

## CSE 312: Foundations of Computing II

## Quiz Section #8: Normal Distribution, Central Limit Theorem, Tail Bounds

Review/Mini-Lecture/Main Theorems and Concepts From Lecture

**Normal (Gaussian, “bell curve”):**  $X \sim N(\mu, \sigma^2)$  if  $X$  has the following probability density function:

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\frac{(x-\mu)^2}{\sigma^2}}, \quad x \in \mathbb{R}$$

$E[X] = \mu$  and  $Var(X) = \sigma^2$ . The “standard normal” random variable is typically denoted  $Z$  and has mean 0 and variance 1. If  $X \sim N(\mu, \sigma^2)$ , then  $Z = \frac{X-\mu}{\sigma} \sim N(0,1)$ . The CDF has no closed form, but we denote the CDF of the standard normal as  $\Phi(z) = F_Z(z) = P(Z \leq z)$ . Note from symmetry of the probability density function about  $z = 0$  that:  $\Phi(-z) = 1 - \Phi(z)$ .

**Standardizing:** Let  $X$  be any random variable (discrete or continuous, not necessarily normal), with  $E[X] = \mu$  and  $Var(X) = \sigma^2$ . If we let  $Y = \frac{X-\mu}{\sigma}$ , then  $E[Y] = 0$  and  $Var(Y) = 1$ .

**Closure of the Normal Distribution:** Let  $X \sim N(\mu, \sigma^2)$ . Then,  $aX + b \sim N(a\mu + b, a^2\sigma^2)$ . That is, linear transformations of normal random variables are still normal.

**“Reproductive” Property of Normals:** Let  $X_1, \dots, X_n$  be independent normal random variables with  $E[X_i] = \mu_i$  and  $Var(X_i) = \sigma_i^2$ . Let  $a_1, \dots, a_n \in \mathbb{R}$  and  $b \in \mathbb{R}$ . Then,

$$X = \sum_{i=1}^n a_i X_i + b \sim N\left(\sum_{i=1}^n a_i \mu_i + b, \sum_{i=1}^n a_i^2 \sigma_i^2\right)$$

There’s nothing special about the parameters – the important result here is that the resulting random variable is still normally distributed.

**Central Limit Theorem (CLT):** Let  $X_1, \dots, X_n$  be iid random variables with  $E[X_i] = \mu$  and  $Var(X_i) = \sigma^2$ . Let  $X = \sum_{i=1}^n X_i$  which has  $E[X] = n\mu$  and  $Var(X) = n\sigma^2$ . Let  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ ,

which has  $E[\bar{X}] = \mu$  and  $Var(\bar{X}) = \frac{\sigma^2}{n}$ .  $\bar{X}$  is called the sample mean. Then, as  $n \rightarrow \infty$ ,  $Y =$

$\frac{\bar{X}-\mu}{\sigma/\sqrt{n}} \sim N(0,1)$  (same as  $\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$ ). Equivalently,  $Y' = \frac{X-n\mu}{\sigma\sqrt{n}} \sim N(0,1)$  (same as

$X \sim N(n\mu, n\sigma^2)$ ). It is no surprise that  $\bar{X}$  has mean  $\mu$  and variance  $\sigma^2/n$  – this can be done with simple calculations. The importance of the CLT is that, for large  $n$ , regardless of what distribution  $X_i$  comes from,  $\bar{X}$  is approximately normally distributed with mean  $\mu$  and variance  $\sigma^2/n$ . Don’t forget the continuity correction, only when  $X_1, \dots, X_n$  are discrete random variables.

**Markov’s Inequality:** Let  $X$  be a non-negative random variable, and  $\alpha \in \mathbb{R}$ . Then,  $P(X \geq \alpha) \leq \frac{E[X]}{\alpha}$ .

**Chebyshev’s Inequality:** Suppose  $Y$  is a random variable with  $E[Y] = \mu$  and  $Var(X) = \sigma^2$ . Then, for any  $\alpha \in \mathbb{R}$ ,  $P(|Y - \mu| \geq \alpha) \leq \frac{\sigma^2}{\alpha^2}$ .

