

CSE 312

Foundations of Computing II

**Lecture 14: Wrapup of Bloom filters +
Continuous RV**

Agenda

- Wrap-up of Bloom Filters ◀
- Continuous Random Variables
- Probability Density Function
- Cumulative Distribution Function
- Expectation and Variance of continuous RVs

Bloom Filters – Main points

- Probabilistic data structure.
- Close cousins of hash tables.
 - But: Ridiculously space efficient
- Occasional errors, specifically false positives.

Bloom Filters

- Stores information about a set of elements $S \subseteq U$.
- Supports two operations:
 1. **add**(x) - adds $x \in U$ to the set S
 2. **contains**(x) – ideally: true if $x \in S$, false otherwise



Instead, relaxed guarantees:

- False → **definitely** not in S
- True → **possibly** in S
[i.e. we could have *false positives*]

Bloom Filters – Why Accept False Positives?

- **Speed** – both **add** and **contains** very very fast.
- **Space** – requires a miniscule amount of space relative to storing all the actual items that have been added.
 - Often just 8 bits per inserted item!
- **Fallback mechanism** – can distinguish false positives from true positives with extra cost
 - Ok if mostly negatives expected + low false positive rate

Bloom Filters – Ingredients

Basic data structure is a $k \times m$ binary array
“the Bloom filter”

- k rows t_1, \dots, t_k , each of size m
- Think of each row as an m -bit vector

k different hash functions $\mathbf{h}_1, \dots, \mathbf{h}_k: U \rightarrow [m]$

Bloom Filters – Three operations

- Set up Bloom filter for $S = \emptyset$

```
function INITIALIZE( $k, m$ )  
  for  $i = 1, \dots, k$ : do  
     $t_i =$  new bit vector of  $m$  0s
```

- Update Bloom filter for $S \leftarrow S \cup \{x\}$

```
function ADD( $x$ )  
  for  $i = 1, \dots, k$ : do  
     $t_i[h_i(x)] = 1$ 
```

- Check if $x \in S$

```
function CONTAINS( $x$ )  
  return  $t_1[h_1(x)] == 1 \wedge t_2[h_2(x)] == 1 \wedge \dots \wedge t_k[h_k(x)] == 1$ 
```

Bloom Filters: Example

Bloom filter t of length $m = 5$ that uses $k = 3$ hash functions

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| t_1 | 0 | 0 | 0 | 0 | 0 |
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function ADD( $x$ )  
  for  $i = 1, \dots, k$ : do  
     $t_i[h_i(x)] = 1$ 
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add("thisisavirus.com")

$h_1(\text{"thisisavirus.com"}) \rightarrow 2$

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Bloom Filters: Contains

```
function CONTAINS( $x$ )
```

```
    return  $t_1[h_1(x)] == 1 \wedge t_2[h_2(x)] == 1 \wedge \dots \wedge t_k[h_k(x)] == 1$ 
```

Returns True if the bit vector t_i for each hash function has bit 1 at index determined by $h_i(x)$,

Returns False otherwise

Bloom Filters: Example

Bloom filter t of length $m = 5$ that uses $k = 3$ hash functions

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function CONTAINS( $x$ )  
  return  $t_1[h_1(x)] == 1 \wedge t_2[h_2(x)] == 1 \wedge \dots \wedge t_k[h_k(x)] == 1$ 
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contains("thisisavirus.com")

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True

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h_1 ("thisisavirus.com") \rightarrow 2

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Since all conditions satisfied, returns **True** (correctly)

Bloom Filters: False Positives

Bloom filter t of length $m = 5$ that uses $k = 3$ hash functions

add("totallynotsuspicious.com")

```
function ADD( $x$ )  
  for  $i = 1, \dots, k$ : do  
     $t_i[h_i(x)] = 1$ 
```

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h_1 ("totallynotsuspicious.com") \rightarrow 1

h_2 ("totallynotsuspicious.com") \rightarrow 0

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Bloom Filters: False Positives

