

CSE 312

# Foundations of Computing II

## Lecture 8: Bayes Rule, Limited Independence



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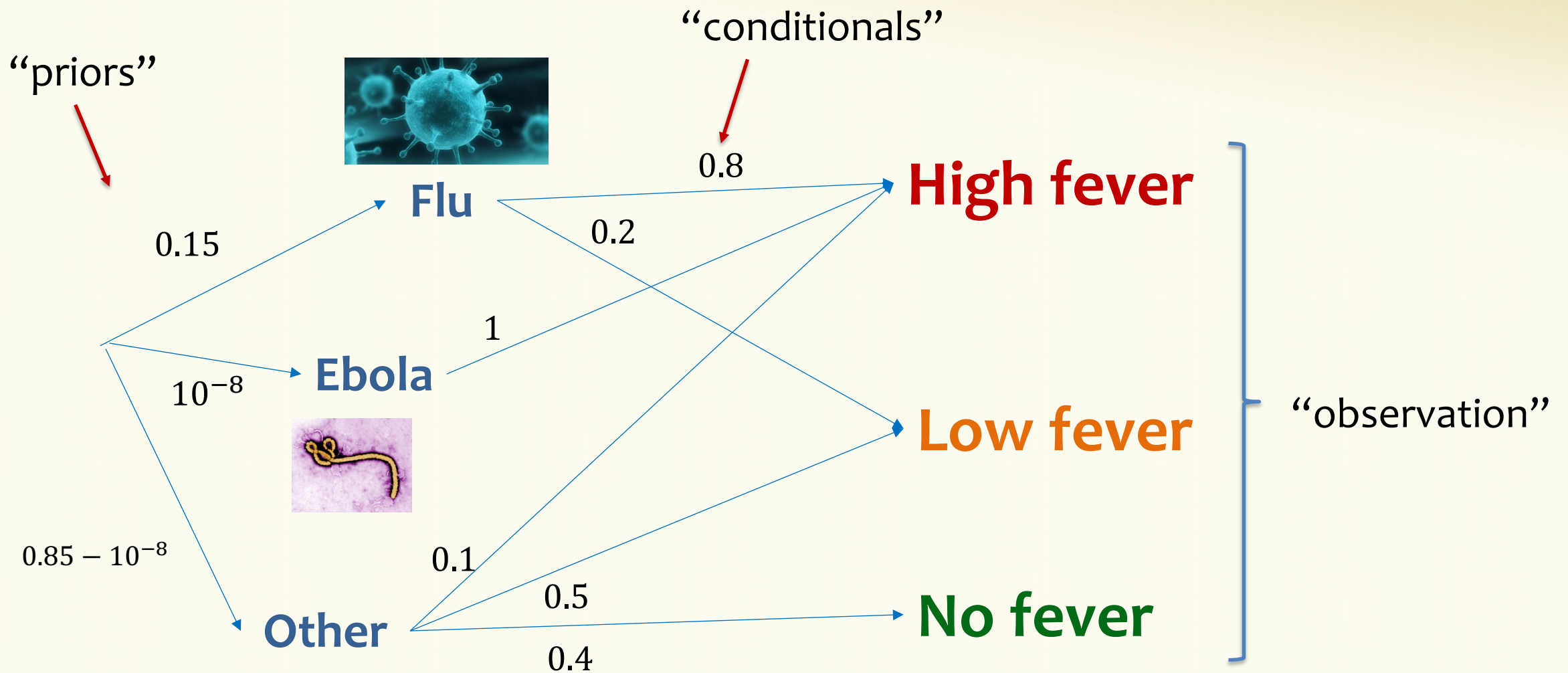
# Today

- Bayes Rule
- Independence of multiple events

# On LaTeX

- Overleaf is not the best approach for using LaTeX
  - Tool for collaborative editing of LaTeX documents.
  - Not needed for class.
  - Has become somewhat unstable.
- LaTeX is free software – you can find several installations, depending on OS.
- Several environment for LaTeX development, your favorite editor often will do.

## 7.1 – Bayes Rule



Assume we observe high fever, what is the probability that the subject has Ebola? Posterior:  $\mathbb{P}(\text{Ebola} | \text{High fever})$

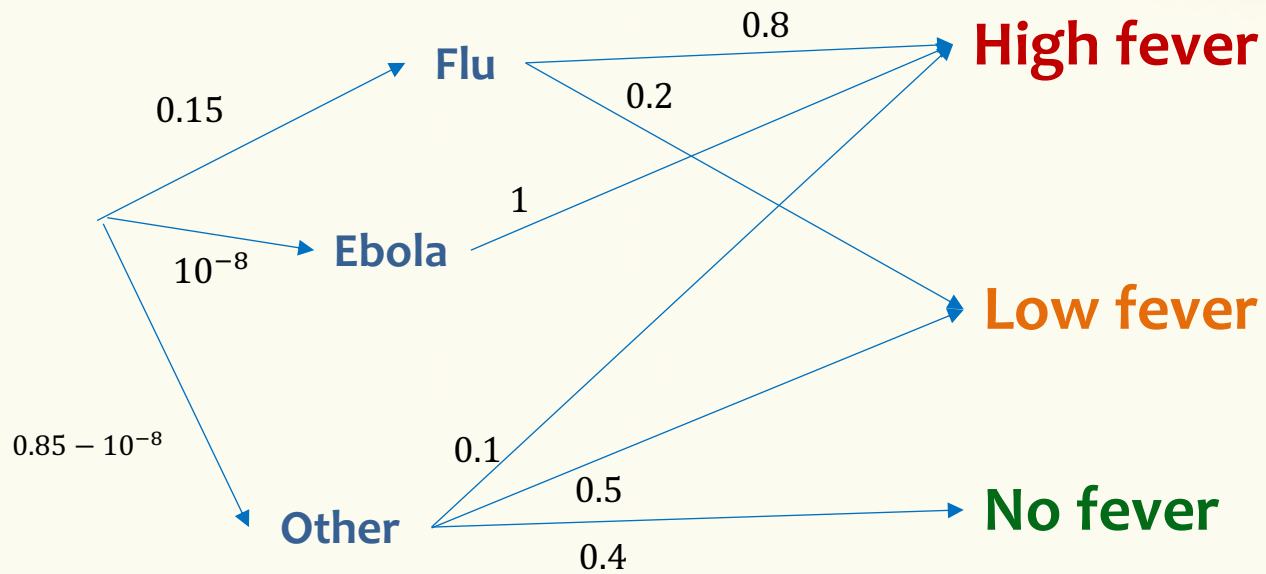
# Bayes Rule

**Theorem. (Bayes Rule)** For events  $\mathcal{A}$  and  $\mathcal{B}$ , where  $\mathbb{P}(\mathcal{A}), \mathbb{P}(\mathcal{B}) > 0$ ,