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**Chernoff Bound (for the Binomial):** Suppose  $X \sim \text{Bin}(n, p)$ . Then, for any  $0 < \delta < 1$ ,

- $P(X > (1 + \delta)\mu) \leq e^{-\frac{\delta^2\mu}{3}}$
- $P(X < (1 - \delta)\mu) \leq e^{-\frac{\delta^2\mu}{2}}$

Exercises

1. Suppose  $Z = X + Y$ , where  $X \perp Y$ .  $Z$  is called the convolution of two random variables.

If  $X, Y, Z$  are discrete,

$$p_Z(z) = P(Z = z) = \sum_x P(X = x \cap Y = z - x) = \sum_x p_X(x)p_Y(z - x)$$

If  $X, Y, Z$  are continuous,

$$F_Z(z) = P(X + Y \leq z) = \int_{-\infty}^{\infty} P(Y \leq z - X \mid X = x)f_X(x)dx = \int_{-\infty}^{\infty} F_Y(z - x)f_X(x)dx$$

Suppose  $X_1 \sim N(\mu_1, \sigma_1^2)$  and  $X_2 \sim N(\mu_2, \sigma_2^2)$ .

a) Find an expression for  $P(X_1 < 2X_2)$  using a similar idea to convolution, in terms of  $F_{X_1}, F_{X_2}, f_{X_1}, f_{X_2}$ . (Your answer will be in the form of a single integral, and requires no calculations – do not evaluate it).

b) Find  $s$ , where  $\Phi(s) = P(X_1 < 2X_2)$  using the “reproductive” property of normal distributions.

2. Suppose  $X_1, \dots, X_n$  are iid  $Poi(\lambda)$  random variables, and let  $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ , the sample mean. How large should we choose  $n$  to be such that  $P\left(\frac{\lambda}{2} \leq \bar{X}_n \leq \frac{3\lambda}{2}\right) \geq 0.99$ ? Use the CLT and give an answer involving  $\Phi^{-1}(\cdot)$ . Then evaluate it exactly when  $\lambda = 1/10$  and using the phi table below.

