CSE 312: Foundations of Computing II Quiz Section #8: Normal Distribution, Central Limit Theorem (solutions)

Review: Main Theorems and Concepts

Standardizing: Let *X* be any random variable (discrete or continuous, not necessarily normal), with $\mathbb{E}[X] = \mu$ and Var (*X*) = σ^2 . If we let $Y = \frac{X - \mu}{\sigma}$, then $\mathbb{E}[Y] = 0$ and Var (*Y*) = 1.

Normal (Gaussian, "bell curve"): $X \sim N(\mu, \sigma^2)$ iff X has the following probability density function:

$$f_X(x) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}}, \ x \in \mathbb{R}$$

 $\mathbb{E}[X] = \mu$ and $\operatorname{Var}(X) = \sigma^2$. The "standard normal" random variable is typically denoted *Z* and has mean 0 and variance 1: if $X \sim N(\mu, \sigma^2)$, then $Z = \frac{X-\mu}{\sigma} \sim N(0, 1)$. The CDF has no closed form, but we denote the CDF of the standard normal as $\Phi(z) = F_Z(z) = \mathbb{P}(Z \le z)$. Note from symmetry of the probability density function about z = 0 that: $\Phi(-z) = 1 - \Phi(z)$.

Closure of the Normal Distribution: Let $X \sim \mathcal{N}(\mu, \sigma^2)$. Then, $aX + b \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$. That is, linear transformations of normal random variables are still normal.

"Reproductive" Property of Normals: Let X_1, \ldots, X_n be independent normal random variables with $\mathbb{E}[X_i] = \mu_i$ and $\operatorname{Var}(X_i) = \sigma_i^2$. Let $a_1, \ldots, a_n \in \mathbb{R}$ and $b \in \mathbb{R}$. Then,

$$X = \sum_{i=1}^{n} (a_i X_i + b) \sim \mathcal{N}\left(\sum_{i=1}^{n} (a_i \mu_i + b), \sum_{i=1}^{n} a_i^2 \sigma_i^2\right)$$

There's nothing special about the parameters – the important result here is that the resulting random variable is still normally distributed.

Central Limit Theorem (CLT): Let X_1, \ldots, X_n be iid random variables with $\mathbb{E}[X_i] = \mu$ and $\operatorname{Var}(X_i) = \sigma^2$. Let $X = \sum_{i=1}^n X_i$, which has $\mathbb{E}[X] = n\mu$ and $\operatorname{Var}(X) = n\sigma^2$. Let $\overline{X} = \frac{1}{n} \sum_{i=1}^n X_i$, which has $\mathbb{E}[\overline{X}] = \mu$ and $\operatorname{Var}(\overline{X}) = \frac{\sigma^2}{n}$. \overline{X} is called the *sample mean*. Then, as $n \to \infty$, \overline{X} approaches the normal distribution $\mathcal{N}(\mu, \frac{\sigma^2}{n})$. Standardizing, this is equivalent to $Y = \frac{\overline{X} - \mu}{\sigma/\sqrt{n}}$ approaching $\mathcal{N}(0, 1)$. Similarly, as $n \to \infty$, X approaches $\mathcal{N}(n\mu, n\sigma^2)$ and $Y' = \frac{X - n\mu}{\sigma\sqrt{n}}$ approaches $\mathcal{N}(0, 1)$.

It is no surprise that \overline{X} has mean μ and variance σ^2/n – this can be done with simple calculations. The importance of the CLT is that, for large *n*, regardless of what distribution X_i comes from, \overline{X} is *approximately normally distributed with mean* μ *and variance* σ^2/n . Don't forget the continuity correction, only when X_1, \ldots, X_n are discrete random variables.

Detailed explanation of CLT

One form of the Central Limit Theorem states that if random variables $X_1, X_2, ..., X_n$ are **independent and identically distributed**, as $n \to \infty$, the **sample mean** $(\frac{1}{n} \cdot (X_1 + X_2 + \cdots + X_n))$ approaches the normal distribution $N(\mu, \sigma^2/n)$. Note that μ, σ^2 are the mean/variance of each individual X_i .