$$\mathbb{P}(\mathcal{B}|\mathcal{A}) = \frac{\mathbb{P}(\mathcal{B}) \cdot \mathbb{P}(\mathcal{A}|\mathcal{B})}{\mathbb{P}(\mathcal{A})}$$

Rev. Thomas Bayes [1701-1761]

Proof:  $\mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}|\mathcal{A}) = \mathbb{P}(\mathcal{A} \cap \mathcal{B})$



$$\begin{aligned}
 \mathbb{P}(\text{Ebola}|\text{High fever}) &= \frac{\mathbb{P}(\text{Ebola}) \cdot \mathbb{P}(\text{High fever}|\text{Ebola})}{\mathbb{P}(\text{High fever})} \\
 &= \frac{10^{-8} \cdot 1}{0.15 \times 0.8 + 10^{-8} \times 1 + (0.85 - 10^{-8}) \times 0.1} \approx 7.4 \times 10^{-8}
 \end{aligned}$$

$$\mathbb{P}(\text{Flu}|\text{High fever}) \approx 0.89$$

$$\mathbb{P}(\text{Other}|\text{High fever}) \approx 0.11$$

Most-likely a-posteriori  
outcome (MLA)

## Bayes Rule – Example

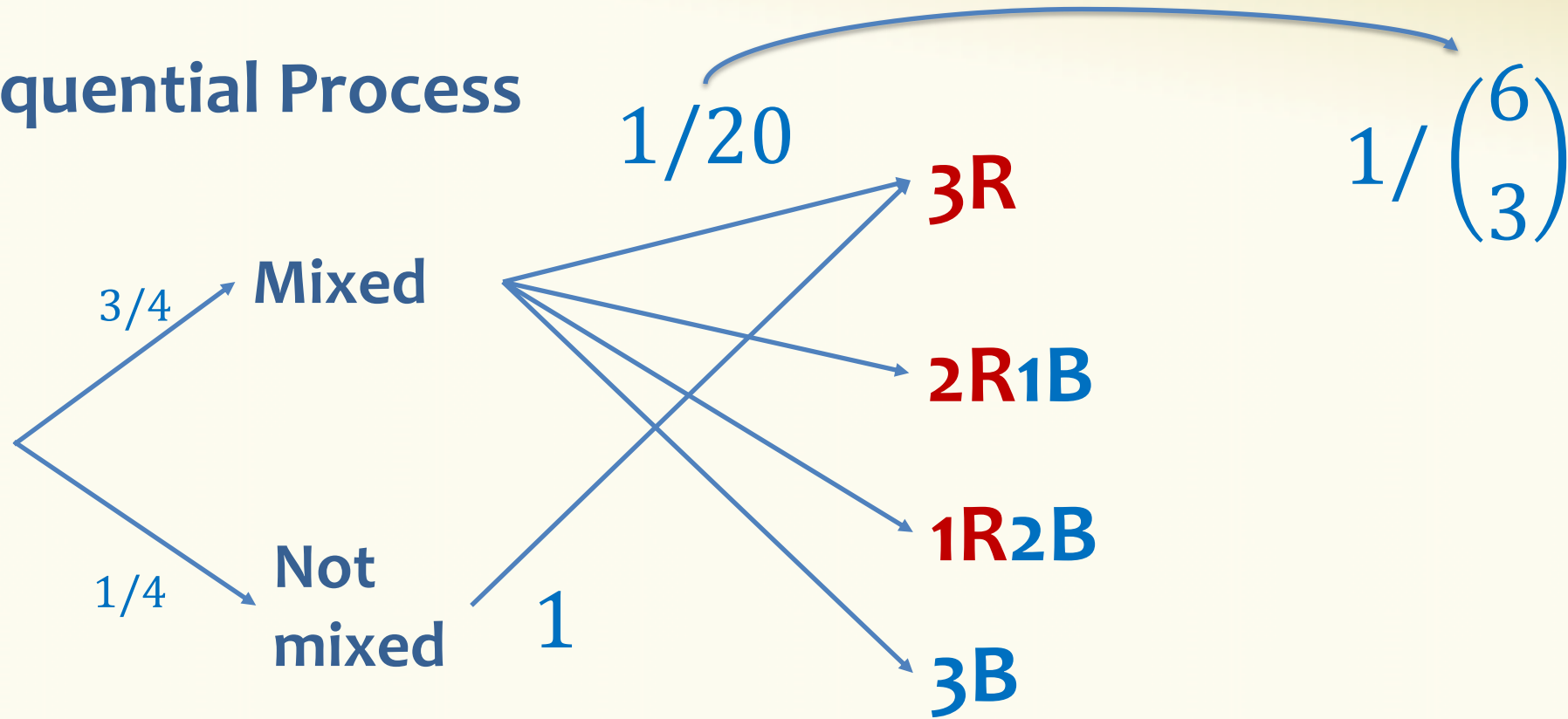
**Setting:** An urn contains 6 balls:

- 3 **red** and 3 **blue** balls w/ probability  $\frac{3}{4}$
- 6 **red** balls w/ probability  $\frac{1}{4}$

We draw three balls at random from the urn.