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```
function ADD( $x$ )  
  for  $i = 1, \dots, k$ : do  
     $t_i[h_i(x)] = 1$ 
```

add("totalynotsuspicious.com")

h_1 ("totalynotsuspicious.com") \rightarrow 1

h_2 ("totalynotsuspicious.com") \rightarrow 0

h_3 ("totalynotsuspicious.com") \rightarrow 4

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h_1 ("totalynotsuspicious.com") \rightarrow 1

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```

contains("verynormalsite.com")

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True

contains("verynormalsite.com")

h_1 ("verynormalsite.com") \rightarrow 2

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True

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Since all conditions satisfied, returns **True** (incorrectly)

Analysis: False positive probability

Question: For an element $x \in U$, what is the probability that **contains**(x) returns true if **add**(x) was never executed before?

Probability over what?! Over the choice of the h_1, \dots, h_k

Assumptions for the analysis:

- Each $h_i(x)$ is uniformly distributed in $[m]$ for all x and i
- Hash function outputs for each h_i are mutually independent (not just in pairs)
- Different hash functions are independent of each other

False positive probability – Events

Assume we perform **add**(x_1), ..., **add**(x_n)
+ **contains**(x) for $x \notin \{x_1, \dots, x_n\}$

False positive iff $\mathbf{h}_i(x) \in \{\mathbf{h}_i(x_1), \dots, \mathbf{h}_i(x_n)\}$. $\forall i = 1, \dots, k$

$P(\text{false positive})$


False positive probability – Events

Assume we perform **add**(x_1), ..., **add**(x_n)
+ **contains**(x) for $x \notin \{x_1, \dots, x_n\}$

Event E_i holds iff $\mathbf{h}_i(x) \in \{\mathbf{h}_i(x_1), \dots, \mathbf{h}_i(x_n)\}$

$$P(\text{false positive}) = P(E_1 \cap E_2 \cap \dots \cap E_k) = \prod_{i=1}^k P(E_i)$$

$\mathbf{h}_1, \dots, \mathbf{h}_k$ independent



False positive probability – Events

Event E_i holds iff $\mathbf{h}_i(x) \in \{\mathbf{h}_i(x_1), \dots, \mathbf{h}_i(x_n)\}$

Event E_i^c holds iff $\mathbf{h}_i(x) \neq \mathbf{h}_i(x_1)$ and ... and $\mathbf{h}_i(x) \neq \mathbf{h}_i(x_n)$

$$P(E_i^c) = \sum_{z=1}^m P(\mathbf{h}_i(x) = z) \cdot P(E_i^c \mid \mathbf{h}_i(x) = z)$$

LTP



False positive probability – Events

Event E_i^c holds iff $\mathbf{h}_i(x) \neq \mathbf{h}_i(x_1)$ and ...
