

CSE 312


Foundations of Computing II

Lecture 22: Maximum Likelihood Estimation (MLE)

Announcement

- Lecture on Wed is cancel.
- No lecture on Friday. Happy thanks giving!
- Pset 7 is due on Wed
- Pset 8 is out on Wed, due on next Friday

Agenda

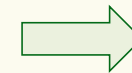
- Chernoff Bound
 - Example: Server Load
 - The Union Bound
- Probability vs statistics 
 - Estimation

Probability vs Statistics

$\text{Ber}(p = 0.5)$



Probability
Given model, predict
data



$P(\underline{THHTHH})$



$\text{Ber}(p = ??)$



Statistics
Given data, predict
model



THHTHH

Recall Formalizing Polls

Population size N , true fraction of voting in favor p , sample size n .

Problem: We don't know p

Polling Procedure

for $i = 1, \dots, n$:

1. Pick uniformly random person to call (prob: $1/N$)
2. Ask them how they will vote

$$X_i = \begin{cases} 1, & \text{voting in favor} \\ 0, & \text{otherwise} \end{cases}$$

Report our estimate of p :

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

What type of r.v. is X_i ?

	$E[X_i]$	$\text{Var}(X_i)$
a. Bernoulli	p	$p(1-p)$

Recall Formalizing Polls

We assume that poll answers $X_1, \dots, X_n \sim \text{Ber}(p)$ i.i.d. for unknown p

Goal: Estimate p

We did this by computing $\hat{p} = \frac{1}{n} \sum_{i=1}^n X_i$

Why is that a good estimate for p ?

More generally ...

In estimation we....

- **Assume:** we know the type of the random variable that we are observing independent samples from
 - We just don't know the parameters, e.g.
 - the bias p of a random coin Bernoulli(p)
 - The arrival rate λ for the Poisson(λ) or Exponential(λ)
 - The mean μ and variance σ of a normal $\mathcal{N}(\mu, \sigma)$
- **Goal:** find the “best” parameters to fit the data

Notation – Parametric Model (discrete case)

Definition. A **(parametric) model** is a family of distributions indexed by a parameter θ , described by a two-argument function

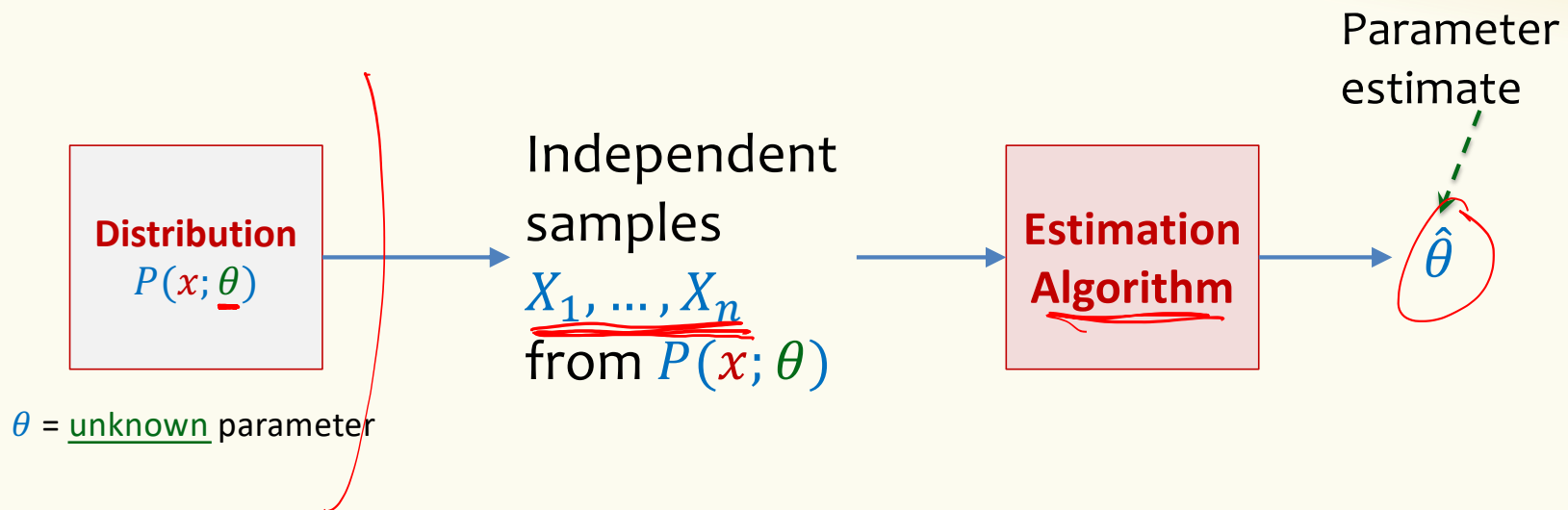
$P(x; \theta)$ = prob. of outcome x when distribution has parameter θ

