# 15-251 Great Theoretical Ideas in Computer Science

### CMU

### Sample Space



weight or probability of t D(t) = p(t) = 0.2

### **Events**

#### Any set $E \subseteq S$ is called an event



## $Pr_{D}[E] = 0.4$

### **Random Variable**

Let S be sample space in a probability distribution A Random Variable is a function from S to reals

**Examples:** 

X = value of white die in a two-dice rollX(3,4) = 3,X(1,6) = 1Y = sum of values of the two diceY(3,4) = 7,Y(1,6) = 7

## Notational Conventions Use letters like A, B, C for events Use letters like X, Y, f, g for R.V.'s R.V. = random variable



#### **Two Views of Random Variables**

Think of a R.V. as

Input to the function is random

A function from S to the reals R

Or think of the induced distribution on R

Randomness is "pushed" to the values of the function

Given a distribution, a random variable transforms it into a distribution on reals

#### **Two dice**

I throw a white die and a black die.

Sample space S =

 $\{(1,1), (1,2), (1,3), (1,4), \}$ (1,5), (1,6), (2,1),(2,2),(2,3), (2,4), (2,5), (2,6), (3,3), (3,4), (3,6), (3,2), (3,1), (3,5), (4,3), (4,5), (4,1), (4,2), (4,4), (4,6), (5,4), (5,1), (5,2), (5,3), (5,6), (5,5), (6,1), (6,2), (6,3), (6,4), (6,5), (6,6) } Probability mass function p.m.f.



#### X = sum of both dice

function with X(1,1) = 2, X(1,2) = 3, ..., X(6,6)=12

### It's a Floor Wax And a Dessert Topping

It's a function on the sample space S

It's a variable with a probability distribution on its values

You should be comfortable with both views

#### **Two dice**

I throw a white die and a black die.

Sample space S =

 $\{(1,1), (1,2), (1,3), (1,4), \}$ (1,5), (1,6), (2,1),(2,2),(2,3), (2,4), (2,5), (2,6), (3,3), (3,4), (3,6), (3,2), (3,1), (3,5), (4,3), (4,5), (4,1), (4,2), (4,4), (4,6), (5,4), (5,1), (5,2), (5,3), (5,6), (5,5), (6,1), (6,2), (6,3), (6,4), (6,5), (6,6) } Probability mass function p.m.f.



#### X = sum of both dice

function with X(1,1) = 2, X(1,2) = 3, ..., X(6,6)=12

#### **From Random Variables to Events**

Note that each event in the induced distribution corresponds to some event in the original one.

For any random variable X and value a, we can define the event A that X = a

 $Pr(A) = Pr(X=a) = Pr(\{t \in S \mid X(t)=a\})$ 

### **Two Coins Tossed** X: {TT, TH, HT, HH} $\rightarrow$ {0, 1, 2} counts # of heads



Pr(X = a) = $Pr({t \in S | X(t) = a})$ 

Pr(X = 1)

 $= \Pr(\{t \in S | X(t) = 1\})$ 

**Distribution** =  $Pr({TH, HT}) = \frac{1}{2}$ 

#### **Two dice**

I throw a white die and a black die.

Sample space S =  $\{(1,1), (1,2), (1,3), (1,4), (1,5), (1,6),$ (2,5), (2,1), (2,2),(2,3), (2,4), (2,6), (3,3),(3,4), (3,5), (3,6), (3,1), (3,2), (4,6), (4,3)(4,4), (4,1), (4,5), (4.2)(5,6), (5.2)(5,3), (5,5), (5,1) (5,4), (6,2), (6,3), (6,5), (6,4), **(6,6)** } (6,1),



X = sum of both dice

Pr(X = 7) = 6/36 = 1/6

### **Definition: Expectation**

The expectation, or expected value of a random variable X is written as E[X], and is

$$E[X] = \sum_{k \in S} Pr(t) X(t) = \sum_{k \in S} k Pr[X = k]$$

X is a function on the sample space S

X has a prob. distribution on its values

(assuming X takes values in the naturals)



#### Quick $\sum_{t \in S} \Pr(t) X(t) = \sum_{k} k \Pr[X = k]$ check: $t \in S$ k

#### $Pr(X = a) = Pr(\{t \in S | X(t) = a\})$

### A Quick Calculation...

What if I flip a coin 3 times? What is the expected number of heads?

 $E[X] = (1/8) \times 0 + (3/8) \times 1 + (3/8) \times 2 + (1/8) \times 3 = 1.5$ 

But Pr[X = 1.5] = 0

Moral: don't always expect the expected. Pr[ X = E[X] ] may be 0 !



#### **Operations on R.V.s**

You can sum them, take differences, or do most other math operations...