and $\mathbf{h}_i(x) \neq \mathbf{h}_i(x_n)$

$$P(E_i^c | \mathbf{h}_i(x) = z) = P(\mathbf{h}_i(x_1) \neq z, \dots, \mathbf{h}_i(x_n) \neq z | \mathbf{h}_i(x) = z)$$

Independence of values
of \mathbf{h}_i on different inputs

$$= P(\mathbf{h}_i(x_1) \neq z, \dots, \mathbf{h}_i(x_n) \neq z)$$

$$= \prod_{j=1}^n P(\mathbf{h}_i(x_j) \neq z)$$

False positive probability – Events

Event E_i^c holds iff $\mathbf{h}_i(x) \neq \mathbf{h}_i(x_1)$ and ...
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Independence of values
of \mathbf{h}_i on different inputs

$$= P(\mathbf{h}_i(x_1) \neq z, \dots, \mathbf{h}_i(x_n) \neq z)$$

$$= \prod_{j=1}^n P(\mathbf{h}_i(x_j) \neq z)$$

Outputs of \mathbf{h}_i uniformly spread

$$= \prod_{j=1}^n \left(1 - \frac{1}{m}\right) = \left(1 - \frac{1}{m}\right)^n$$


$$\longrightarrow P(E_i^c) = \sum_{z=1}^m P(\mathbf{h}_i(x) = z) \cdot P(E_i^c | \mathbf{h}_i(x) = z) = \left(1 - \frac{1}{m}\right)^n$$

False positive probability – Events

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Event E_i^c holds iff $\mathbf{h}_i(x) \neq \mathbf{h}_i(x_1)$ and ... and $\mathbf{h}_i(x) \neq \mathbf{h}_i(x_n)$

$$P(E_i^c) = \left(1 - \frac{1}{m}\right)^n$$


$$\text{FPR} = \prod_{i=1}^k (1 - P(E_i^c)) = \left(1 - \left(1 - \frac{1}{m}\right)^n\right)^k$$

False Positivity Rate – Example

$$\text{FPR} = \left(1 - \left(1 - \frac{1}{m} \right)^n \right)^k$$

e.g., $n = 5,000,000$

$k = 30$

$m = 2,500,000$



FPR = 1.28%

Comparison with Hash Tables - Space

- Google storing 5 million URLs, each URL 40 bytes.
- Bloom filter with $k = 30$ and $m = 2,500,000$

Hash Table

(optimistic)

$$5,000,000 \times 40B = 200MB$$

Bloom Filter

$$2,500,000 \times 30 = 75,000,000 \text{ bits}$$

$$< 10 \text{ MB}$$

Time

- Say avg user visits **102,000** URLs in a year, of which **2,000** are malicious.
- **0.5** seconds to do lookup in the database, **1ms** for lookup in Bloom filter.
- Suppose the false positive rate is **3%**

$$1\text{ms} + \frac{100000 \times 0.03 \times 500\text{ms} + 2000 \times 500\text{ms}}{102000} \approx 25.51\text{ms}$$

1ms ↑ Bloom filter lookup

↑ false positives

↑ 0.5 seconds DB lookup

↑ malicious URLs

total URLs

Bloom Filters typical of...

... randomized algorithms and randomized data structures.

- **Simple**
- **Fast**
- **Efficient**
- **Elegant**
- **Useful!**

Agenda

