

Ex: Independent samples $x_1, x_2, \dots, x_n \sim N(\mu, \sigma^2)$, where $\theta_1 = \mu$ and $\theta_2 = \sigma^2$ are both unknown.

$$L(x_1, x_2, \dots, x_n | \theta_1, \theta_2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\theta_2}} e^{-\frac{(x_i - \theta_1)^2}{2\theta_2}}$$

$$\ln L(x_1, x_2, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n \left(-\frac{1}{2} \ln(2\pi\theta_2) - \frac{(x_i - \theta_1)^2}{2\theta_2} \right)$$

$$\frac{\partial}{\partial \theta_1} \ln L(x_1, x_2, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n + \frac{2(x_i - \theta_1)}{2\theta_2} = \boxed{\sum_{i=1}^n \frac{x_i - \theta_1}{\theta_2}} = 0$$

$$\sum_{i=1}^n (x_i - \hat{\theta}_1) = 0$$

$$\sum_{i=1}^n x_i - n\hat{\theta}_1 = 0$$

$$\hat{\theta}_1 = \frac{1}{n} \sum_{i=1}^n x_i, \text{ the sample mean}$$

Is it a max?

$$\frac{\partial^2}{\partial \theta_2} \ln L(x_1, x_2, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n -\frac{1}{\theta_2} = \frac{-n}{\theta_2} < 0$$

so $\ln L$ is concave downward everywhere.

$$\frac{\partial}{\partial \theta_2} \ln L(x_1, x_2, \dots, x_n | \theta_1, \theta_2) = \sum_{i=1}^n \left(-\frac{2\pi}{2 \cdot 2\pi\theta_2} + \frac{(x_i - \theta_1)^2}{2\theta_2^2} \right)$$

$$= \sum_{i=1}^n \left(-\frac{1}{2\theta_2} + \frac{(x_i - \theta_1)^2}{2\theta_2^2} \right) = 0$$

$$\sum_{i=1}^n \left(-\hat{\theta}_2 + \frac{(x_i - \hat{\theta}_1)^2}{\hat{\theta}_2} \right) = 0$$

$$-n\hat{\theta}_2 + \sum_{i=1}^n (x_i - \hat{\theta}_1)^2 = 0$$

$$\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\theta}_1)^2$$

2nd derivative test shows that it is negative at $\theta_2 = \hat{\theta}_2$.

$\hat{\theta}_2$ is a "sample variance".

$\text{Var}(X) = E[(X - \mu)^2]$ where $\mu = E[X]$.

MLE of population variance θ^2 is a sample variance of x_1, x_2, \dots, x_n .

Bias

Defn: An estimator $\hat{\theta}$ of θ is unbiased iff $E[\hat{\theta}] = \theta$.

Ex: For $\hat{\theta}_1$ of $N(\theta_1, \theta_2)$, is $\hat{\theta}_1$ unbiased?

$$\begin{aligned} E[\hat{\theta}_1] &= E\left[\frac{1}{n} \sum_{i=1}^n x_i\right] = \frac{1}{n} \sum_{i=1}^n E[x_i] \\ &= \frac{1}{n} \sum_{i=1}^n \mu = \frac{1}{n} \cdot n\mu = \mu = \theta_1 \end{aligned}$$

So $\hat{\theta}_1$ is an unbiased estimator.

Is $\hat{\theta}_2$ an unbiased estimator of θ_2 ? No

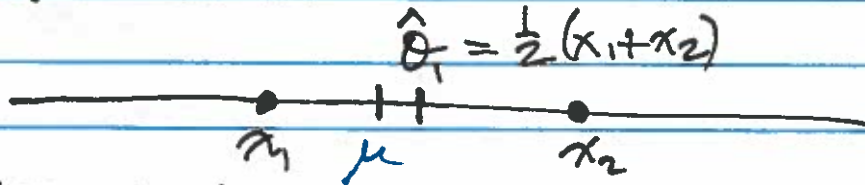
$$\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\theta}_1)^2 \text{ ~~has~~ maximizes likelihood}$$

$$\hat{\theta}_2' = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\theta}_1)^2 \text{ is an unbiased}$$

estimator of θ_2 .
 $\hat{\theta}_2$ and $\hat{\theta}_2'$ differ by a factor of $\frac{n-1}{n}$.
 As n gets large, they get closer.

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why might $\hat{\sigma}_2$ be biased?
 $n=2$



$$\hat{\sigma}_2 = \frac{1}{2}((x_1 - \hat{\sigma}_1)^2 + (x_2 - \hat{\sigma}_1)^2)$$

$$< \frac{1}{2}((x_1 - \mu)^2 + (x_2 - \mu)^2)$$

so $\hat{\sigma}_2$ underestimates σ^2 .