$P_\theta(x)$ [i.e., every θ defines a different distribution $\sum_x P(x; \theta) = 1$]

Examples

- “Bernoullis”: $P(x; \theta = p) = \begin{cases} p & \underline{x = 1} \\ 1 - p & \underline{x = 0} \end{cases}$
- “Geometrics”: $P(i; \theta = p) = \underline{(1 - p)^{i-1} p}$ for $i \in \mathbb{N}$

Statistics: Parameter Estimation – Workflow



Example: coin flip distribution with unknown $\theta =$ probability of heads

Observation: HTTHHHTHTTTTHTTTTHT

Goal: Estimate θ

Example

Suppose we have a mystery coin with some probability p of coming up heads. We flip the coin 8 times, independent of other flips, and see the following sequence flips

TTHTHTTH

Given this data, what would you estimate p is?

rachel.312
Poll: pollev.com/paulbeame028

- a. $1/2$
- b. $5/8$
- c. $3/8$
- d. $1/4$

Agenda

- Idea: Estimation
- Maximum Likelihood Estimation (example: mystery coin) ◀
- Continuous MLE

Likelihood

Say we see outcome HHTHH.

You tell me your best guess about the value of the unknown parameter θ (a.k.a. p) is $4/5$. Is there some way that you can argue “objectively” that this is the best estimate?