Why is the variance σ^2/n ? By the properties of variance and independence,

$$Var(\frac{1}{n}(X_1 + X_2 + \dots + X_n)) = \frac{1}{n^2} Var(X_1 + X_2 + \dots + X_n)$$

= $\frac{1}{n^2} (Var(X_1) + Var(X_2) + \dots + Var(X_n))$
= $\frac{1}{n^2} (n \cdot \sigma^2) = \frac{\sigma^2}{n}$

However, in many applications, you will be asked to approximate the **sum** of the random variables $X_1 + X_2 + \cdots + X_n$. By the CLT, if these variables are independent and identically distributed, this sum is approximately normal for large *n*. Since we're using the normal distribution to approximate the sum, we want the mean and variance of the Normal to be the same as the sum's, so the sum is approximated by $N(n\mu, n\sigma^2)$.

If you standardize this (by subtracting the mean and then dividing by the standard deviation of the overall sum), you'll get the formula that was shown in class:

$$\frac{X_1 + X_2 + \dots + X_n - n\mu}{\sqrt{n\sigma^2}} \to N(0,1)$$

There's one important catch. The normal distribution can take on any real value. However, if you're using the normal distribution to approximate anything where the random variable can only be an integer (such as the binomial), you'll have to apply the **continuity correction**. For example, let's say we want to find P(X > 5), and X is discrete. Then, what we really want is for X to be 6, 7, 8, etc. So when using the normal distribution to approximate this, we want to find the probability that the value returned by the normal distribution **rounds to** 6, 7, 8, ... so we should calculate P(X > 5.5).

To summarize, here are the steps you should take for Central Limit Theorem problems.

- 1. Check that all random variables are independent and identically distributed.
- 2. Find the mean and variance of the normal distribution that should be used. Set the mean/variance of this normal distribution to be the same mean and variance as the sum or sample mean.
- 3. Apply the continuity correction, if we're approximating a discrete distribution.
- 4. Standardize the random variable: subtract the mean from both sides and divide both sides by the standard deviation of the normal distribution.
- 5. Look up the value in the Φ table.

Exercises

The Φ table is on the last page for use in these exercises.

1. (From last week) Prove the memorylessness property for the exponential distribution $\text{Exp}(\lambda)$: If *s* and *t* are nonnegative, then $\mathbb{P}(X > s + t \mid X > s) = \mathbb{P}(X > t)$.

$$\mathbb{P}(X > s + t \mid X > s) = \frac{\mathbb{P}(X > s + t \cap X > s)}{\mathbb{P}(X > s)}$$
$$= \frac{\mathbb{P}(X > s + t)}{\mathbb{P}(X > s)}$$
$$= \frac{e^{-\lambda(s+t)}}{e^{-\lambda s}}$$
$$= e^{-\lambda t}$$
$$= \mathbb{P}(X > t)$$

2. Let $X \sim \mathcal{N}(50, 5)$. What is the probability that X is greater than 45 and less than 52?

$$\mathbb{P}(45 < X < 52) = \mathbb{P}\left(\frac{45 - 50}{\sqrt{5}} < \frac{X - 50}{\sqrt{5}} < \frac{52 - 50}{\sqrt{5}}\right) \approx \mathbb{P}\left(-2.24 < \frac{X - 50}{\sqrt{5}} < 0.89\right)$$
$$= \Phi(0.89) - \Phi(-2.24) = \Phi(0.89) - (1 - \Phi(2.24)) = \Phi(0.89) + \Phi(2.24) - 1$$
$$\approx 0.8133 + 0.9875 - 1 = 0.8008$$

3. A factory produces X_i gadgets on day *i*, where the X_i are independent and identically distributed random variables, each with mean 5 and variance 9.

Let $Y_n = X_1 + X_2 + \cdots + X_n$. By the Central Limit Theorem, if *n* is large enough then Y_n is distributed approximately as $\mathcal{N}(5n, 9n)$. (Prove that these are the correct parameters for the normal distribution.)