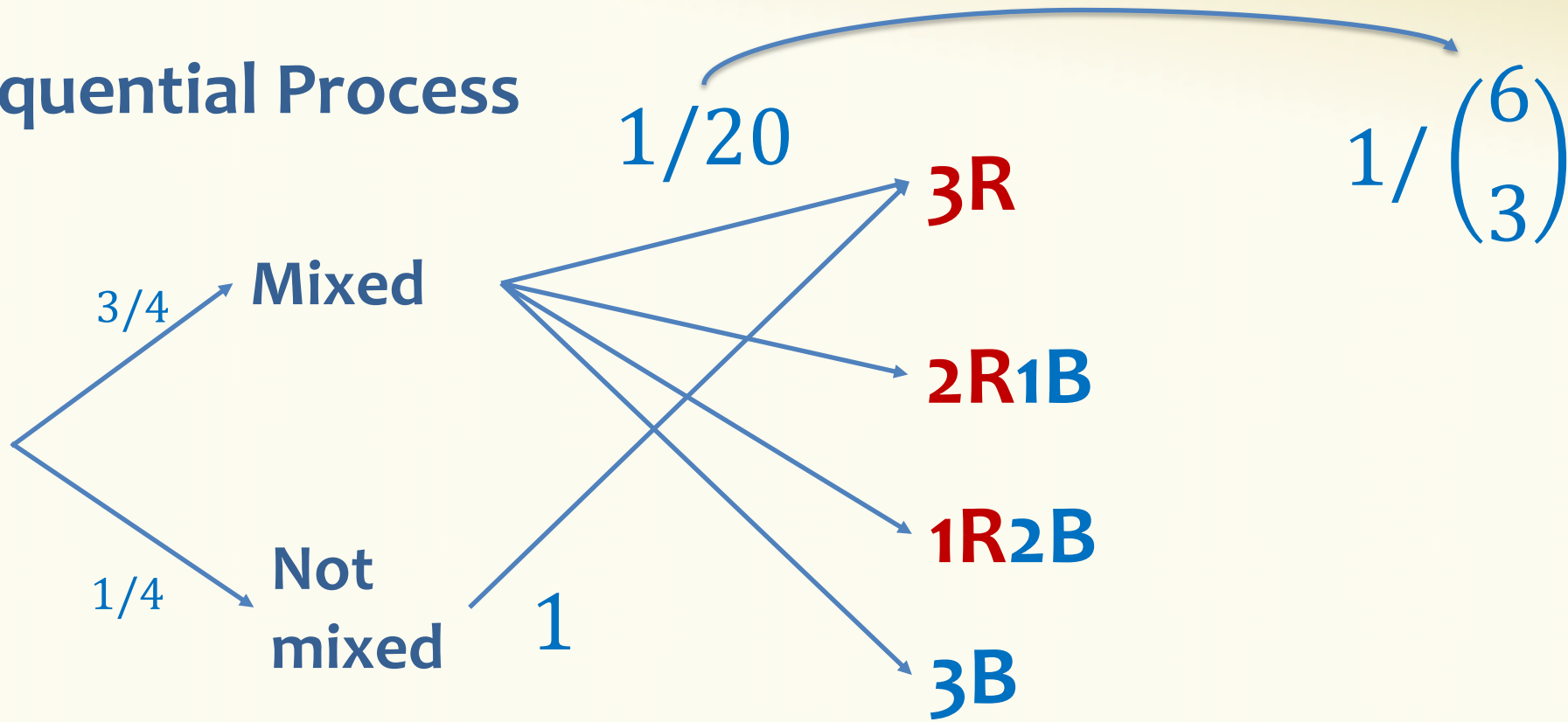
*All three balls are red. What is the probability that the remaining (undrawn) balls are all blue?*

## Sequential Process



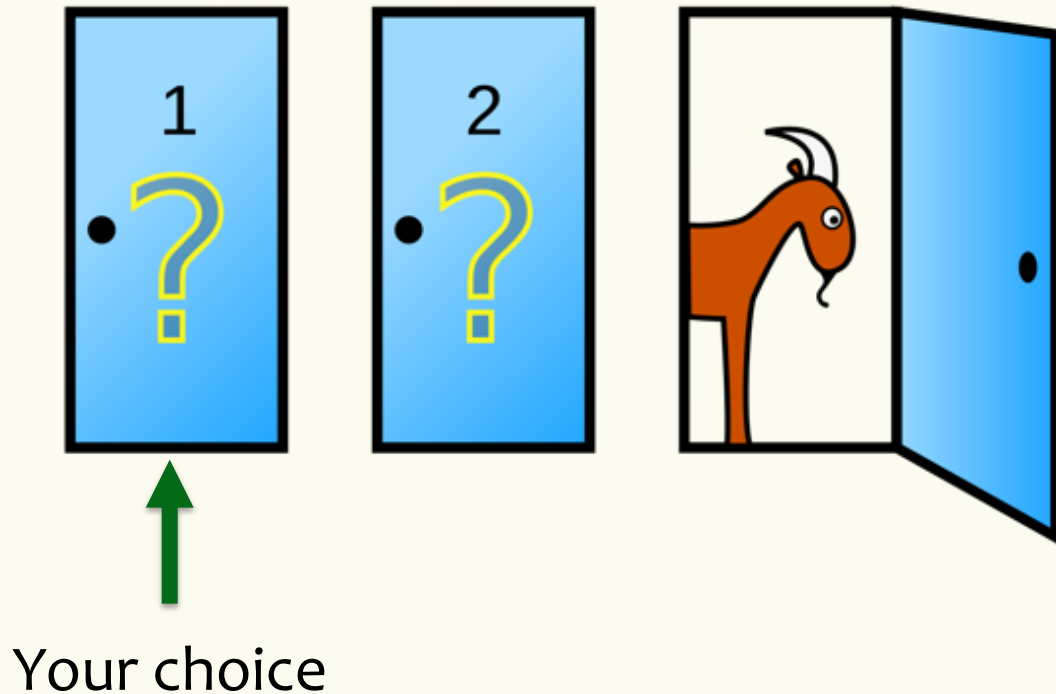
Wanted:  $\mathbb{P}(\text{Mixed} | \mathbf{3R})$