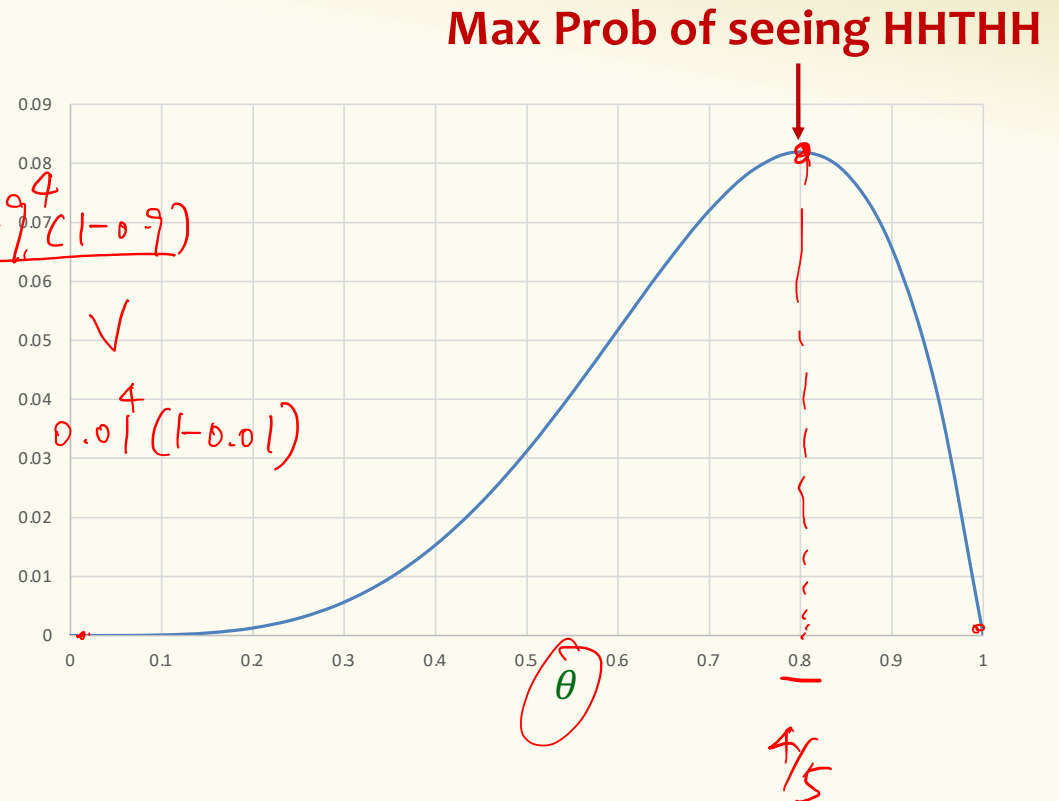
Likelihood

Say we see outcome HHTHH. $\theta=0.9 \Rightarrow 0.9^4 (1-0.9)$

$$\mathcal{L}(HHTHH | \theta) = \theta^4 (1 - \theta)$$

Probability of observing the outcome HHTHH if $\theta =$ prob. of heads. $\theta=0.01 \Rightarrow 0.01^4 (1-0.01)$

For a fixed outcome HHTHH, this is a function of θ .



Likelihood of Different Observations

(Discrete case)

Definition. The **likelihood** of independent observations x_1, \dots, x_n is

$$\mathcal{L}(x_1, \dots, x_n | \theta) = \prod_{i=1}^n P(x_i; \theta)$$

Maximum Likelihood Estimation (MLE). Given data x_1, \dots, x_n , find $\hat{\theta}$ such that $\mathcal{L}(x_1, \dots, x_n | \hat{\theta})$ is maximized!

$$\hat{\theta} = \operatorname{argmax}_{\theta} \mathcal{L}(x_1, \dots, x_n | \theta)$$

Usually: Solve $\frac{\partial \mathcal{L}(x_1, \dots, x_n | \theta)}{\partial \theta} = 0$ or $\frac{\partial \ln \mathcal{L}(x_1, \dots, x_n | \theta)}{\partial \theta} = 0$ [+check it's a max!]

Likelihood vs. Probability

- Fixed θ : **probability** $\prod_{i=1}^n P(x_i; \theta)$ that dataset x_1, \dots, x_n is sampled by distribution with parameter θ
 - A function of x_1, \dots, x_n
- Fixed x_1, \dots, x_n : **likelihood** $\mathcal{L}(x_1, \dots, x_n | \theta)$ that parameter θ explains dataset x_1, \dots, x_n .
 - A function of θ

These notions are the same number if we fix both x_1, \dots, x_n and θ , but different role/interpretation

Example – Coin Flips

$$\hat{\theta} = \frac{n_H}{n}$$

Observe: Coin-flip outcomes x_1, \dots, x_n , with n_H heads, n_T tails

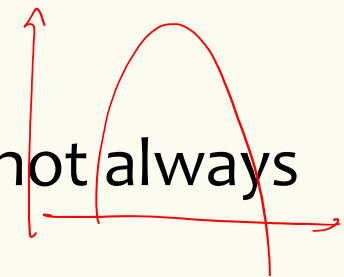
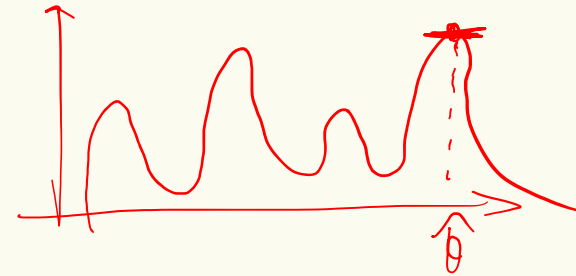
– i.e., $n_H + n_T = n$

Goal: estimate θ = prob. heads.

$$\mathcal{L}(x_1, \dots, x_n | \theta) = \theta^{n_H} (1 - \theta)^{n_T}$$

$$\frac{\partial}{\partial \theta} \mathcal{L}(x_1, \dots, x_n | \theta) = ???$$

$$\frac{\partial}{\partial \theta} \theta^{n_H} (1 - \theta)^{n_T}$$



While it is possible to compute this derivative, it's not always nice since we are working with products.

Log-Likelihood

We can save some work if we use the **log-likelihood** instead of the likelihood directly.