(a) Approximate the probability that the total number of gadgets produced in 100 days is less than 440.

$$\mathbb{P}(Y_{100} < 440) = \mathbb{P}(Y_{100} < 439.5) = \mathbb{P}\left(\frac{Y_{100} - 500}{30} < \frac{439.5 - 500}{30}\right) \approx \Phi\left(\frac{439.5 - 500}{30}\right)$$
$$= \Phi(-2.02) = 1 - \Phi(2.02) \approx 1 - 0.9783 = 0.0217$$

(b) Approximate the greatest value of *n* such that $P(X_1 + X_2 + \dots + X_n \ge 5n + 200) \le 0.05$.

$$\mathbb{P}(Y_n \ge 5n + 200) = \mathbb{P}(Y_n \ge 5n + 199.5) = \mathbb{P}\left(\frac{Y_n - 5n}{3\sqrt{n}} \ge \frac{199.5}{3\sqrt{n}}\right) = 1 - \Phi\left(\frac{199.5}{3\sqrt{n}}\right) \le 0.05$$

Therefore, $\Phi\left(\frac{199.5}{3\sqrt{n}}\right) \ge 0.95$. What value of z has $\Phi(z) = 0.95$ in the standard normal table? It is between z = 1.64 and z = 1.65. Choosing z = 1.645, then, we need to solve the equation $\frac{199.5}{3\sqrt{n}} = 1.645$. The solution is $n \approx 1634$.

4. (a) A fair coin is tossed 50 times. Use the Central Limit Theorem to estimate the probability that fewer than 20 of those tosses come up heads.

Let
$$X \sim Bin(50, 0.5)$$
. Then $\mathbb{E}[X] = 50 \times 0.5 = 25$ and $Var(X) = 50 \times 0.5 \times 0.5 = 12.5$.
 $\mathbb{P}(X < 20) = \mathbb{P}(X < 19.5) = \mathbb{P}\left(\frac{X - 25}{\sqrt{12.5}} < \frac{19.5 - 25}{\sqrt{12.5}}\right) \approx \mathbb{P}\left(\frac{X - 25}{\sqrt{12.5}} < -1.56\right) \approx \Phi(-1.56)$
 $= 1 - \Phi(1.56) \approx 0.0594$

(b) A fair coin is tossed until it comes up heads for the 20th time. Use the Central Limit Theorem to estimate the probability that more than 50 tosses are needed. (Hint: you will need the mean and variance of a geometric random variable, which you can find in Example 2.15 of the text.)

Let $X_1, X_2, \dots, X_{20} \sim \text{Geo}(0.5)$ be independent and let $X = \sum_{i=1}^{20} X_i$. Then $\mathbb{E}[X] = \sum_{i=1}^{20} \mathbb{E}[X_i] = 40$ and $\text{Var}(X) = \sum_{i=1}^{20} \text{Var}(X_i) = \sum_{i=1}^{20} \frac{1-1/2}{(1/2)^2} = 40$.

$$\mathbb{P}(X > 50) = \mathbb{P}(X > 50.5) = \mathbb{P}\left(\frac{X - 40}{\sqrt{40}} > \frac{50.5 - 40}{\sqrt{40}}\right) \approx \mathbb{P}\left(\frac{X - 40}{\sqrt{40}} > 1.66\right) \approx 1 - \Phi(1.66) \approx 0.0485$$

(c) Compare your answers from parts (a) and (b). Why are they close but not exactly equal?

Suppose the coin is tossed until it has been tossed at least 50 times and heads has come up at least 20 times. If the first 50 tosses have fewer than 20 heads, then the 20th head requires more than 50 tosses. Conversely, if the 20th head requires more than 50 tosses, then the first 50 tosses have fewer than 20 heads. Thus, the exact probabilities in parts (a) and (b) must be identical. The reason we did not get the exact same answers in parts (a) and (b) is that the Central Limit Theorem says that the Normal distribution is only an approximation (except in the limit).