## Sequential Process



$$\mathbb{P}(\text{Mixed}|\mathbf{3R}) = \frac{\mathbb{P}(\text{Mixed})\mathbb{P}(\mathbf{3R}|\text{Mixed})}{\mathbb{P}(\mathbf{3R})} = \frac{\frac{3}{4} \times \frac{1}{20}}{\frac{3}{4} \times \frac{1}{20} + \frac{1}{4} \times 1} \approx 0.13$$

# The Monty Hall Problem

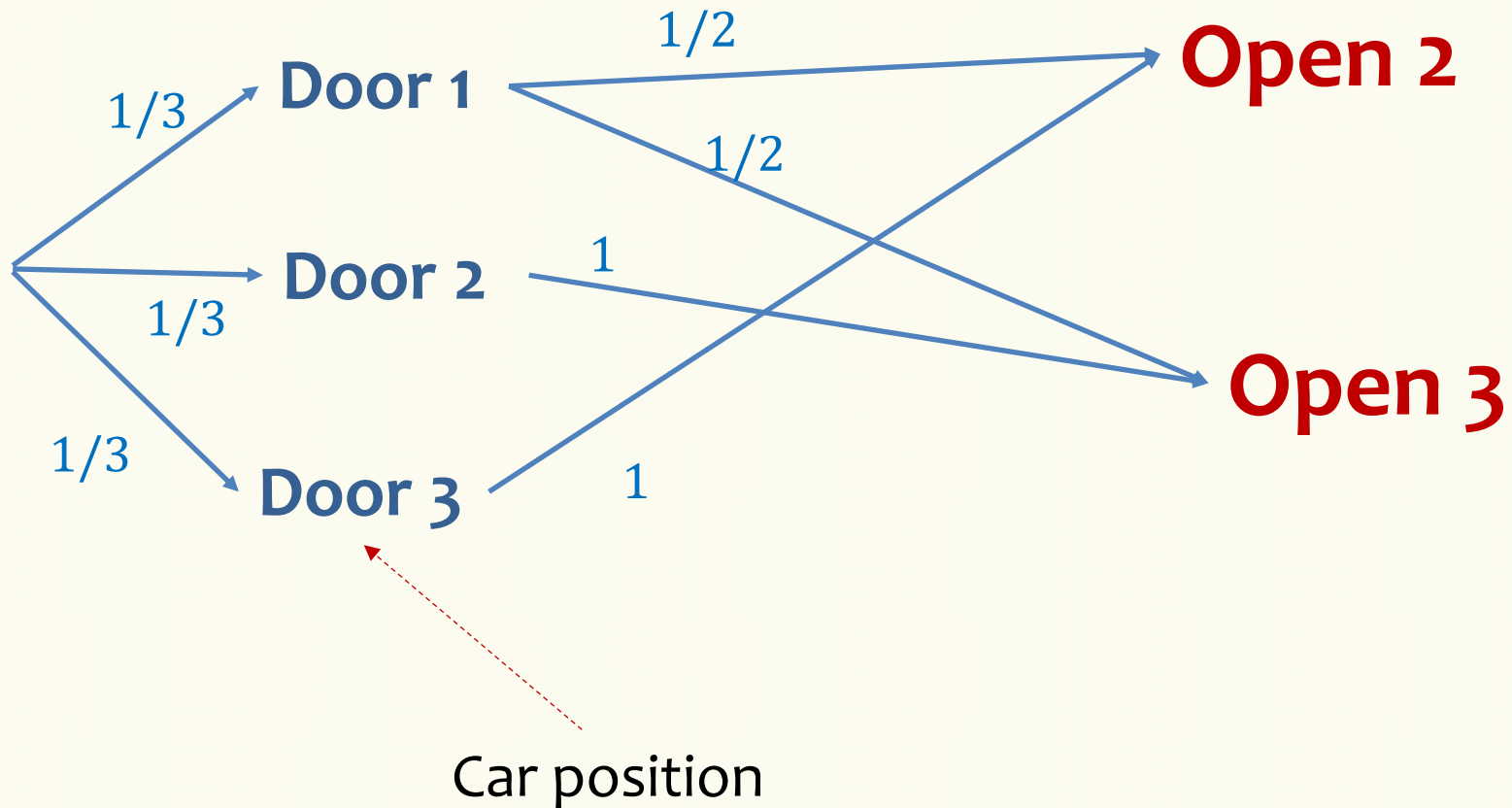


*Suppose you're on a game show, and you're given the choice of three doors: Behind one door is a car; behind the others, goats. You pick a door, say No. 1, and the host, who knows what's behind the doors, opens another door, say No. 3, which has a goat. He then says to you, "Do you want to pick door No. 2?" Is it to your advantage to switch your choice?*

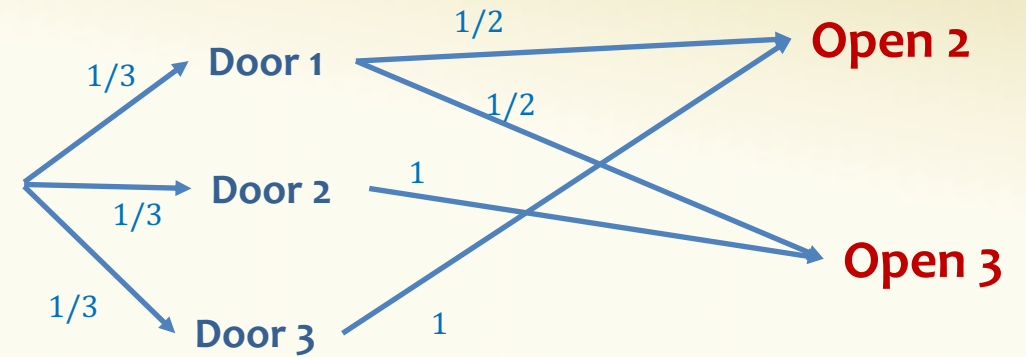
What would you do?