- Wrap-up of Bloom Filters
- Continuous Random Variables ◀
- Probability Density Function
- Cumulative Distribution Function
- Expectation and Variance of continuous r.v.

Often we want to model experiments where the outcome is not discrete.

Hope you enjoyed the zoo!



$X \sim \text{Unif}(a, b)$

$$P(X = k) = \frac{1}{b - a + 1}$$
$$\mathbb{E}[X] = \frac{a + b}{2}$$
$$\text{Var}(X) = \frac{(b - a)(b - a + 2)}{12}$$

$X \sim \text{Ber}(p)$

$$P(X = 1) = p, P(X = 0) = 1 - p$$
$$\mathbb{E}[X] = p$$
$$\text{Var}(X) = p(1 - p)$$

$X \sim \text{Bin}(n, p)$

$$P(X = k) = \binom{n}{k} p^k (1 - p)^{n - k}$$
$$\mathbb{E}[X] = np$$
$$\text{Var}(X) = np(1 - p)$$

$X \sim \text{Geo}(p)$

$$P(X = k) = (1 - p)^{k - 1} p$$
$$\mathbb{E}[X] = \frac{1}{p}$$
$$\text{Var}(X) = \frac{1 - p}{p^2}$$

$X \sim \text{NegBin}(r, p)$

$$P(X = k) = \binom{k - 1}{r - 1} p^r (1 - p)^{k - r}$$
$$\mathbb{E}[X] = \frac{r}{p}$$
$$\text{Var}(X) = \frac{r(1 - p)}{p^2}$$

$X \sim \text{Poisson}(\lambda)$

$$P(X = i) = e^{-\lambda} \cdot \frac{\lambda^i}{i!}$$

$$\mathbb{E}[X] = \lambda$$

$$\text{Var}(X) = \lambda$$

$X \sim \text{HypGeo}(N, K, n)$

$$P(X = k) = \frac{\binom{K}{k} \binom{N - K}{n - k}}{\binom{N}{n}}$$
$$\mathbb{E}[X] = \frac{n(K)}{N}$$
$$\text{Var}(X) = \frac{n(N - K)(N - n)}{N^2(N - 1)}$$

Agenda

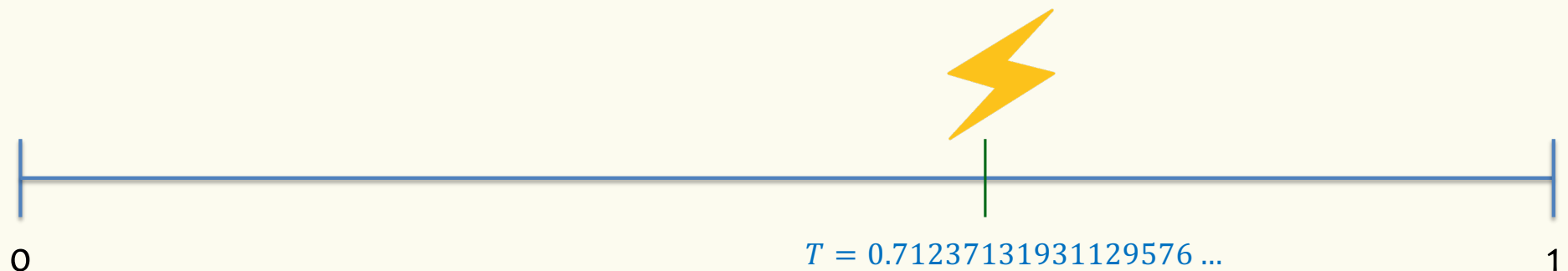
- Wrap-up of Bloom Filters
- Continuous Random Variables ◀
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- Expectation and Variance of continuous r.v.

Often we want to model experiments where the outcome is not discrete.

Example – Lightning Strike

Lightning strikes a pole within a one-minute time frame

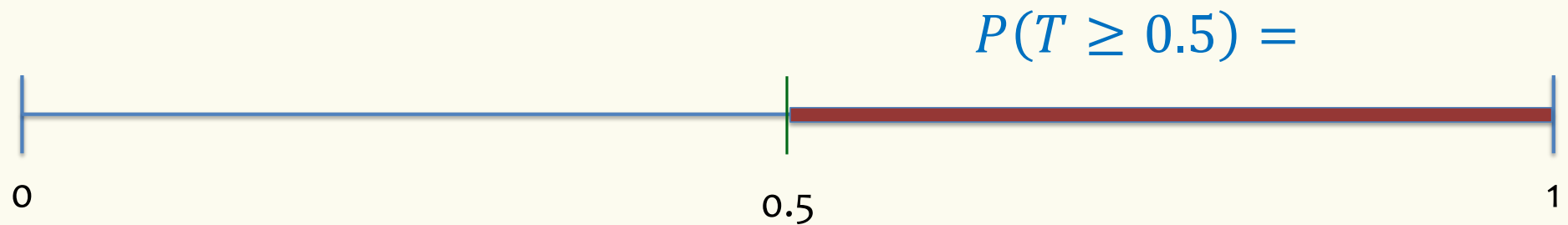
- T = time of lightning strike
- Every time within $[0,1]$ is equally likely
 - Time measured with infinitesimal precision.



The outcome space is not discrete

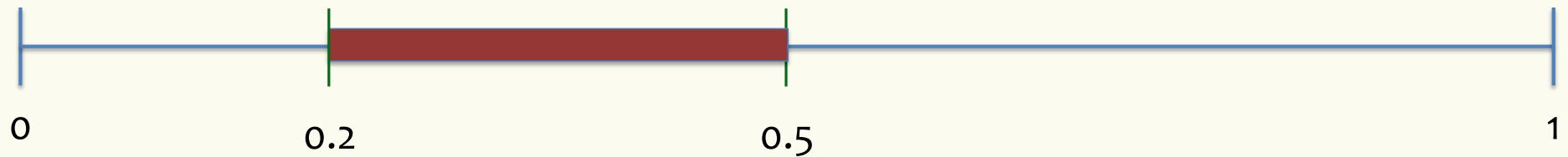
Lightning strikes a pole within a one-minute time frame

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Lightning strikes a pole within a one-minute time frame

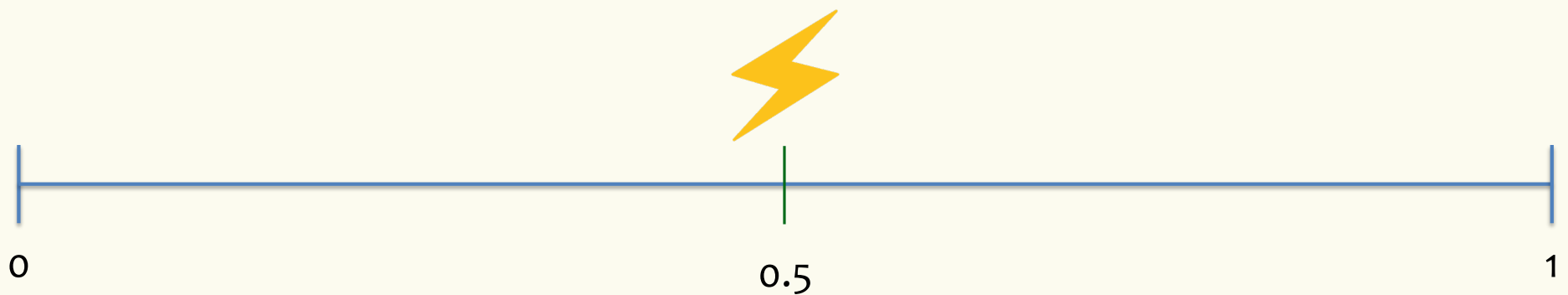
- T = time of lightning strike
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$$P(0.2 \leq T \leq 0.5) =$$

Lightning strikes a pole within a one-minute time frame

- T = time of lightning strike
- Every point in time within $[0,1]$ is equally likely



$$P(T = 0.5) =$$

Bottom line

