

Reminders

$$X \sim N(\mu, \sigma^2) \quad Y = aX + b \quad Y \sim N(a\mu + b, a^2\sigma^2)$$

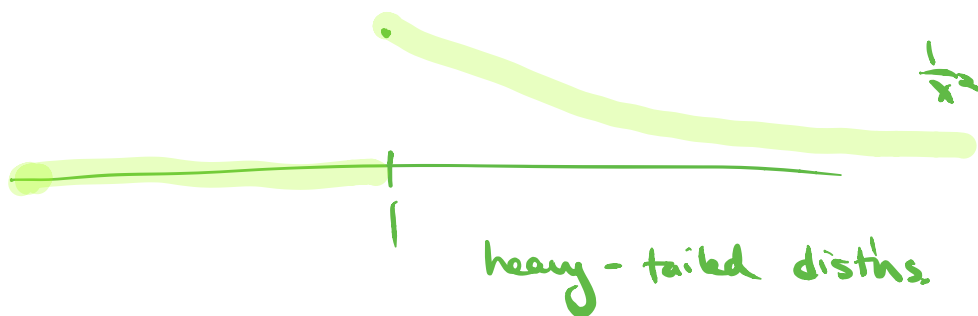
X_1, X_2, \dots, X_n indep normal r.v.'s
w/ $\mu_1, \mu_2, \dots, \mu_n$ $\sigma_1^2, \dots, \sigma_n^2$

$$Z = \sum_{i=1}^n X_i \quad Z \sim N\left(\sum_{i=1}^n \mu_i, \sum_{i=1}^n \sigma_i^2\right)$$

Example:
$$F_X(x) = \begin{cases} 1 - \frac{1}{x} & x \geq 1 \\ 0 & \text{o.w.} \end{cases}$$

$$f_X(x) = \begin{cases} \frac{1}{x^2} & x \geq 1 \\ 0 & x < 1 \end{cases}$$

$$E(X) = \int_1^{\infty} x f_X(x) dx = \int_1^{\infty} \frac{x}{x^2} dx = \ln x \Big|_1^{\infty}$$

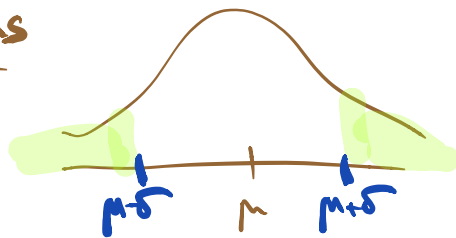


in what follows, I will assume
finite mean & variance

Tail bounds and limit theorems

Markov's Inequality $X \geq 0$

$$\Pr(X \geq \alpha E(X)) \leq \frac{1}{\alpha}$$



Chebyshev's Inequality

Y mean μ , var σ^2

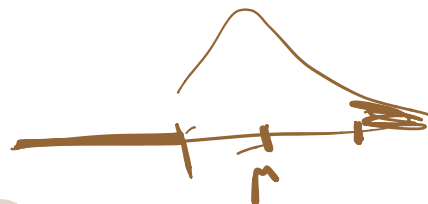
$$\Pr(|Y - E(Y)| \geq \delta) \leq \frac{\sigma^2}{\delta^2}$$

by Markov's inequality

$$\Pr\left(\underbrace{(Y - E(Y))^2}_{\geq \delta^2} \geq \alpha \underbrace{E[(Y - E(Y))^2]}_{\sigma^2}\right) \leq \frac{1}{\alpha}$$

$$\Pr\left(|Y - E(Y)| \geq \underbrace{\sqrt{\alpha} \sigma}_{\delta}\right) \leq \frac{1}{\alpha}$$

$$\sqrt{\alpha} \sigma = \delta \Rightarrow \frac{1}{\alpha} = \frac{\sigma^2}{\delta^2}$$



Law of Large Numbers

X_1, \dots, X_n are independent, identically distributed, iid

with $E(X_i) = \mu$
 $\text{Var}(X_i) = \sigma^2$

then $\frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{n \rightarrow \infty} \mu$

mean μ , variance $\frac{\sigma^2}{n}$

$$\text{Var}(\sum X_i) = n\sigma^2$$

linearity of var
for indep. r.v.'s.

$$\Pr\left(\left|\frac{1}{n}\sum X_i - \mu\right| \geq \alpha\right)$$

$$\leq \frac{\text{Var}\left(\frac{1}{n}\sum X_i\right)}{\alpha^2} = \frac{\sigma^2}{n\alpha^2} \xrightarrow{n \rightarrow \infty} 0$$

