## Lecture 13: Conditional Expectation and Variance

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We discuss the conditional expectation and variance.

Given an event $E$, the conditional expectation is exactly what you would expect after conditioning:

$$
\mathbb{E}[X \mid E]=\sum_{x \in E} p(X=x \mid E) \cdot x .
$$

Fact 1. If $E$ is an event and $X$ is a random variable, then

$$
\mathbb{E}[X]=p(E) \cdot \mathbb{E}[X \mid E]+p\left(E^{c}\right) \cdot \mathbb{E}\left[X \mid E^{c}\right] .
$$

Proof.

$$
\begin{aligned}
\mathbb{E}[X] & =\sum_{x} p(X=x) \cdot x \\
& =\sum_{x}\left(p(E) \cdot p(X=x \mid E)+p\left(E^{c}\right) \cdot p\left(X=x \mid E^{c}\right)\right) \cdot x \\
& =p(E) \sum_{x} p(X=x \mid E) \cdot x+p\left(E^{c}\right) \sum_{x} p\left(X=x \mid E^{c}\right) \cdot x \\
& =p(E) \cdot \mathbb{E}[X \mid E]+p\left(E^{c}\right) \cdot \mathbb{E}\left[X \mid E^{c}\right] .
\end{aligned}
$$

## Example

You toss a coin until you see heads. How many coin tosses do you expect to see?

Let $X$ be the number of coin tosses. Then we see that

$$
\mathbb{E}[X]=\sum_{i=1}^{\infty} p(X=i) \cdot i=\sum_{i=1}^{\infty}(1 / 2)^{i} \cdot i .
$$

We could calculate this using some identities from calculus.
We have

$$
(1-x)\left(1+x+x^{2}+\ldots+x^{r}\right)=\left(1-x^{r+1}\right),
$$

so we can express

$$
1+x+x^{2}+\ldots+x^{r}=\frac{1-x^{r+1}}{1-x}
$$

If $|x|<1$, then as $r$ goes to infinity, we get

$$
1+x+x^{2}+\ldots=\frac{1}{1-x}
$$

Now, the expression we care about is
$\mathbb{E}[X]=\sum_{i=1}^{\infty}(1 / 2)^{i} \cdot i=(1 / 2) \sum_{i=1}^{\infty}(1 / 2)^{i-1} \cdot i=x\left(1+2 x+3 x^{2}+4 x^{3}+\ldots\right)$,
where here $x=1 / 2$. The infinite sum in this expression is exactly the derivative of

$$
1+x+x^{2}+\ldots
$$

So, taking the derivative of $(1 /(1-x))$, we get that it should be $1 /(1-x)^{2}$. Plugging this calculation in gives:

$$
\mathbb{E}[X]=(1 / 2) \cdot \frac{1}{(1-(1 / 2))^{2}}=2 .
$$

But there is an easier way to calculate all of this using conditional expectation. Let $E$ denote the event that the first coin toss is heads. Then we have

$$
\mathbb{E}[X]=p(E) \mathbb{E}[X \mid E]+p\left(E^{c}\right) \mathbb{E}\left[X \mid E^{c}\right] .
$$

Now $\mathbb{E}[X \mid E]=1$, since under this event there are no more coin tosses. Conditioned on $E^{c}$, the expected number of heads is exactly 1 more than the expected number of heads for $X$. So, we get:

$$
\mathbb{E}[X]=(1 / 2)+(1 / 2)(\mathbb{E}[X]+1),
$$

which gives $\mathbb{E}[X]=2$.

## Variance

As we saw in the earlier examples, a random variable can be very far from its expectation. One way to measure how far a random variable is typically from its expectation is to measure its variance.

Suppose $X$ is a random variable with $\mathbb{E}[X]=0$. Then the variance of $X$ is

$$
\operatorname{Var}[X]=\mathbb{E}\left[X^{2}\right] .
$$

When it is large then $|X|$ is typically large and the value of $X$ is usually far from its expectation. For general $X$, we define its variance as

$$
\operatorname{Var}[X]=\mathbb{E}\left[(X-\mu)^{2}\right],
$$

where here $\mu=\mathbb{E}[X]$.
The variance of a random variable is always non-negative.