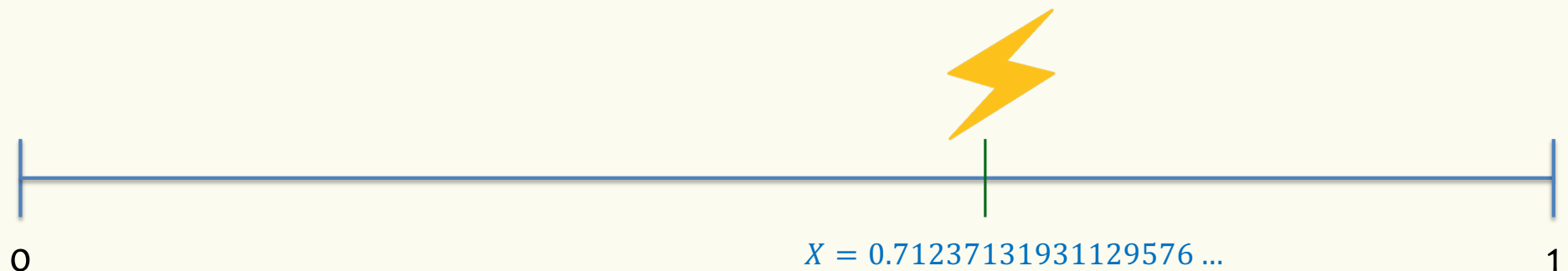
- This gives rise to a different type of random variable
- $P(T = x) = 0$ for all $x \in [0,1]$
- Yet, somehow we want
 - $P(T \in [0,1]) = 1$
 - $P(T \in [a, b]) = b - a$
 - ...
- How do we model the behavior of T ?

First try: A discrete approximation

Example – Lightning Strike

Lightning strikes a pole within a one-minute time frame

- X = time of lightning strike
- Every time within $[0,1]$ is equally likely
 - Time measured with infinitesimal precision.

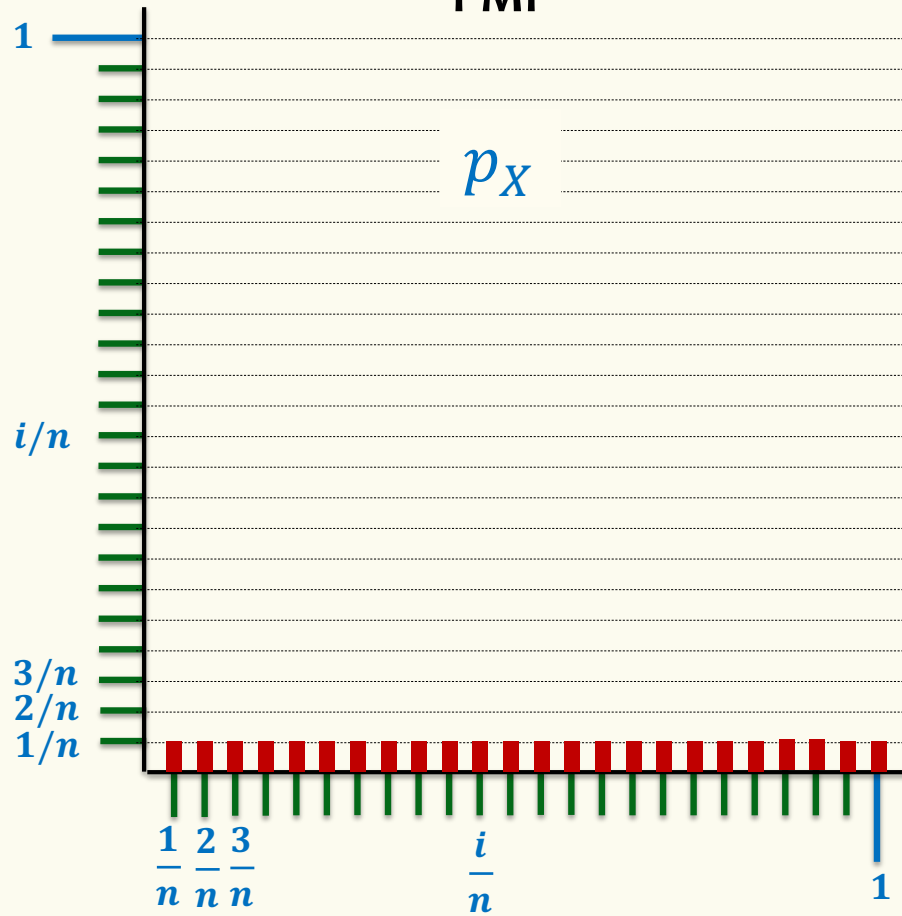


Discrete approximation?

A Discrete Approximation

Probability Mass Function

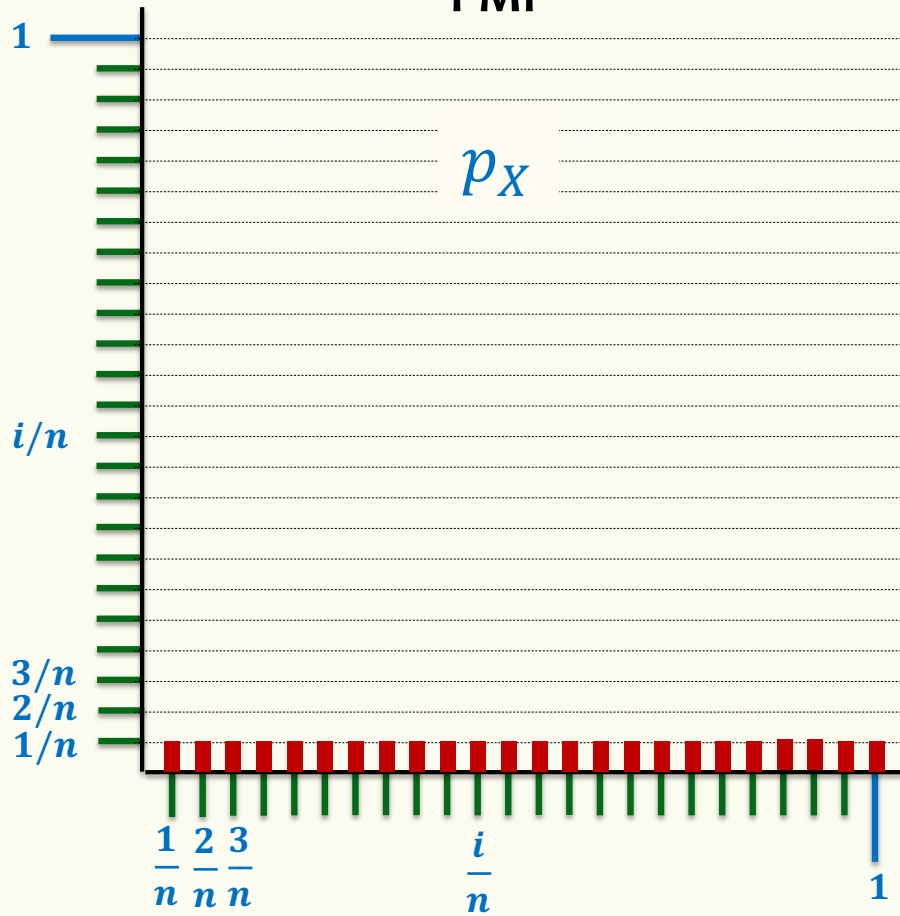
PMF



A Discrete Approximation

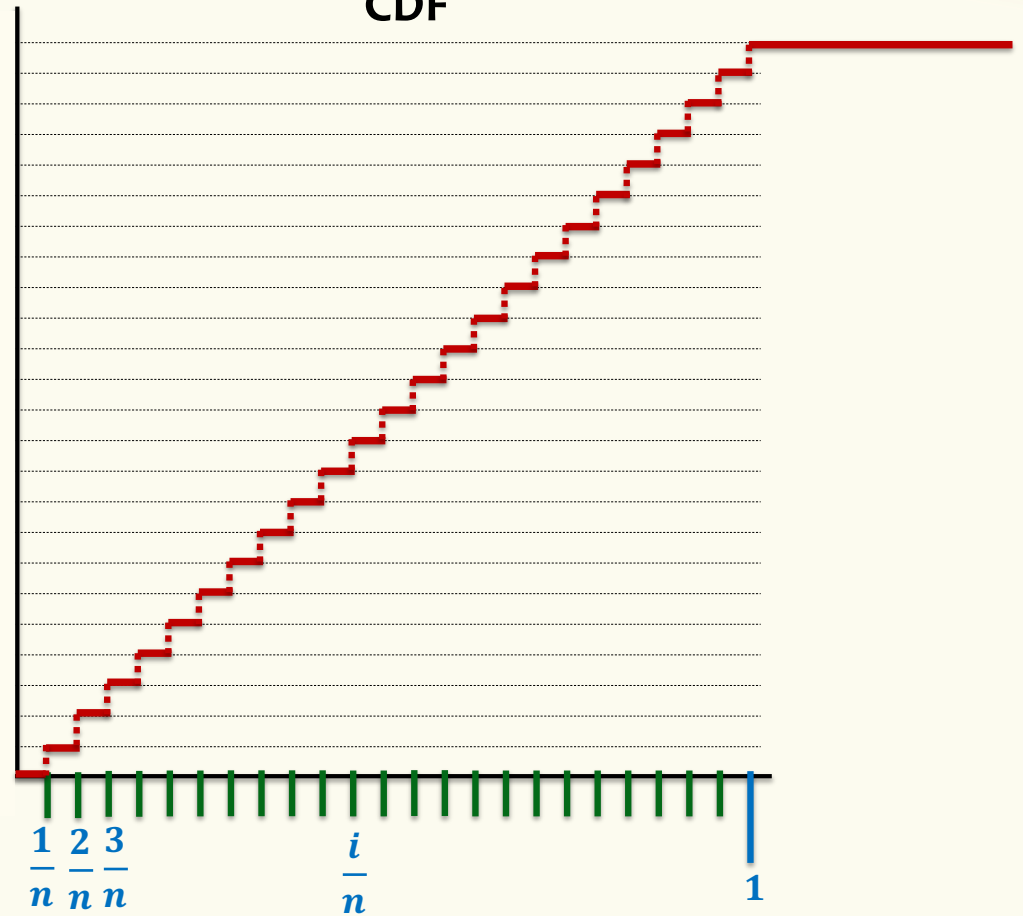
Probability Mass Function

PMF

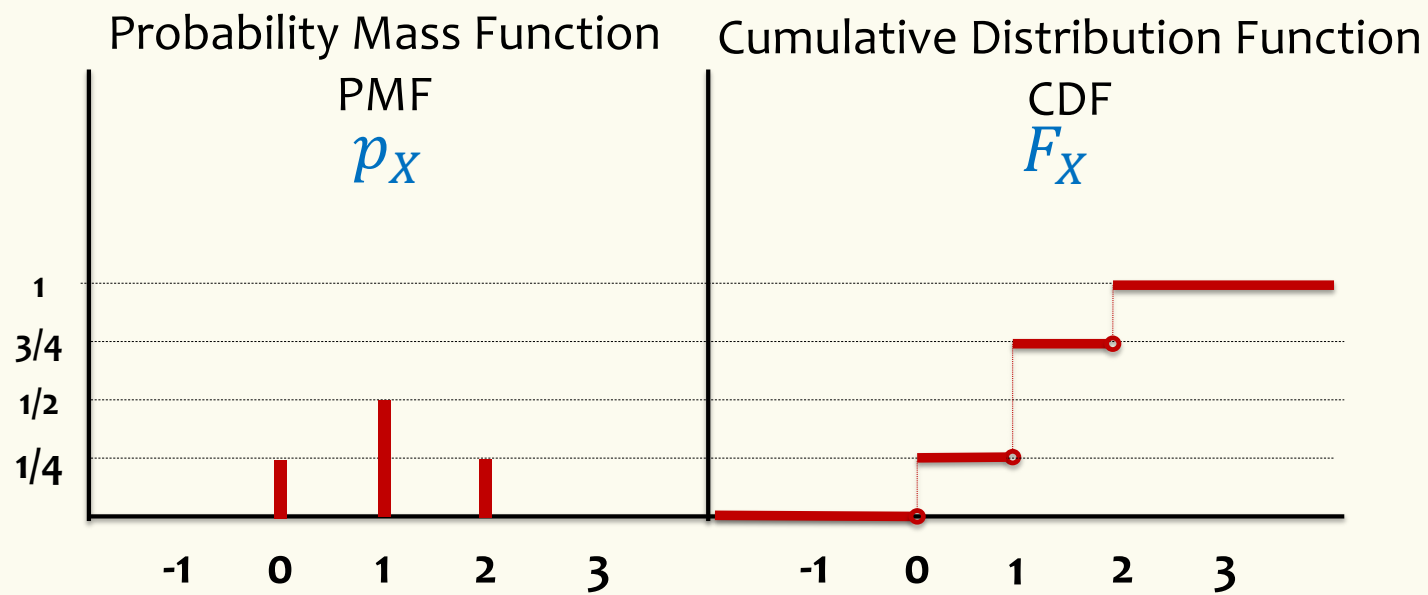


Cumulative Distribution Function

CDF

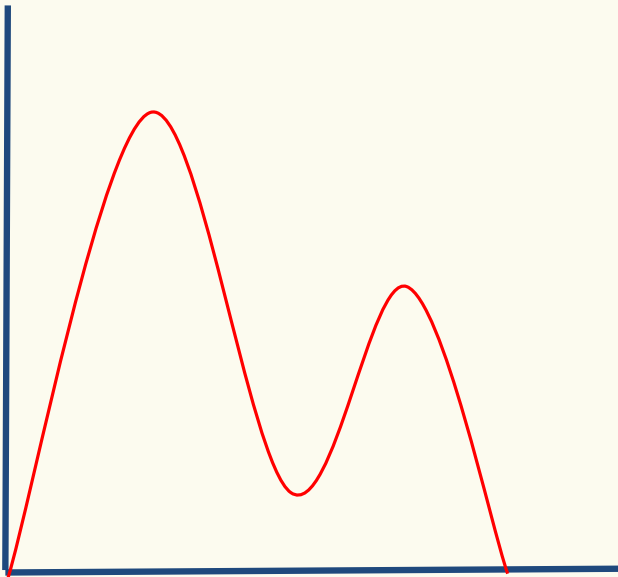


Recall: Cumulative Distribution Function (CDF)

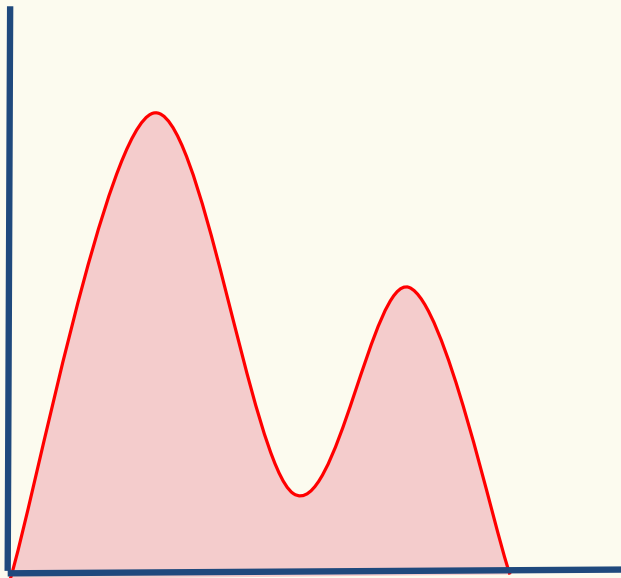


Definition. A **continuous random variable** X is defined by a **probability density function** (PDF) $f_X: \mathbb{R} \rightarrow \mathbb{R}$, such that

Non-negativity: $f_X(x) \geq 0$ for all $x \in \mathbb{R}$



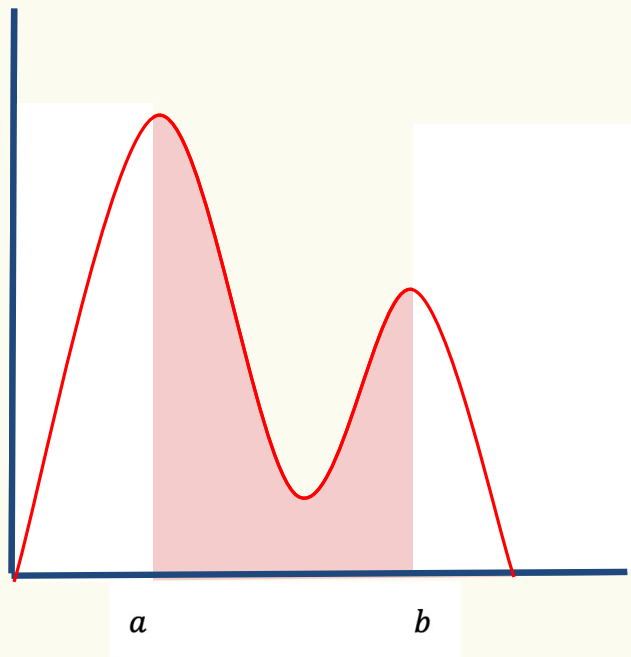
Probability Density Function - Intuition



Non-negativity: $f_X(x) \geq 0$ for all $x \in \mathbb{R}$

Normalization: $\int_{-\infty}^{+\infty} f_X(x) dx = 1$

Probability Density Function - Intuition

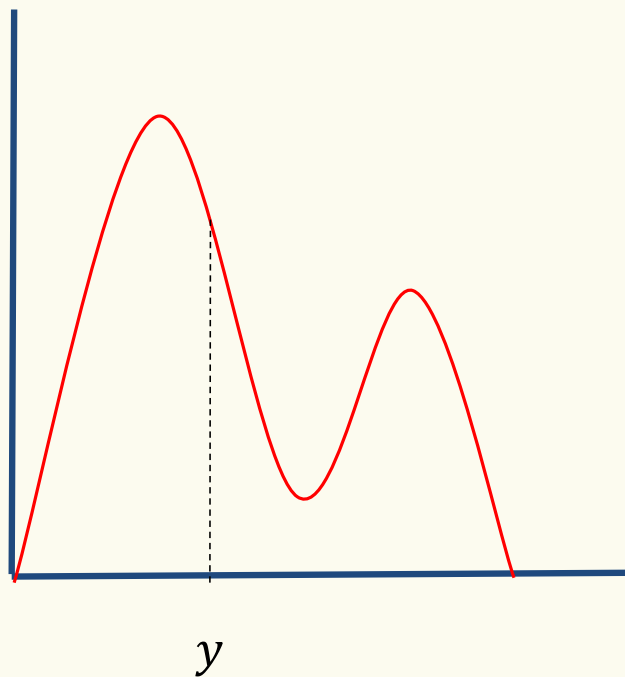


Non-negativity: $f_X(x) \geq 0$ for all $x \in \mathbb{R}$

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$$F(b) - F(a) = P(a \leq X \leq b) = \int_a^b f_X(x) dx$$

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Density \neq Probability

$$f_X(y) \neq 0 \quad P(X = y) = 0$$

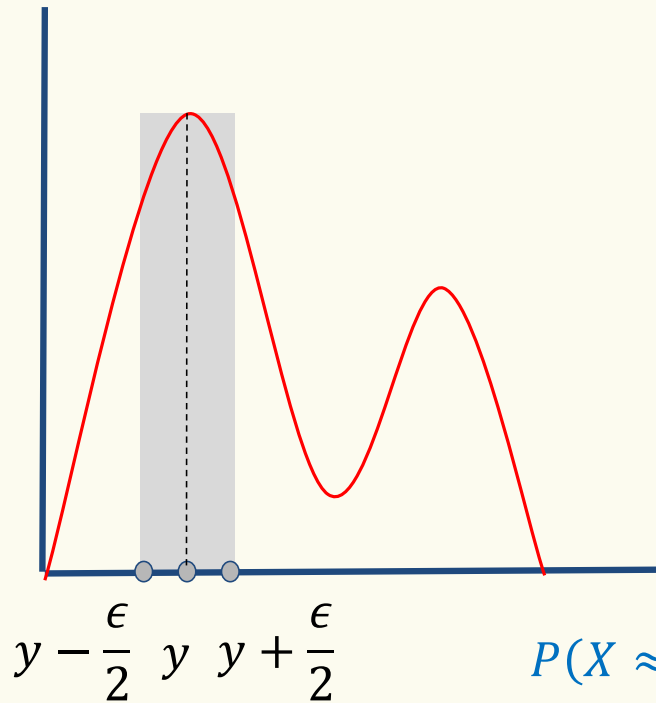
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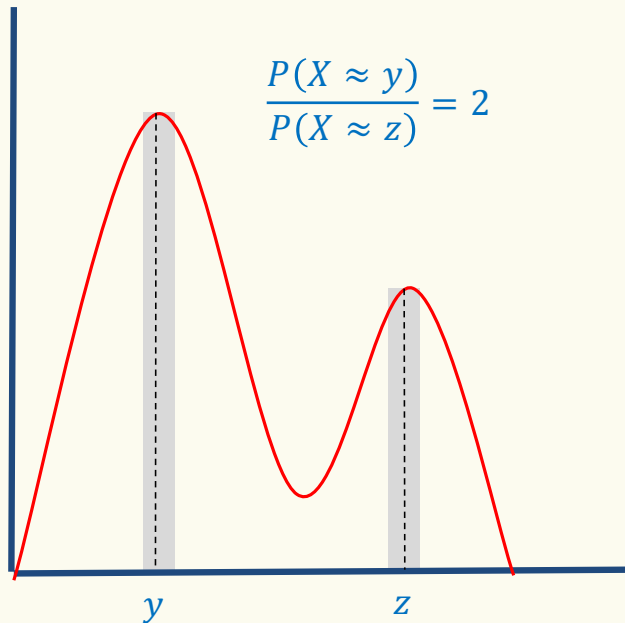
$$P(X = y) = P(y \leq X \leq y) = \int_y^y f_X(x) dx = 0$$



$$P(X \approx y) \approx P\left(y - \frac{\epsilon}{2} \leq X \leq y + \frac{\epsilon}{2}\right) = \int_{y - \frac{\epsilon}{2}}^{y + \frac{\epsilon}{2}} f_X(x) dx \approx \epsilon f_X(y)$$

What $f_X(x)$ measures: The local **rate** at which probability accumulates

Probability Density Function - Intuition



Non-negativity: $f_X(x) \geq 0$ for all $x \in \mathbb{R}$