Chebyshev's Inequality

$$\begin{aligned} \text{Var}\left(\frac{1}{n}\sum X_i\right) &= \frac{1}{n^2} \text{Var}(\sum X_i) \\ &= \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n} \end{aligned}$$

Central Limit Theorem

X_1, \dots, X_n iid $E(X_i) = \mu$ $\text{Var}(X_i) = \sigma^2$

$$\frac{X_1 + \dots + X_n - n\mu}{\sqrt{n\sigma^2}} \xrightarrow{n \rightarrow \infty} N(0, 1)$$

$$\Pr\left(\frac{X_1 + \dots + X_n - n\mu}{\sqrt{n\sigma^2}} \leq a\right) \xrightarrow{n \rightarrow \infty} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^a e^{-\frac{x^2}{2}} dx$$

$\Phi(a)$

$$\frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{n \rightarrow \infty} N\left(\mu, \frac{\sigma^2}{n}\right)$$

Polling

Population of size N

p approve of Trump
 $1-p$ don't

Want to determine p by polling

call n people (each indep, random draw from population)

$$X_i = \begin{cases} 1 & \text{if } i\text{th person approves} \\ 0 & \text{o.w.} \end{cases}$$

$$E(X_i) = p$$
$$\text{Var}(X_i) = p(1-p)$$

$\frac{1}{n} \sum_{i=1}^n X_i$ fraction of people we call that approve.

$$\sim N\left(p, \frac{p(1-p)}{n}\right)$$

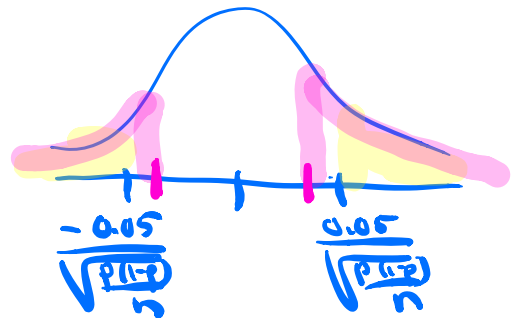
\hat{X} : sample mean.

$$\Pr(|\hat{X} - p| > \underbrace{0.05}_{\epsilon})$$

$$= \Pr\left(\left|\frac{\hat{X} - p}{\sqrt{\frac{p(1-p)}{n}}}\right| > \frac{0.05}{\sqrt{\frac{p(1-p)}{n}}}\right)$$

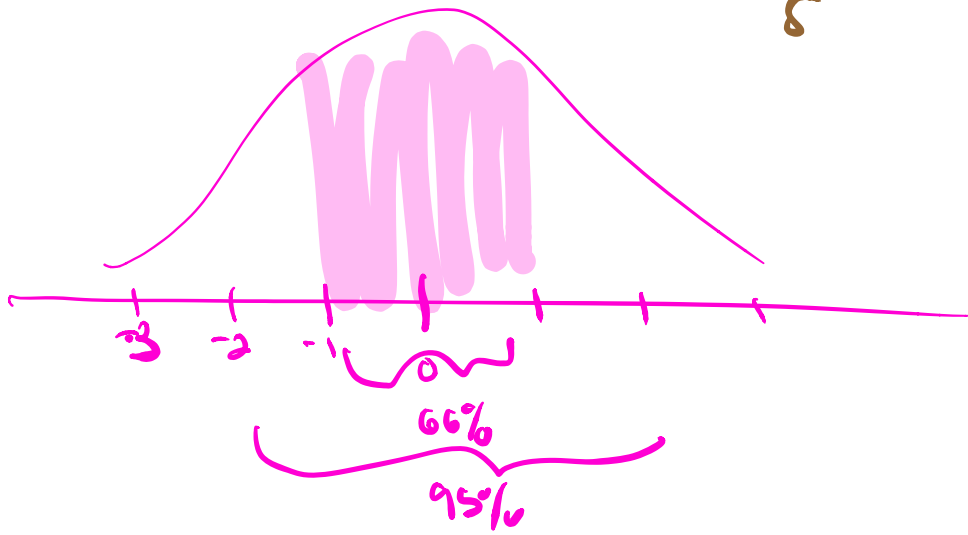
$$= \Pr\left(|Z| > \frac{0.05}{\sqrt{\frac{p(1-p)}{n}}}\right)$$

\uparrow
 $N(0,1)$



$$\leq \Pr(|Z| > \underbrace{0.1\sqrt{n}}_{0.05}) \leq 0.05$$

$\frac{0.05}{\sqrt{\frac{p(1-p)}{n}}}$



if

$$0.1\sqrt{n} = 2$$

$$\sqrt{n} = 20$$

$$n = 400$$