# Monty Hall

Say you picked (without loss of generality) **Door 1**



# Monty Hall

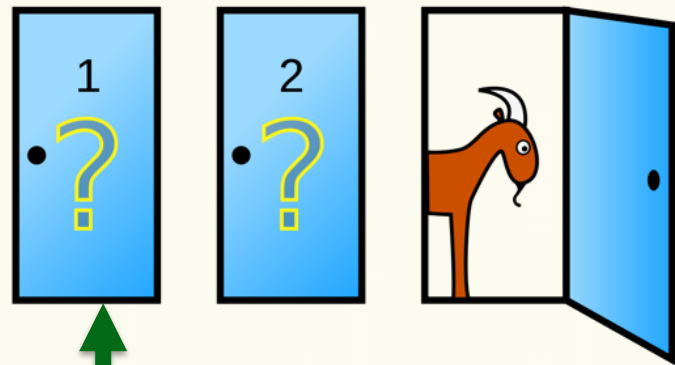


$$\begin{aligned}\mathbb{P}(\text{Door 1}|\text{Open 3}) &= \frac{\mathbb{P}(\text{Door 1})\mathbb{P}(\text{Open 3}|\text{Door 1})}{\mathbb{P}(\text{Open 3})} \\ &= \frac{\frac{1}{3} \times \frac{1}{2}}{\frac{1}{3} \times \frac{1}{2} + \frac{1}{3} \times 1} = \frac{\frac{1}{6}}{\frac{1}{3} + \frac{1}{3}} = \frac{1}{6} \times \frac{3}{2} = \frac{1}{4}\end{aligned}$$

$$\mathbb{P}(\text{Door 2}|\text{Open 3}) = 1 - \mathbb{P}(\text{Door 1}|\text{Open 3}) = \frac{3}{4}$$

# Monty Hall

Bottom line: Always swap!



Your  
choice

## 7.2 – More on Independence

# Independence – Recall

**Definition.** Two events  $\mathcal{A}$  and  $\mathcal{B}$  are (statistically) **independent** if

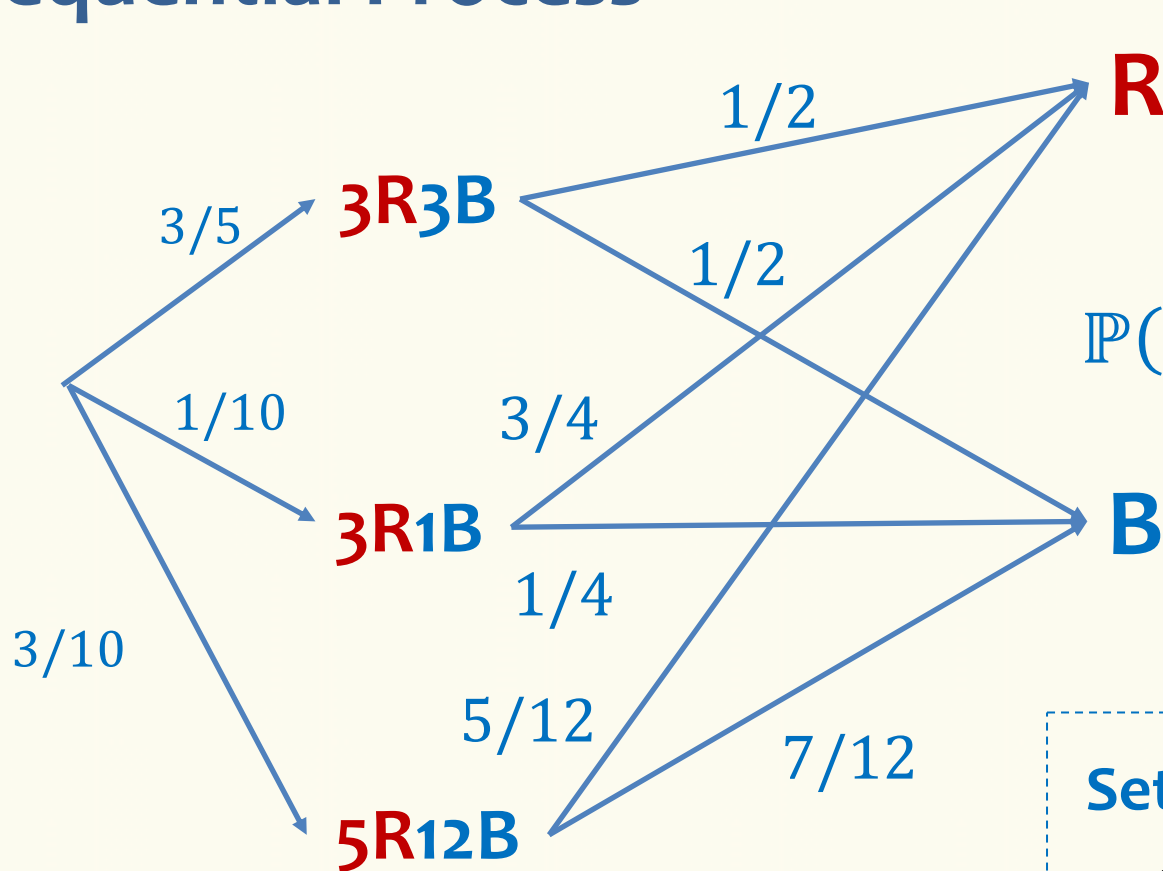
$$\mathbb{P}(\mathcal{A} \cap \mathcal{B}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}).$$

**“Equivalently.”**  $\mathbb{P}(\mathcal{A}|\mathcal{B}) = \mathbb{P}(\mathcal{A}).$

It is important to understand that independence is a property of probabilities of outcomes, not of the root cause generating these events.

This can be very counterintuitive!