Normalization: $\int_{-\infty}^{+\infty} f_X(x) dx = 1$

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$$\frac{P(X \approx y)}{P(X \approx z)} \approx \frac{\epsilon f_X(y)}{\epsilon f_X(z)} = \frac{f_X(y)}{f_X(z)} \quad 61$$

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Cumulative Distribution Function

Definition. The **cumulative distribution function (cdf)** of X is

$$F_X(a) = P(X \leq a) = \int_{-\infty}^a f_X(x) dx$$

By the fundamental theorem of Calculus $f_X(x) = \frac{d}{dx} F_X(x)$

From Discrete to Continuous

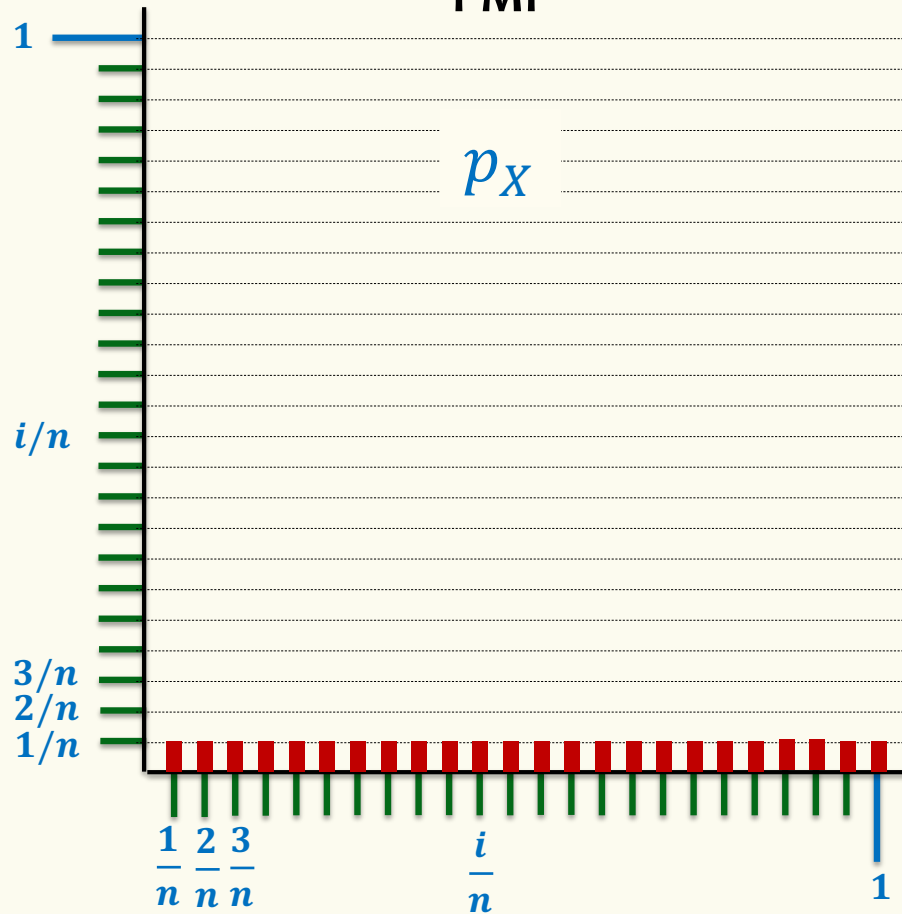
| | Discrete | Continuous |
|----------------------|-----------------------------------|---|
| PMF/PDF | $p_X(x) = P(X = x)$ | $f_X(x) \neq P(X = x) = 0$ |
| CDF | $F_X(x) = \sum_{t \leq x} p_X(t)$ | $F_X(x) = \int_{-\infty}^x f_X(t) dt$ |
| Normalization | $\sum_x p_X(x) = 1$ | $\int_{-\infty}^{\infty} f_X(x) dx = 1$ |



A Discrete Approximation

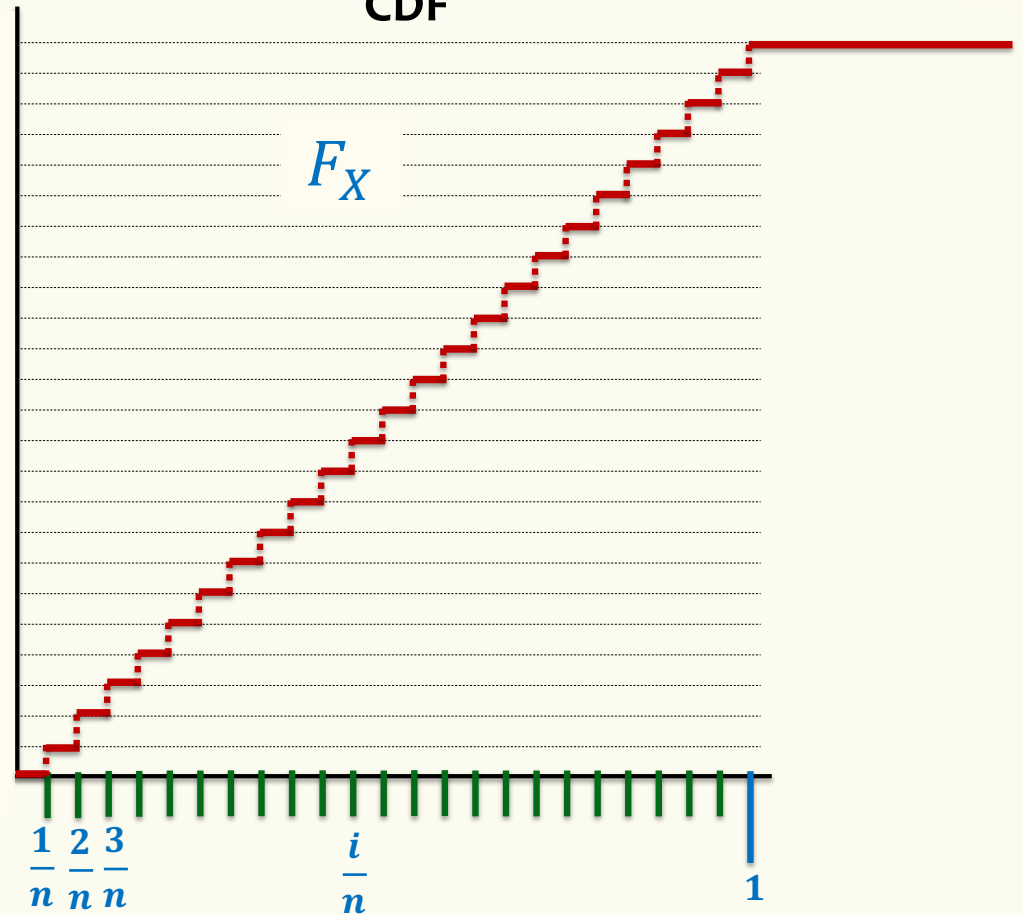
Probability Mass Function

PMF



Cumulative Distribution Function

CDF

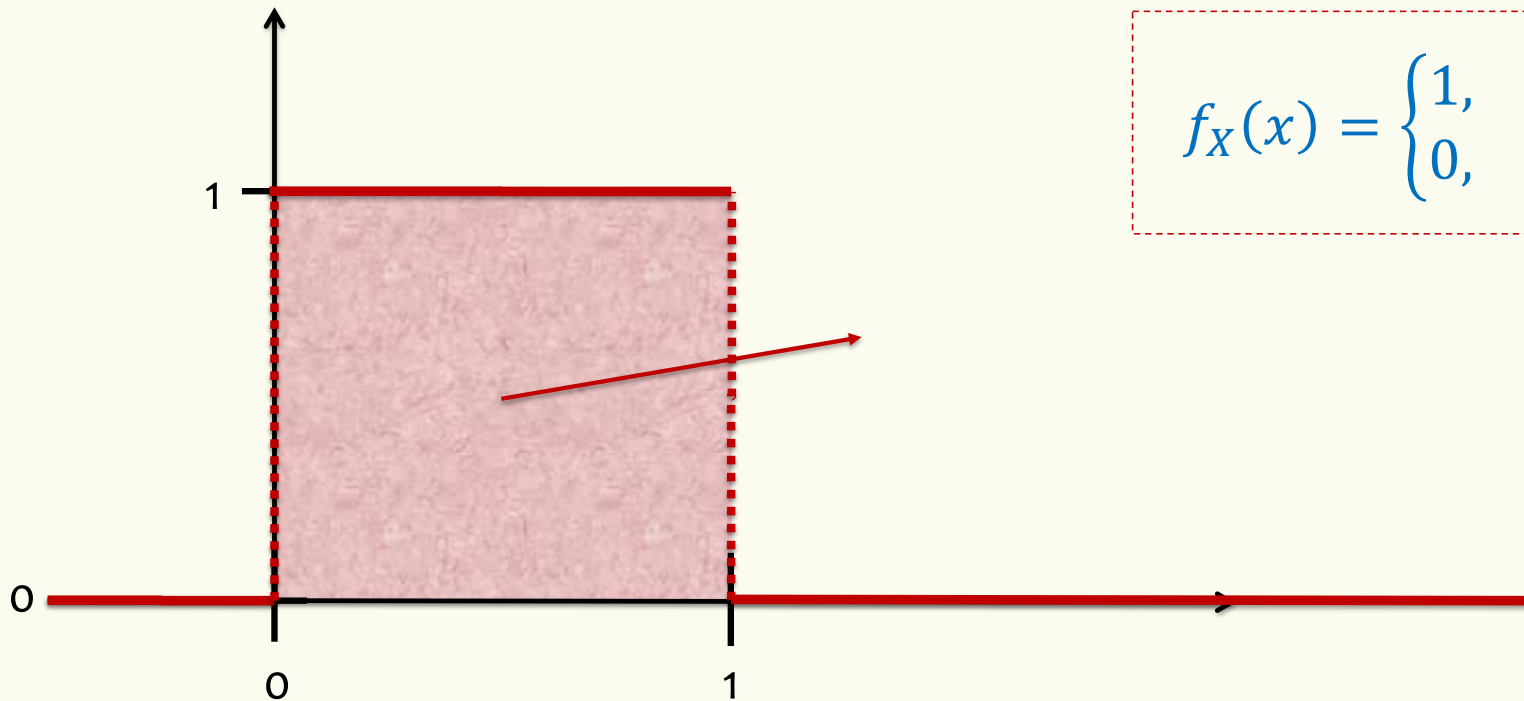


PDF of Uniform RV

$X \sim \text{Unif}(0,1)$

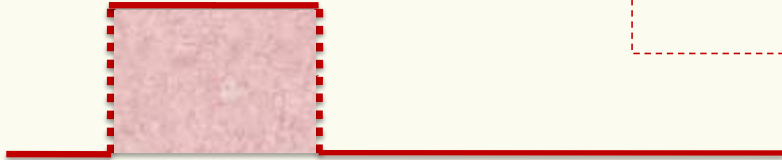
$$F_X(x) = P(X \leq x) = \begin{cases} 0 & x \leq 0 \\ x & 0 \leq x \leq 1 \\ 1 & 1 \leq x \end{cases}$$

$$f_X(x) = \begin{cases} 1, & x \in [0,1] \\ 0, & x \notin [0,1] \end{cases}$$



$X \sim \text{Unif}(0,1)$

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Normalization: $\int_{-\infty}^{+\infty} f_X(x) dx = 1$

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$$\frac{P(X \approx y)}{P(X \approx z)} \approx \frac{\epsilon f_X(y)}{\epsilon f_X(z)} = \frac{f_X(y)}{f_X(z)}$$

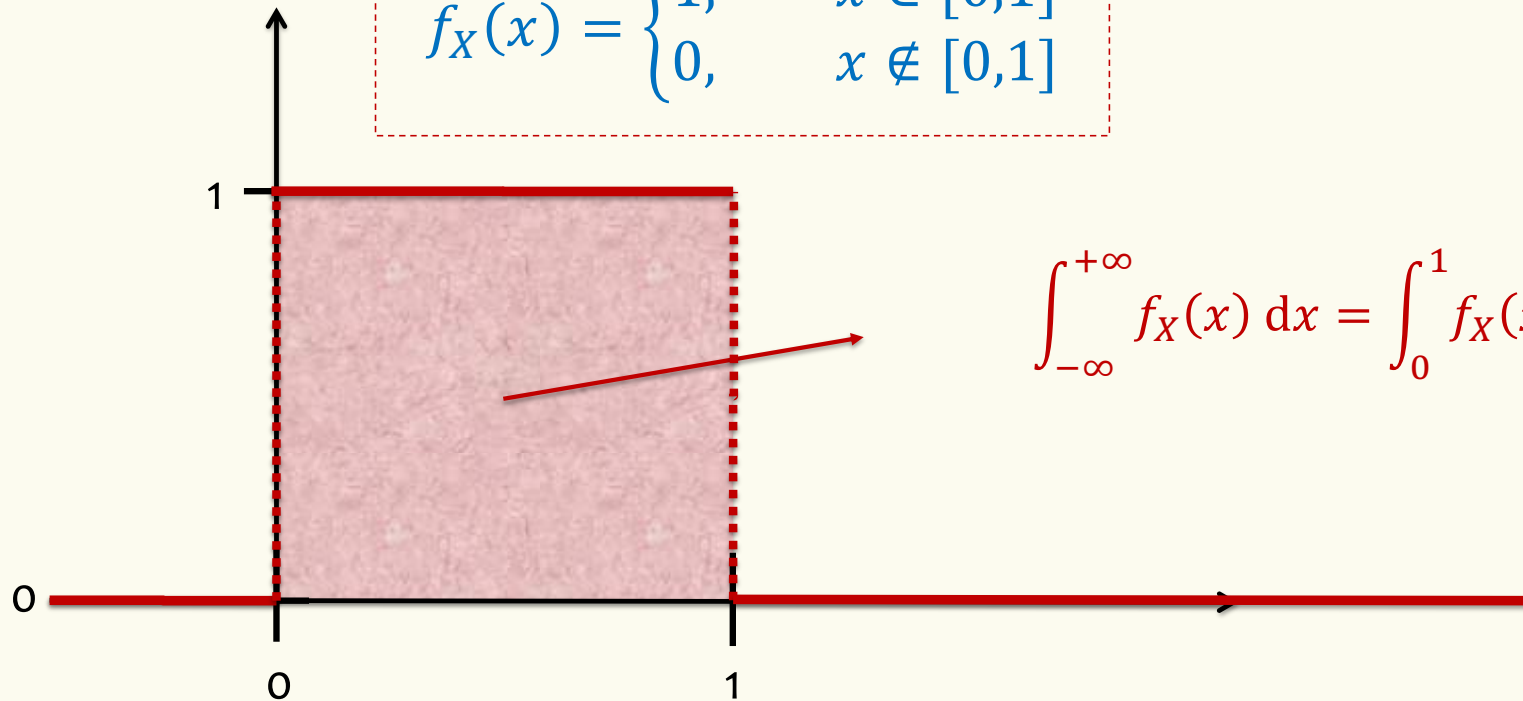
PDF of Uniform RV

$$X \sim \text{Unif}(0,1)$$

Non-negativity: $f_X(x) \geq 0$ for all $x \in \mathbb{R}$

Normalization: $\int_{-\infty}^{+\infty} f_X(x) dx = 1$

$$f_X(x) = \begin{cases} 1, & x \in [0,1] \\ 0, & x \notin [0,1] \end{cases}$$



$$\int_{-\infty}^{+\infty} f_X(x) dx = \int_0^1 f_X(x) dx = 1 \cdot 1 = 1$$

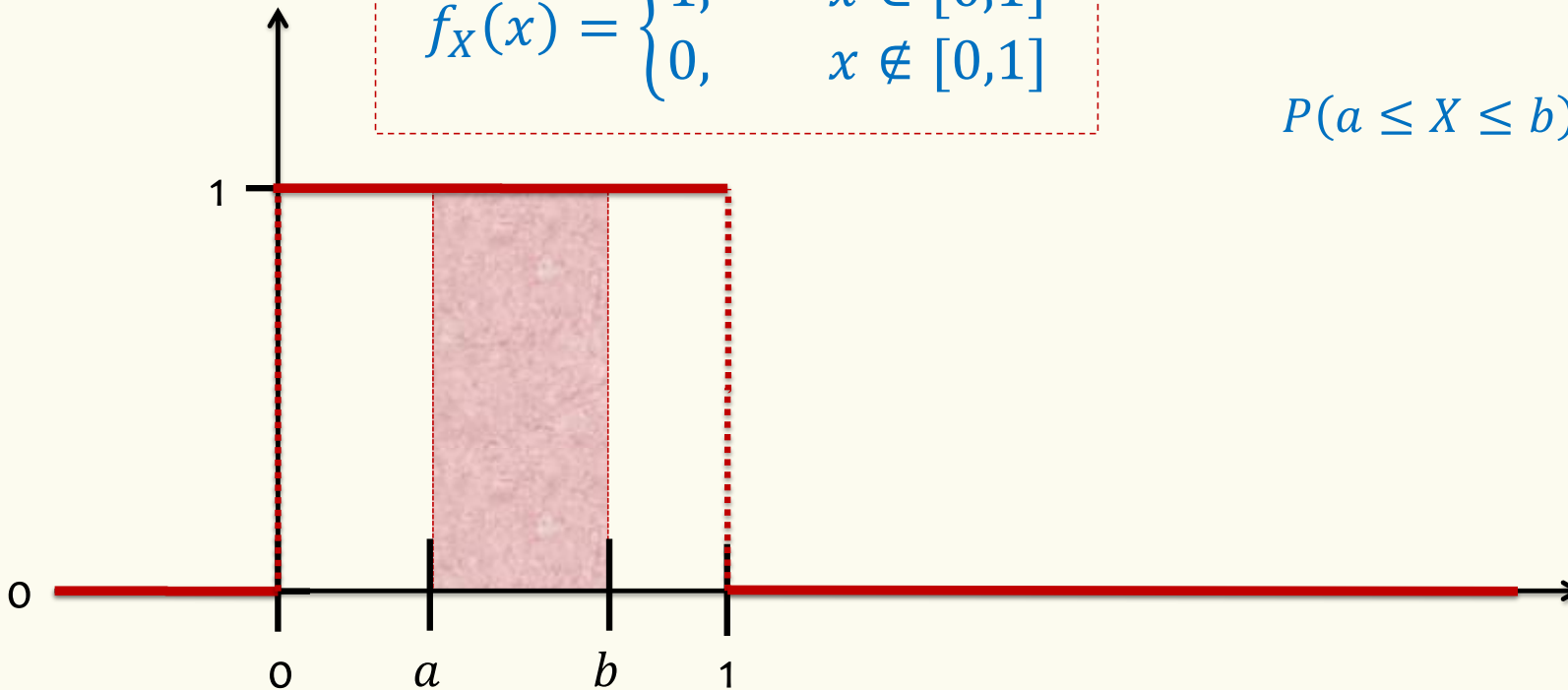
Probability of Event

$X \sim \text{Unif}(0,1)$

Non-negativity: $f_X(x) \geq 0$ for all $x \in \mathbb{R}$

Normalization: $\int_{-\infty}^{+\infty} f_X(x) dx = 1$

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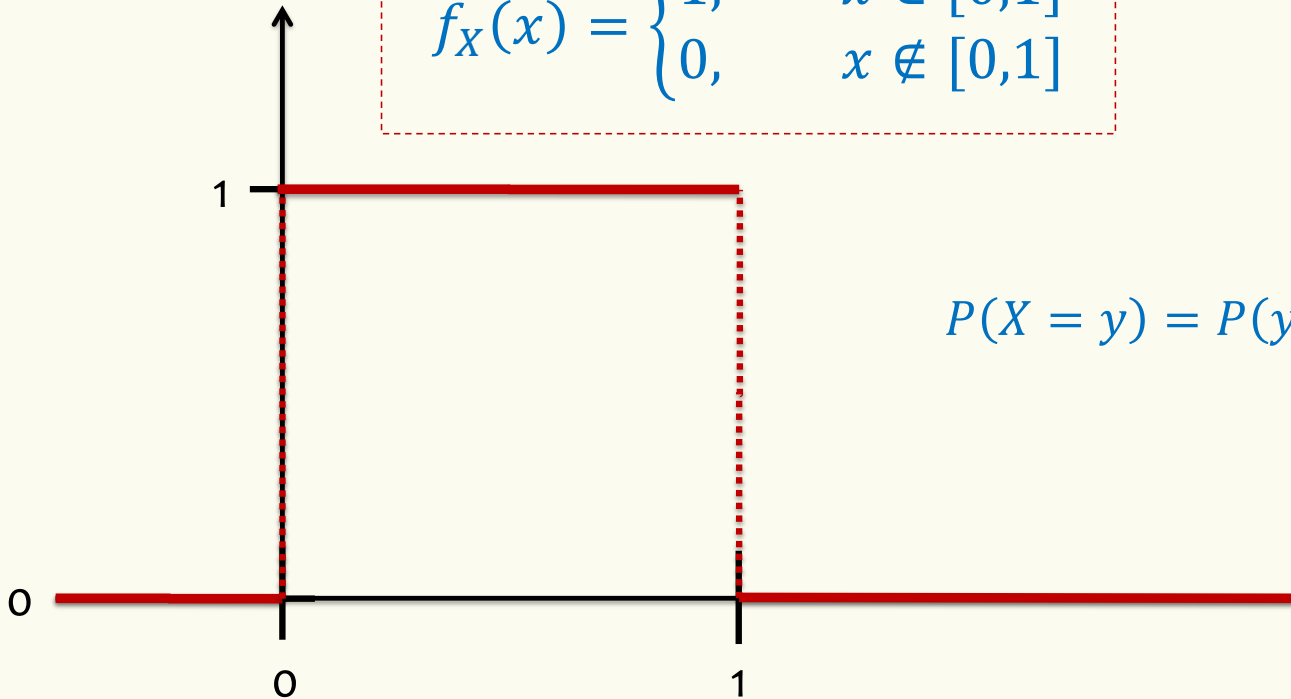


$$P(a \leq X \leq b) = \int_a^b f_X(x) dx$$

Probability of Event

