

## Balls, Urns, and the Supreme Court

- Supreme Court case: *Berghuis v. Smith*  
*If a group is underrepresented in a jury pool, how do you tell?*
    - Article by Erin Miller – Friday, January 22, 2010
    - Thanks to Josh Falk for pointing out this article
- Justice Breyer [Stanford Alum] opened the questioning by invoking the binomial theorem. He hypothesized a scenario involving “**an urn with a thousand balls, and sixty are red, and nine hundred forty are black, and then you select them at random... twelve at a time.**” According to Justice Breyer and the binomial theorem, if the red balls were black jurors then “**you would expect... something like a third to a half of juries would have at least one black person**” on them.
- Justice Scalia’s rejoinder: “We don’t have any urns here.”

## Justice Breyer Meets CS109

- Should model this combinatorially
    - Ball draws not independent trials (balls not replaced)
  - Exact solution:
 
$$P(\text{draw 12 black balls}) = \binom{940}{12} / \binom{1000}{12} \approx 0.4739$$

$$P(\text{draw } \geq 1 \text{ red ball}) = 1 - P(\text{draw 12 black balls}) \approx 0.5261$$
  - Approximation using Binomial distribution
    - Assume  $P(\text{red ball})$  constant for every draw =  $60/1000$
    - $X = \#$  red balls drawn.  $X \sim \text{Bin}(12, 60/1000 = 0.06)$
    - $P(X \geq 1) = 1 - P(X = 0) \approx 1 - 0.4759 = 0.5240$
- In Breyer’s description, should actually expect just over half of juries to have at least one black person on them*

Demo

## From Discrete to Continuous

- So far, all random variables we saw were *discrete*
  - Have finite or countably infinite values (e.g., integers)
  - Usually, values are binary or represent a count
- Now it’s time for *continuous* random variables
  - Have (uncountably) infinite values (e.g., real numbers)
  - Usually represent measurements (arbitrary precision)
    - Height (centimeters), Weight (lbs.), Time (seconds), etc.
- Difference between how many and how much
- Generally, it means replace  $\sum_{x=a}^b f(x)$  with  $\int_a^b f(x)dx$

## Continuous Random Variables

- $X$  is a **Continuous Random Variable** if there is function  $f(x) \geq 0$  for  $-\infty \leq x \leq \infty$ , such that:
 
$$P(a \leq X \leq b) = \int_a^b f(x)dx$$
- $f$  is a Probability Density Function (PDF) if:
 
$$P(-\infty < X < \infty) = \int_{-\infty}^{\infty} f(x)dx = 1$$

## Probability Density Functions

- Say  $f$  is a **Probability Density Function** (PDF)
 
$$P(-\infty < X < \infty) = \int_{-\infty}^{\infty} f(x)dx = 1$$
  - $f(x)$  is **not** a probability, it is probability/units of  $X$
  - Not meaningful without some subinterval over  $X$ 

$$P(X = a) = \int_a^a f(x)dx = 0$$
- Contrast with Probability Mass Function (PMF) in discrete case:  $p(a) = P(X = a)$ 
  - where  $\sum_{i=1}^{\infty} p(x_i) = 1$  for  $X$  taking on values  $x_1, x_2, x_3, \dots$

## Cumulative Distribution Functions

- For a continuous random variable  $X$ , the **Cumulative Distribution Function** (CDF) is:

$$F(a) = P(X < a) = P(X \leq a) = \int_{-\infty}^a f(x) dx$$

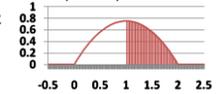
- Density  $f$  is derivative of CDF  $F$ :  $f(a) = \frac{d}{da} F(a)$

- For continuous  $f$  and small  $\varepsilon$ :

$$P(a - \frac{\varepsilon}{2} \leq X \leq a + \frac{\varepsilon}{2}) = \int_{a-\varepsilon/2}^{a+\varepsilon/2} f(x) dx \approx \varepsilon f(a)$$

