P(X = j \wedge Y = k) = \binom{n}{j} \binom{n}{k} 2^{-2n} = P(X = j)P(Y = k)$$

IMPORTANT FACTS ABOUT INDEPENDENT RANDOM VARIABLES

Theorem: If X & Y are independent, then $E[X \cdot Y] = E[X] \cdot E[Y]$

Theorem: If X and Y are independent, then

$$\text{Var}[X + Y] = \text{Var}[X] + \text{Var}[Y]$$

Corollary: If $X_1 + X_2 + \dots + X_n$ are mutually independent then

$$\text{Var}[X_1 + X_2 + \dots + X_n] = \text{Var}[X_1] + \text{Var}[X_2] + \dots + \text{Var}[X_n]$$

$E[XY]$ FOR INDEPENDENT RANDOM VARIABLES

- **Theorem:** If X & Y are independent, then $E[X \cdot Y] = E[X] \cdot E[Y]$
- **Proof:**

Let $x_i, y_i, i = 1, 2, \dots$ be the possible values of X, Y .

$$\begin{aligned} E[X \cdot Y] &= \sum_i \sum_j x_i \cdot y_j \cdot P(X = x_i \wedge Y = y_j) \quad \leftarrow \text{independence} \\ &= \sum_i \sum_j x_i \cdot y_j \cdot P(X = x_i) \cdot P(Y = y_j) \\ &= \sum_i x_i \cdot P(X = x_i) \cdot \left(\sum_j y_j \cdot P(Y = y_j) \right) \\ &= E[X] \cdot E[Y] \end{aligned}$$

Note: *NOT* true in general; see earlier example $E[X^2] \neq E[X]^2$

VARIANCE OF A SUM OF INDEPENDENT R.V.S

Theorem: If X and Y are independent, then

$$\text{Var}[X + Y] = \text{Var}[X] + \text{Var}[Y]$$

Proof:

$$\begin{aligned}\text{Var}[X + Y] &= E[(X + Y)^2] - (E[X + Y])^2 \\ &= E[X^2 + 2XY + Y^2] - (E[X] + E[Y])^2 \\ &= E[X^2] + 2E[XY] + E[Y^2] - ((E[X])^2 + 2E[X]E[Y] + (E[Y])^2) \\ &= E[X^2] - (E[X])^2 + E[Y^2] - (E[Y])^2 + 2(E[XY] - E[X]E[Y]) \\ &= \text{Var}[X] + \text{Var}[Y] + 2(E[X]E[Y] - E[X]E[Y]) \\ &= \text{Var}[X] + \text{Var}[Y]\end{aligned}$$

x

→ using independence
of X & Y

INDEPENDENT VS DEPENDENT R.V.S

- Dependent r.v.s can reinforce/cancel/correlate in arbitrary ways.
- Independent r.v.s are, well, independent.

Example:

$$Z = X_1 + X_2 + \dots + X_n$$

X_i is indicator r.v. with probability 1/2 of being 1.

versus

$$W = n X_1$$

$$E(Z)$$

$$E(W)$$

*(Handwritten notes: E(Z) = n * 0.5, E(W) = n * 0.5)*