Definition. The **log-likelihood** of independent observations

x_1, \dots, x_n is

$$\ln \mathcal{L}(x_1, \dots, x_n | \theta) = \ln \prod_{i=1}^n P(x_i; \theta) = \sum_{i=1}^n \ln P(x_i; \theta)$$

Useful log properties

$$\begin{aligned}\ln(ab) &= \ln(a) + \ln(b) \\ \ln(a/b) &= \ln(a) - \ln(b) \\ \ln(a^b) &= \underline{b \cdot \ln(a)}\end{aligned}$$

Example – Coin Flips

Observe: Coin-flip outcomes x_1, \dots, x_n , with n_H heads, n_T tails

– i.e., $n_H + n_T = n$

Goal: estimate θ = prob. heads.

$$\mathcal{L}(x_1, \dots, x_n | \theta) = \theta^{n_H} (1 - \theta)^{n_T}$$

$$\ln \mathcal{L}(x_1, \dots, x_n | \theta) = \underline{n_H \ln \theta} + \underline{n_T \ln(1 - \theta)}$$

$$\rightarrow \frac{\partial}{\partial \theta} \ln \mathcal{L}(x_1, \dots, x_n | \theta) = \underline{n_H \cdot \frac{1}{\theta}} - \underline{n_T \cdot \frac{1}{1 - \theta}}$$

Want value $\hat{\theta}$ of θ s.t. $\frac{\partial}{\partial \theta} \ln \mathcal{L}(x_1, \dots, x_n | \theta) = 0$

So we need $\underline{n_H \cdot \frac{1}{\hat{\theta}} - n_T \cdot \frac{1}{1 - \hat{\theta}}} = 0$

Solving gives

$$\hat{\theta} = \frac{n_H}{n}$$

General Recipe

1. **Input** Given n i.i.d. samples x_1, \dots, x_n from parametric model with parameter θ .
2. **Likelihood** Define your likelihood $\mathcal{L}(x_1, \dots, x_n | \theta)$.
 - For discrete $\mathcal{L}(x_1, \dots, x_n | \theta) = \prod_{i=1}^n P(x_i; \theta)$ ←
3. **Log** Compute $\ln \mathcal{L}(x_1, \dots, x_n | \theta)$
4. **Differentiate** Compute $\frac{\partial}{\partial \theta} \ln \mathcal{L}(x_1, \dots, x_n | \theta)$
5. **Solve for $\hat{\theta}$** by setting derivative to 0 and solving for max.

Generally, you need to do a second derivative test to verify it is a maximum, but we won't ask you to do that in CSE 312.

Brain Break



Agenda

- Idea: Estimation
- Maximum Likelihood Estimation (example: mystery coin)
- Continuous MLE ◀

The Continuous Case

Given n (independent) samples x_1, \dots, x_n from (continuous) parametric model $f(x_i; \theta)$ which is now a family of densities

Definition. The likelihood of independent observations x_1, \dots, x_n is

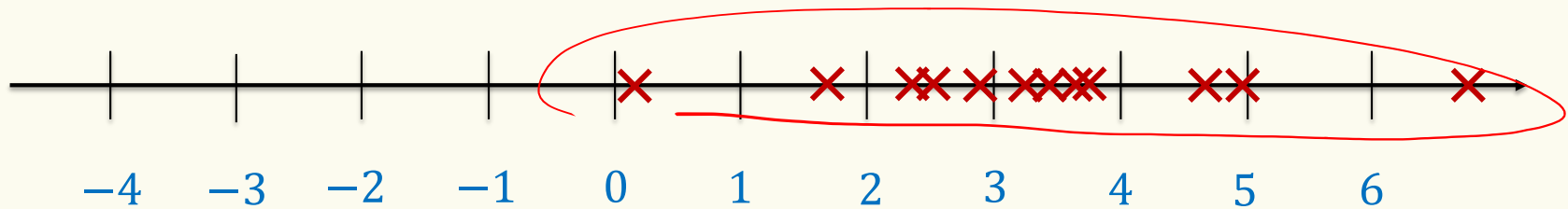
$$\mathcal{L}(x_1, \dots, x_n | \theta) = \prod_{i=1}^n f(x_i; \theta)$$

Density function! (Why?)