## Simple Example

- $X$  is continuous random variable (CRV) with PDF:

$$f(x) = \begin{cases} C(4x - 2x^2) & \text{when } 0 < x < 2 \\ 0 & \text{otherwise} \end{cases}$$


- What is  $C$ ?

$$\int_0^2 C(4x - 2x^2) dx = 1 \Rightarrow C \left( 2x^2 - \frac{2x^3}{3} \right) \Big|_0^2 = 1$$

$$C \left( \left( 8 - \frac{16}{3} \right) - 0 \right) = 1 \Rightarrow C \frac{8}{3} = 1 \Rightarrow C = \frac{3}{8}$$

- What is  $P(X > 1)$ ?

$$\int_1^2 f(x) dx = \int_1^2 \frac{3}{8} (4x - 2x^2) dx = \frac{3}{8} \left( 2x^2 - \frac{2x^3}{3} \right) \Big|_1^2 = \frac{3}{8} \left[ \left( 8 - \frac{16}{3} \right) - \left( 2 - \frac{2}{3} \right) \right] = \frac{1}{2}$$

## Disk Crashes

- $X$  = hours before your disk crashes

$$f(x) = \begin{cases} \lambda e^{-x/100} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

- First, determine  $\lambda$  to have actual PDF

- Good integral to know:  $\int e^u du = e^u$

$$1 = \int \lambda e^{-x/100} dx = -100\lambda \int \frac{1}{100} e^{-x/100} dx = -100\lambda e^{-x/100} \Big|_0^\infty = 100\lambda \Rightarrow \lambda = \frac{1}{100}$$

- What is  $P(50 < X < 150)$ ?

$$F(150) - F(50) = \int_{50}^{150} \frac{1}{100} e^{-x/100} dx = -e^{-x/100} \Big|_{50}^{150} = -e^{-3/2} + e^{-1/2} \approx 0.383$$

- What is  $P(X < 10)$ ?

$$F(10) = \int_0^{10} \frac{1}{100} e^{-x/100} dx = -e^{-x/100} \Big|_0^{10} = -e^{-1/10} + 1 \approx 0.095$$

## Expectation and Variance

For discrete RV  $X$ :

$$E[X] = \sum_x x p(x)$$

$$E[g(X)] = \sum_x g(x) p(x)$$

$$E[X^n] = \sum_x x^n p(x)$$

For continuous RV  $X$ :

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx$$

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) f(x) dx$$

$$E[X^n] = \int_{-\infty}^{\infty} x^n f(x) dx$$

For both discrete and continuous RVs:

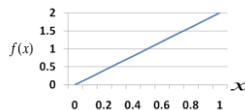
$$E[aX + b] = aE[X] + b$$

$$\text{Var}(X) = E[(X - \mu)^2] = E[X^2] - (E[X])^2$$

$$\text{Var}(aX + b) = a^2 \text{Var}(X)$$

## Linearly Increasing Density

- $X$  is a continuous random variable with PDF:

$$f(x) = \begin{cases} 2x & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$


- What is  $E[X]$ ?

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx = \int_0^1 2x^2 dx = \frac{2}{3} x^3 \Big|_0^1 = \frac{2}{3}$$

- What is  $\text{Var}(X)$ ?

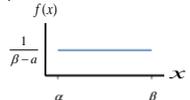
$$E[X^2] = \int_{-\infty}^{\infty} x^2 f(x) dx = \int_0^1 2x^3 dx = \frac{1}{2} x^4 \Big|_0^1 = \frac{1}{2}$$

$$\text{Var}(X) = E[X^2] - (E[X])^2 = \frac{1}{2} - \left( \frac{2}{3} \right)^2 = \frac{1}{18}$$