# Sequential Process



$$\mathbb{P}(R | 3R3B) = \frac{1}{2}$$

$$\mathbb{P}(R) = \frac{3}{5} \times \frac{1}{2} + \frac{1}{10} \times \frac{3}{4} + \frac{3}{10} \times \frac{5}{12} = \frac{1}{2}$$

Independent!  $\mathbb{P}(R) = \mathbb{P}(R | 3R3B)$

Are **R** and **3R3B** independent?

**Setting:** An urn contains:

- 3 **red** and 3 **blue** balls w/ probability  $\frac{3}{4}$
- 3 **red** and 1 **blue** balls w/ probability  $\frac{1}{10}$
- 5 **red** and 12 **blue** balls w/ probability  $\frac{3}{10}$

We draw a ball at random from the urn.

# Independence – Multiple Events

**Definition.** Two events  $\mathcal{A}$  and  $\mathcal{B}$  are (statistically) **independent** if

$$\mathbb{P}(\mathcal{A} \cap \mathcal{B}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}).$$

**Equivalently.**  $\mathbb{P}(\mathcal{A}|\mathcal{B}) = \mathbb{P}(\mathcal{A})$ .

If we have more than two events, interesting phenomena can happen.

# Example – Two Coin Tosses

*“first coin is heads”*

$$\mathcal{A} = \{HH, HT\}$$

$$\mathbb{P}(\mathcal{A}) = \frac{1}{2}$$

*“second coin is heads”*

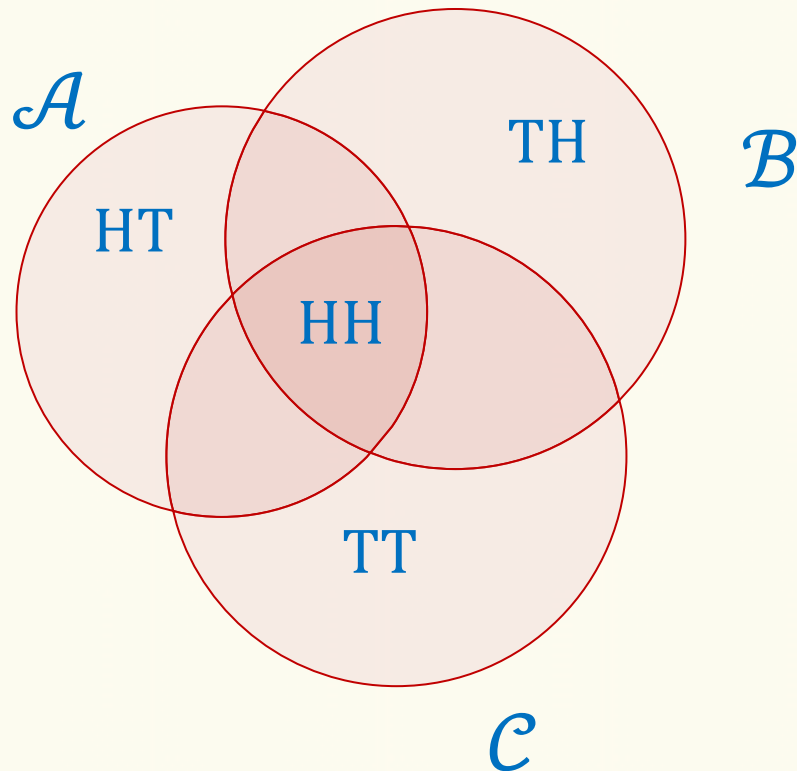
$$\mathcal{B} = \{HH, TH\}$$

$$\mathbb{P}(\mathcal{B}) = \frac{1}{2}$$

*“equal outcomes”*

$$\mathcal{C} = \{HH, TT\}$$

$$\mathbb{P}(\mathcal{C}) = \frac{1}{2}$$



$$\mathbb{P}(\mathcal{A} \cap \mathcal{B}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}) = \frac{1}{4}.$$

$$\mathbb{P}(\mathcal{A} \cap \mathcal{C}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{C}) = \frac{1}{4}.$$

$$\mathbb{P}(\mathcal{B} \cap \mathcal{C}) = \mathbb{P}(\mathcal{B}) \cdot \mathbb{P}(\mathcal{C}) = \frac{1}{4}.$$

Every **pair** of events is independent

# Pairwise Independence

**Definition.** The events  $\mathcal{A}_1, \dots, \mathcal{A}_n$  are **pairwise-independent** if for all distinct  $i, j \in [n]$ ,

$$\mathbb{P}(\mathcal{A}_i \cap \mathcal{A}_j) = \mathbb{P}(\mathcal{A}_i) \cdot \mathbb{P}(\mathcal{A}_j).$$

As we will see next week, pairwise independence is very powerful in computer science.

