

Statistics: analyzing & understanding data

Common approach: use parametric model of data

$\text{Bin}(p)$, $\text{Poi}(\lambda)$, $\text{Exp}(\lambda)$, $N(\mu, \sigma^2)$, $\text{Uni}[a, b]$

Use $\vec{\theta}$ to denote unknown parameters

Goal: Given $\underbrace{x_1, x_2, \dots, x_n}_{\text{indep samples}}$ from parametric model, determine best choice of parameters $\vec{\theta}$

Approach: Find MLE, most likely choice of $\vec{\theta}$

$$L(x_1, \dots, x_n | \vec{\theta}) = \prod_{i=1}^n f(x_i | \vec{\theta})$$

likelihood function

density cont.
p.m.f. cont.

$$\text{LL}(\vec{x} | \vec{\theta}) = \log(L(\vec{x} | \vec{\theta})) = \sum_{i=1}^n \log[f(x_i | \vec{\theta})]$$

log-likelihood function

choose $\vec{\theta}$ to maximize $L(\vec{x} | \vec{\theta})$

\Leftrightarrow maximize $LL(\vec{x} | \vec{\theta})$

1 parameter

compute $\frac{dLL}{d\theta}$

set $\frac{dLL}{d\theta} = 0$

Solve

verify soln is max (2nd deriv < 0)

Multiple parameters

$\frac{\partial LL}{\partial \theta_1} = 0$

$\frac{\partial LL}{\partial \theta_2} = 0$

⋮

$\frac{\partial LL}{\partial \theta_k} = 0$

check max
Hessian -ve definite

$\frac{\partial^2 f}{\partial x_i \partial x_j}$

Example :

k_1, \dots, k_n samples from geometric dist'n, param θ

Find MLE of θ

$$L(\vec{k} | \theta) = \prod_{i=1}^n (1-\theta)^{k_i-1} \theta$$

$$p(k_i) = (1-\theta)^{k_i-1} \theta$$

$$LL(\vec{k} | \theta) = \sum_{i=1}^n (k_i-1) \log(1-\theta) + n \log \theta$$

$$\frac{d}{d\theta} \text{LL}(\vec{k} | \theta) = -\sum_i \frac{(k_i - 1)}{(1-\theta)} + \frac{n}{\theta}$$

$$\frac{d}{d\theta} \text{LL}(\vec{k} | \theta) = 0 \Rightarrow \frac{n}{\hat{\theta}} = \sum_i \frac{(k_i - 1)}{(1-\theta)}$$

$$\equiv (1-\hat{\theta})n = \hat{\theta} \sum_i (k_i - 1)$$

$$\equiv \hat{\theta} = \frac{n}{\sum_i k_i}$$

X_1, \dots, X_n samples from $\cup [0, \theta]$

↑
unknown

Find MLE of θ

$$f(x_i) = \begin{cases} \frac{1}{\theta} & 0 \leq x_i \leq \theta \\ 0 & \text{o.w.} \end{cases}$$

$$L(\vec{x} | \theta) = \frac{1}{\theta^n} \quad \text{if all } x_i \in [0, \theta]$$

0 otherwise

$$L(\vec{x} | \theta) \downarrow \text{as } \theta \uparrow \Rightarrow \hat{\theta} = \max(x_i)$$

Estimator unbiased if $E(\hat{\theta}) = \theta$
 ↑
 true value

Last time for mean of normal dist'n

$$E(\hat{\theta}) = E\left(\frac{\sum X_i}{n}\right) = \mu \quad \text{unbiased.}$$

for $\mathbb{U}[0, \Theta]$ $\hat{\theta} = \max(x_1, \dots, x_n)$

$$E(\max(X_1, \dots, X_n)) = ? \quad \Pr(\hat{\theta} \leq x) = \prod_{i=1}^n \Pr(X_i \leq x)$$

\downarrow
 $\mathbb{U}[0, \Theta]$

$$= \begin{cases} \frac{x^n}{\Theta^n} & 0 \leq x \leq \Theta \\ 0 & \text{otherwise} \end{cases}$$

$$f_{\Theta}(x) = \begin{cases} \frac{n x^{n-1}}{\Theta^n} & 0 \leq x \leq \Theta \\ 0 & \text{o.w.} \end{cases}$$

$$E(\hat{\theta}) = \int_0^\Theta x f_\Theta(x) dx = \frac{n}{\Theta^n} \int_0^\Theta x \cdot x^{n-1} = \frac{n}{n+1} \Theta$$

