

## Maximum Likelihood Estimation

Statistics: about analyzing & understanding data

Common approach: use parametric model of data  
Bin( $n, p$ ), Poi( $\lambda$ ), Exp( $\lambda$ ),  $N(\mu, \sigma^2)$ , ...

How to estimate params of distn from samples?

Approach: Find MLE, most likely choice of parameters.

Ex: outcome of  $n$  coin tosses.  $x_1, \dots, x_n$  each  $x_i \in \{H, T\}$   
assuming indep Bin( $p$ ) what's  $p$ ?

$$L(x_1, \dots, x_n | \theta) = \prod_{i=1}^n f(x_i | \theta)$$

unknown parameters (in this ex, prob of H's)  $\approx$  Prob of seeing  $x_i$  if param is  $\theta$   
discrete case p.m.f.  
cont case p.d.f.

$x_1, \dots, x_n$   $n_1$  H's  $n_0$  T's  $n = n_0 + n_1$

$$L(x_1, \dots, x_n | \theta) = \theta^{n_1} (1-\theta)^{n_0}$$

Find  $\theta$  that maximizes this

$$(LL) \log \text{Likelihood}(x_1, \dots, x_n | \theta) = \log L(x_1, \dots, x_n | \theta) = n_1 \log \theta + n_0 \log(1-\theta)$$

$$\frac{d}{d\theta} LL(\theta) = \frac{n_1}{\theta} - \frac{n_0}{1-\theta}$$

at max  
 $\hat{\theta}$

$$\frac{n_1}{\hat{\theta}} = \frac{n_0}{1-\hat{\theta}}$$

$$n_1(1-\hat{\theta}) = n_0 \hat{\theta}$$

$$\hat{\theta} = \frac{n_1}{n_0 + n_1} = \frac{n_1}{n}$$

In general

$$L(x_1, \dots, x_n | \theta) = \prod_{i=1}^n f(x_i | \theta)$$

$$LL = \sum_{i=1}^n \log f(x_i | \theta)$$

find max  $LL(x_1, \dots, x_n | \hat{\theta})$

$$\frac{d}{d\theta} LL(\vec{x} | \theta) = 0$$

solve for  $\hat{\theta}$

$$\frac{\partial}{\partial \theta_1} LL(\vec{x} | \vec{\theta}) = 0$$

$$\frac{\partial}{\partial \theta_2} LL(\vec{x} | \vec{\theta}) = 0$$

⋮

samples from a geometric distn w/ <sup>unknown</sup> param  $\theta$

$k_1, k_2, \dots, k_n$

$$L(k_1, \dots, k_n | \theta) = \prod_{i=1}^n \underbrace{(1-\theta)^{k_i-1}} \theta$$

$$LL(k_1, \dots, k_n | \theta) = \sum_{i=1}^n [(k_i-1) \log(1-\theta) + \log \theta]$$

$$\frac{d}{d\theta} LL = \sum_{i=1}^n \left[ \frac{-(k_i-1)}{1-\theta} + \frac{1}{\theta} \right]$$

set this = 0 to solve  $\hat{\theta}$

$$\sum_{i=1}^n \frac{k_i-1}{1-\theta} = \frac{n}{\theta}$$

$$\sum_{i=1}^n (k_i-1) \hat{\theta} = n(1-\hat{\theta})$$

$$\sum_{i=1}^n k_i \hat{\theta} = n$$

$$\hat{\theta} = \frac{n}{\sum_{i=1}^n k_i}$$

$x_1, \dots, x_n$

samples from

$N(\theta_1, \theta_2)$   
↑      ↑  
mean   variance.

$$L(x_1, \dots, x_n | \theta_1, \theta_2) = \prod_{i=1}^n f(x_i | \theta_1, \theta_2)$$

$$= \prod_{i=1}^n \left[ \frac{1}{\sqrt{2\pi\theta_2}} e^{-\frac{(x_i - \theta_1)^2}{2\theta_2}} \right]$$

density of  $N(\theta_1, \theta_2)$  at  $x_i$

$$LL(\vec{x} | \theta_1, \theta_2) = \sum_{i=1}^n \left[ -\frac{1}{2} \ln(2\pi\theta_2) - \frac{(x_i - \theta_1)^2}{2\theta_2} \right]$$

$$\frac{\partial LL}{\partial \theta_1} = 0$$

$$\frac{\partial LL}{\partial \theta_2} = 0$$

$$\Rightarrow \hat{\theta}_1, \hat{\theta}_2$$

$$\frac{\partial LL}{\partial \theta_1} = \sum_{i=1}^n \frac{x(x_i - \theta_1)}{\theta_2}$$

set = 0

$$\sum_{i=1}^n \frac{x_i}{\theta_2} = \sum_{i=1}^n \frac{\theta_1}{\theta_2}$$

$$\hat{\theta}_1 = \frac{\sum_{i=1}^n x_i}{n}$$

sample mean

$$\frac{\partial LL}{\partial \theta_2} = \sum_{i=1}^n \left[ -\frac{1}{2} \frac{2\pi}{2\pi\theta_2} + \frac{1}{\theta_2} \frac{(x_i - \theta_1)^2}{\theta_2^2} \right]$$

set = 0

$$\frac{n}{2\theta_2} = \frac{1}{\theta_2^2} \sum_{i=1}^n (x_i - \hat{\theta}_1)^2$$

$$\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\theta}_1)^2$$

sample variance

$$E\left(\frac{\sum X_i}{n}\right) = \text{true mean.}$$

An estimator  $\hat{\theta} = g(x_1, \dots, x_n)$  is called unbiased if

$$E(\hat{\theta}) = \text{true parameter}$$

$$\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\theta}_1)^2$$

sample variance

Question: If  $X_i$  are iid  
distr. w/ mean  $\mu$ , var  $\sigma^2$   
is  $E[\text{sample var}] = \sigma^2$ ?

$$E(\hat{\theta}_2(x_1, \dots, x_n)) = E\left[\frac{1}{n} \sum_{i=1}^n \left(x_i - \underbrace{\frac{\sum x_i}{n}}_{\bar{X}}\right)^2\right]$$

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n (x_i - \bar{X})^2 &= \frac{1}{n} \sum_{i=1}^n (x_i^2 - 2x_i\bar{X} + \bar{X}^2) \\ &= \frac{1}{n} \sum_{i=1}^n x_i^2 - 2 \frac{1}{n} \sum_{i=1}^n x_i \bar{X} + \bar{X}^2 \\ &= \frac{1}{n} \sum_{i=1}^n x_i^2 - 2\bar{X}^2 + \bar{X}^2 = \frac{1}{n} \sum_{i=1}^n x_i^2 - \bar{X}^2 \end{aligned}$$

$$E\left[\frac{1}{n} \sum_{i=1}^n x_i^2 - \bar{X}^2\right] = \frac{1}{n} \sum_{i=1}^n E(x_i^2) - \frac{1}{n^2} E\left(\sum x_i\right)^2$$

$\swarrow$   
 $\text{Var}(x_i) = E(x_i^2) - (E(x_i))^2$   
 $\searrow$

$$= \frac{1}{n} n(\sigma^2 + \mu^2)$$

$x_1, \dots, x_n$

i.i.d.  $f, \sigma^2$

$$E \left[ \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2 \right] = ?$$

random exp.

$x_1, \dots, x_n$  1 exp

estimate variance

$$\frac{1}{n} \sum_{i=1}^n \left( x_i - \left( \frac{\sum x_i}{n} \right) \right)^2$$