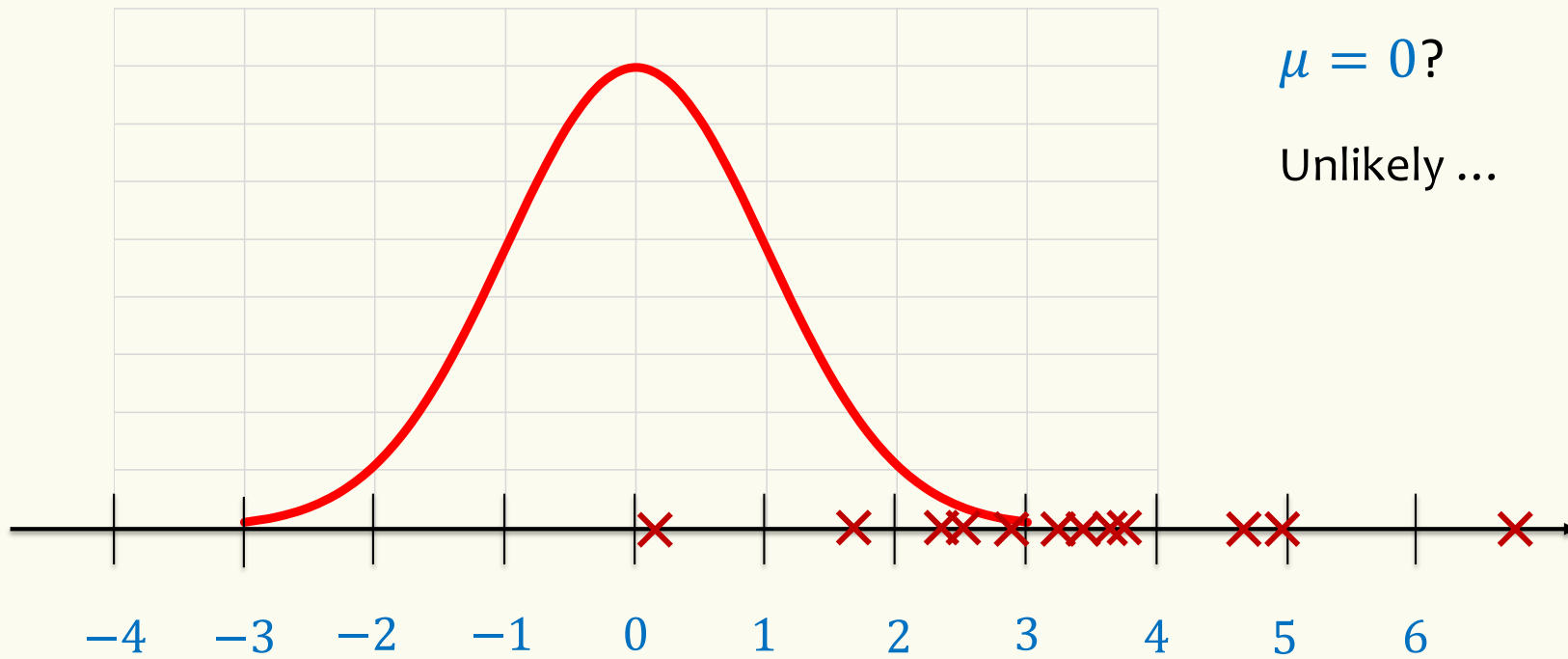
Why density?

- Density \neq probability, but:
 - For maximizing likelihood, **we really only care about relative likelihoods**, and density captures that
 - has desired property that likelihood increases with better fit to the model

n samples $x_1, \dots, x_n \in \mathbb{R}$ from Gaussian $\mathcal{N}(\mu, 1)$. Most likely μ ?
[i.e., we are given the promise that the variance is 1]



n samples $x_1, \dots, x_n \in \mathbb{R}$ from Gaussian $\mathcal{N}(\mu, 1)$. Most likely μ ?