$\Rightarrow \hat{\theta}$ not unbiased, but it is asymptotically unbiased

$$\lim_{n \rightarrow \infty} E(\hat{\theta}_n) = \Theta$$

Estimator **consistent** if $\hat{\Theta}_n \xrightarrow[n \rightarrow \infty]{\text{in probability}} \Theta$ no matter what Θ is

$$\Pr(\max(X_1, \dots, X_n) \leq \Theta - \varepsilon) = \left(\frac{\Theta - \varepsilon}{\Theta}\right)^n = \left(1 - \frac{\varepsilon}{\Theta}\right)^n \rightarrow 0 \text{ as } n \rightarrow \infty$$

(under mild technical conditions, MLE always consistent)

iid observations $X_1, \dots, X_n \sim U[\Theta, \Theta+1]$

Find MLE of Θ

$$f_{X_i}(x_i) = \begin{cases} 1 & \Theta \leq x_i \leq \Theta + 1 \\ 0 & \text{o.w.} \end{cases}$$

$$L(\vec{x} | \Theta) = \begin{cases} 1 & \Theta \leq \min x_i \leq \max x_i \leq \Theta + 1 \\ 0 & \text{otherwise} \end{cases}$$

Any $\hat{\Theta} \in [\max X_i - 1, \min X_i]$ equally good

consistent since $\min X_i \rightarrow \Theta$

$\max X_i \rightarrow \Theta + 1$

If choose midpoint $\hat{\Theta} = \frac{1}{2} [\max X_i + \min X_i - 1]$
it's unbiased

$$E(\hat{\theta}) = \frac{1}{2} \left[E(\max X_i) + E(\min X_i) - 1 \right] = \theta$$

$\theta + 1 - \frac{1}{n+1}$ $\theta + \frac{1}{n+1}$

Grade dist'n for 312

A	$\frac{1}{2}$
B	μ
C	2μ
F	$\frac{1}{2} - 3\mu$

MLE for μ when samples are x_A, x_B, x_C, x_F
 (Assume we see precise grade each person got)

$$L(x_A, x_B, x_C, x_F | \mu) = \left(\frac{1}{2}\right)^{x_A} \mu^{x_B} (2\mu)^{x_C} \left(\frac{1}{2} - 3\mu\right)^{x_D}$$

$$\log L(\vec{x} | \mu) = x_A \log\left(\frac{1}{2}\right) + x_B \log(\mu) + x_C \log(2\mu) + x_D \log\left(\frac{1}{2} - 3\mu\right)$$

Set $\frac{d \log L}{d \mu} = 0$, solve for $\hat{\mu}$

$$\frac{x_B}{\hat{\mu}} + \frac{2x_C}{2\hat{\mu}} - \frac{3x_D}{\frac{1}{2} - 3\hat{\mu}} = 0$$

$$\frac{(x_B + x_C)}{\hat{\mu}} = \frac{3x_D}{\frac{1}{2} - 3\hat{\mu}}$$

$$\left(\frac{1}{2} - 3\hat{\mu}\right)(x_B + x_C) = 3\hat{\mu}x_D$$

$$\frac{x_B + x_C}{2} = 3\hat{\mu}(x_B + x_C + x_D)$$

$$\hat{\mu} = \frac{x_B + x_C}{6(x_B + x_C + x_D)}$$

Same problem, but suppose we don't see x_A & x_B

We just see x_{AB} = # students that got A or B

Expectation-Maximization Algorithm EM

alg for finding MLE of parameters when there are hidden vars
or computational issues with MLE

Idea:

If we knew z_A, z_B (where $z_A + z_B = x_{AB}$) could compute $\hat{\mu}$

If we knew $\hat{\mu}$ → could determine $E(z_A), E(z_B)$

EM: Starts w/ guess for μ : μ^0

Repeatedly $t=0, 1, \dots$

Expectation: (compute $E(z_A | \mu^t, x_{AB})$)

$$\bar{z}_A = E(z_A | \mu^t, x_{AB}) = \frac{x_{AB} \frac{1}{2}}{\frac{1}{2} + \mu^t}$$

$$\bar{z}_B = E(z_B | \mu^t, x_{AB}) = x_{AB} - E(z_A | \mu^t, x_{AB})$$

Maximization:

$$LL(z_A, z_B, x_c, x_D | \mu)$$

$$= z_A \log\left(\frac{1}{2}\right) + z_B \log(\mu) + x_c \log(2\mu) + x_D (\log\left(\frac{1}{2} - 3\mu\right))$$

recall: $\log L(\vec{x} | \mu) = x_A \log\left(\frac{1}{2}\right) + x_1(\mu) + x_c \log(2\mu) + x_D \log\left(\frac{1}{2} - 3\mu\right)$

$$\begin{aligned} E & \left[LL(z_A, z_B, x_c, x_D | \mu) \right] \\ &= \bar{z}_A \log\left(\frac{1}{2}\right) + \bar{z}_B \log(\mu) + x_c \log(2\mu) + x_D (\log\left(\frac{1}{2} - 3\mu\right)) \end{aligned}$$

Find μ th - maximizes $E[LL(z_A, z_B, x_c, x_D | \mu)]$

$$\text{call this } \mu^{t+1} = \frac{\bar{z}_B + x_c}{6(\bar{z}_B + x_c + x_D)}$$

until convergence

$$\text{recall } \hat{M} = \frac{x_B + x_C}{6(x_B + x_C + x_D)}$$

Summary of structure of EM alg

- Samples x_1, \dots, x_n Missing data z_1, \dots, z_m

Parameters: $\theta_1, \dots, \theta_k$

Initialize $\vec{\theta}^0$

repeat until convergence $t=0, \dots$

Expectation:

compute $E(z_i | \vec{x}, \vec{\theta}^+)$

Maximization:

Set $\vec{\theta}^{++}$ to

maximize $E(\text{Likelihood}(\vec{x}, \vec{z} | \vec{\theta}))$

Mixture of Normals

Initialize $\vec{\theta}^0 = (\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \tau_1)$ $\tau_2 = 1 - \tau_1$

Repeat until convergence:

Expectation:

$$\forall i \quad E(z_{ii} | x_i; \vec{\theta}^+) = \frac{f_1(x_i; \vec{\theta}^+) \tau_1}{f_1(x_i; \vec{\theta}^+) \tau_1 + f_2(x_i; \vec{\theta}^+) \tau_2}$$

Maximization :

$$L(\vec{x}, \vec{z} | \vec{\theta}) = \prod_{i=1}^n \left(\tau_1 f_1(x_i; \vec{\theta}) \right)^{z_{ii}} \left(\tau_2 f_2(x_i; \vec{\theta}) \right)^{z_{i2}}$$

$$= \prod_{i=1}^n \left(\frac{\tau_1}{\sqrt{2\pi\sigma_1^2}} \right)^{z_{ii}} \left(\frac{\tau_2}{\sqrt{2\pi\sigma_2^2}} \right)^{z_{i2}} e^{-\frac{z_{ii}(x_i - \mu_1)^2}{2\sigma_1^2}} e^{-\frac{z_{i2}(x_i - \mu_2)^2}{2\sigma_2^2}}$$

$$E_{\vec{z}} (\log L(\vec{x}, \vec{z} | \vec{\theta})) =$$

$$E \left(\sum_{i=1}^n z_{ii} \left[\ln \left(\frac{\tau_1}{\sqrt{2\pi\sigma_1^2}} \right) - \frac{(x_i - \mu_1)^2}{2\sigma_1^2} \right] \right)$$

$$+ \sum_{i=1}^n z_{i2} \left[\ln \left(\frac{\tau_2}{\sqrt{2\pi\sigma_2^2}} \right) - \frac{(x_i - \mu_2)^2}{2\sigma_2^2} \right]$$

$$= \sum_{i=1}^n E(z_{ii}) \left[\ln \left(\frac{\tau_1}{\sqrt{2\pi\sigma_1^2}} \right) - \frac{(x_i - \mu_1)^2}{2\sigma_1^2} \right]$$

$$+ \sum_{i=1}^n E(z_{i2}) \left[\ln \left(\frac{\tau_2}{\sqrt{2\pi\sigma_2^2}} \right) - \frac{(x_i - \mu_2)^2}{2\sigma_2^2} \right]$$

find \vec{g}^{t+1} to maximize this

using $E(z_{ii}) E(z_{ia})$
computed before

Turns out $M_j = \frac{\sum_i E(z_{ij}) x_i}{\sum_i E(z_{ij})}$