## Example – Two Coin Tosses

*“first coin is heads”*

$$\mathcal{A} = \{HH, HT\}$$

$$\mathbb{P}(\mathcal{A}) = \frac{1}{2}$$

*“second coin is heads”*

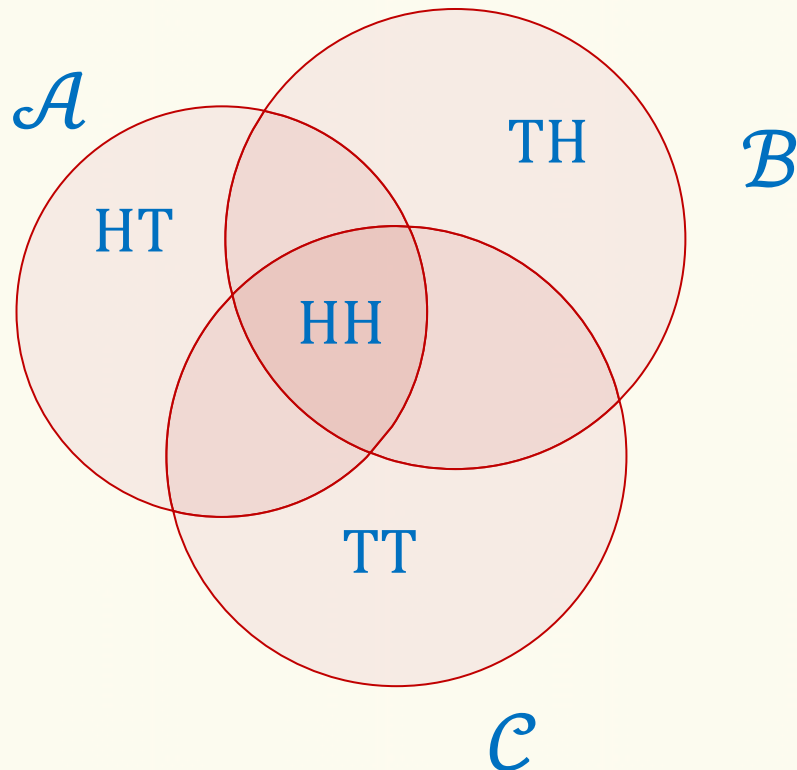
$$\mathcal{B} = \{HH, TH\}$$

$$\mathbb{P}(\mathcal{B}) = \frac{1}{2}$$

*“equal outcomes”*

$$\mathcal{C} = \{HH, TT\}$$

$$\mathbb{P}(\mathcal{C}) = \frac{1}{2}$$



$$\mathbb{P}(\mathcal{A} \cap \mathcal{B}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}) = \frac{1}{4}.$$

$$\mathbb{P}(\mathcal{A} \cap \mathcal{C}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{C}) = \frac{1}{4}.$$

$$\mathbb{P}(\mathcal{B} \cap \mathcal{C}) = \mathbb{P}(\mathcal{B}) \cdot \mathbb{P}(\mathcal{C}) = \frac{1}{4}.$$

$\mathcal{A}, \mathcal{B}, \mathcal{C}$  are  
pairwise  
independent

# Independence – Multiple Events

**Definition.** The events  $\mathcal{A}_1, \dots, \mathcal{A}_n$  are **independent** if for every  $k \leq n$  and  $1 \leq j_1 < j_2 < \dots < j_k \leq n$ ,

$$\mathbb{P}(\mathcal{A}_{j_1} \cap \mathcal{A}_{j_2} \cap \dots \cap \mathcal{A}_{j_k}) = \mathbb{P}(\mathcal{A}_{j_1}) \cdot \mathbb{P}(\mathcal{A}_{j_2}) \cdots \mathbb{P}(\mathcal{A}_{j_k}).$$

**Fact.** Pairwise independence does not imply independence!

Proof by counterexample\*! (see next slide)

**Fact.** Independence implies pairwise-independence.

Trivial by definition, use  $k = 2$

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\* Giving a counterexample is always sufficient to disprove an implication.

## Example – Two Coin Tosses

*“first coin is heads”*

$$\mathcal{A} = \{HH, HT\}$$

$$\mathbb{P}(\mathcal{A}) = \frac{1}{2}$$

*“second coin is heads”*

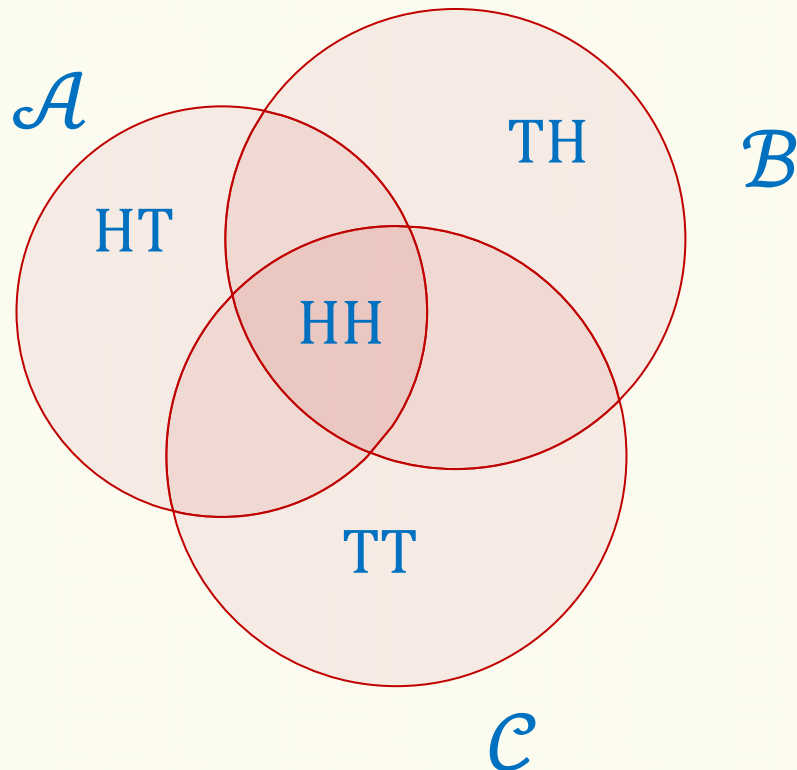
$$\mathcal{B} = \{HH, TH\}$$

$$\mathbb{P}(\mathcal{B}) = \frac{1}{2}$$

*“equal outcomes”*

$$\mathcal{C} = \{HH, TT\}$$

$$\mathbb{P}(\mathcal{C}) = \frac{1}{2}$$



$$\mathbb{P}(\mathcal{A} \cap \mathcal{B} \cap \mathcal{C}) = \mathbb{P}(\text{HH}) = \frac{1}{4}.$$

$$\frac{1}{4} \neq \frac{1}{2} \times \frac{1}{2} \times \frac{1}{2}.$$

$\mathcal{A}, \mathcal{B}, \mathcal{C}$  are not independent

## Example – Two Coin Tosses

<i>“first coin is heads”</i>	$\mathcal{A} = \{\text{HH}, \text{HT}\}$	$\mathbb{P}(\mathcal{A}) = \frac{1}{2}$
<i>“second coin is heads”</i>	$\mathcal{B} = \{\text{HH}, \text{TH}\}$	$\mathbb{P}(\mathcal{B}) = \frac{1}{2}$
<i>“equal outcomes”</i>	$\mathcal{C} = \{\text{HH}, \text{TT}\}$	$\mathbb{P}(\mathcal{C}) = \frac{1}{2}$