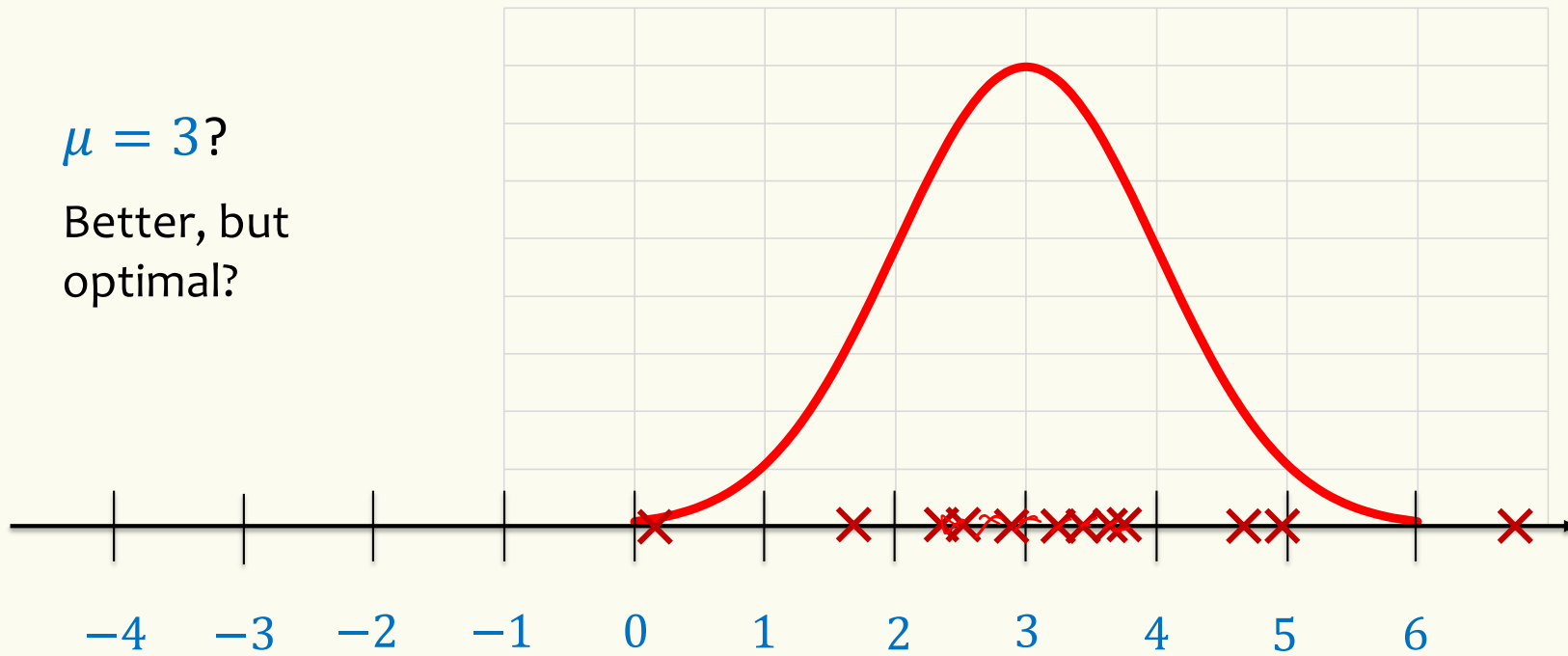
$\mu = 0$?

Unlikely ...

n samples $x_1, \dots, x_n \in \mathbb{R}$ from Gaussian $\mathcal{N}(\mu, 1)$. Most likely μ ?

$\mu = 3$?

Better, but
optimal?



Example – Gaussian Parameters

Normal outcomes x_1, \dots, x_n , known variance $\sigma^2 = 1$ but *unknown* mean μ

Goal: estimate $\theta = \text{mean}$

Next time:

$$\hat{\theta} = \frac{\sum_i^n x_i}{n}$$

In other words, MLE is the *sample mean* of the data.

$$\ln \frac{1}{x}$$

$$g \circ (f(x))$$

~~⊗~~

General Recipe

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2. **Likelihood** Define your likelihood $\mathcal{L}(x_1, \dots, x_n | \theta)$.
 - For discrete $\mathcal{L}(x_1, \dots, x_n | \theta) = \prod_{i=1}^n P(x_i; \theta)$
 - For continuous $\mathcal{L}(x_1, \dots, x_n | \theta) = \prod_{i=1}^n f(x_i; \theta)$
3. **Log** Compute $\ln \mathcal{L}(x_1, \dots, x_n | \theta)$
4. **Differentiate** Compute $\frac{\partial}{\partial \theta} \ln \mathcal{L}(x_1, \dots, x_n | \theta)$
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