

Maximum Likelihood Estimators for the Normal Distribution  $N(\theta_1, \theta_2)$ . ( $\theta_1 = \mu, \theta_2 = \sigma^2$ )

Given independent samples  $x_1, x_2, \dots, x_n \sim N(\theta_1, \theta_2)$ .

$$L(x_1, x_2, \dots, x_n | \theta_1, \theta_2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\theta_2}} e^{-(x_i - \theta_1)^2 / (2\theta_2)}$$

$$\ln L(x_1, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n \left( -\frac{1}{2} \ln(2\pi\theta_2) - \frac{(x_i - \theta_1)^2}{2\theta_2} \right)$$

$$\frac{\partial}{\partial \theta_1} \ln L(x_1, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n \frac{2(x_i - \theta_1)}{2\theta_2} = \sum_{i=1}^n \frac{x_i - \theta_1}{\theta_2} = 0$$

$$\sum_{i=1}^n (x_i - \hat{\theta}_1) = 0$$

$$\sum_{i=1}^n x_i = n \hat{\theta}_1$$

$$\hat{\theta}_1 = \frac{1}{n} \sum_{i=1}^n x_i$$

To show that  $\theta_1 = \hat{\theta}_1$  is a maximum:

$$\frac{\partial^2}{\partial \theta_1^2} \ln L(x_1, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n -\frac{1}{\theta_2} < 0$$

so  $\ln L$  is concave downward everywhere and  $\theta_1 = \hat{\theta}_1$  is a maximum.

So the maximum likelihood estimator for  $\theta_1 = \mu$  is the sample mean  $\frac{1}{n} \sum_{i=1}^n x_i$ .

$$\begin{aligned} \frac{\partial}{\partial \theta_2} \ln L(x_1, \dots, x_n | \theta_1, \theta_2) &= \sum_{i=1}^n \left( -\frac{1}{2} \cdot \frac{2\pi}{2\pi\theta_2} + \frac{2(x_i - \theta_1)^2}{(2\theta_2)^2} \right) \\ &= \sum_{i=1}^n \left( -\frac{1}{2\theta_2} + \frac{(x_i - \theta_1)^2}{2\theta_2^2} \right) = 0 \end{aligned}$$

$$\sum_{i=1}^n \left( -\hat{\theta}_2 + (x_i - \hat{\theta}_1)^2 \right) = 0$$

$$\sum_{i=1}^n (x_i - \hat{\theta}_1)^2 = n\hat{\theta}_2$$

$$\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\theta}_1)^2$$

Recall that  $\text{Var}(Y) = E[(Y - \mu)^2]$ , where  $\mu = E[Y]$

$\hat{\theta}_2$  is called a sample variance.

To show that  $\hat{\theta}_2 = \hat{\theta}_2$  ~~is~~ maximizes  $\ln L$ , ~~take~~ compute  $\frac{\partial^2}{\partial \theta_2^2} \ln L$  and show that it is negative at  $\hat{\theta}_2 = \hat{\theta}_2$ .

In general, if there are two parameters  $\theta_1$  and  $\theta_2$  in the distribution, when you set the two derivatives to 0 you will get 2 equations in 2 unknowns,  $\hat{\theta}_1$  and  $\hat{\theta}_2$ . Solve simultaneously.

Bias

~~Defn~~ Defn: An estimator  $\hat{\theta}$  of  $\theta$  is unbiased if  $E[\hat{\theta}] = \theta$ .

$\hat{\theta}$  is a function of  $x_1, x_2, \dots, x_n$ . If we think of  $x_1, \dots, x_n$  as random variables, then  $\hat{\theta}$  is also a random variable and it has an expected value. It's desirable that  $E[\hat{\theta}] = \theta$ .

For the estimator  $\hat{\theta}_1$  of  $N(\theta_1, \theta_2)$ :  $x_i \sim N(\theta_1, \theta_2)$

$$E[\hat{\theta}_1] = E\left[\frac{1}{n} \sum_{i=1}^n x_i\right] = \frac{1}{n} \sum_{i=1}^n E[x_i] = \frac{1}{n} \sum_{i=1}^n \theta_1 = \frac{1}{n} n \theta_1$$

$$= \theta_1, \text{ so } \hat{\theta}_1 \text{ is unbiased.}$$