	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
0.0	0.5000	0.5040	0.5080	0.5120	0.5160	0.5199	0.5239	0.5279	0.5319	0.5359
0.1	0.5398	0.5438	0.5478	0.5517	0.5557	0.5596	0.5636	0.5675	0.5714	0.5753
0.2	0.5793	0.5832	0.5871	0.5910	0.5948	0.5987	0.6026	0.6064	0.6103	0.6141
0.3	0.6179	0.6217	0.6255	0.6293	0.6331	0.6368	0.6406	0.6443	0.6480	0.6517
0.4	0.6554	0.6591	0.6628	0.6664	0.6700	0.6736	0.6772	0.6808	0.6844	0.6879
0.5	0.6915	0.6950	0.6985	0.7019	0.7054	0.7088	0.7123	0.7157	0.7190	0.7224
0.6	0.7257	0.7291	0.7324	0.7357	0.7389	0.7422	0.7454	0.7486	0.7517	0.7549
0.7	0.7580	0.7611	0.7642	0.7673	0.7704	0.7734	0.7764	0.7794	0.7823	0.7852
0.8	0.7881	0.7910	0.7939	0.7967	0.7995	0.8023	0.8051	0.8078	0.8106	0.8133
0.9	0.8159	0.8186	0.8212	0.8238	0.8264	0.8289	0.8315	0.8340	0.8365	0.8389
1.0	0.8413	0.8438	0.8461	0.8485	0.8508	0.8531	0.8554	0.8577	0.8599	0.8621
1.1	0.8643	0.8665	0.8686	0.8708	0.8729	0.8749	0.8770	0.8790	0.8810	0.8830
1.2	0.8849	0.8869	0.8888	0.8907	0.8925	0.8944	0.8962	0.8980	0.8997	0.9015
1.3	0.9032	0.9049	0.9066	0.9082	0.9099	0.9115	0.9131	0.9147	0.9162	0.9177
1.4	0.9192	0.9207	0.9222	0.9236	0.9251	0.9265	0.9279	0.9292	0.9306	0.9319
1.5	0.9332	0.9345	0.9357	0.9370	0.9382	0.9394	0.9406	0.9418	0.9429	0.9441
1.6	0.9452	0.9463	0.9474	0.9484	0.9495	0.9505	0.9515	0.9525	0.9535	0.9545
1.7	0.9554	0.9564	0.9573	0.9582	0.9591	0.9599	0.9608	0.9616	0.9625	0.9633
1.8	0.9641	0.9649	0.9656	0.9664	0.9671	0.9678	0.9686	0.9693	0.9699	0.9706
1.9	0.9713	0.9719	0.9726	0.9732	0.9738	0.9744	0.9750	0.9756	0.9761	0.9767
2.0	0.9772	0.9778	0.9783	0.9788	0.9793	0.9798	0.9803	0.9808	0.9812	0.9817
2.1	0.9821	0.9826	0.9830	0.9834	0.9838	0.9842	0.9846	0.9850	0.9854	0.9857
2.2	0.9861	0.9864	0.9868	0.9871	0.9875	0.9878	0.9881	0.9884	0.9887	0.9890
2.3	0.9893	0.9896	0.9898	0.9901	0.9904	0.9906	0.9909	0.9911	0.9913	0.9916
2.4	0.9918	0.9920	0.9922	0.9925	0.9927	0.9929	0.9931	0.9932	0.9934	0.9936
2.5	0.9938	0.9940	0.9941	0.9943	0.9945	0.9946	0.9948	0.9949	0.9951	0.9952
2.6	0.9953	0.9955	0.9956	0.9957	0.9959	0.9960	0.9961	0.9962	0.9963	0.9964
2.7	0.9965	0.9966	0.9967	0.9968	0.9969	0.9970	0.9971	0.9972	0.9973	0.9974
2.8	0.9974	0.9975	0.9976	0.9977	0.9977	0.9978	0.9979	0.9979	0.9980	0.9981
2.9	0.9981	0.9982	0.9982	0.9983	0.9984	0.9984	0.9985	0.9985	0.9986	0.9986
3.0	0.9987	0.9987	0.9987	0.9988	0.9988	0.9989	0.9989	0.9989	0.9990	0.9990

Table 1: Cumulative distribution function of the standard normal N(0, 1)