$X \sim \text{Unif}(0,1)$

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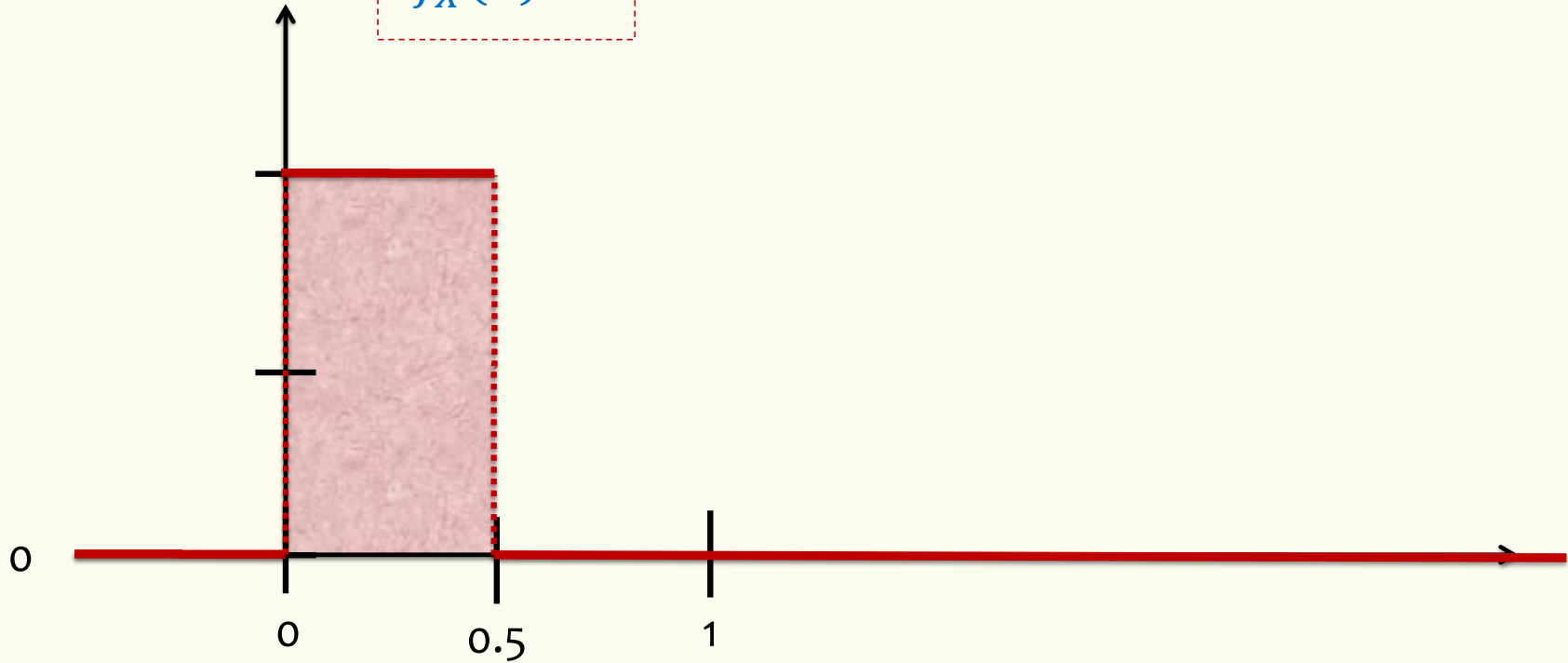
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$$\frac{P(X \approx y)}{P(X \approx z)} \approx \frac{\epsilon f_X(y)}{\epsilon f_X(z)} = \frac{f_X(y)}{f_X(z)}$$

PDF of Uniform RV

$X \sim \text{Unif}(0,0.5)$

$$f_X(x) =$$



PDF of Uniform RV

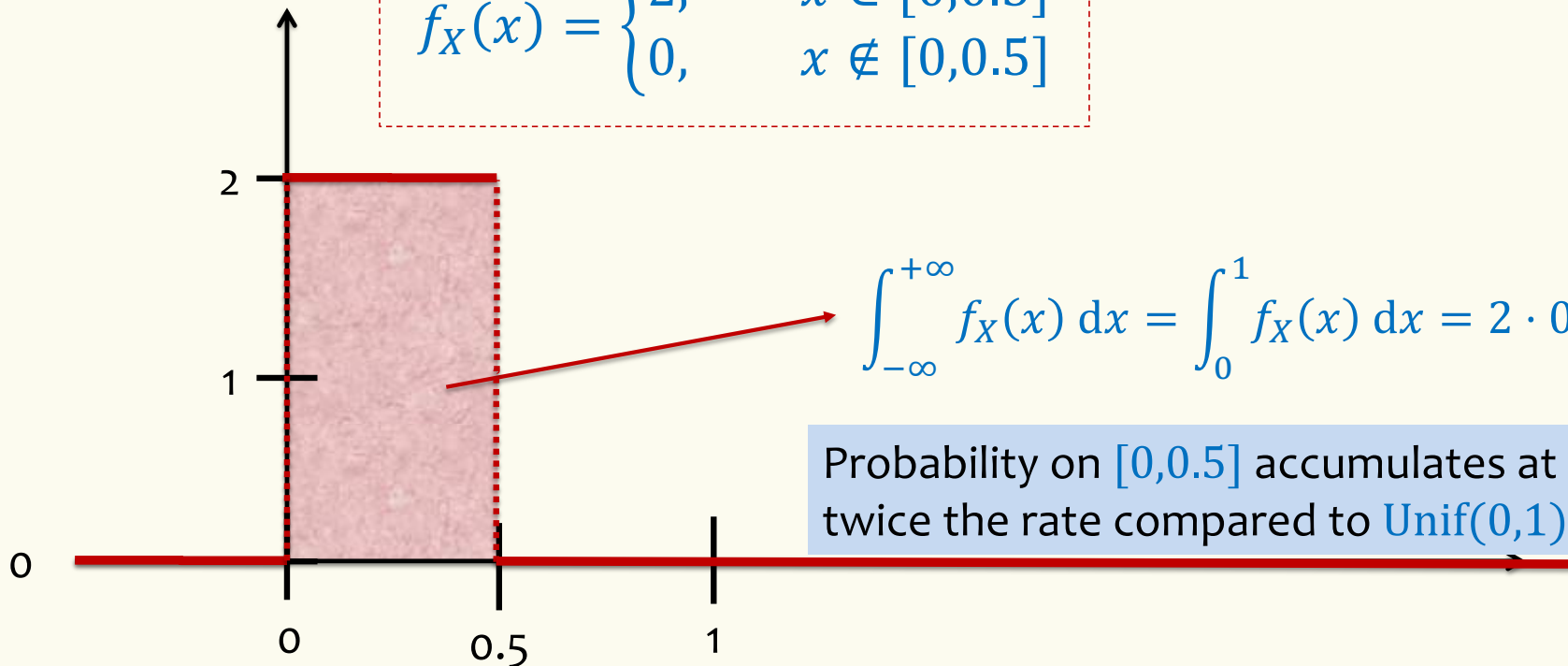
$$X \sim \text{Unif}(0,0.5)$$



Density \neq Probability

$f_X(x) \gg 1$ is possible!

$$f_X(x) = \begin{cases} 2, & x \in [0,0.5] \\ 0, & x \notin [0,0.5] \end{cases}$$

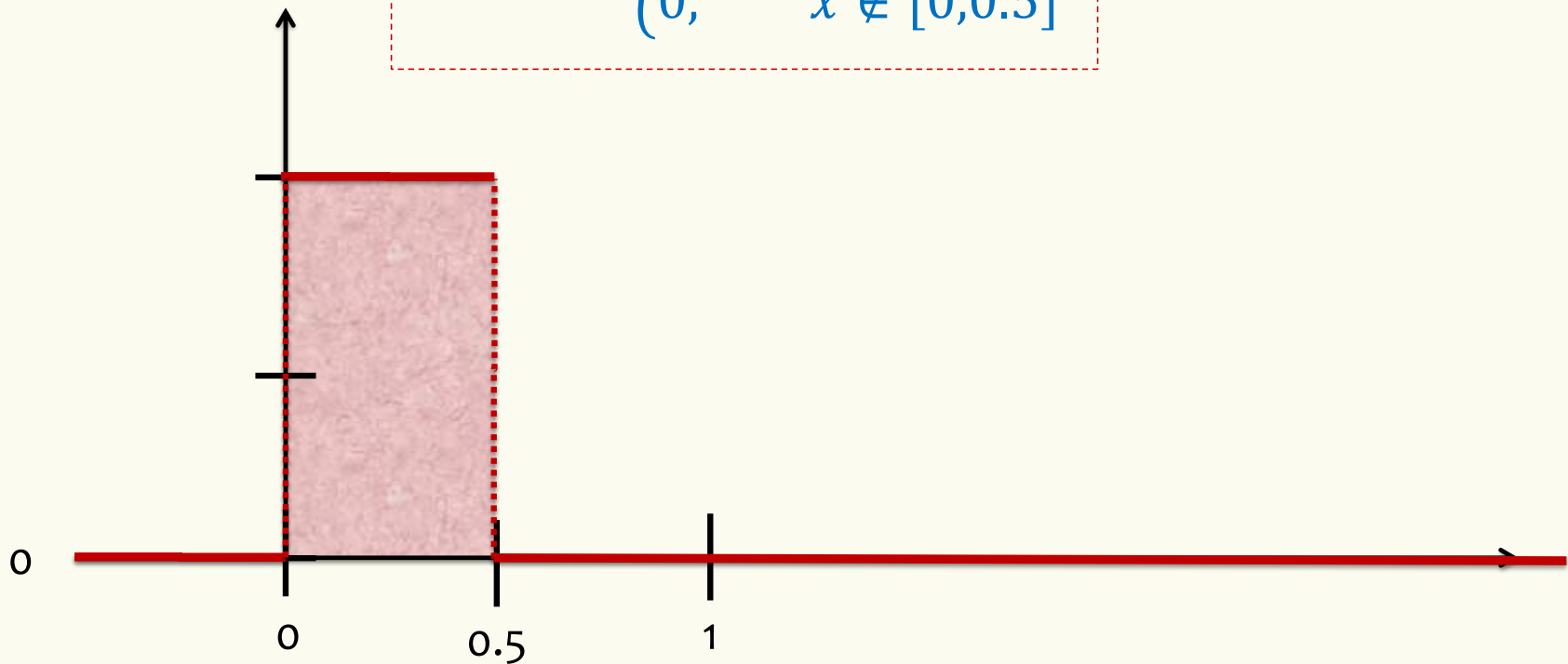


Probability on $[0,0.5]$ accumulates at twice the rate compared to $\text{Unif}(0,1)$

PDF of Uniform RV

$X \sim \text{Unif}(0,0.5)$

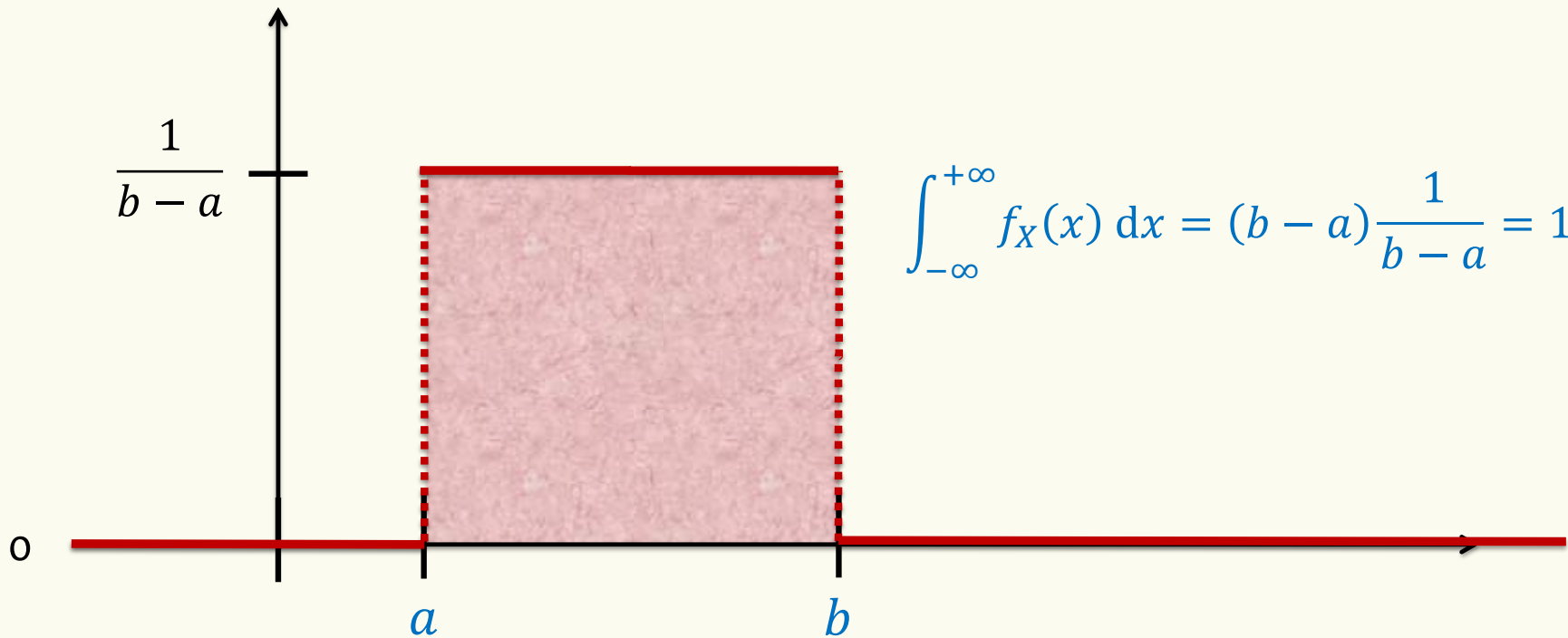
$$f_X(x) = \begin{cases} 2, & x \in [0,0.5] \\ 0, & x \notin [0,0.5] \end{cases}$$



Uniform Distribution

$X \sim \text{Unif}(a, b)$

$$f_X(x) = \begin{cases} \frac{1}{b-a} & x \in [a, b] \\ 0 & \text{else} \end{cases}$$



$$\int_{-\infty}^{+\infty} f_X(x) dx = (b-a) \frac{1}{b-a} = 1$$

Cumulative Distribution Function

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By the fundamental theorem of Calculus $f_X(x) = \frac{d}{dx} F_X(x)$

From Discrete to Continuous

| | Discrete | Continuous |
|----------------------|-----------------------------------|---|
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Cumulative Distribution Function

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$$F_X(a) = P(X \leq a) = \int_{-\infty}^a f_X(x) dx$$


By the fundamental theorem of Calculus $f_X(x) = \frac{d}{dx} F_X(x)$

Therefore: $P(X \in [a, b]) = F_X(b) - F_X(a)$

F_X is monotone increasing, since $f_X(x) \geq 0$. That is $F_X(c) \leq F_X(d)$ for $c \leq d$

$$\lim_{a \rightarrow -\infty} F_X(a) = P(X \leq -\infty) = 0 \quad \lim_{a \rightarrow +\infty} F_X(a) = P(X \leq +\infty) = 1$$

Agenda

- Wrap-up of Poisson RVs
- Continuous Random Variables
- Probability Density Function
- Cumulative Distribution Function
- Expectation and Variance of continuous r.v. 

Expectation of a Continuous RV

Definition. The **expected value** of a continuous RV X is defined as

$$\mathbb{E}[X] = \int_{-\infty}^{+\infty} f_X(x) \cdot x \, dx$$

Fact. $\mathbb{E}[aX + bY + c] = a\mathbb{E}[X] + b\mathbb{E}[Y] + c$

← Proof follows same ideas as discrete case