$\mathcal{A}, \mathcal{B}, \mathcal{C}$  are not independent

**Important:** The formal notion matches the intuition, namely

- If  $\mathcal{A}$  and  $\mathcal{B}$  have happened, we know both coins are heads.
- Therefore,  $\mathcal{C}$  must have happened, i.e.,  $\mathbb{P}(\mathcal{C}|\mathcal{A} \cap \mathcal{B}) = 1$

## Example – Three Coin Tosses

*“first coin is heads”*

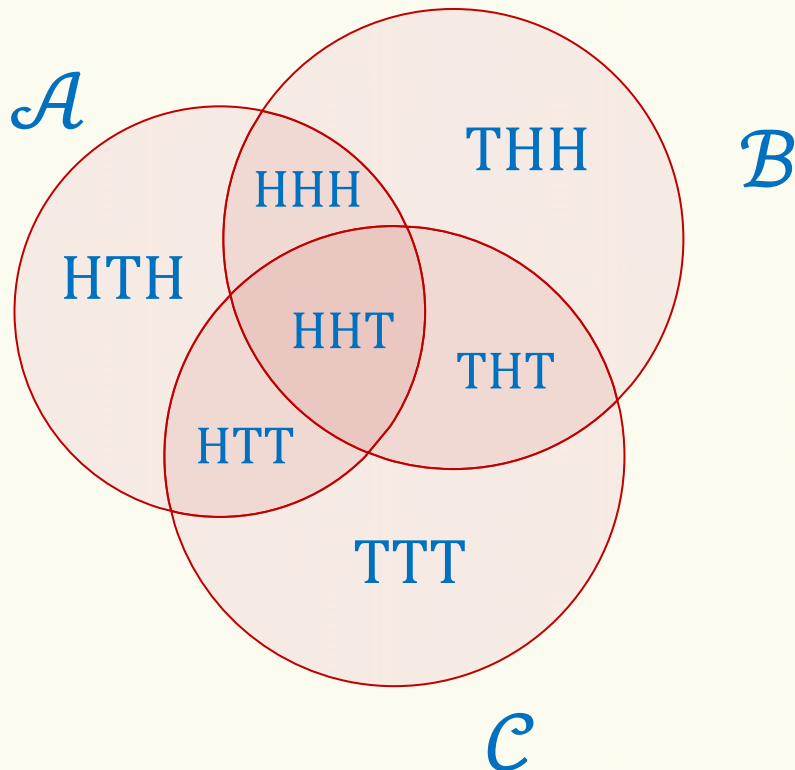
*“second coin is heads”*

*“third coin is tails”*

$$\mathcal{A} = \{HHH, HHT, HTH, HTT\} \quad \mathbb{P}(\mathcal{A}) = \frac{1}{2}$$

$$\mathcal{B} = \{HHH, HHT, THH, THT\} \quad \mathbb{P}(\mathcal{B}) = \frac{1}{2}$$

$$\mathcal{C} = \{HHT, HTT, THT, TTT\} \quad \mathbb{P}(\mathcal{C}) = \frac{1}{2}$$



$$\mathbb{P}(\mathcal{A} \cap \mathcal{B} \cap \mathcal{C}) = \mathbb{P}(\text{HHT}) = \frac{1}{8}.$$

$$= \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}) \cdot \mathbb{P}(\mathcal{C})$$

$$\mathbb{P}(\mathcal{A} \cap \mathcal{B}) = \frac{1}{4} = \frac{1}{2} \cdot \frac{1}{2} = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B})$$

$$\text{Similarly: } \mathbb{P}(\mathcal{A} \cap \mathcal{C}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{C})$$

$$\mathbb{P}(\mathcal{B} \cap \mathcal{C}) = \mathbb{P}(\mathcal{B}) \cdot \mathbb{P}(\mathcal{C})$$

$\rightarrow \mathcal{A}, \mathcal{B}, \mathcal{C}$  are independent

# Independence & Conditioning

Conditioning can break independence.

*“first coin is heads”*  $\mathcal{A} = \{\text{HH}, \text{HT}\}$   $\mathbb{P}(\mathcal{A}) = \frac{1}{2}$

*“second coin is tails”*  $\mathcal{B} = \{\text{HT}, \text{TT}\}$   $\mathbb{P}(\mathcal{B}) = \frac{1}{2}$

*“equal outcomes”*  $\mathcal{C} = \{\text{HH}, \text{TT}\}$   $\mathbb{P}(\mathcal{C}) = \frac{1}{2}$

$$\mathbb{P}(\mathcal{A} \cap \mathcal{B}) = \mathbb{P}(\mathcal{A}) \cdot \mathbb{P}(\mathcal{B}) = \frac{1}{4}. \quad \mathbb{P}(\mathcal{A} \cap \mathcal{B} | \mathcal{C}) = \mathbb{P}(\mathcal{A} | \mathcal{C}) \cdot \mathbb{P}(\mathcal{B} | \mathcal{C})?$$

$\mathbb{P}(\mathcal{A} \cap \mathcal{B} | \mathcal{C}) = 0$  b/c if both outcomes are equal, we cannot have  $\mathcal{A} \cap \mathcal{B}$

$$\mathbb{P}(\mathcal{A} | \mathcal{C}) = \frac{\mathbb{P}(\mathcal{A} \cap \mathcal{C})}{\mathbb{P}(\mathcal{C})} = \frac{\mathbb{P}(\text{HH})}{\mathbb{P}(\mathcal{C})} = \frac{1}{4} \times \frac{2}{1} = \frac{1}{2} \quad \mathbb{P}(\mathcal{B} | \mathcal{C}) = \frac{1}{2}$$