E.g., (X + Y)(t) = X(t) + Y(t)

 $(X^{*}Y)(t) = X(t) * Y(t)$ 

 $(X^{\vee})(t) = X(t)^{\vee(t)}$ 

**Random variables** and expectations allow us to give elegant solutions to problems that seem really really messy...

If I randomly put 100 letters into 100 addressed envelopes, on average how many letters will end up in their correct envelopes?

On average, in class of size m, how many pairs of people will have the same birthday?

Pretty messy with direct counting...

The new tool is called "Linearity of Expectation"

### Linearity of Expectation

# If Z = X+Y, then E[Z] = E[X] + E[Y]

Even if X and Y are not independent

# **By Induction** $E[X_1 + X_2 + ... + X_n] =$ $E[X_1] + E[X_2] + .... + E[X_n]$ The expectation of the sum The sum of the expectations

### Expectation of a Sum of r.v.s = Sum of their Expectations

even when r.v.s are not independent!

Expectation of a Product of r.v.s = Product of their Expectations

**ONLY** when r.v.s are independent!

### Independence for r.v.s

Two random variables X and Y are independent if for every a,b, the events X=a and Y=b are independent

How about the case of X=1st die, Y=2nd die?

Let's test our Linearity of Expectation chops... If I randomly put 100 letters into 100 addressed envelopes, on average how many letters will end up in their correct envelopes?

Hmm...

 $\sum_{k} k Pr(exactly k letters end up in correct envelopes)$ 

 $= \sum_{k} k (\dots aargh!!\dots)$ 

#### **Indicator Random Variables**

For any event A, can define the "indicator random variable" for event A:

$$X_{A}(t) = \begin{cases} 1 & \text{if } t \in A \\ 0 & \text{if } t \notin A \end{cases}$$

$$E[X_A] = 1 \times Pr(X_A = 1) = Pr(A)$$



#### **Use Linearity of Expectation**

Let A<sub>i</sub> be the event the i<sup>th</sup> letter ends up in its correct envelope

Let X<sub>i</sub> be the "indicator" R.V. for A<sub>i</sub>



1 if A<sub>i</sub> occurs 0 otherwise  $X_i =$ Let  $Z = X_1 + \dots + X_{100}$ We are asking for **E**[**Z**]  $E[X_i] = Pr(A_i) = 1/100$ So E[Z] = 1

So, in expectation, 1 letter will be in the same correct envelope

Pretty neat: it doesn't depend on how many letters!

**A**Question: were the X<sub>i</sub> independent?

No! E.g., think of n=2

#### **Use Linearity of Expectation**

**General approach:** 

View thing you care about as expected value of some RV

Write this RV as sum of simpler RVs

Solve for their expectations and add them up!

### Example #2

We flip n coins of bias p. What is the expected number of heads?

We could do this by summing

$$\sum_{k} k \operatorname{Pr}(X = k) = \sum_{k} k {n \choose k} p^{k} (1-p)^{n-k}$$
$$= n.p$$

But now we know a better way!

### Linearity of Expectation!

Let X = number of heads when n independent coins of bias p are flipped

**Break X into n simpler RVs:** 

 $X_{i} = \begin{cases} 1 & \text{if the } i^{\text{th}} \text{ coin is heads} \\ 0 & \text{if the } i^{\text{th}} \text{ coin is tails} \end{cases}$ 

 $\mathsf{E}[\mathsf{X}] = \mathsf{E}[\Sigma_i \mathsf{X}_i] = \Sigma_i \mathsf{E}[\mathsf{X}_i] = \Sigma_i \mathsf{p} = \mathsf{np}$ 

On average, in class of size m, how many pairs of people will have the same birthday?

 $\sum_{k} k Pr(exactly k collisions)$ 

 $=\sum_{k} k (\dots aargh!!!!)$ 

Use linearity of expectation

Suppose we have m people each with a uniformly chosen birthday from 1 to 365

X = number of pairs of people with the same birthday

E[X] = ?

X = number of pairs of people with the same birthday

E[X] = ?

Use m(m-1)/2 indicator variables, one for each pair of people

 $X_{jk}$  = 1 if person j and person k have the same birthday; else 0

E[X<sub>jk</sub>] = (1/365) 1 + (1 – 1/365) 0 = 1/365 X = number of pairs of people with the same birthday

X<sub>jk</sub> = 1 if person j and person k have the same birthday; else 0

 $E[X_{jk}] = (1/365) 1 + (1 - 1/365) 0$ = 1/365

 $E[X] = E[\sum_{1 \le j < k \le m} X_{jk}]$  $= \sum_{1 \le j < k \le m} E[X_{jk}]$  $= m(m-1)/2 \times 1/365$ 

E.g., setting m = 23, we get E[# pairs with same birthday ] = 0.691...

How does this compare to Pr[ at least one pair has same birthday]?