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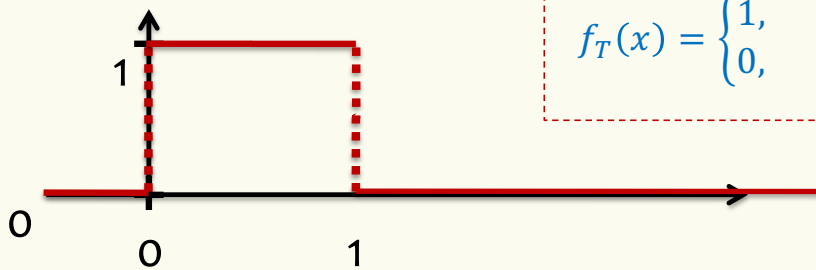
Proofs follow same ideas as discrete case

Definition. The **variance** of a continuous RV X is defined as

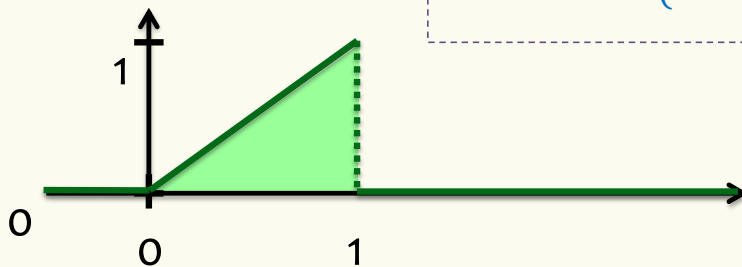
$$\text{Var}(X) = \int_{-\infty}^{+\infty} f_X(x) \cdot (x - \mathbb{E}[X])^2 \, dx = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

Expectation of a Continuous RV

Example. $T \sim \text{Unif}(0,1)$



$$f_T(x) = \begin{cases} 1, & x \in [0,1] \\ 0, & x \notin [0,1] \end{cases}$$



$$f_T(x) \cdot x = \begin{cases} x, & x \in [0,1] \\ 0, & x \notin [0,1] \end{cases}$$

Definition.

$$\mathbb{E}[X] = \int_{-\infty}^{+\infty} f_X(x) \cdot x \, dx$$

$$\mathbb{E}[T] = \underbrace{\frac{1}{2} 1^2}_{\text{Area of triangle}} = \frac{1}{2}$$

Uniform Density – Expectation

$$X \sim \text{Unif}(a, b)$$

$$f_X(x) = \begin{cases} \frac{1}{b-a} & x \in [a, b] \\ 0 & \text{else} \end{cases}$$

$$\mathbb{E}[X] = \int_{-\infty}^{+\infty} f_X(x) \cdot x \, dx$$

$$\begin{aligned} &= \frac{1}{b-a} \int_a^b x \, dx = \frac{1}{b-a} \left(\frac{x^2}{2} \right) \Big|_a^b = \frac{1}{b-a} \left(\frac{b^2 - a^2}{2} \right) \\ &= \frac{(b-a)(a+b)}{2(b-a)} = \frac{a+b}{2} \end{aligned}$$

Uniform Density – Variance

$$X \sim \text{Unif}(a, b)$$

$$f_X(x) = \begin{cases} \frac{1}{b-a} & x \in [a, b] \\ 0 & \text{else} \end{cases}$$

$$\mathbb{E}[X^2] = \int_{-\infty}^{+\infty} f_X(x) \cdot x^2 \, dx$$

$$= \frac{1}{b-a} \int_a^b x^2 \, dx = \frac{1}{b-a} \left(\frac{x^3}{3} \right) \Big|_a^b = \frac{b^3 - a^3}{3(b-a)}$$

$$= \frac{(b-a)(b^2 + ab + a^2)}{3(b-a)} = \frac{b^2 + ab + a^2}{3}$$

Uniform Density – Variance

$$\mathbb{E}[X^2] = \frac{b^2 + ab + a^2}{3}$$

$$\mathbb{E}[X] = \frac{a + b}{2}$$

$$X \sim \text{Unif}(a, b)$$

$$\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

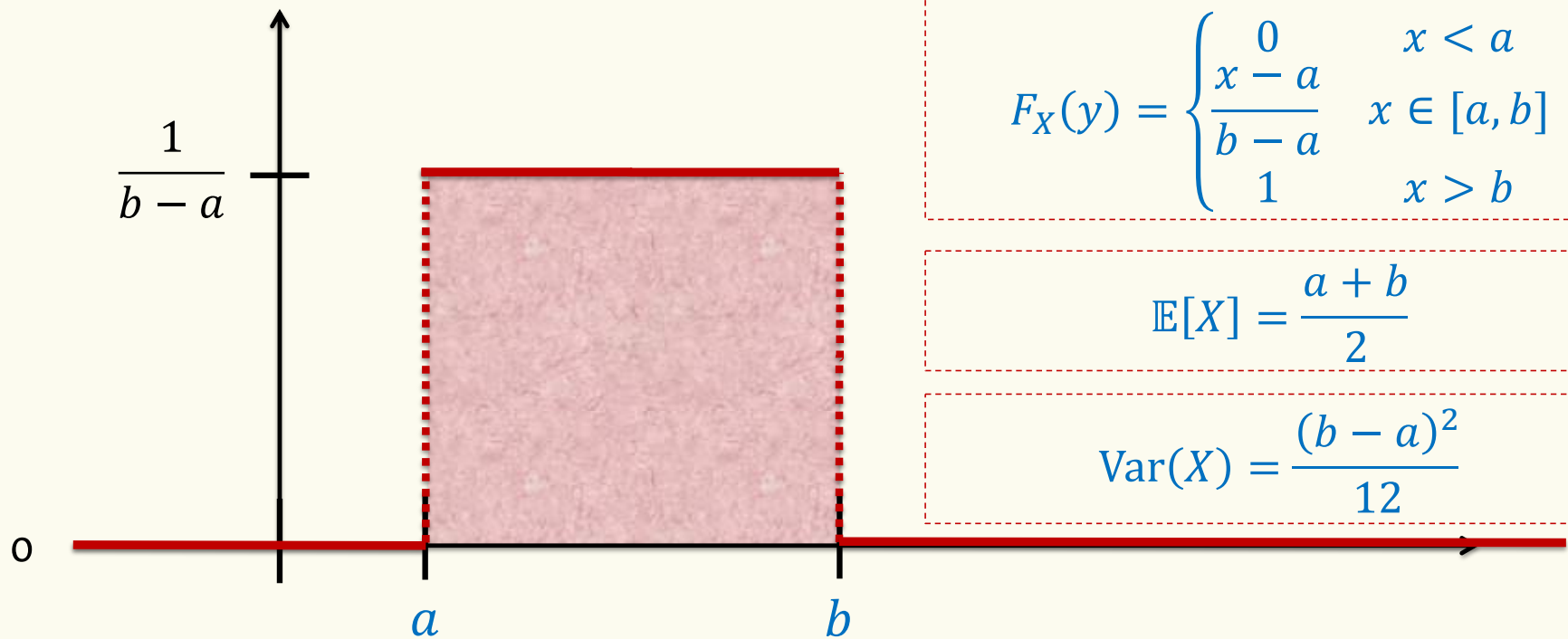
$$= \frac{b^2 + ab + a^2}{3} - \frac{a^2 + 2ab + b^2}{4}$$

$$= \frac{4b^2 + 4ab + 4a^2}{12} - \frac{3a^2 + 6ab + 3b^2}{12}$$

$$= \frac{b^2 - 2ab + a^2}{12} = \frac{(b - a)^2}{12}$$

Uniform Distribution Summary

$X \sim \text{Unif}(a, b)$



$$f_X(x) = \begin{cases} \frac{1}{b-a} & x \in [a, b] \\ 0 & \text{else} \end{cases}$$

$$F_X(y) = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & x \in [a, b] \\ 1 & x > b \end{cases}$$

$$\mathbb{E}[X] = \frac{a+b}{2}$$

$$\text{Var}(X) = \frac{(b-a)^2}{12}$$