## Uniform Random Variable

- $X$  is a **Uniform Random Variable**:  $X \sim \text{Uni}(\alpha, \beta)$

- Probability Density Function (PDF):

$$f(x) = \begin{cases} \frac{1}{\beta - \alpha} & \alpha < x < \beta \\ 0 & \text{otherwise} \end{cases}$$


$$P(a \leq x \leq b) = \int_a^b f(x) dx = \frac{b - a}{\beta - \alpha}$$

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx = \int_{\alpha}^{\beta} \frac{x}{\beta - \alpha} dx = \frac{\beta^2 - \alpha^2}{2(\beta - \alpha)} = \frac{\alpha + \beta}{2}$$

$$\text{Var}(X) = \frac{(\beta - \alpha)^2}{12}$$

## Fun with the Uniform Distribution

- $X \sim \text{Uni}(0, 20)$

$$f(x) = \begin{cases} \frac{1}{20} & 0 < x < 20 \\ 0 & \text{otherwise} \end{cases}$$

- $P(X < 6)$ ?

$$P(x < 6) = \int_0^6 \frac{1}{20} dx = \frac{6}{20}$$

- $P(4 < X < 17)$ ?

$$P(4 < x < 17) = \int_4^{17} \frac{1}{20} dx = \frac{17}{20} - \frac{4}{20} = \frac{13}{20}$$

## Riding the Marguerite Bus

- Say the Marguerite bus stops at the Gates bldg. at 15 minute intervals (2:00, 2:15, 2:30, etc.)
  - Passenger arrives at stop uniformly between 2-2:30pm
  - $X \sim \text{Uni}(0, 30)$

- $P(\text{Passenger waits} < 5 \text{ minutes for bus})$ ?

- Must arrive between 2:10-2:15pm or 2:25-2:30pm

$$P(10 < X < 15) + P(25 < x < 30) = \int_{10}^{15} \frac{1}{30} dx + \int_{25}^{30} \frac{1}{30} dx = \frac{5}{30} + \frac{5}{30} = \frac{1}{3}$$

- $P(\text{Passenger waits} > 14 \text{ minutes for bus})$ ?

- Must arrive between 2:00-2:01pm or 2:15-2:16pm

$$P(0 < X < 1) + P(15 < x < 16) = \int_0^1 \frac{1}{30} dx + \int_{15}^{16} \frac{1}{30} dx = \frac{1}{30} + \frac{1}{30} = \frac{1}{15}$$

## When to Leave For Class

- Biking to a class on campus

- Leave  $t$  minutes before class starts
- $X$  = travel time (minutes).  $X$  has PDF:  $f(x)$
- If early, incur cost:  $c/\text{min}$ . If late, incur cost:  $k/\text{min}$ .

$$\text{Cost: } C(X, t) = \begin{cases} c(t - X) & \text{if } x < t \\ k(X - t) & \text{if } x \geq t \end{cases}$$

- Choose  $t$  (when to leave) to minimize  $E[C(X, t)]$ :

$$E[C(X, t)] = \int_0^t C(X, t) f(x) dx + \int_t^\infty C(X, t) f(x) dx = \int_0^t c(t - x) f(x) dx + \int_t^\infty k(x - t) f(x) dx$$

## Minimization via Differentiation

- What to minimize w.r.t.  $t$ :

$$E[C(X, t)] = \int_0^t c(t - x) f(x) dx + \int_t^\infty k(x - t) f(x) dx$$

- Differentiate  $E[C(X, t)]$  w.r.t.  $t$ , and set = 0 (to obtain  $t^*$ ):
  - Leibniz integral rule:

$$\frac{d}{dt} \int_{f_1(t)}^{f_2(t)} g(x, t) dx = \frac{df_2(t)}{dt} g(f_2(t), t) - \frac{df_1(t)}{dt} g(f_1(t), t) + \int_{f_1(t)}^{f_2(t)} \frac{\partial g(x, t)}{\partial t} dx$$

$$\frac{d}{dt} E[C(X, t)] = c(t - t)f(t) + \int_0^t cf(x) dx - k(t - t)f(t) - \int_t^\infty kf(x) dx$$

$$0 = cF(t^*) - k[1 - F(t^*)] \Rightarrow F(t^*) = \frac{k}{c + k}$$