## Example

For example, let $X, Y, Z$ be random variables such that

$$
\begin{gathered}
X= \begin{cases}1000 & \text { with probability } 1 / 2 \\
-1000 & \text { with probability } 1 / 2\end{cases} \\
Y= \begin{cases}1 & \text { with probability } \frac{n-1}{n} \\
-(n-1) & \text { with probability } \frac{1}{n}\end{cases} \\
Z= \begin{cases}0 & \text { with probability } 1 .\end{cases}
\end{gathered}
$$

Then all have expectation 0 . We have

$$
\begin{gathered}
\operatorname{Var}[X]=(1 / 2)(1000)^{2}+(1 / 2)(1000)^{2}=1000000, \\
\operatorname{Var}[Y]=\frac{n-1}{n} \cdot 1^{2}+\frac{1}{n}(n-1)^{2}=\frac{n-1+(n-1)^{2}}{n}=\frac{n^{2}-n}{n}=n-1, \\
\operatorname{Var}[Z]=0 .
\end{gathered}
$$

Fact 2. If $X$ is $a$ random variable, and $Y=a X+b$, for constants $a, b$, then $\operatorname{Var}[Y]=a^{2} \cdot \operatorname{Var}[X]$.

Proof.

$$
\begin{aligned}
\operatorname{Var}[Y] & =\mathbb{E}\left[(Y-\mathbb{E}[Y])^{2}\right] \\
& =\mathbb{E}\left[(a X+b-a \mathbb{E}[X]-b)^{2}\right] \\
& =\mathbb{E}\left[a^{2}(X-\mathbb{E}[X])^{2}\right]=a^{2} \cdot \operatorname{Var}[X] .
\end{aligned}
$$

This means that we can always shift $X$ to $X-\mathbb{E}[X]$ without changing its variance.

In general, if $X, Y$ are two random variables, then $\operatorname{Var}[X+Y]$ is not the same as $\operatorname{Var}[X]+\operatorname{Var}[Y]$. In fact, the variance of $X+Y$ might even be lower than the variance of either $X$ or $Y$.

$$
\begin{aligned}
& \text { For example, if } X=-Y \text {, then } \\
& \operatorname{Var}[X+Y]=0 .
\end{aligned}
$$

Fact 3. When $X$ and $Y$ are independent, $\operatorname{Var}[X+Y]=\operatorname{Var}[X]+\operatorname{Var}[Y]$.
Proof.

$$
\operatorname{Var}[X+Y]=\operatorname{Var}[X+Y-\mathbb{E}[X]-\mathbb{E}[Y]]=\operatorname{Var}\left[X^{\prime}+Y^{\prime}\right]
$$

where here $X^{\prime}=X-\mathbb{E}[X]$ and $Y^{\prime}=Y-\mathbb{E}[Y]$. Note that $\operatorname{Var}[X]=$ $\operatorname{Var}\left[X^{\prime}\right]$ and $\operatorname{Var}[Y]=\operatorname{Var}\left[Y^{\prime}\right]$. Then we have

$$
\begin{array}{ll}
=\mathbb{E}\left[\left(X^{\prime}+Y^{\prime}\right)^{2}\right] & \\
=\mathbb{E}\left[X^{\prime 2}+Y^{\prime 2}+2 X^{\prime} Y^{\prime}\right] & \\
=\mathbb{E}\left[X^{\prime 2}\right]+\mathbb{E}\left[Y^{\prime 2}\right]+\mathbb{E}\left[2 X^{\prime} Y^{\prime}\right] & \text { by linearity of expectation } \\
=\operatorname{Var}[X]+\operatorname{Var}[Y]+2 \cdot \mathbb{E}\left[X^{\prime}\right] \cdot \mathbb{E}\left[Y^{\prime}\right] & \\
=\sin [X, Y \text { of expectation } \\
=\operatorname{Var}[X]+\operatorname{Var}[Y] . & \text { since } \mathbb{E}\left[X^{\prime}\right]=0
\end{array}
$$