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| <b>Z</b>   | <b>0.00</b> | <b>0.01</b> | <b>0.02</b> | <b>0.03</b> | <b>0.04</b> | <b>0.05</b> | <b>0.06</b> | <b>0.07</b> | <b>0.08</b> | <b>0.09</b> |
|------------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|
| <b>0.0</b> | 0.5000      | 0.5040      | 0.5080      | 0.5120      | 0.5160      | 0.5199      | 0.5239      | 0.5279      | 0.5319      | 0.5359      |
| <b>0.1</b> | 0.5398      | 0.5438      | 0.5478      | 0.5517      | 0.5557      | 0.5596      | 0.5636      | 0.5675      | 0.5714      | 0.5753      |
| <b>0.2</b> | 0.5793      | 0.5832      | 0.5871      | 0.5910      | 0.5948      | 0.5987      | 0.6026      | 0.6064      | 0.6103      | 0.6141      |
| <b>0.3</b> | 0.6179      | 0.6217      | 0.6255      | 0.6293      | 0.6331      | 0.6368      | 0.6406      | 0.6443      | 0.6480      | 0.6517      |
| <b>0.4</b> | 0.6554      | 0.6591      | 0.6628      | 0.6664      | 0.6700      | 0.6736      | 0.6772      | 0.6808      | 0.6844      | 0.6879      |
| <b>0.5</b> | 0.6915      | 0.6950      | 0.6985      | 0.7019      | 0.7054      | 0.7088      | 0.7123      | 0.7157      | 0.7190      | 0.7224      |
| <b>0.6</b> | 0.7257      | 0.7291      | 0.7324      | 0.7357      | 0.7389      | 0.7422      | 0.7454      | 0.7486      | 0.7517      | 0.7549      |
| <b>0.7</b> | 0.7580      | 0.7611      | 0.7642      | 0.7673      | 0.7704      | 0.7734      | 0.7764      | 0.7794      | 0.7823      | 0.7852      |
| <b>0.8</b> | 0.7881      | 0.7910      | 0.7939      | 0.7967      | 0.7995      | 0.8023      | 0.8051      | 0.8078      | 0.8106      | 0.8133      |
| <b>0.9</b> | 0.8159      | 0.8186      | 0.8212      | 0.8238      | 0.8264      | 0.8289      | 0.8315      | 0.8340      | 0.8365      | 0.8389      |
| <b>1.0</b> | 0.8413      | 0.8438      | 0.8461      | 0.8485      | 0.8508      | 0.8531      | 0.8554      | 0.8577      | 0.8599      | 0.8621      |
| <b>1.1</b> | 0.8643      | 0.8665      | 0.8686      | 0.8708      | 0.8729      | 0.8749      | 0.8770      | 0.8790      | 0.8810      | 0.8830      |
| <b>1.2</b> | 0.8849      | 0.8869      | 0.8888      | 0.8907      | 0.8925      | 0.8944      | 0.8962      | 0.8980      | 0.8997      | 0.9015      |
| <b>1.3</b> | 0.9032      | 0.9049      | 0.9066      | 0.9082      | 0.9099      | 0.9115      | 0.9131      | 0.9147      | 0.9162      | 0.9177      |
| <b>1.4</b> | 0.9192      | 0.9207      | 0.9222      | 0.9236      | 0.9251      | 0.9265      | 0.9279      | 0.9292      | 0.9306      | 0.9319      |
| <b>1.5</b> | 0.9332      | 0.9345      | 0.9357      | 0.9370      | 0.9382      | 0.9394      | 0.9406      | 0.9418      | 0.9429      | 0.9441      |
| <b>1.6</b> | 0.9452      | 0.9463      | 0.9474      | 0.9484      | 0.9495      | 0.9505      | 0.9515      | 0.9525      | 0.9535      | 0.9545      |
| <b>1.7</b> | 0.9554      | 0.9564      | 0.9573      | 0.9582      | 0.9591      | 0.9599      | 0.9608      | 0.9616      | 0.9625      | 0.9633      |
| <b>1.8</b> | 0.9641      | 0.9649      | 0.9656      | 0.9664      | 0.9671      | 0.9678      | 0.9686      | 0.9693      | 0.9699      | 0.9706      |
| <b>1.9</b> | 0.9713      | 0.9719      | 0.9726      | 0.9732      | 0.9738      | 0.9744      | 0.9750      | 0.9756      | 0.9761      | 0.9767      |
| <b>2.0</b> | 0.9772      | 0.9778      | 0.9783      | 0.9788      | 0.9793      | 0.9798      | 0.9803      | 0.9808      | 0.9812      | 0.9817      |
| <b>2.1</b> | 0.9821      | 0.9826      | 0.9830      | 0.9834      | 0.9838      | 0.9842      | 0.9846      | 0.9850      | 0.9854      | 0.9857      |
| <b>2.2</b> | 0.9861      | 0.9864      | 0.9868      | 0.9871      | 0.9875      | 0.9878      | 0.9881      | 0.9884      | 0.9887      | 0.9890      |
| <b>2.3</b> | 0.9893      | 0.9896      | 0.9898      | 0.9901      | 0.9904      | 0.9906      | 0.9909      | 0.9911      | 0.9913      | 0.9916      |
| <b>2.4</b> | 0.9918      | 0.9920      | 0.9922      | 0.9924      | 0.9927      | 0.9929      | 0.9931      | 0.9932      | 0.9934      | 0.9936      |
| <b>2.5</b> | 0.9938      | 0.9940      | 0.9941      | 0.9943      | 0.9945      | 0.9946      | 0.9948      | 0.9949      | 0.9951      | 0.9952      |
| <b>2.6</b> | 0.9953      | 0.9955      | 0.9956      | 0.9957      | 0.9958      | 0.9960      | 0.9961      | 0.9962      | 0.9963      | 0.9964      |
| <b>2.7</b> | 0.9965      | 0.9966      | 0.9967      | 0.9968      | 0.9969      | 0.9970      | 0.9971      | 0.9972      | 0.9973      | 0.9974      |
| <b>2.8</b> | 0.9974      | 0.9975      | 0.9976      | 0.9977      | 0.9977      | 0.9978      | 0.9979      | 0.9979      | 0.9980      | 0.9981      |
| <b>2.9</b> | 0.9981      | 0.9982      | 0.9982      | 0.9983      | 0.9984      | 0.9984      | 0.9985      | 0.9985      | 0.9986      | 0.9986      |